

1. *In the genetic screening example, calculate the (unconditional) probability of testing positive, and of having the genetic variant.*

We have:

$$\begin{aligned}\mathbb{P}[\text{variant}] &= 0.01 \\ \mathbb{P}[\text{no variant}] &= 0.99 \\ \mathbb{P}[\text{test +ve} \mid \text{variant}] &= 0.8 \text{ (80\% sensitivity)} \\ \mathbb{P}[\text{test -ve} \mid \text{variant}] &= 0.2 \\ \mathbb{P}[\text{test -ve} \mid \text{no variant}] &= 0.95 \text{ (95\% specificity)} \\ \mathbb{P}[\text{test +ve} \mid \text{no variant}] &= 0.05\end{aligned}$$

So:

$$\begin{aligned}\mathbb{P}[\text{test +ve}] &= \mathbb{P}[\text{test+ve} \mid \text{variant}] * \mathbb{P}[\text{variant}] + \mathbb{P}[\text{test+ve} \mid \text{no variant}] * \mathbb{P}[\text{no variant}] \\ &= 0.8 * 0.01 + 0.05 * 0.99 = 0.0575, \text{ i.e. } 5.75\% \\ \mathbb{P}[\text{have variant}] &= 0.01, \text{ i.e. } 1\%\end{aligned}$$

Compare the probability of having the genetic variant given a positive test to the unconditional probability of having the genetic variant: did getting a positive test result change much?

$\mathbb{P}[\text{have variant} \mid \text{+ve result}]$ is 14%, up by a factor of 14 from $\mathbb{P}[\text{have variant}] = 1\%$. So the amount you should be worried based on a positive test results is a *lot* more than you would be with no test at all. But it's still a (fairly) small amount.

2. *In the ASE example, verify that when $Y=0$ (i.e. no expression observed in BY) the classical analysis give a point estimate $=0$ and estimated standard error $=0$. Why are both these values implausible?*

The point estimates is the sample mean, \bar{Y} , which is zero. The standard error is $\frac{\sqrt{\bar{Y}(1-\bar{Y})}}{\sqrt{n}}$, which is again zero. These are silly values! The first would imply that Brewers' Yeast can *never* produce expression values at some particular location, i.e. ruling it out – but based on a small sample size and not using any other information source. The zero standard error indicates we would *never* expect this estimate to change, regardless of how many times we ran the experiment. Expressing complete confidence in a somewhat extreme situation, based on this data alone, would not be creditable.

3. *Human twins can be either monozygotic (from a single egg) or dizygotic (from two eggs). Among all sets of twins, approximately one third are monozygotic, and the ratio of males to females can be taken to be 50:50, regardless of monozygosity.*

Suppose it is known that someone will have twins, e.g., from detection of two heartbeats, and a sonogram indicates there are twin girls. What is the probability that they are monozygotic?

Hint: make a Venn diagram for events G : the twins are both girls and M : the twins are monozygotic. We want $\Pr[M \mid G]$, which is related to $\Pr[G \mid M]$.

Writing out all the algebra:

$$\mathbb{P}[G \mid M] = 1/2$$

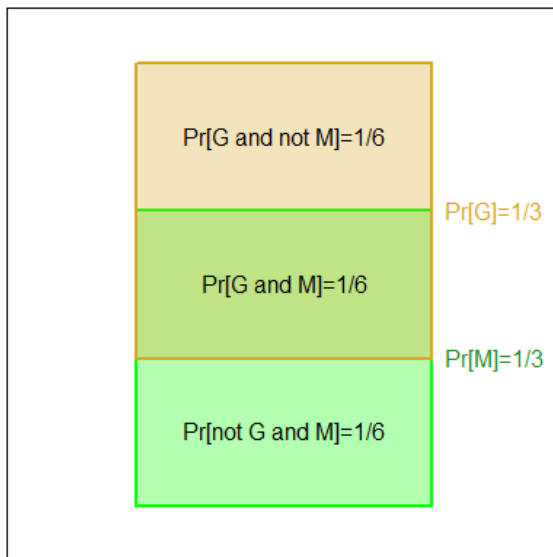
$$\mathbb{P}[G] = \mathbb{P}[G|M] \cdot \mathbb{P}[M] + \mathbb{P}[G|\text{not } M] \cdot \mathbb{P}[\text{not } M] = 1/2 \cdot 1/3 + 1/4 \cdot 2/3 = 1/3$$

$$\mathbb{P}[M] = 1/3$$

$$\mathbb{P}[M|G] = \mathbb{P}[G|M] \cdot \mathbb{P}[M] / \mathbb{P}[G] = 1/2 \cdot 1/3 / 1/3 = 1/2$$

So belief that there are identical twins should go up from 1/3 to 1/2 based on the sonogram. (Or use [cell-free DNA testing](#) and find out for sure!)

As a Venn diagram: (there are lots of ways to lay this out: the total area of the diagram is 1)



Or if you prefer to work with the sexes of both babies:

GG, not M: $1/4 \cdot 2/3 = 1/6$	GG, M: $1/2 \cdot 1/3 = 1/6$
GB, not M: $1/4 \cdot 2/3 = 1/6$	
BG, not M: $1/4 \cdot 2/3 = 1/6$	
BB, not M: $1/4 \cdot 2/3 = 1/6$	BB, M: $1/2 \cdot 1/3 = 1/6$

Code

```
par(mar=c(0,0,0,0)+0.1)
plot(0,0,type="n", xlab="", ylab="", axes=FALSE, xlim=0:1, ylim=0:1, asp=1)
```

```

box()
rect(0.5-sqrt(1/3)/2, 0.5-3*sqrt(1/3)/4, 0.5+sqrt(1/3)/2, 0.5+sqrt(1/3)/4,
col=adjustcolor("green",alpha=0.3), border="green", lwd=2)
rect(0.5-sqrt(1/3)/2, 0.5-sqrt(1/3)/4, 0.5+sqrt(1/3)/2, 0.5+3*sqrt(1/3)/4,
col=adjustcolor("goldenrod",alpha=0.3), border="goldenrod", lwd=2)
text(x=0.5, y=0.5+c(1,0,-1)*sqrt(1/3)/2,
c("Pr[G and not M]=1/6", "Pr[G and M]=1/6", "Pr[not G and M]=1/6"))
text(x=0.5 + sqrt(1/3)/2, y=0.5 + sqrt(1/3)/4, "Pr[G]=1/3", col="goldenrod",
pos=4)
text(x=0.5 + sqrt(1/3)/2, y=0.5 - sqrt(1/3)/4, "Pr[M]=1/3",
col="forestgreen", pos=4)

par(mar=c(0,0,0,0)+0.1)
plot(0,0,type="n", xlab="", ylab="", axes=FALSE, xlim=0:1, ylim=0:1, asp=1)
rect(rep(0,4), seq(0,0.75, 0.25), rep(2/3,4), seq(0.25, 1, 0.25))
text(rep(1/3,4), seq(1/8,7/8,1/4), c("BB, not M: 1/4 * 2/3 = 1/6", "BG, not
M: 1/4 * 2/3 = 1/6", "GB, not M: 1/4 * 2/3 = 1/6", "GG, not M: 1/4 * 2/3 =
1/6"))
rect(rep(2/3,2), c(0,0.5), rep(1,4), c(0.5, 1))
text(rep(5/6,4), c(0.25, 0.75), c("BB, M:\n1/2 * 1/3 = 1/6", "GG, M:\n1/2 *
1/3 = 1/6"))
rect(2/3, 0, 1, 1, , col=adjustcolor("green",alpha=0.3), border="green",
lwd=2)
polygon(x=c(0,2/3,2/3,1,1,0), y=c(3/4,3/4,1/2,1/2,1,1),
col=adjustcolor("goldenrod",alpha=0.3), border="goldenrod", lwd=2)

```