# Online Appendix for "What Does the Yield Curve Tell Us About Exchange Rate Predictability?" (Chen and Tsang 2011)

Version: June 2011

This Online Appendix contains the following Sections, Tables, and Figures:

- Section OA1: Out-of-Sample Forecasting discussion and results (Tables OA1 & OA2 for main sample and recent sample, respectively)
- Section OA2: Standard Errors for Overlapping Data using Simulated Critical Values
- Tables OA3 OA5: in-sample regressions with the inclusion of lagged dependent variable
  - These results demonstrate that the explanatory power of the relative factors is beyond the information contained in the time series of exchange rates themselves.
- Table OA6: Correlation between SPF Forecasts and the US Factors
- Tables OA7 & OA8: Predictive regression results using Most Recent Data and from Sub-Samples (Canada and the UK)
- Figures OA1 (a)-(c): Juxtaposing Rejections of the Interest Differential Model and the Factors
- Figures OA2 (a)&(b): Rolling Test of the Interest Differential Restrictions
- Figures OA3 (a)&(b): Recursive Adjusted R-squares with a 5-Year Rolling Window
- Figures OA4 (a)-(c): Rolling Clark-West Estimate with 90% Confidence Interval for Comparing Model with Factors to Model with Interest Differential

#### OA1. Out-Sample Forecasting

To complement the in-sample predictive regression results in the main text, this section presents some illustrative comparisons for the out-of-sample forecasting performance of the factors model relative to that of two simpler models: the random walk model and the interest differential model eq.(19) in text. We note that pseudo out-of-sample forecast comparisons involve a different set of considerations from model evaluations using in-sample regressions. Specifications with good in-sample fits commonly fail to deliver good out-of-sample performance. It is also well-known that imposing parameter restrictions, even wrong ones, can lead to smaller forecast errors ("principle of parsimony").<sup>1</sup> Furthermore, inherent instabilities, choices of window size and sub-sample periods all contribute to the fragility of any conclusive results in the forecasting literature, except that the simplest univariate specifications often deliver the lowest root-mean-squared forecast errors (RMSEs). See, for example, Clark and McCracken (2009) and Rossi and Inoue (2011) and references therein for a full discussion.

Against this backdrop, we present selected results as illustrative support for the aforementioned observations. We use rolling windows of various sizes (from 4 to 9 years), and construct out-of-sample forecasts for one- to four-quarters ahead for the three models.<sup>2</sup> In Table OA1, we report the RMSEs ratios of the factor model against the random walk (OA1a) and the interest differential model (OA1b). We also report the p-values based on the Clark-West (2006) predictability test, which accounts for the upward shift of RMSEs in the factor model due to

<sup>&</sup>lt;sup>1</sup> Parameter estimation error is one key reason, among others. If the marginal explanatory power associated with the additional parameters is low enough, in finite samples the extra estimation noise may raise the forecast error variance by more than the amount the extra information lowers it.

<sup>&</sup>lt;sup>2</sup> The first regression uses the first  $\tau + m$  observations, and makes a forecast for the exchange rate change from  $\tau + m$  to  $\tau + 2m$ , where  $\tau$  is window size. The second regression moves forward over time by one period and make another forecast, and so on. At the end of the rolling process, we calculate the root mean squared forecast error (RMSE) for our model, and compare it with RMSE produced by a drift-less random walk and by the interest differential model.

estimation noise.<sup>3</sup> We observe results consistent with the literature discussed above. First, our factor model, being the most general, tends to deliver larger RMSEs than the two more restricted models. However, using the Clark-West (2006) statistics, we are sometimes able to reject the null hypothesis of equal forecast performance in favor of our factor model, especially for Canada. The results are sensitive to window sizes, though overall, the models are mostly statistically indistinguishable. In Tables OA2a and OA2b, we report parallel results using more recent data samples, Jan 1991- May 2011, which show similar patterns.

#### OA2. Standard Errors for Overlapping Data: Simulated Critical Values

Following Parker and Julliard (2005), we set up a Monte Carlo experiment under the null hypothesis that the exchange rate follows a random walk. First, we sample with replacement from the 1-month exchange rate returns and create a series of size equal to our sample of "white noise" under the null. Using this re-sampled 1-month exchange rate change series, we generate the 1, 3, 6, 12, 18, and 24 month-ahead exchange rate changes as our LHS variables. We then regress these variables on the relative factors and keep the *t*-statistics. We repeat the three steps 2,000 times and use the critical values from the distributions of the *t*-statistics to do our inference. The setup for the excess currency return regression is similar except we use the actual yields to create the re-sampled excess returns. The rationale behind the experiment is that, if exchange rate is truly unpredictable as a random walk, the Monte Carlo results will tell us the probability that the predictability we find is spurious.

<sup>&</sup>lt;sup>3</sup> Under the null of equal predictability, the sample RMSE of the factor model is expected to be greater than those of the more restricted models. The Clark and West (2006) test statistic adjusts for this upward shift in the sample MSFE. Their simulations show that the inference made using asymptotically normal critical values gives properly-sized tests for rolling regressions.

Table OA1(a): RMSE Ratios & Clark-West p-Va	lues
Factors Model vs. Random Walk Model	

	Can	ada	Japan		U	K
		Windows	size = 48 mont	ths		
1	1.050*	0.075	1.093	0.441	1.017	0.224
3	1.097*	0.037	1.250	0.198	1.047*	0.049
6	1.094*	0.059	1.300	0.243	1.177*	0.058
12	1.024*	0.006	1.730*	0.052	1.398	0.278
		Windows	size = 60 mont	ths		
1	1.036*	0.068	1.091	0.245	1.025	0.481
3	1.084*	0.060	1.251+	0.042	1.065	0.266
6	1.071*	0.078	1.319	0.286	1.116	0.266
12	1.080*	0.020	1.659*	0.068	1.112	0.386
		Windows	size = 72 mont	ths		
1	1.017*	0.059	1.060	0.207	1.024	0.413
3	1.060*	0.085	1.155+	0.063	1.014	0.248
6	1.045*	0.069	1.182	0.345	1.049	0.323
12	1.077*	0.006	1.457*	0.047	1.013	0.248
		Windows	size = 84 mont	ths		
1	1.009*	0.047	1.049	0.244	1.026	0.362
3	1.047*	0.058	1.125	0.115	1.018	0.328
6	1.029*	0.030	1.143	0.462	1.040	0.378
12	1.047*	0.001	1.191*	0.073	1.052	0.460
		Windows	size = 96 mont	ths		
1	1.007*	0.036	1.032	0.342	1.029	0.116
3	1.016*	0.023	1.074	0.299	0.972*	0.058
6	0.995*	0.012	1.079	0.311	0.947*	0.043
12	1.063*	0.003	1.155	0.124	0.961	0.148
		Window s	ize = 108 mon	ths		
1	1.005*	0.047	1.025	0.441	1.010	0.494
3	1.014*	0.035	1.090	0.258	0.994*	0.084
6	1.030*	0.066	1.095	0.370	1.001	0.221
12	1.187	0.144	1.154	0.203	1.159	0.369

 $1985m8 - 2005m7^+$ 

*Note:* The 1st column under each country reports the RMSE-ratios for the factor model over the RW model, and the 2nd column are the p-values based on the Clark-West test. The sign \* means the factor model is preferred to the RW model at 10% level, and the sign + means the RW model is preferred at 10% level. Data for the UK starts at 1992m10.

	1	Canada	J	apan		UK
		Winde	ow size = 48 mo	onths		
3	1.062	0.361	1.114	0.163	1.001*	0.030
6	1.050	0.404	1.147	0.165	1.139	0.446
12	0.944*	0.011	1.020	0.401	1.164	0.136
		Winde	ow size = 60 mo	onths		
3	1.054	0.339	1.106+	0.050	1.034*	0.077
6	0.999	0.104	1.162	0.167	1.040	0.217
12	0.989*	0.040	1.234+	0.097	1.067	0.402
		Winde	ow size = 72 mo	onths		
3	1.033	0.282	1.083	0.285	0.995	0.127
6	0.997	0.126	1.175	0.158	1.109 +	0.058
12	1.151	0.251	1.412+	0.017	1.007	0.334
		Winde	ow size = 84 mo	onths		
3	1.045	0.460	1.089	0.118	0.988	0.147
6	1.019	0.244	1.137+	0.052	1.046	0.263
12	1.237	0.416	1.206+	0.019	1.066+	0.031
		Winde	ow size = 96 mo	onths		
3	1.036	0.489	1.056	0.275	0.982	0.110
6	1.037	0.455	1.076	0.239	1.016	0.498
12	1.218	0.486	1.136+	0.037	1.134+	0.027
		Windo	ow size = 108 m	onths		
3	1.036	0.422	1.074	0.174	0.992	0.202
6	1.046	0.320	1.069	0.343	1.067 +	0.065
12	1.184	0.115	1.077	0.180	$1.161 \pm$	0.004

# Table OA1(b): RMSE Ratios & Clark-West p-Values Factors Model vs. Interest Differential Model

 $1985m8 - 2005m7^+$ 

*Note:* The 1st column under each country reports the RMSE-ratios for the factor model over the interestdifferential model, and the 2nd column are the p-values based on the Clark-West test. The sign \* means the factor model is preferred to the interest differential model at 10% level, and the sign + means the interest differential model is preferred at 10% level. Data for the UK starts at 1992m10.

1//1111 2011113										
	Canada UK									
	Window size = 48 months									
1	1.132	0.221	0.906	0.215						
3	1.154*	0.043	0.915*	0.043						
6	1.198	0.151	1.026*	0.017						
12	1.315	0.415	1.240*	0.097						
	Window	v size = 60 mo	nths							
1	1.148	0.186	0.925	0.481						
3	1.198*	0.051	0.948*	0.071						
6	1.216*	0.081	1.088	0.192						
12	1.406	0.350	1.263	0.384						
Window size = 72 months										
1	1.166	0.212	0.922	0.381						
3	1.206*	0.090	0.978	0.284						
6	1.225	0.108	1.129	0.392						
12	1.465+	0.097	1.292	0.250						
	Window	v size = 84 mo	nths							
1	1.203	0.449	0.939	0.489						
3	1.236	0.251	0.999	0.474						
6	1.300	0.378	1.163	0.254						
12	1.405	0.153	1.383+	0.049						
	Window	v size = 96 mo	nths							
1	1.212	0.493	0.957	0.290						
3	1.235	0.325	1.030	0.321						
6	1.302	0.425	1.224+	0.089						
12	1.397+	0.098	1.433+	0.029						
	Window	size = 108 mc	onths							
1	1.237	0.418	0.979	0.234						
3	1.271	0.475	1.068	0.217						
6	1.388	0.043	1.259 +	0.044						
12	$1.558 \pm$	0.001	$1.535 \pm$	0.009						

# Table OA2(a): RMSE Ratios & Clark-West p-Values Factors Model vs. Random Walk Model

1991m1 – 2011m5

Note: The 1st column under each country reports the Clark-West statistics and the 2nd column are the p-values based on the Clark-West test. The sign \* means our model is preferred to the random walk model at 10% level, and the sign + means the random walk model is preferred at 10% level. Exchange rate data for Canada are available till May 2011, while yields data are available only till Feb 2011.

Γ	able	OA2	(b)	: RMSE	Ratios	&	Clark-V	West <sub>1</sub>	p-V	alu	ies
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#### Factors Model vs. Interest Differential Model

	Can	ada	U	К					
	Window	size = 48 mon	ths						
3	1.028*	0.085	1.055	0.123					
6	1.034*	0.042	1.049*	0.066					
12	1.111	0.278	1.129	0.197					
	Window	v size = 60 mo	nths						
3	1.055	0.242	1.034	0.101					
6	1.029*	0.067	1.073	0.345					
12	1.139	0.293	1.086	0.372					
	Window size = 72 months								
3	1.030	0.229	1.044	0.197					
6	1.000	0.143	1.055	0.423					
12	1.221+	0.001	1.043	0.231					
	Window	v size = 84 mo	nths						
3	1.036	0.485	1.032	0.229					
6	1.048	0.462	1.035	0.355					
12	1.143+	0.018	1.054	0.293					
	Window	v size = 96 mo	nths						
3	1.033	0.358	1.019	0.134					
6	1.044	0.389	1.037	0.303					
12	1.091	0.251	1.084	0.467					
	Window size = 108 months								
3	1.025	0.371	1.030	0.349					
6	1.053	0.255	1.057	0.305					
12	1.108	0.120	1.138+	0.021					

#### 1991m1 – 2011m5

Note: The 1st column under each country reports the Clark-West statistics and the 2nd column are the p-values based on the Clark-West test. The sign \* means our model is preferred to the interest differential model at 10% level, and the sign + means the random walk model is preferred at 10% level. Exchange rate data for Canada are available till May 2011, while yields data are available only till Feb 2011.

	m=1	m=3	m=6	m=12	m=18	m=24
L <sup>R</sup>	-3.472*	-2.724*	-1.650	-1.265	-1.057	-0.657
$t/\sqrt{m}$	-2.291	-1.770	-1.074	-0.822	-0.602	-0.336
S <sup>R</sup>	-0.732	-0.628	-0.528	-0.255	-0.127	-0.006
$t/\sqrt{m}$	-1.149	-0.972	-0.812	-0.385	-0.164	-0.007
C <sup>R</sup>	-0.931*	-0.865*	-0.633	-0.512	-0.347	-0.235
$t/\sqrt{m}$	-1.977	-1.813	-1.305	-1.022	-0.547	-0.334
Lagged LHS	0.026	-0.015	0.016	0.031	0.014	0.002
$t/\sqrt{m}$	0.399	-0.407	0.591	1.638	0.760	0.081
N. obs.	238	234	228	216	204	192

### Table OA3: Canadian regressions with the inclusion of Lagged Dependent Variable

a) Exchange Rate  $\frac{1200(s_{t+m}-s_t)}{m} = \beta_{m,0} + \beta_{m,1}L_t^R + \beta_{m,2}S_t^R + \beta_{m,3}C_t^R + \beta_{m,4}\frac{1200(s_t-s_{t-m})}{m} + u_{t+m}$ 

b) Excess Return  $i_t^{m*} - i_t^m + \frac{1200(s_{t+m}-s_t)}{m} = \gamma_{m,0} + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + i_{t-m}^m - i_{t-m}^m + \frac{1200(s_t-s_{t-m})}{m} + v_{t+m}$ 

	m=3	m=6	m=12	m=18	m=24
L <sup>R</sup>	-2.247	-2.535	-2.045	-2.109	-1.845
$t/\sqrt{m}$	-1.202	-1.631	-1.261	-1.151	-0.929
S <sup>R</sup>	-1.609*	-1.145*	-0.868	-0.775	-0.631
$t/\sqrt{m}$	-1.913	-1.669	-1.231	-0.964	-0.712
C <sup>R</sup>	-0.467	-0.736	-0.701	-0.637	-0.509
$t/\sqrt{m}$	-0.739	-1.505	-1.354	-0.985	-0.736
Lagged LHS	-0.096	0.126	0.291	0.085	-0.134
$t/\sqrt{m}$	-0.716	0.781	1.323	0.275	-0.314
N. obs.	132	210	216	204	192

*Note:* The sample is Aug 1985-Jul 2005. Exchange rate *s* is log(USD/CAD). The row  $t/\sqrt{m}$  reports the rescaled *t*-statistics for the estimates (see text for details). Estimates for the constant term are omitted, and \* indicates significance level of 10% or below.

a) Exchan	ge Rate $\frac{1200(s_{t+n})}{m}$	$\frac{n-s_t}{s_{m,0}} = \beta_{m,0} + \beta_r$	$_{n,1}L_t^R + \beta_{m,2}S_t^R +$	$\beta_{m,3}C_t^R + \beta_{m,4}\frac{1}{2}$	$\frac{200(s_t-s_{t-m})}{m}+u_t.$	+m
	m=1	m=3	m=6	m=12	m=18	m=24
L <sup>R</sup>	-3.636	-2.357	-2.362	-2.770	-2.359	-1.604
$t/\sqrt{m}$	-1.206	-0.751	-0.812	-1.118	-0.998	-0.839
S <sup>R</sup>	-3.572*	-3.586*	-3.712*	-2.642*	-2.747*	-2.395*
$t/\sqrt{m}$	-2.298	-2.194	-2.424	-2.053	-2.134	-2.341
C <sup>R</sup>	0.299	0.696	0.402	-0.284	-0.098	-0.267
$t/\sqrt{m}$	0.265	0.601	0.376	-0.314	-0.102	-0.342
Lagged LHS	-0.010	-0.004	-0.032	0.000	-0.014	-0.018*
$t/\sqrt{m}$	-0.150	-0.111	-1.311	-0.009	-1.177	-2.233
N. obs.	238	234	228	216	204	192

#### Table OA4: Japanese regressions with the inclusion of Lagged Dependent Variable

b) Excess Return  $i_t^{m*} - i_t^m + \frac{1200(s_{t+m} - s_t)}{m} = \gamma_{m,0} + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + i_{t-m}^{m*} - i_{t-m}^m + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + i_{t-m}^{m*} - i_{t-m}^m + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + i_{t-m}^{m*} - i_{t-m}^m + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + i_{t-m}^m + \gamma_{m,1}L_t^R + \gamma_{m,3}C_t^R + \gamma_$  $\frac{\frac{1200(s_t-s_{t-m})}{m} + v_{t+m}}{m}$ 

	m=3	m=6	m=12	m=18	m=24
L <sup>R</sup>	-6.202	-3.666	-3.779	-3.602	-3.017
$t/\sqrt{m}$	-1.330	-1.240	-1.514	-1.522	-1.555
S <sup>R</sup>	-4.824*	-4.957*	-3.372*	-3.417*	-2.924*
$t/\sqrt{m}$	-2.197	-3.150	-2.571	-2.624	-2.827
C <sup>R</sup>	0.747	0.289	-0.515	-0.361	-0.554
$t/\sqrt{m}$	0.522	0.269	-0.570	-0.376	-0.702
Lagged LHS	-0.067	-0.164	-0.005	-0.230	-0.373*
$t/\sqrt{m}$	-0.443	-1.143	-0.027	-1.203	-2.187
N. obs.	103	216	216	204	192

*Note:* The sample is Aug 1985-Jul 2005. Exchange rate s is  $\log(\text{USD/JPN})$ . The row  $t/\sqrt{m}$  reports the rescaled *t*-statistics for the estimates (see text for details). Estimates for the constant term are omitted, and \* indicates significance level of 10% or below.

	m				m	
	m=1	m=3	m=6	m=12	m=18	m=24
LR	-3.150*	-4.053*	-3.139*	-2.615*	-1.495	-1.173
$t/\sqrt{m}$	-1.641	-2.345	-2.005	-1.727	-0.940	-0.766
SR	-1.796*	-2.312*	-1.969*	-1.433*	-0.763	-0.466
$t/\sqrt{m}$	-1.797	-2.527	-2.296	-1.720	-0.934	-0.597
CR	-0.755	-1.131*	-1.007*	-0.820*	-0.429	-0.253
$t/\sqrt{m}$	-1.383	-2.301	-2.230	-1.862	-0.945	-0.584
Lagged LHS	0.002	-0.014	-0.007	-0.001	-0.009	-0.008
$t/\sqrt{m}$	0.029	-0.432	-0.310	-0.090	-0.657	-0.681
N. obs.	214	210	204	192	180	168

### Table OA5: UK regressions with the inclusion of Lagged Dependent Variable

a) Exchange Rate  $\frac{1200(s_{t+m}-s_t)}{m} = \beta_{m,0} + \beta_{m,1}L_t^R + \beta_{m,2}S_t^R + \beta_{m,3}C_t^R + \beta_{m,4}\frac{1200(s_t-s_{t-m})}{m} + u_{t+m}$ 

b) Excess Return  $i_t^{m*} - i_t^m + \frac{1200(s_{t+m}-s_t)}{m} = \gamma_{m,0} + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + i_{t-m}^m - i_{t-m}^m + \frac{1200(s_t-s_{t-m})}{m} + v_{t+m}$ 

	m=3	m=6	m=12	m=18	m=24
L <sup>R</sup>	-2.929	-3.829*	-3.733*	-2.632*	-2.255
$t/\sqrt{m}$	-0.907	-1.720	-2.340	-1.649	-1.487
S <sup>R</sup>	-5.065*	-2.890*	-2.254*	-1.382*	-1.009
$t/\sqrt{m}$	-2.095	-1.803	-2.355	-1.680	-1.295
C <sup>R</sup>	0.803	-0.781	-1.068*	-0.727	-0.545
$t/\sqrt{m}$	0.544	-1.005	-2.192	-1.592	-1.258
Lagged LHS	-0.257	-0.087	0.002	-0.177	-0.256
$t/\sqrt{m}$	-1.297	-0.481	0.011	-0.715	-0.858
N. obs.	57	122	175	180	168

*Note*: The sample is Aug 1985-Jul 2005. Exchange rate *s* is log(USD/GBP). The row  $t/\sqrt{m}$  reports the rescaled *t*-statistics for the estimates (see text for details). Estimates for the constant term are omitted, and \* indicates significance level of 10% or below.

## Table OA6: Correlation between SPF Forecasts and the US Factors

Horizon <i>m</i>	$\beta_{1m}$	$\beta_{2m}$	$\beta_{3m}$	Adj. R-Sq.
3	-0.244* (0.071)	-0.061 (0.076)	0.063 (0.055)	0.102
6	-0.187* (0.049)	-0.070 (0.053)	-0.035 (0.038)	0.214
9	-0.136* (0.034)	-0.066* (0.037)	-0.046 (0.026)	0.306
12	-0.165* (0.034)	-0.145* (0.036)	-0.091* (0.026)	0.563

<u>Real GDP Growth</u>:  $E_t \Delta y_{t+m} = \beta_{0m} + \beta_{1m}L_t + \beta_{2m}S_t + \beta_{3m}C_t + u_{mt}$ 

<u>CPI Inflation</u>:  $E_t \pi_{t+m} = \beta_{0m} + \beta_{1m}L_t + \beta_{2m}S_t + \beta_{3m}C_t + u_{kt}$ 

Horizon <b>m</b>	$\beta_{1m}$	$\beta_{2m}$	$\beta_{3m}$	Adj. R-Sq.
2	0.458*	0.154*	0.050	0.608
5	(0.041)	(0.044)	(0.032)	0.098
6	0.449*	0.118*	0.072*	0.756
0	(0.035)	(0.038)	(0.028)	0.750
0	0.448*	0.094*	0.080*	0.768
	(0.034)	(0.037)	(0.027)	0.700
12	0.461*	0.086*	0.081*	0 779
12	(0.034)	(0.036)	(0.026)	0.119

<u>Anxiety Index</u>:  $A_{t+m} = \beta_{0m} + \beta_{1m}L_t + \beta_{2m}S_t + \beta_{3m}C_t + u_{mt}$ 

Horizon <b>m</b>	$\beta_{1m}$	$\beta_{2m}$	$\beta_{3m}$	Adj. R-Sq.
3	3.399*	1.676	-2.150*	0.148
5	(0.914)	(0.981)	(0.711)	0.140
6	2.585*	2.235*	-1.117*	0 246
0	(0.541)	(0.581)	(0.421)	0.210
9	2.004*	2.212*	-0.175	0.530
-	(0.312)	(0.335)	(0.243)	0.000
12	1.731*	1.754*	0.555*	0.585
12	(0.304)	(0.326)	(0.236)	

Note: The quarterly sample is drawn from the last month of each quarter from the sample Aug 1985-Jul 2005. See the website <u>http://www.philadelphiafed.org/research-and-data/real-time-center/survey-of-professional-forecasters</u> for details about the survey data.

# Table OA7: Predicting the Canadian-US Exchange Rate and Excess ReturnsMost Recent Data and Sub-Sample Results

	(a) Exchange	Rate $\frac{1200(s_{t+m}-m)}{m}$	$\frac{-s_t}{2} = \beta_{m,0} + \beta_n$	$_{n,1}L_t^R + \beta_{m,2}S_t^R$	$+\beta_{m,3}C_t^R+u_t$	:+ <i>m</i>
	m=1	m=3	m=6	m=12	m=18	m=24
$L^{R}$	-2.436	-2.008	-0.790	-0.558	0.028	0.373
$t/\sqrt{m}$	-1.202	-0.802	-0.317	-0.224	0.011	0.133
<b>S</b> <sup>R</sup>	-0.366	-0.239	-0.217	0.231	0.344	0.428
$t/\sqrt{m}$	-0.338	-0.213	-0.194	-0.208	0.287	0.342
C <sup>R</sup>	-2.425*	-1.812	-1.094	-1.510	-1.563	-1.603
$t/\sqrt{m}$	-2.112	-1.527	-0.925	-1.279	-1.230	-1.209
N. obs.	168	168	168	168	168	168

# (Jan 1991 – Dec 2004)

(Jan 2005 – May 2011)

	m=1	m=3	m=6	m=12
$L^{R}$	-13.341	-4.174	-9.278	9.369
$t/\sqrt{m}$	-0.726	-0.254	-0.551	0.459
S <sup>R</sup>	10.954	15587*	14.071*	7.942
$t/\sqrt{m}$	1.231	1.959	1.716	0.956
C <sup>R</sup>	-3.068	-4.403	-4.360	0.672
$t/\sqrt{m}$	-0.841	-1.349	-1.295	0.166
N. obs.	74	74	71	65

	m=3	m=6	m=12	m=24
$L^{R}$	-2.999	-1.774	-1.591	-0.650
$t/\sqrt{m}$	-1.196	-0.712	-0.719	-0.233
S <sup>R</sup>	-1.159	-1.043	-0.483	-0.099
$t/\sqrt{m}$	-1.034	-0.936	-1.504	-0.079
C <sup>R</sup>	-1.885	-1.232	-1.763	-1.900
$t/\sqrt{m}$	-1.587	-1.043	-1.494	-1.438
N. obs.	168	168	168	168

#### (Jan 1991 – Dec 2004)

(c) Excess Return  $i_t^{m*} - i_t^m + \frac{1200(s_{t+m} - s_t)}{m} = \gamma_{m,0} + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + v_{t+m}$ 

(Jan 2005 – May 2011)

(d) Excess Return  $i_t^{m*} - i_t^m + \frac{1200(s_{t+m} - s_t)}{m} = \gamma_{m,0} + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + v_{t+m}$ 

	m=3	m=6	m=12
$L^{R}$	-5.165	-10.290	8.388
$t/\sqrt{m}$	-0.314	-0.612	0.412
S <sup>R</sup>	14.638*	13.253	7.280
$t/\sqrt{m}$	1.839	1.618	0.879
O <sup>R</sup>			
Ск	-4.482	-4.504	0.446
$t/\sqrt{m}$	-1.372	-1.339	0.111
N. obs.	74	71	65

*Note*: Exchange rate *s* is log(USD/CAD). The row  $t/\sqrt{m}$  reports the re-scaled *t*-statistics for the estimates (see text for details). Estimates for the constant term are omitted, and \* indicates significance level of 10% or below. Yields data (of 3-month, 6-month, and 1 to 10-year maturities) for the three countries are respectively from the Federal Reserve Board, Bank of England, and Bank of Canada. Exchange rates are from the FRED database at the St Louis Fed. Exchange rate data for Canada are available till May 2011, while yields data are available only till Feb 2011.

# Table OA8: Predicting the UK-US Exchange Rate and Excess Returns

# Most Recent Data and Sub-Sample Results

	(a) Exchang	ge Rate $\frac{1200(s_{t+1})}{m}$	$\frac{(m-s_t)}{m} = \beta_{m,0} + \beta_{m,0}$	$\beta_{m,1}L_t^R + \beta_{m,2}S_t^R + \beta_{m$	$+\beta_{m,3}C_t^R+u_{t+m}$	
	m=1	m=3	m=6	m=12	m=18	m=24
L <sup>R</sup>	-2.545	-2.983	-3.272	-3.576*	-3.042	-2.017
$t/\sqrt{m}$	-0.825	-1.143	-1.382	-1.853	-1.312	-0.836
SR	-0.720	-1.462	-2.013	-1.559	-0.993	-0.448
$t/\sqrt{m}$	-0.357	-0.857	-1.300	-1.046	-0.655	-0.284
C <sup>R</sup>	-1.782	-1.139	-0.926	-1.235	-1.316	-1.227
$t/\sqrt{m}$	-1.361	-1.028	-0.921	-1.277	-1.337	-1.200
N. obs.	147	147	147	147	147	147

# (Oct 1992 – Dec 2004)

### (Jan 2005 – May 2011)

(~)	m	F 11,0 · F	<i>π</i> ,1-ι · <i>Ρπ</i> ,2-ι ·	
	m=1	m=3	m=6	m=12
L <sup>R</sup>	8.075	1.396	2.666	5.962
$t/\sqrt{m}$	0.608	0.106	0.225	0.502
SR	10.525*	12.894*	12.647*	7.818*
$t/\sqrt{m}$	3.429	3.789	4.107	2.715
CR	3.058	1.699	1.958	2.989
$t/\sqrt{m}$	1.010	0.566	0.720	1.253
N. obs.	76	74	71	65

(b)	Exchange Rate	$\frac{1200(s_{t+m}-s_t)}{m}$	$=\beta_{m,0}$	$+ \beta_{m,1}L_t^R$	+β	$m_{m,2}S_t^R$	$+ \beta_{m,3}$	$C_t^R$ -	$\vdash u_{t+r}$
	0	m	1 110,0	1 110,1 0		11 <i>0,</i> 2 C	1 110,0	c	

	m=3	m=6	m=12	m=24
L <sup>R</sup>	-3.948	-4.258*	-4.626*	-3.031
$t/\sqrt{m}$	-1.514	-1.792	-2.020	-1.255
S <sup>R</sup>	-2.361	-2.988*	-2.311	-0.985
$t/\sqrt{m}$	-1.384	-1.900	-1.543	-0.624
C <sup>R</sup>	-1.218	-1.109	-1.478	-1.514
$t/\sqrt{m}$	-1.100	-1.100	-1.521	-1.477
l. obs.	147	147	147	147

#### (Oct 1992 – Dec 2004)

#### (Jan 2005 – May 2011)

(d) Excess Return  $i_t^{m*} - i_t^m + \frac{1200(s_{t+m} - s_t)}{m} = \gamma_{m,0} + \gamma_{m,1}L_t^R + \gamma_{m,2}S_t^R + \gamma_{m,3}C_t^R + v_{t+m}$ 

	m=3	m=6	m=12
L <sup>R</sup>	0.399	1.677	4.977
$t/\sqrt{m}$	0.030	0.141	0.421
S <sup>R</sup>	11.973*	11.807*	7.110*
$t/\sqrt{m}$	3.523	3.834	2.483
C <sup>R</sup>	1.610	1.812	2.783
$t/\sqrt{m}$	0.380	0.667	1.173
N. obs.	74	71	65

*Note:* Exchange rate s is log(USD/GBP). The row  $t/\sqrt{m}$  reports the re-scaled t-statistics for the estimates (see text for details). Estimates for the constant term are omitted, and \* indicates significance level of 10% or below. Yields data (of 3-month, 6-month, and 1 to 10-year maturities) for the three countries are respectively from the Federal Reserve Board, Bank of England, and Bank of Canada. Exchange rates are from the FRED database at the St Louis Fed.

Figure OA1: Juxtaposing Rejections of the Interest Differential Model and the Factors

A) Canada



Note: The sample period is Aug 1985 – Jul 2005. We are simply plotting the factors against the F-statistic of the rejection of the interest differential model.

B) Japan



Note: The sample period is Aug 1985 – Jul 2005. We are simply plotting the factors against the F-statistic of the rejection of the interest differential model.



Note: The sample period is Aug 1985 – Jul 2005. We are simply plotting the factors against the F-statistic of the rejection of the interest differential model.

#### Figure OA2(a): Rolling Test of the Interest Differential Restrictions (Canada)

(Jan 1991– May 2011)



**3-Month Horizon** 

Note: The solid line plots the F-statistic for the null hypothesis that the restriction imposed by the UIP on the N-S factors is correct. The red dotted line is the Monte Carlo 10% critical value, accounting for small sample bias and persistence of the data. Exchange rate data for Canada are available till May 2011, while yields data are available only till Feb 2011. For more details see the Appendix.

# Figure OA2(b): Rolling Test of the Interest Differential Restrictions (United Kingdom)

(Jan 1991 - May 2011)

**3-Month Horizon** 



Note: The solid line plots the F-statistic for the null hypothesis that the restriction imposed by the UIP on the N-S factors is correct. The red dotted line is the Monte Carlo 10% critical value, accounting for small sample bias and persistence of the data. For more details see the Appendix.

Figure OA3(a): Recursive Adjusted R-squares with a 5-Year Rolling Window (Canada) 3-Month Horizon



Note: The sample period is Jan 1991 – May 2011. See the paper for the calculation of the R-square and the correction of bias. Exchange rate data for Canada are available till May 2011, while yields data are available only till Feb 2011. Blue solid line (factor model); red dotted line (interest differential model).

Figure OA3(b): Recursive Adjusted R-squares with a 5-Year Rolling Window (United Kingdom)



Note: The sample period is Jan 1991 – May 2011. See the paper for the calculation of the R-square and the correction of bias. Blue solid line (factor model); Red dotted line (interest differential model)

Figure OA4(a): Rolling Clark-West Estimate with 90% Confidence Interval for Comparing Model with Factors to Model with Interest Differential (Canada)



1-month

3-month





Note: The sample period is Aug 1985-Jul 2005. We are simply doing the Clark-West recursively using a 5-year window, based on recursive regressions with factors or interest differential using, again, a 5-year window. A test statistic (solid line) significantly above zero means the model with factors is preferred to the model interest differential, and vice versa.

Figure OA4(b): Rolling Clark-West Estimate with 90% Confidence Interval for Comparing Model with Factors to Model with Interest Differential (Japan)



1-month





Note: The sample period is Aug 1985-Jul 2005. We are simply doing the Clark-West recursively using a 5-year window, based on recursive regressions with factors or interest differential using, again, a 5-year window. A test statistic (solid line) significantly above zero means the model with factors is preferred to the model interest differential, and vice versa.

Figure OA4(c): Rolling Clark-West Estimate with 90% Confidence Interval for Comparing Model with Factors to Model with Interest Differential (United Kingdom)



1-month





12-month



Note: The sample period is Aug 1985-Jul 2005. We are simply doing the Clark-West recursively using a 5-year window, based on recursive regressions with factors or interest differential using, again, a 5-year window. A test statistic (solid line) significantly above zero means the model with factors is preferred to the model interest differential, and vice versa.