## A short note on the mixture of experts

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The mixture of expert is a popular approach in statistics and machine learning. It is similar but different from the usual mixture model (and the mixture of regression). Here we give a gentle introduction about this idea. For readers who are interested in more details, I would recommend the following book chapter:

Gormley, I. C., & Frühwirth-Schnatter, S. (2019). Mixture of experts models. Handbook of Mixture Analysis, 271-307.

Let  $Y \in \mathbb{R}$  be a continuous random variable that is our primary response variable and  $Z \in \{1, 2, \dots, K\}$  be a discrete/categorical variable and  $X \in \mathbb{R}^d$  be a multivariate covariate. We only observe (X, Y) and Z is unobserved; here Z is often refers to the latent class label or the label of an *expert*. In mixture models or mixture of experts, we often use a parametric form of the conditional densities. Depending on the relation among X, Y, Z, there are 4 popular *mixture-type* models:

• Mixture model. In the usual mixture model, there is no covariate X so we only observe Y. The mixture model can be written as a graphical model with a direct arrow  $Z \rightarrow Y$ . Suppose we observe both (Y,Z), then

$$p(y,z) = p(y|z)p(z) = p_z(y)\pi_z \Rightarrow p(y) = \sum_k p_k(y)\pi_k,$$

where  $p_k(y)$  is the conditional distribution of *Y* given Z = k and  $\pi_k = P(Z = k)$  is the proportion of the *k*-th component. Let  $\theta_k$  be the parameter of  $p_k(y)$ , then the marginal distribution is

$$p(y; \mathbf{\theta}) = \sum_{k} p(y; \mathbf{\theta}_{k}) \pi_{k}$$

which is the usual mixture model. The Gaussian mixture model is that each  $p(y;\theta_k)$  is a Gaussian, i.e.,  $p(y;\theta_k) = p(y;\mu_k,\sigma_k^2)$ , where  $\mu_k$  and  $\sigma_k^2$  is the mean and variance of *k*-th component.

• Mixture of expert. In the mixture of expert, the model can be expressed as a graphical model with two arrows  $X \to Z$  and  $Z \to Y$ . Note that Z is unobserved—we only observe X,Y. In this case,

$$p(x,y,z) = p(y|z)p(z|x)p(x) = p_z(y)\pi_z(x)p(x) \Rightarrow p(y,z|x) = p_z(y)\pi_z(x)$$
$$\Rightarrow p(y|x) = \sum_k p_k(y)\pi_k(x).$$

Namely, in the mixture of expert, the density of *Y* at each component remains the same across different *X*. What changes with respect to *X* is the proportion  $\pi_k(x)$ .

In this case, we need parameters for both  $p_k(y)$  and  $\pi_k(x)$ , which leads to

$$p(y|x; \theta, \eta) = \sum_{k} p(y; \theta_k) \pi_k(x; \eta).$$

A popular model is place a Gaussian model over  $p(y; \theta_k)$  and a logistic model of  $\pi_k(x; \eta)$ , i.e.,

$$\pi_k(x;\eta) = \frac{\exp(\eta_{0,k} + \eta_{1,k}^T x)}{\sum_m \exp(\eta_{0,m} + \eta_{1,m}^T x)}.$$

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Mixture of regression. The mixture of regression (a.k.a. regression mixture) looks very similar to the mixture of expert from a graphical perspective. The mixture of regression has two arrows: *X* → *Y* and *Z* → *Y*. This, the difference compared to the mixture of expert is that the arrow *X* → *Z* becomes *X* → *Y*. In this case,

$$p(x,y,z) = p(y|x,z)p(z)p(x) = p_z(y|x)\pi_z p(x) \Rightarrow p(y,z|x) = p_z(y|x)\pi_z$$
$$\Rightarrow p(y|x) = \sum_k p_k(y|x)\pi_k.$$

In particular, the conditional mean (regression function) becomes

$$m(x) = \mathbb{E}(Y|X=x) = \int \sum_{k} p_k(y|x) \pi_k dy = \sum_{k} \pi_k \cdot m_k(x),$$

where  $m_k(x) = \mathbb{E}(Y|Z = k, X = x)$  is the regression function of the *k*-th component. So the regression function is written as a mixture of several regression function. Note that the proportion  $\pi_k$  is independent of *X*.

• Mixture of expert regression. The mixture of expert and the mixture of regression can be combined into the mixture of expert regression. It corresponds to the graph with three arrows:  $X \rightarrow Y, X \rightarrow Z$ , and  $Z \rightarrow Y$ . In this case,

$$p(x,y,z) = p(y|x,z)p(z|x)p(x) = p_z(y|x)\pi_z(x)p(x) \Rightarrow p(y,z|x) = p_z(y|x)\pi_z(x)$$
$$\Rightarrow p(y|x) = \sum_k p_k(y|x)\pi_k(x).$$

The conditional mean (regression function) is

$$m(x) = \mathbb{E}(Y|X=x) = \int \sum_{k} p_k(y|x) \pi_k(x) dy = \sum_{k} \pi_k(x) \cdot m_k(x).$$

So it is the mixture of regression with the proportion  $\pi_k(x)$  being allowed to change with respect to x.

## **1** Mixture of expert

In what follows, we will focus on the mixture of expert. Recall that in the mixture of expert,

$$p(y|x;\theta,\eta) = \sum_{k} p(y;\theta_k) \pi_k(x;\eta)$$

and what we observe is

$$(X_1,Y_1),\cdots,(X_n,Y_n).$$

The goal is to estimate  $\theta$  and  $\eta$  from the observed data.

A simple way to estimate these parameters is based on the maximum likelihood (ML) approach. For a single observation  $X_i, Y_i$ , the likelihood function is

$$L(\theta,\eta|X_i,Y_i) = \sum_k p(Y_i;\theta_k)\pi_k(X_i;\eta)$$

and the log-likelihood is

$$\ell(\boldsymbol{\theta},\boldsymbol{\eta}|X_i,Y_i) = \log\left(\sum_k p(Y_i;\boldsymbol{\theta}_k)\pi_k(X_i;\boldsymbol{\eta})\right)$$

The MLE (maximum likelihood estimator) is

$$\widehat{\Theta}, \widehat{\eta} = \operatorname{argmax}_{\theta, \eta} \frac{1}{n} \sum_{i=1}^{n} \ell(\theta, \eta | X_i, Y_i)$$

## 2 EM algorithm

Although the MLE is well-defined, it is often hard to compute due to the fact that it does not have a closed-form in general. So we often need to use the EM-algorithm to numerically find the MLE. An introduction on the procedure of the EM is given in: http://faculty.washington.edu/yenchic/19A\_stat535/ Lec13\_EM\_SGD.pdf. Starting with an initial guess ( $\theta^{(0)}, \eta^{(0)}$ ), the EM algorithm creates a sequence of parameters

$$(\theta^{(0)}, \eta^{(0)}), (\theta^{(1)}, \eta^{(1)}), \cdots, (\theta^{(t)}, \eta^{(t)}), (\theta^{(t+1)}, \eta^{(t+1)}), \cdots$$

such that the likelihood function

$$\sum_{i=1}^{n} \ell(\theta^{(t+1)}, \eta^{(t+1)} | X_i, Y_i) \ge \sum_{i=1}^{n} \ell(\theta^{(t)}, \eta^{(t)} | X_i, Y_i).$$

A key quantity in the EM algorithm is the complete-data likelihood–the likelihood function when the latent variable *Z* is also observed:

$$L_{\mathsf{comp}}(\boldsymbol{\theta},\boldsymbol{\eta}|\boldsymbol{X},\boldsymbol{Y},\boldsymbol{Z}) = \prod_{k} [p(\boldsymbol{Y};\boldsymbol{\theta}_{k})\boldsymbol{\pi}_{k}(\boldsymbol{X};\boldsymbol{\eta})]^{I(\boldsymbol{Z}=k)}$$

and  $\ell_{comp}(\theta, \eta | X, Y, Z) = \log L_{comp}(\theta, \eta | X, Y, Z)$ . Given a complete-data likelihood and a previous parameter  $(\theta^{(t)}, \eta^{(t)})$ , we define the *Q* function in the EM algorithm:

$$Q(\theta, \eta; \theta^{(t)}, \eta^{(t)} | X, Y) = \mathbb{E}(\ell_{\mathsf{comp}}(\theta, \eta | X, Y, Z) | X, Y; \theta^{(t)}, \eta^{(t)})$$

$$= \mathbb{E}\left(\sum_{k} I(Z = k) \log[p(Y; \theta_{k})\pi_{k}(X; \eta)] \middle| X, Y; \theta^{(t)}, \eta^{(t)}\right)$$

$$= \sum_{k} \omega_{k}(X, Y; \theta^{(t)}, \eta^{(t)}) \log[p(Y; \theta_{k})\pi_{k}(X; \eta)],$$

$$\omega_{k}(X, Y; \theta^{(t)}, \eta^{(t)}) = P(Z = k | X, Y; \theta^{(t)}, \eta^{(t)})$$

$$= \frac{p(Y; \theta^{(t)}_{k})\pi_{k}(X; \eta^{(t)})}{\sum_{m} p(Y; \theta^{(t)}_{m})\pi_{m}(X; \eta^{(t)})}.$$

With this, we can write down the E-step and the M-step in the algorithm:

• E-step. Compute

$$\omega_k(X_i, Y_i; \boldsymbol{\theta}^{(t)}, \boldsymbol{\eta}^{(t)}) = \frac{p(Y_i; \boldsymbol{\theta}_k^{(t)}) \pi_k(X_i; \boldsymbol{\eta}^{(t)})}{\sum_m p(Y_i; \boldsymbol{\theta}_m^{(t)}) \pi_m(X_i; \boldsymbol{\eta}^{(t)})}$$

and

$$Q_n(\theta, \eta; \theta^{(t)}, \eta^{(t)}) = \frac{1}{n} \sum_{i=1}^n Q(\theta, \eta; \theta^{(t)}, \eta^{(t)} | X_i, Y_i)$$
$$Q(\theta, \eta; \theta^{(t)}, \eta^{(t)} | X_i, Y_i) = \sum_k \omega_k(X_i, Y_i; \theta^{(t)}, \eta^{(t)}) \left(\log p(Y_i; \theta_k) + \log \pi_k(X_i; \eta)\right).$$

• **M-step.** We update  $\theta$ ,  $\eta$  using

$$\mathbf{\theta}^{(t+1)}, \mathbf{\eta}^{(t+1)} = \operatorname{argmax}_{\mathbf{\theta},\mathbf{\eta}} Q_n(\mathbf{\theta},\mathbf{\eta};\mathbf{\theta}^{(t)},\mathbf{\eta}^{(t)}).$$

A nice property of this maximization is that  $\theta$  and  $\eta$  can be maximized separately and each componentwise parameter  $\theta_k$  can also be maximized individually:

$$\theta_k^{(t+1)} = \operatorname{argmax}_{\theta_k} Q_{k,n}(\theta_k; \theta^{(t)}, \eta^{(t)})$$
$$Q_{k,n}(\theta_k; \theta^{(t)}, \eta^{(t)}) = \frac{1}{n} \sum_{i=1}^n \omega_k(X_i, Y_i; \theta^{(t)}, \eta^{(t)}) \log p(Y_i; \theta_k)$$

and

$$\eta^{(t+1)} = \operatorname{argmax}_{\eta} Q_n(\eta; \theta^{(t)}, \eta^{(t)})$$
$$Q_n(\eta; \theta^{(t)}, \eta^{(t)}) = \frac{1}{n} \sum_{i=1}^n \sum_k \omega_k(X_i, Y_i; \theta^{(t)}, \eta^{(t)}) \log \pi_k(Y_i; \eta).$$

Note that the EM algorithm suffers from the same problem of being stuck in a local maximum. Thus, multiple random initializations are often needed to increase the chance of getting the MLE.

## **3** Remarks

Here are some remarks about the mixture of expert method.

• Common choice of the parametric model. A popular choice is  $p(y; \theta_k) = \phi(y; \mu_k, \sigma_k^2)$ , where  $\phi(y; \mu, \sigma^2)$  is the normal density with mean  $\mu$  and variance  $\sigma^2$  and  $\pi_k(x; \eta) = \frac{\exp(\tilde{x}^T \eta_k)}{\sum_m \exp(\tilde{x}^T \eta_m)}$ , where  $\tilde{x} = (1, x) \in \mathbb{R}^{d+1}$  is the augmented covariate with the interception term. Note that although here we assume a univariate response  $Y \in \mathbb{R}$ , the whole model can be easily generalized to multivariate response  $Y \in \mathbb{R}^p$ .

• Identifiability. Model identifiability is often a problem in the mixture model and so is the mixture of experts. The label switching would occur if we do not place constraint over parameters  $\theta_1, \dots, \theta_k$ . Consider a simple mixture of experts model with two experts:

 $p(y|x;\theta,\eta) = p(y;\theta_1)\pi_1(x;\eta) + p(y;\theta_2)\pi_2(x;\eta) = p(y;\theta_1')\pi_1(x;\eta) + p(y;\theta_2')\pi_2(x;\eta)$ 

if we choose  $\theta'_1 = \theta_2$  and  $\theta'_2 = \theta_1$ . Thus,  $(\theta', \eta) \neq (\theta, \eta)$  but the probability model is the same. This also implies that the MLE will not be unique (since we can permute the parameters). A common approach to resolve this is to enforce some ordering among parameters.

- Asymptotic theory. The asymptotic theory follows from the regular MLE theory and we can construct confidence intervals using either a sandwich estimator of the underlying variance or a bootstrap approach.
- Choice of number of experts *K*. In general, the choice of number of experts is similar to the problem of choosing the number of mixture components in a mixture model. Common approaches such as AIC, BIC are often used. Note that if the problem is written as a prediction problem (given *X*, we use mixture of expert to predict *Y*), we may also use the cross-validation approach.
- **Bayesian approach and variational inference.** It is possible to use a Bayesian approach in the mixture of experts. The following paper discussed this idea along with variational inference:

Bishop, C. M., & Svenskn, M. (2002, August). Bayesian hierarchical mixtures of experts. In Proceedings of the Nineteenth conference on Uncertainty in Artificial Intelligence (pp. 57-64).