# Two insights from nonparametric statistics on cosmology research 

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## Nonparametric Statistics

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- The spirit of nonparametrics also appears in other problem such as causal inference, graphical models, and the analysis of missing data (in particular, imputation).
- It offers a flexible way to investigate the underlying structures.
- Examples in today's talk

1. Density estimation and the discovery of large-scale structures.
2. Analysis of bias from using the best fit (imputation).

Part 1: Density estimation and detection of large-scale structures.

## Cosmic Web: What Does Our Universe Look Like



Credit: Millennium Simulation

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## The Data

Here is what our data looks like:


## Filament finding problem

- In simulations, we saw that there are clear filamentary structures.
- In the real data, we also saw some weakly filamentary forms in the distribution.
- How to recover filaments from the data is an open problem.


## Consensus about Filaments

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But there are some common properties that a filament should have (Bond et al. 1996):

- It is a curve-like structure.
- It characterizes high (matter) density area.
- It shows connectivity of the matter distribution.


## Density Ridges

We formalize the notion of filaments as density ridges.

## Example: Ridges in Mountains



Credit: Google

## Example: Ridges in Smooth Functions



## Example: Ridges in Smooth Functions



## Ridges: Local Modes in Subspace



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## Finding ridges

- Original data.
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- Thresholding (denoising).
- Ridge finding.



## Example for Estimated Density Ridges



## Example for Estimated Density Ridges



## 3D Example for Estimated Ridges



Blue curves: density ridges.
Red points: density local modes.

## SDSS: Comparing to Clusters

- Blue: filaments. Red: galaxy clusters (redMaPPer).



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## SDSS: Filament Effects VS Environments

Do filaments have an extra effect other than environments?

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Do filaments have an extra effect other than environments?
$\longrightarrow$ Yes!


## SDSS: Color



Similar pattern also appears for other galaxy properties such as brightness, size, and age.

## The Alignment of a Galaxy along a Filament - 1

Theorists have conjectured about the alignment of galaxy along nearby filaments.


We now try to test such a conjecture.

## The Alignment of a Galaxy along a Filament - 2

We can easily define the orientation of filament because it is a curve.
For each galaxy, we measure its orientation by fitting an ellipse.


We are interested in the inner product between the major axis of a galaxy and the orientation of the nearest filament.

## Excess Probability Density



Y -axis: the ratio of observed angular distribution versus a uniform distribution over [0, 90] deg.
If no alignment, the ratio should be 1 .

## Part 2: The danger of using the best fit

## A common value-added data

|  | Mass | M. err | Age | A. err | RA | dec | redshift | others |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Galaxy 1 | $M_{1}$ | $E_{M, 1}$ | $A_{1}$ | $E_{A, 1}$ | $\mathrm{RA}_{1}$ | $\operatorname{dec}_{1}$ | $Z_{1}$ | $O_{1}$ |
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| $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ |

- Blue variables: directly observed using the telescope.
- Red variables: unobserved, inferred from the observables.
- Q: is it reliable to use the inferred variables (often best fitted) to make scientific conclusion?


## A motivating example: detecting filament effects

Here we attempted to analyze the effect from filaments on the age-mass relation (regression coefficient).
We use the best fitted mass and age from the data.


## Model prediction

- For a galaxy, let $O$ denotes all its observed profiles.
- A model that associates $O$ with age $(A)$ and mass $(M)$ can be viewed as a distribution/likelihood of $A, M$ given $O$ :

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- In many data products, we have predicted mass and age of each galaxy. Generally, the predicted mass and age are

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- In our previous analysis, we were computing the association between age and mass via the best fitted value $\widehat{A}, \widehat{M}$.
- Namely, we are ignoring the uncertainty of $A, M$ in our analysis. Will this be okay?


## Examining simulations

- To investigate the effect of ignoring the uncertainty, we use simulation data.
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- So our simulation data can be summarized as

$$
\left(M_{1}, A_{1}, O_{1}\right), \cdots,\left(M_{n}, A_{n}, O_{n}\right)
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- To predict $M, A$ from $O$, we consider a simple linear regression and a 50-nearest neighbor (NN) regression.


## Simulations: updated data

|  | Original |  | Linear model |  | $50-\mathrm{NN}$ |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Galaxy 1 | $M_{1}$ | $A_{1}$ | $\widehat{M}_{L M, 1}$ | $\widehat{A}_{L M, 1}$ | $\widehat{M}_{k N N, 1}$ | $\widehat{A}_{k N N, 1}$ |
| Galaxy 2 | $M_{2}$ | $A_{2}$ | $\widehat{M}_{L M, 2}$ | $\widehat{A}_{L M, 2}$ | $\widehat{M}_{k N N, 2}$ | $\widehat{A}_{k N N, 2}$ |
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| $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ |

- The true regression coefficient is obtained by regressing $Y=A_{i}$ with $X=M_{i}$.
- Question: if we use the best fitted/predicted values from the linear model or nonparametric model, will we obtain a similar regression coefficients?


## Failure of best fitted method

## MBII simulation



No! It fails miserably!

## Why the best fitted values fail?

- Often the best predicted value of $A, M$ from $O$ is the conditional mean $\mathbb{E}(A \mid O)$ and $\mathbb{E}(M \mid O)$.
- Regressing $A$ with $M$ is different from regressing $\mathbb{E}(A \mid O)$ with $\mathbb{E}(M \mid O)$ !


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- Take the covariance as an example, by law of total covariance,

$$
\operatorname{Cov}(A, M)=\operatorname{Cov}(\mathbb{E}(A \mid O), \mathbb{E}(M \mid O))+\mathbb{E}(\operatorname{Cov}(\mathrm{A}, \mathrm{M} \mid \mathrm{O}))
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- The first term is what we compute when using the best fitted value.
- But it ignores the second term!


## A simple remedy: random imputation

- Here is a simple remedy: instead of using the best fitted value, we take a random draw from the conditional density $p(A, M \mid O)$ (known as random imputation) ${ }^{1}$.

MBII simulation


[^0]
## Why random imputation works? - 1

- In the ideal case where we get to observe the age and mass directly, our data can be summarized as IID random vectors

$$
\left(M_{1}, A_{1}, O_{1}\right), \cdots,\left(M_{n}, A_{n}, O_{n}\right) \sim p(m, a, o)
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- A measure of association between age and mass can often be written as $\theta(M, A)$ and we are interested in the population average

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\theta=\mathbb{E}(\theta(M, A))=\int \theta(m, a) p(m, a) d m d a
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$$

where $p(m, a)$ is the joint density of $M, A$.

- In practice, mass and age are missing, what we observe are IID random elements

$$
O_{1}, \cdots, O_{n} \sim p(o)
$$

## Why random imputation works? - 2

- The decomposition

$$
p(m, a, o)=p(m, a \mid o) p(o)
$$

implies that if we augment the $i$-th galaxy with random numbers $\left(M_{i}^{*}, A_{i}^{*}\right)$ from $p\left(m, a \mid O_{i}\right)$, the triplet can be viewed from

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- Thus, as long as we independently draw mass and age from the conditional density, we obtain a dataset that behaves like a fully observed data.
- Then we can use

$$
\left(M_{1}^{*}, A_{1}^{*}\right), \cdots,\left(M_{n}^{*}, A_{n}^{*}\right)
$$

to accurately estimate $\theta=\mathbb{E}(\theta(M, A))$.

## Why random imputation works? (Visually)

Mass-age (Linear best fit)


## Why random imputation works? (Visually)

Mass-age (kNN best fit)


## Why random imputation works? (Visually)

Mass-age (kNN Random Imputation)


## Multiple imputation and Monte Carlo errors

- Although the above procedure gives us an unbiased estimator, it may suffer from the Monte Carlo errors if we only impute the unobserved entries once.
- In general, we should repeat this imputation multiple times, creating multiple imputed data, and compute the final estimates. ${ }^{2}$

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## Multiple imputation and Monte Carlo errors

- Although the above procedure gives us an unbiased estimator, it may suffer from the Monte Carlo errors if we only impute the unobserved entries once.
- In general, we should repeat this imputation multiple times, creating multiple imputed data, and compute the final estimates. ${ }^{2}$
- Luckily, in most Astronomy survey, the sample size is large so the Monte Carlo errors are small.

[^2]
## What if we only know the marginal error? - 1

Recall the original dataset:

|  | Mass | M. err | Age | A. err | RA | dec | redshift | others |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
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- We do have the errors that represents the marginal distribution of $p(M \mid O)$ and $p(A \mid O)$.
- If this is all we have, can we make a better inference?
- Note: $E_{M, 1}$ can be viewed as the SD of $p(M \mid O)$.


## What if we only know the marginal error? - 2

- If we assume that $M \mid O$ and $A \mid O$ follow a normal distribution, then the above table gives us information about $M_{i} \mid O_{i}$ and $A_{i} \mid O_{i}:$

$$
M_{i}\left|O_{i} \sim N\left(M_{i}, E_{M, i}\right), \quad A_{i}\right| O_{i} \sim N\left(A_{i}, E_{A, i}\right)
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- It seems that we can generate from the distribution $p(m, a \mid o)$ using this information.
- Actually, we CANNOT-we still need to know the (conditional) correlation between the two random variables.
- Namely, we need $\operatorname{Cor}\left(A_{i}, M_{i} \mid O_{i}\right)$ to reconstruct $p(a, m \mid o)$.


## Sensitivity analysis: a partial solution

- Here is a simple method to roughly investigate the effect-we assume a single number for all correlations and evaluate how it influences our final result.


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Distribution of Mass-Age Correlation


## Sensitivity analysis: a graphical illustration

Mass-age (same correlation)


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- And the error due to the imputation is way higher than the estimation errors, which means that we should not ignore this effect.


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- As can be seen from the analysis on SDSS, the problem is very severe!
- We saw that the effect (from filaments) may reverse the direction if we incorrectly specify the correlation.
- And the error due to the imputation is way higher than the estimation errors, which means that we should not ignore this effect.
- Note: here we assume that the correlation is the same across different galaxies, but in simulation, we know that they are not the same.


## Challenges of random imputation - 1

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- Moreover, this idea works only if the normal distribution assumption is correct!


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- However, this would increase the number of variables a lot-if we have $k$ inferred variables, we would have $\binom{k}{2}$ correlations.
- Moreover, this idea works only if the normal distribution assumption is correct!
- The normal distribution may not be correct in practice, so even if we have all correlations, our estimate may still be inaccurate.


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- As long as the sample size is sufficiently large, such procedure will give us a reliable estimate (Monte Carlo error will not be an issue).


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- Suppose that we have $k$ inferred variables, this would only require adding additionally $k$ variables to the original data.
- As long as the sample size is sufficiently large, such procedure will give us a reliable estimate (Monte Carlo error will not be an issue).
- Of course, this idea relies on the assumption that the conditional density is correct, which is another strong assumption.


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1. Statistical methods offers new exciting tools in Astronomy.
2. A good tool allows us to detect weak signals.
3. When we are using multiple derived variables, we need to be careful.
4. Using the best fitted values may result in bias in the estimation.
5. Random imputation offers a solution to this problem.

MBII simulation


## Thank you!

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## Algorithm

1. Rawdata


## Algorithm

1. Rawdata
2. Density Reconstruction


## Algorithm

1. Rawdata
2. Density Reconstruction
3. Thresholding


## Algorithm

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1. Rawdata
}
2. Density Reconstruction
3. Thresholding
4. Ridge Recovery


## Bandwidth Selection




## Bandwidth Selection




## Bandwidth Selection


$L_{1}$ distance are like the area of the shady regions. We estimate this distance by data splitting or the bootstrap. Reference: Chen et al. 'Optimal Ridge Detection using Coverage Risk' (NIPS 2015).

## General Ridges

We can generalize ridges to higher dimensions. Pick
$V_{r}(x)=\left[v_{r+1}(x), \cdots, v_{d}(x)\right]$.
We define

$$
r \text {-Ridge }(p)=\left\{x: V_{r}(x) V_{r}(x)^{T} \nabla p(x)=0, \lambda_{r+1}(x)<0\right\}
$$

$V_{r}(x)$ is a $d \times(d-r)$ matrix. There are $d-r$ constraints.
By Implicit Function Theorem, $r$-ridges are $r$-manifolds.
In Astronomy, $r=2$ can be used to model 'Cosmic Sheets (Walls)'.
$r=0$ coincides with the definition of local modes.

## Density Ridges on the SDSS data



## Density Ridges on the SDSS data



## Curse of Number Density

## Curse of Number Density



## Curse of Number Density



## Curse of Number Density



## SDSS: Red and Blue Galaxies

- Redshift range: $0.05<z<0.20$ (main sample galaxy).
- Color cut: $(g-r)=0.8$.


## SDSS: Red and Blue Galaxies

Filaments, MGS


Clusters (redMaPPer), MGS


## SDSS: Size for Galaxies

1. Size: Effective Radii.
2. Data: LOWZ ( $0.20<z<0.43$ )
3. Partitioning galaxies into three groups according to their size.


## SDSS: Size for Galaxies

Filaments, LOWZ


Clusters (redMaPPer), LOWZ


## Age for Galaxies



## Age for Galaxies



## Age for Galaxies





## Comparison: Voronoi Model



## Comparison: Voronoi Model

Ridges and all galaxies


## Comparison: Voronoi Model

Ridges and Clusters (Voronoi)


## Comparison: Voronoi Model

Ridges and Filaments (Voronoi)


## Comparison: Voronoi Model

## Ridges and Walls (Voronoi)



## Comparison: Voronoi Model

Ridges and Voids (Voronoi)



[^0]:    ${ }^{1}$ Estimated by the $50-\mathrm{NN}$ in this case.

[^1]:    ${ }^{2}$ known as the multiple imputation.

[^2]:    ${ }^{2}$ known as the multiple imputation.

