

STATISTICAL INFERENCE USING GEOMETRIC FEATURES

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Ryan Tibshirani

Collaborators

Astronomy



Shirley Ho



Peter Freeman



Rachel Mandelbaum

Biology



Ran Kafri
(University of Toronto)

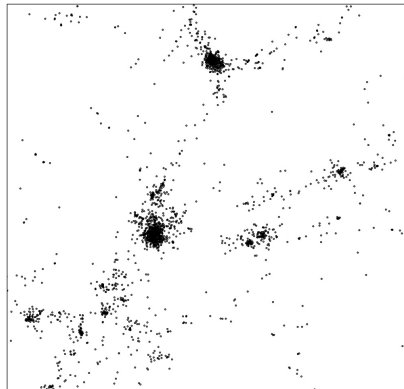


Miriam Ginzberg
(Harvard)



Shixuan Liu
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What are Geometric Features?

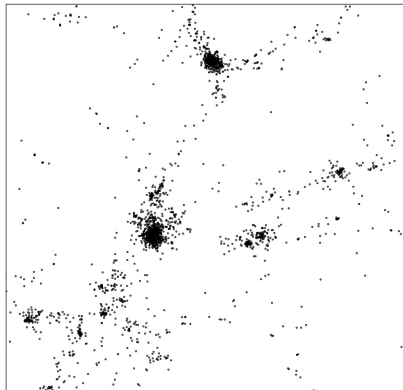


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$$X_1, \dots, X_n \sim p,$$

p is a probability density function.



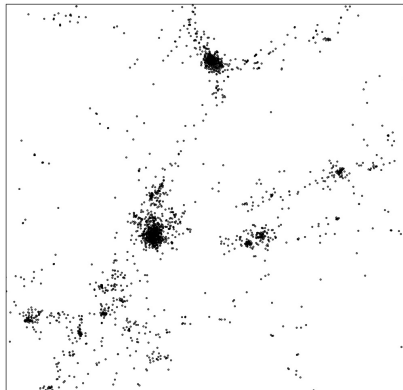
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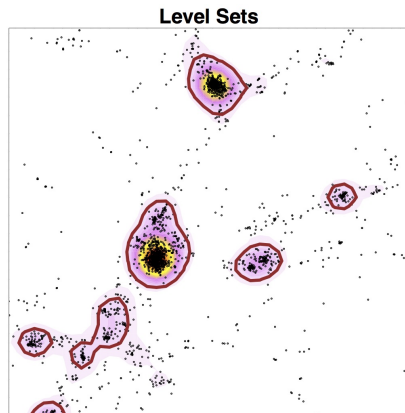
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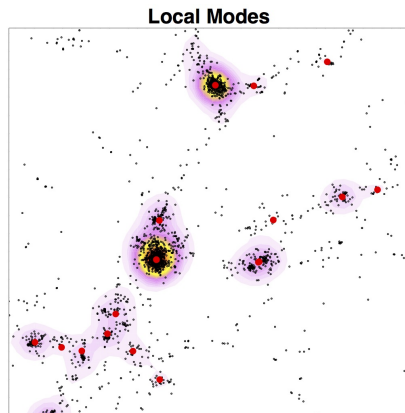
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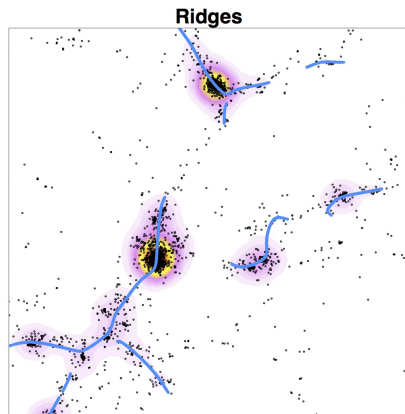
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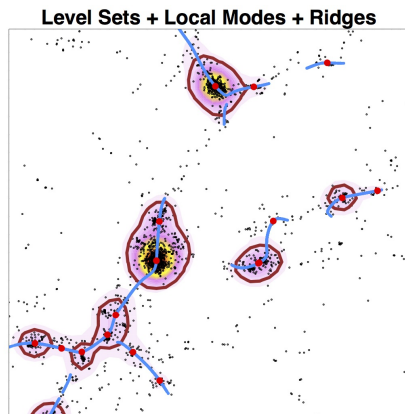
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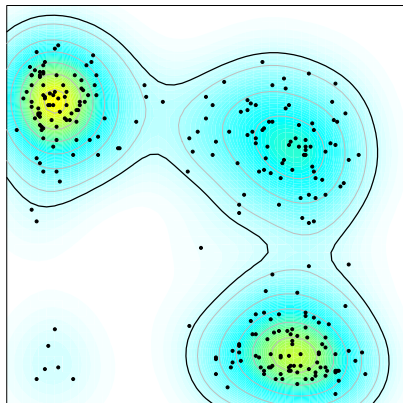
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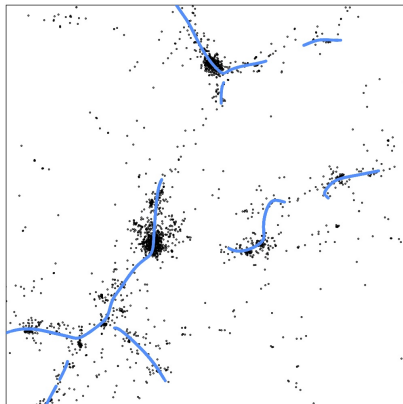
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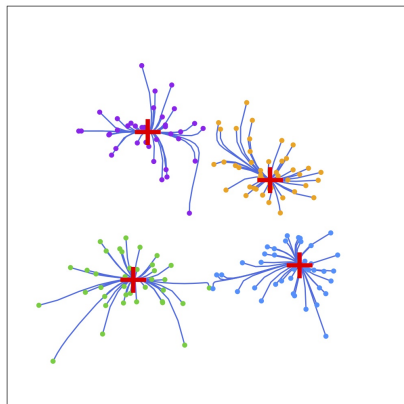
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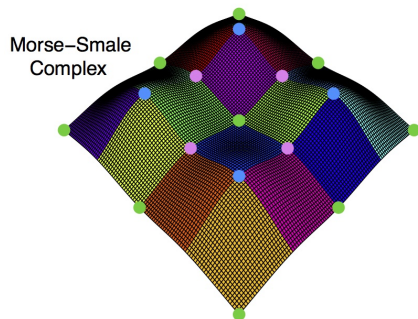
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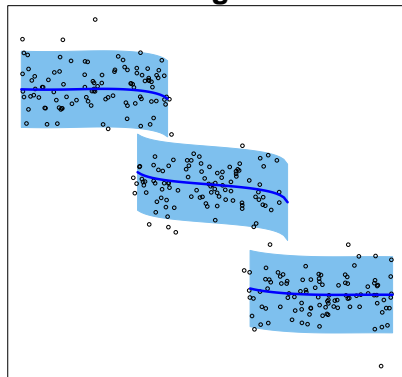


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Modal Regression



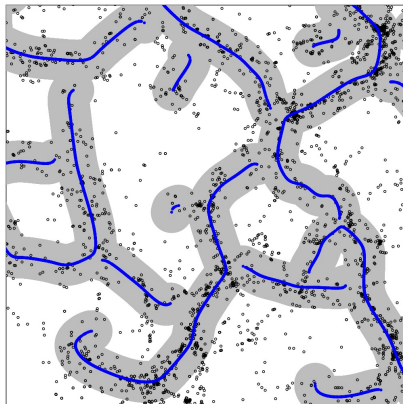
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List of Papers

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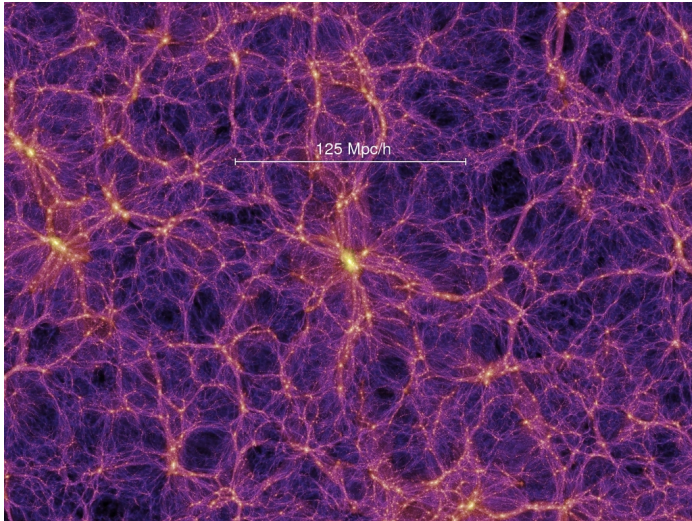
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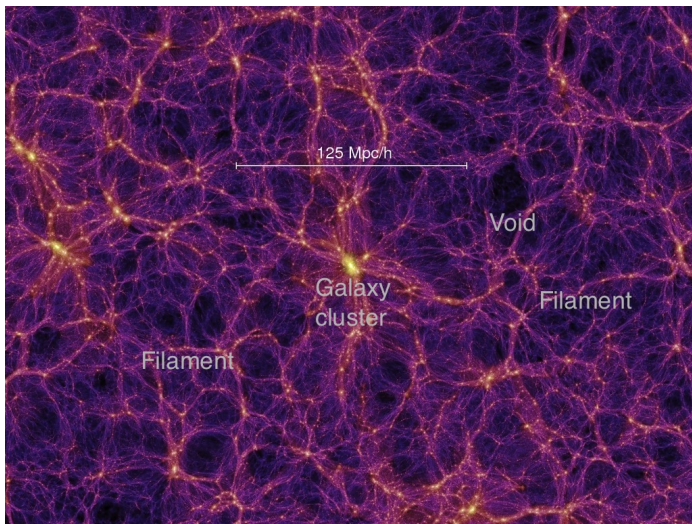
DENSITY RIDGES

Example: Cosmology



Credit: Millennium Simulation

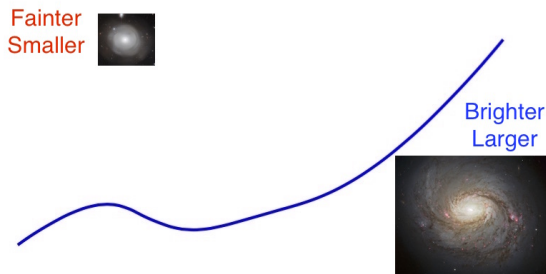
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The Importance of Filaments

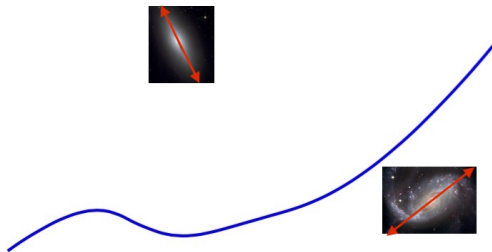
- A galaxy's brightness, size, and mass are associated with the distance to filaments.



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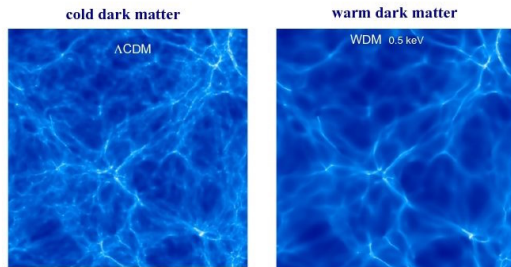
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The Importance of Filaments

- A galaxy's brightness, size, and mass are associated with the distance to filaments.
- A galaxy's alignment is associated with filaments.
- Filaments can be used to test cosmological theories.



- Credit: Kavli Institute for Cosmology, Cambridge

Density Ridges

We formalize the notion of filaments as *density ridges*.

Early work on ridges is in image analysis ([Eberly 1996](#), [Damon 1999](#)).

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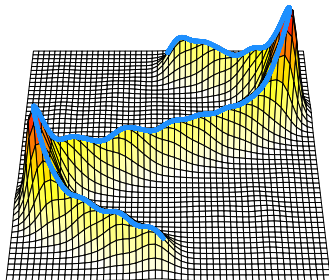
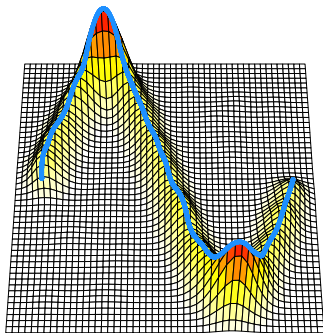
→ In our work, we derive the asymptotic theory for ridge estimators and propose methods for constructing confidence sets.

Example: Ridges in Mountains

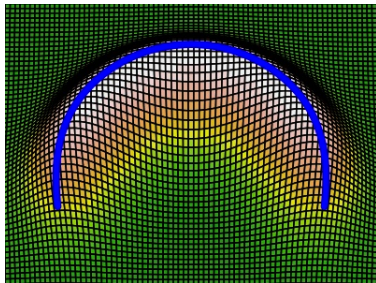
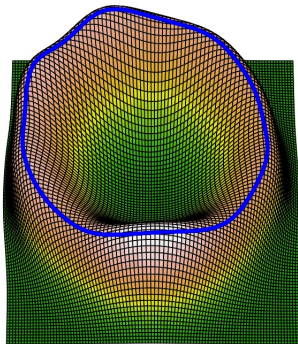


Credit: Google

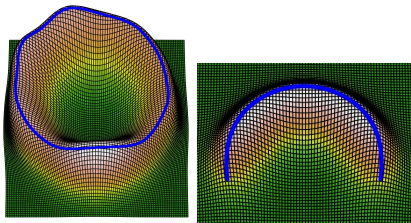
Example: Ridges in Smooth Functions



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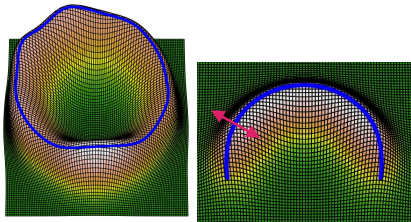


Ridges: Local Modes in Subspace



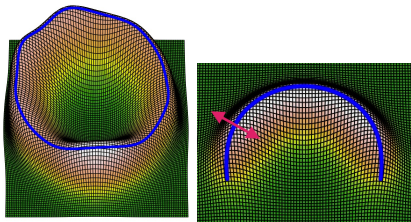
A generalized local mode in a specific 'subspace'.

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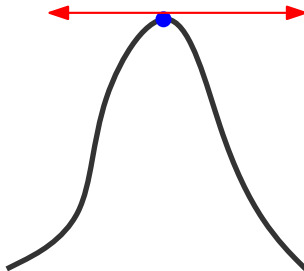


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- Local modes:

$$\text{Mode}(p) = \{x : \nabla p(x) = 0, \lambda_1(x) < 0\}.$$

Dimension of Ridges

The dimension of a ridge is 1.

This is because ridges are points satisfying $V(x)V(x)^T \nabla p(x) = 0$.

$V(x)V(x)^T$ has rank $d - 1$, so there are $d - 1$ effective constraints.

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Note that there are higher dimensional ridges but in this talk, we will focus on 1 dimensional ridges.

Estimator and Algorithm

We use the plug-in estimate:

$$\widehat{R}_h = \text{Ridge}(\widehat{p}_h),$$

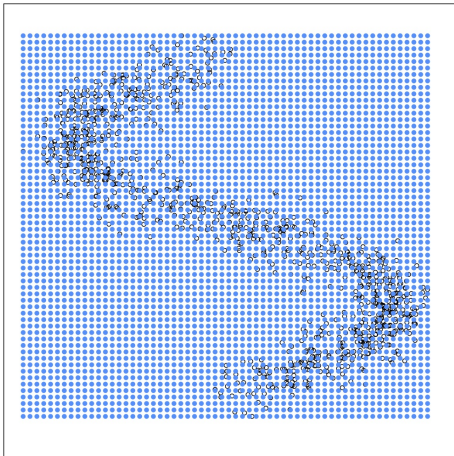
where $\widehat{p}_h = \frac{1}{nh^d} \sum_{i=1}^n K\left(\frac{x-X_i}{h}\right)$ is the kernel density estimator (KDE).

h is the smoothing bandwidth, which controls the amount of smoothing.

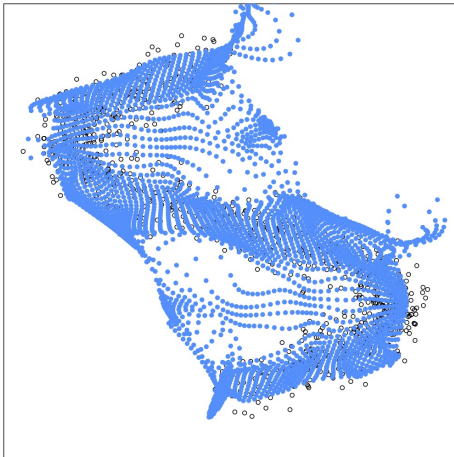
- In general, finding ridges from a given function is hard.
- The Subspace Constraint Mean Shift¹ (SCMS) algorithm allows us to find \widehat{R}_h , ridges of the KDE.

¹Ozertem, Umut, and Deniz Erdogmus. "Locally defined principal curves and surfaces." JMLR (2011).

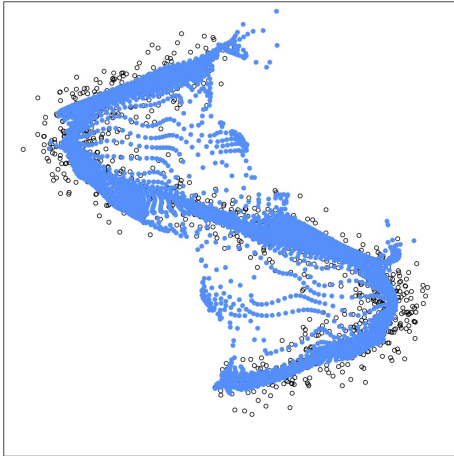
SCMS: Ridge Recovery Algorithm



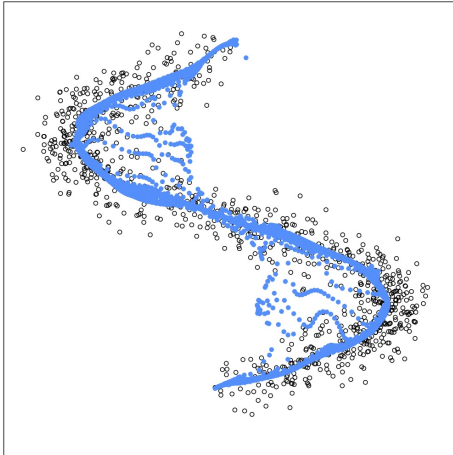
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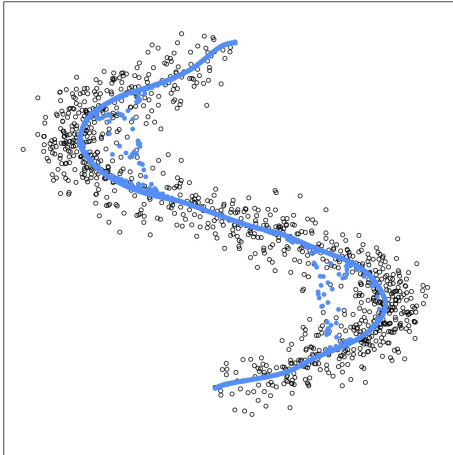
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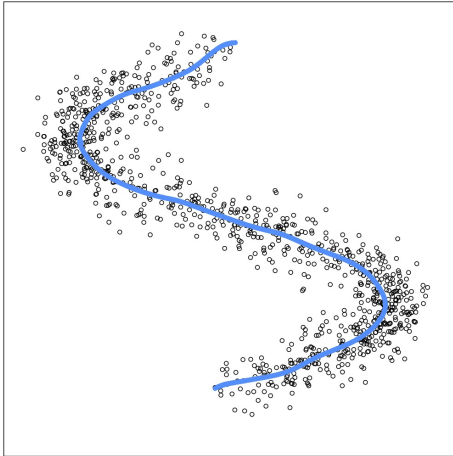
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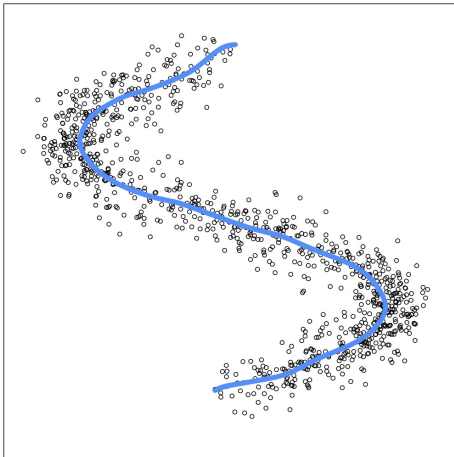
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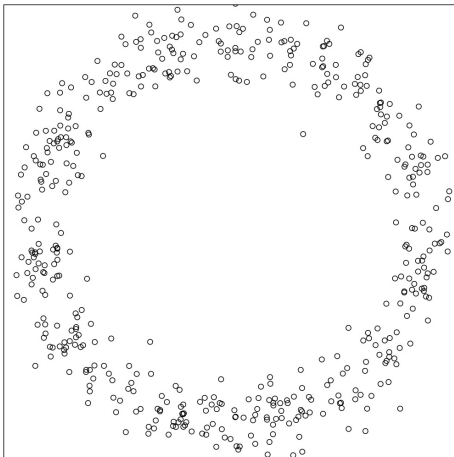


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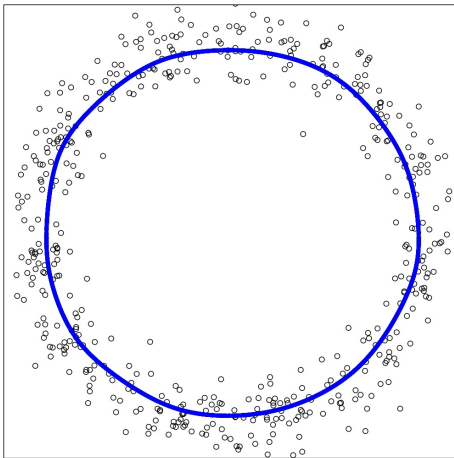


SCMS moves blue mesh points by gradient ascent and a projection.

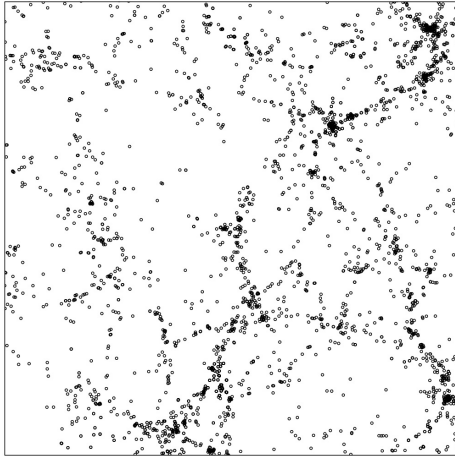
Example for Estimated Density Ridges



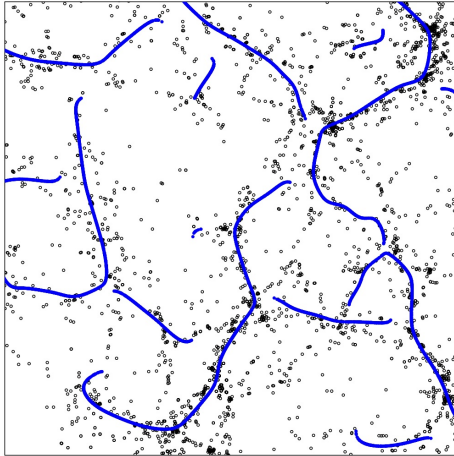
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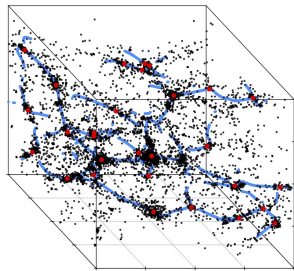
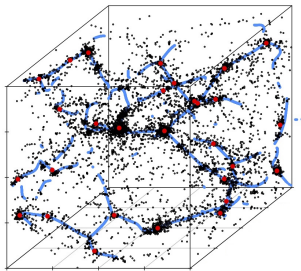
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3D Example for Estimated Ridges



Blue curves: density ridges.

Red points: density local modes.

Statistical Inference: Confidence Sets

Having estimators is not enough for statistical inference.

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We need confidence sets for density ridges.

Namely, we want to find a set $C_{1-\alpha,n}$ from the data such that

$$\mathbb{P}(R \subset C_{1-\alpha,n}) \geq 1 - \alpha.$$

Smoothed Density Ridges

In particular, we focus on making inference for the smoothed ridges $R_h = \text{Ridge}(p_h)$.

p_h is the smoothed density function:

$$p_h(x) = p \otimes K_h(x) = \mathbb{E}(\widehat{p}_h(x)), \quad K_h(x) = \frac{1}{h^d} K\left(\frac{x}{h}\right),$$

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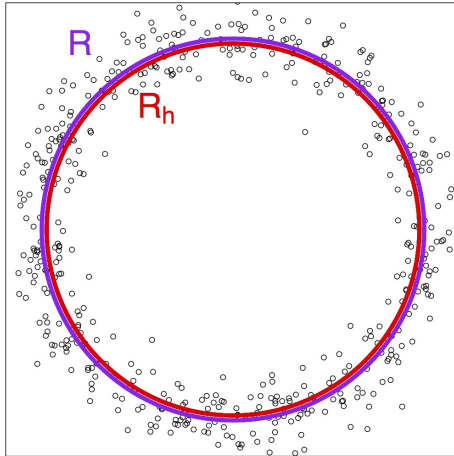
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where \otimes denotes the convolution.

- The advantages of R_h over R :
 - Always well-defined.
 - Topologically similar.
 - We can undersmooth so that inference for R_h is also valid for R .

Ridges VS Smoothed Ridges



Useful Metric: Hausdorff Distance

We introduce a useful metric—the *Hausdorff distance* for sets:

$$\text{Haus}(A, B) = \max \left\{ \sup_{x \in A} d(x, B), \sup_{x \in B} d(x, A) \right\},$$

where $d(x, A) = \inf_{y \in A} \|x - y\|$ is the projection distance from point x to a set A .

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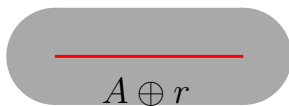
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- Haus is an L_∞ metric for sets.

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Then we have the following inclusion property:

$$A \subset B \oplus \text{Haus}(A, B), \quad B \subset A \oplus \text{Haus}(A, B).$$

Confidence Sets

We can use the Hausdorff distance and \oplus operation to construct confidence sets.

Let F_n be the CDF for $\text{Haus}(\widehat{R}_h, R_h)$ and $t_{1-\alpha} = F_n^{-1}(1 - \alpha)$ be the $1 - \alpha$ quantile.

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- It can be shown that

$$\mathbb{P} \left(R_h \subset \widehat{R}_h \oplus t_{1-\alpha} \right) \geq 1 - \alpha.$$

→ This follows from the property

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- We need to find the distribution F_n .

Asymptotic Theory

Need: F_n , the CDF of $\text{Haus}(\widehat{R}_h, R_h)$.

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Key observation:

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Theorem (Chen, Genovese, and Wasserman (2015))

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- Good news: we have the limiting distribution.
 - Bad news: the limiting distribution involves **unknown quantities**.
- A solution: the bootstrap.

Bootstrap Confidence Set

- Bootstrap sample \implies bootstrap ridges \widehat{R}_h^* .
- Repeat B times, we obtain B bootstrap ridges $\widehat{R}_h^{*(1)}, \dots, \widehat{R}_h^{*(B)}$.
- Compute the CDF estimator \widehat{F}_n by

$$\widehat{F}_n(t) = \frac{1}{B} \sum_{\ell=1}^B I \left(\text{Haus}(\widehat{R}_h^{*(\ell)}, \widehat{R}_h) < t \right)$$

- Choose $\widehat{t}_{1-\alpha}$ be the $1 - \alpha$ quantile for \widehat{F}_n .
- The confidence set is

$$C_{1-\alpha, n} = \widehat{R}_h \oplus \widehat{t}_{1-\alpha}$$

Bootstrap Consistency

We proved that

$$\begin{aligned}\sqrt{nh^{d+2}}\text{Haus}(\widehat{R}_h^*, \widehat{R}_h) &\approx \sup \{\text{Gaussian process on } \widehat{R}_h\} \\ &\approx \sup \{\text{Gaussian process on } R_h\} \\ &\approx \sqrt{nh^{d+2}}\text{Haus}(\widehat{R}_h, R_h).\end{aligned}$$

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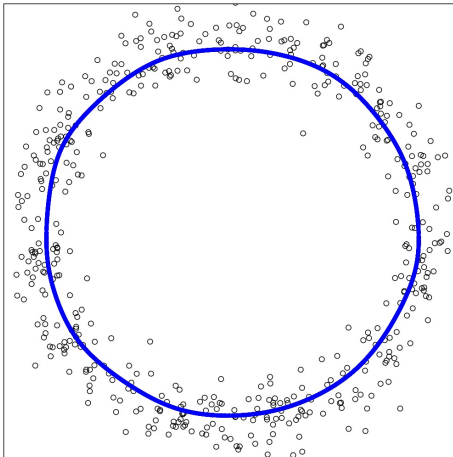
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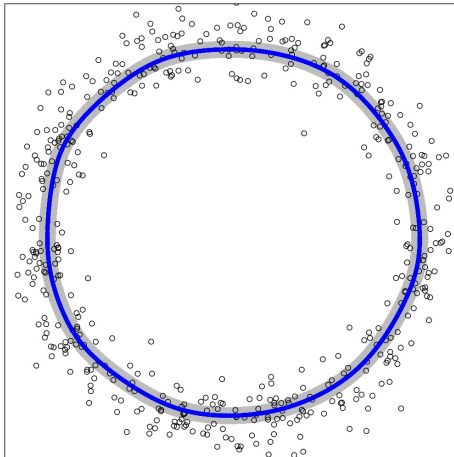
Under regularity conditions and $\frac{\log n}{nh^{d+8}} \rightarrow 0$,

$$\mathbb{P}\left(R_h \subset \widehat{R}_h \oplus \widehat{t}_{1-\alpha}\right) = 1 - \alpha + O\left(\left(\frac{\log^7 n}{nh^{d+2}}\right)^{1/8}\right).$$

Example of Confidence Sets

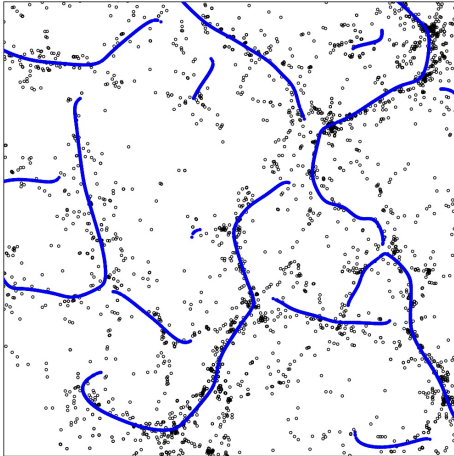


Example of Confidence Sets

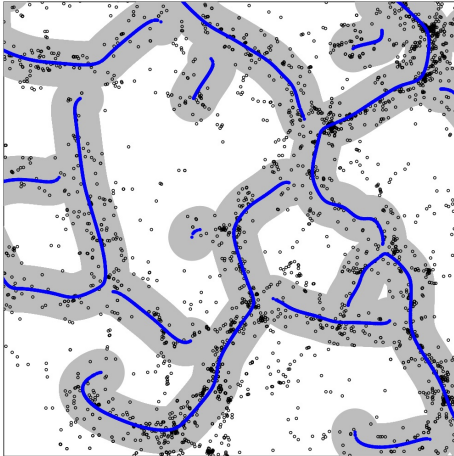


We have checked the coverage by simulation.

Example of Confidence Sets

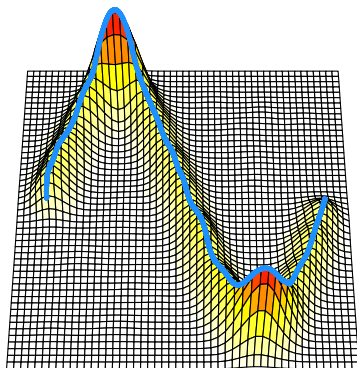


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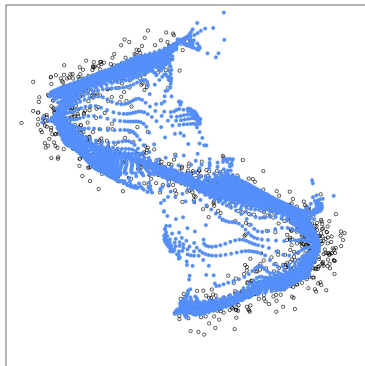
Summary for Density Ridges

- Ridges of the density function.



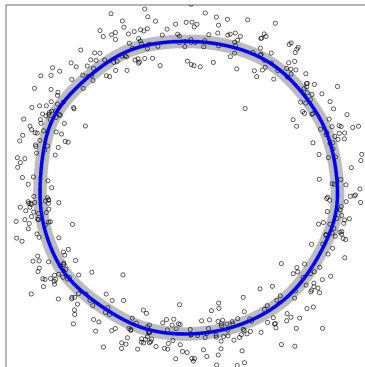
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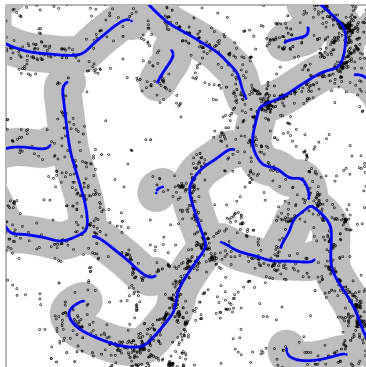
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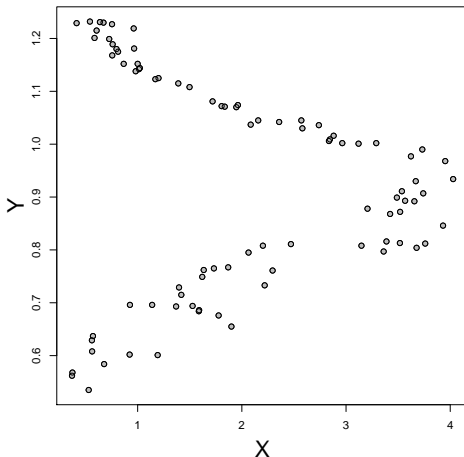
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- An algorithm for the estimator.
- Confidence sets.
- Applications in Astronomy.

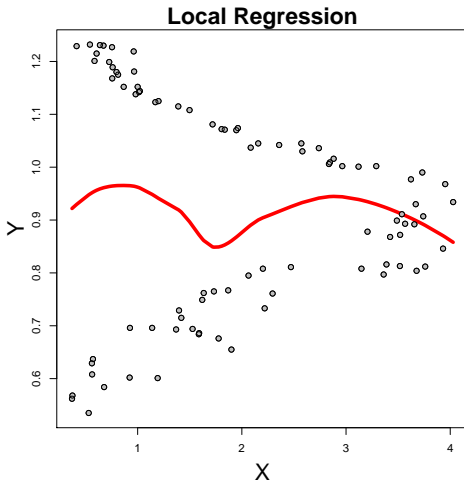


MODAL REGRESSION

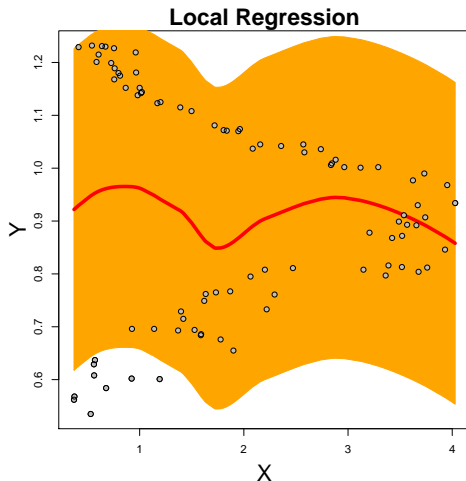
A Motivating Example for Modal Regression



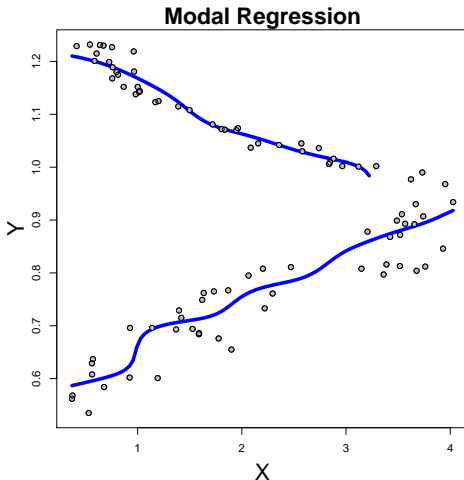
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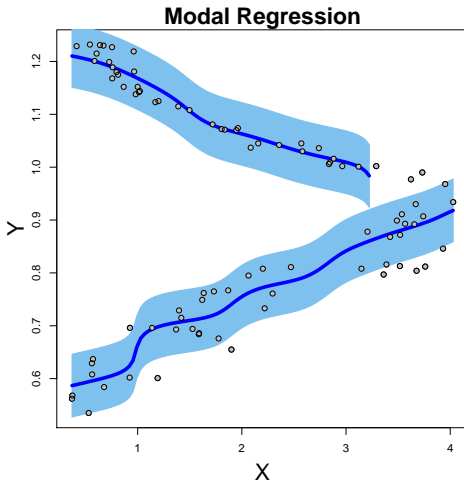
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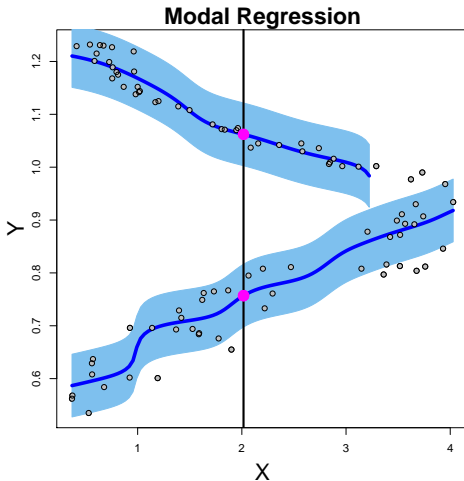
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A Motivating Example for Modal Regression



A Motivating Example for Modal Regression



Modal regression first appeared in [Sager and Thisted \(1982\)](#), [Lee \(1989\)](#), and [Scott \(1992\)](#).

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In most of the above work, they consider the mode of the conditional density function.

→ In our work, we consider the multiple local modes of the conditional density function.

Definition for Modal Regression

We assume $x \in \mathbb{K} \subset \mathbb{R}^d$, where \mathbb{K} is a compact set.

- Modal function—the conditional (local) **modes**:

$$M(x) = \text{Mode}(Y|X = x) = \left\{ y : \frac{d}{dy}p(y|x) = 0, \frac{d^2}{dy^2}p(y|x) < 0 \right\}.$$

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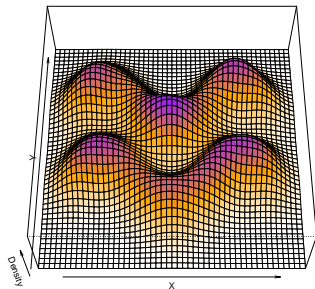
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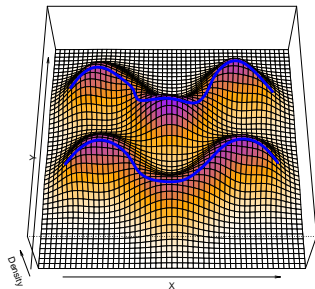
- $M(x)$ is a multi-valued function.
- An equivalent expression:

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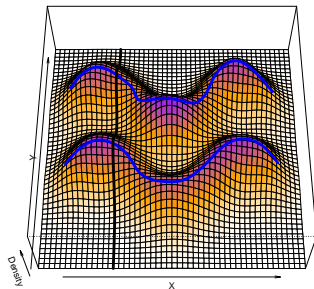
Conditional Local Modes



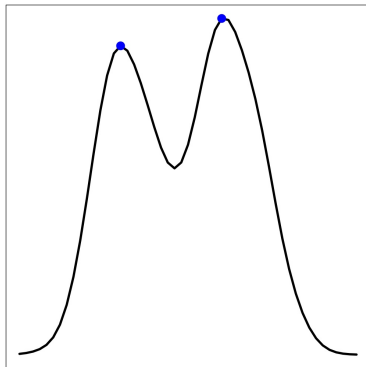
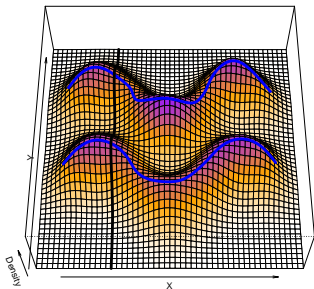
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Estimator for Modal Regression

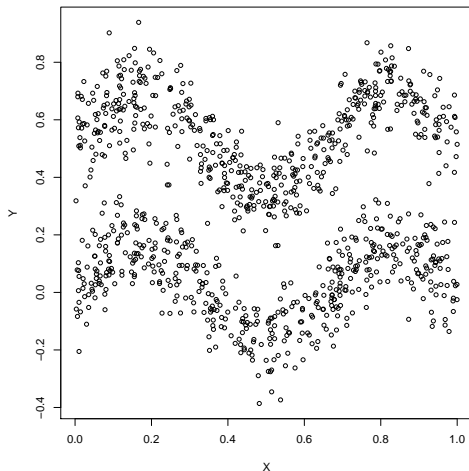
- Our estimator is the plug-in from the KDE:

$$\hat{M}_n(x) = \left\{ y : \frac{\partial}{\partial y} \hat{p}_n(x, y) = 0, \frac{\partial^2}{\partial y^2} \hat{p}_n(x, y) < 0 \right\}.$$

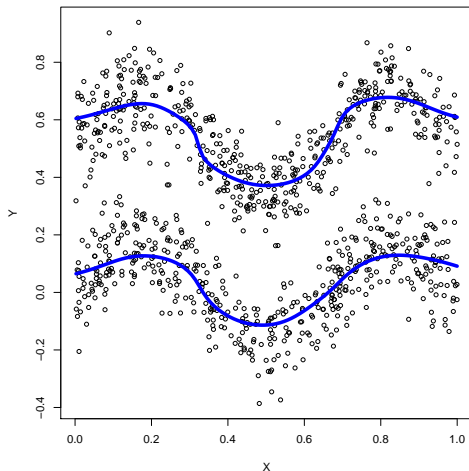
- Partial mean shift²: a simple algorithm for computing $\hat{M}_n(x)$, the plug-in estimator of the KDE, from the data.

²Einbeck, Jochen, and Gerhard Tutz. "Modelling beyond regression functions: an application of multimodal regression to speed-flow data." JRSSC (2006)

Example for Modal Regression



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Losses of Modal regression

To measure the errors, we consider the following two losses:

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- the *uniform* loss

$$\Delta_n = \sup_x \Delta_n(x) = \sup_x \text{Haus}(\widehat{M}_n(x), M(x)).$$

Illustration for Losses

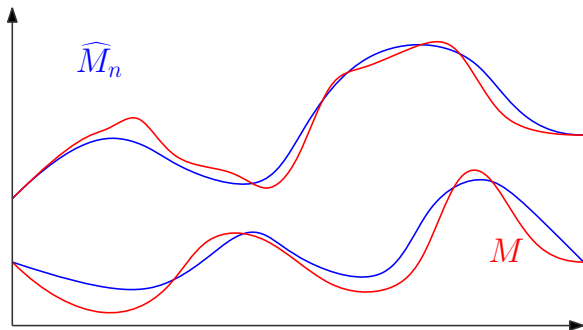


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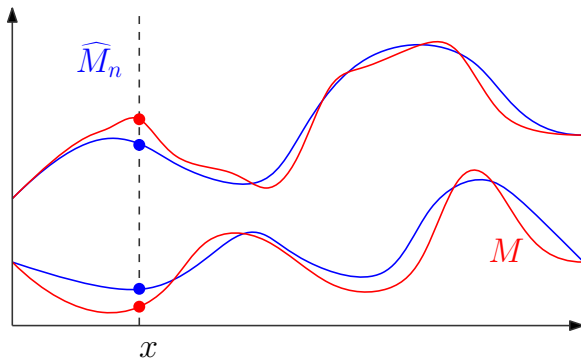


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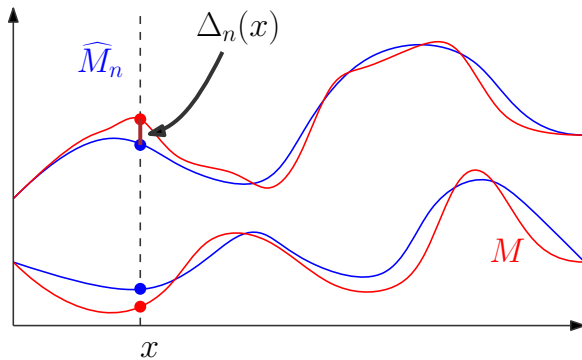
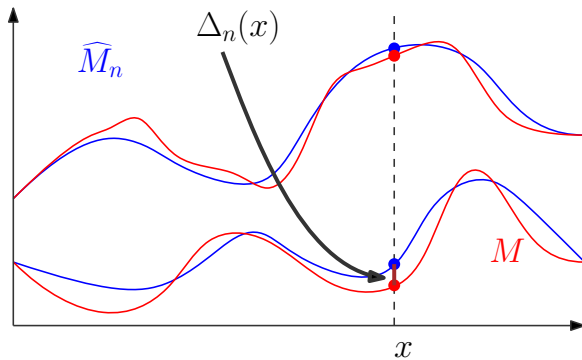


Illustration for Losses



Rate of Convergence

Both the pointwise and the uniform losses obey the common nonparametric rate:

Theorem

Under regularity conditions and $\frac{\log n}{nh^{d+3}} \rightarrow 0$,

$$\Delta_n(x) = O(h^2) + O_P\left(\sqrt{\frac{1}{nh^{d+3}}}\right)$$
$$\Delta_n = O(h^2) + O_P\left(\sqrt{\frac{\log n}{nh^{d+3}}}\right).$$

Risk = Bias + $\sqrt{\text{Variance}}$.

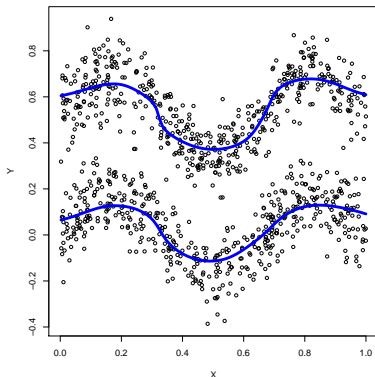
$d + 3 = d + 1 + 2 = \dim(X) + \dim(Y) + \text{gradient}.$

Confidence Sets

We can construct confidence sets using the uniform loss and the bootstrap.

Reason: the uniform loss Δ_n is an L_∞ metric for modal regression.

Bootstrap consistency follows in a similar way as density ridges.

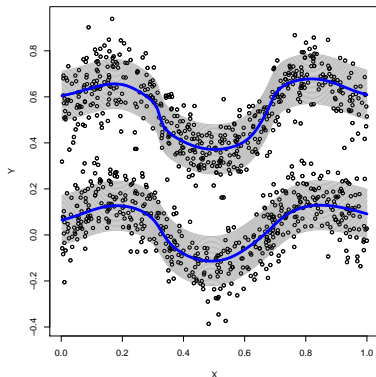


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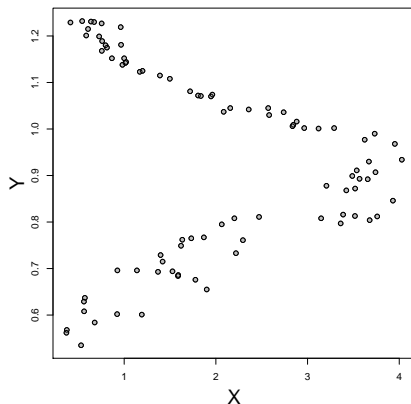
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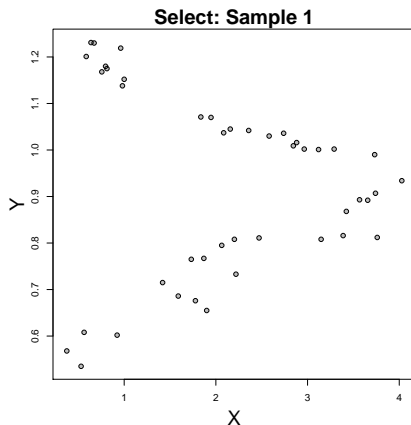
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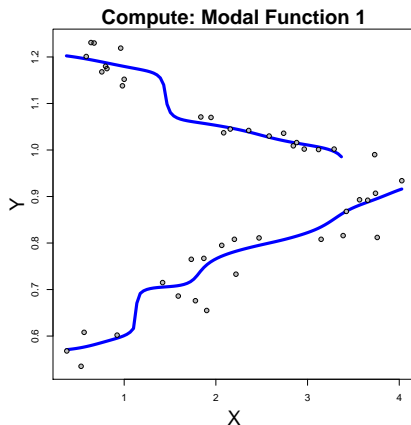
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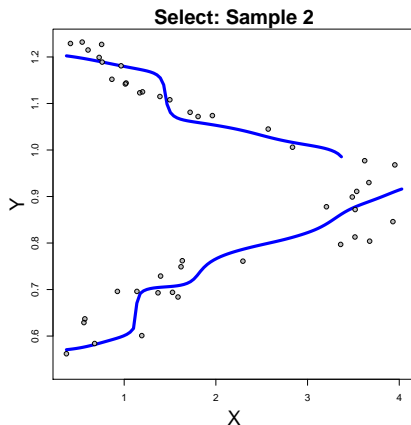
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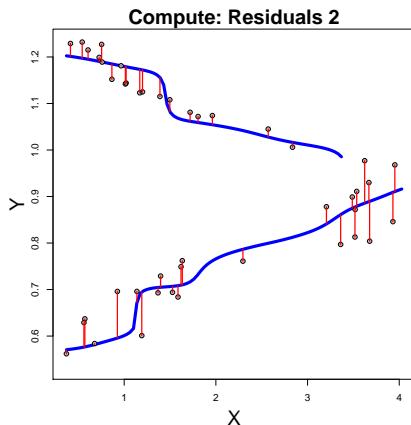
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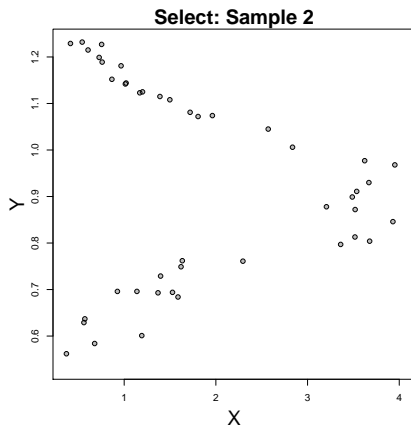
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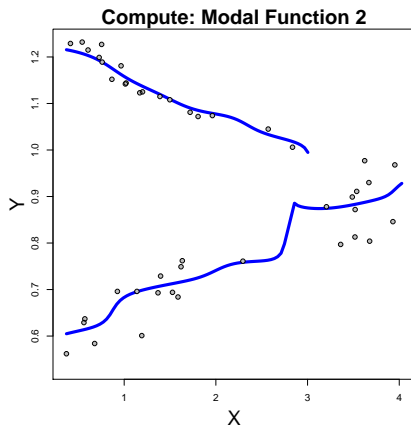
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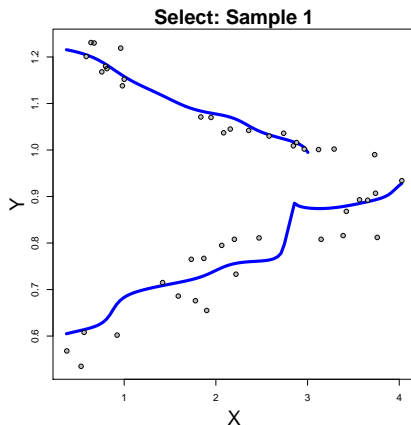
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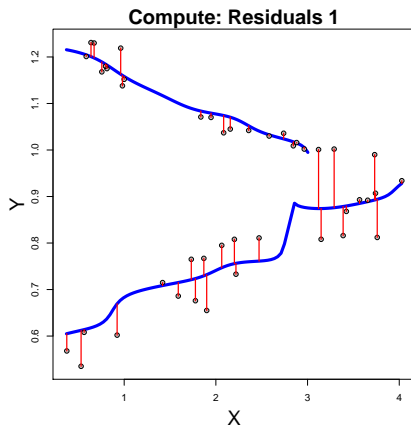
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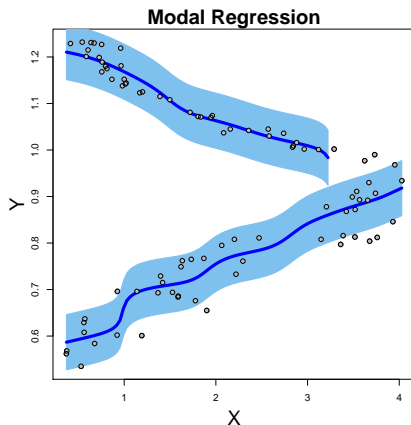
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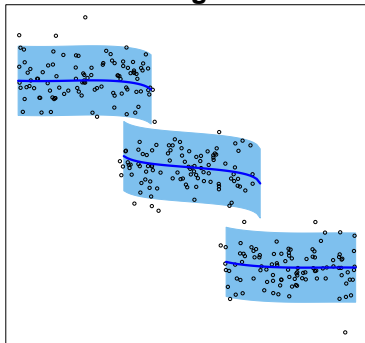
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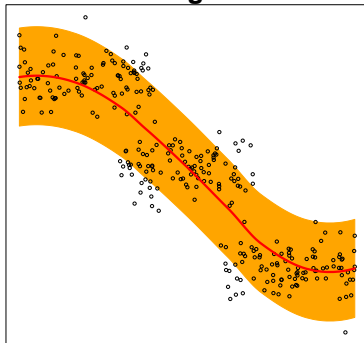


Examples of Prediction Sets

Modal Regression

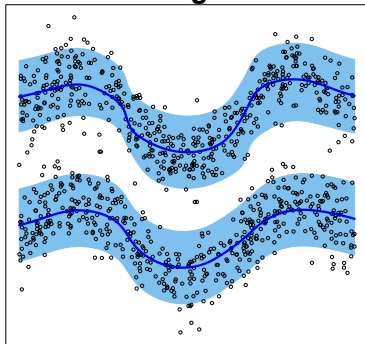


Local Regression

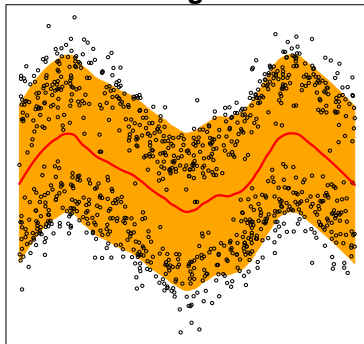


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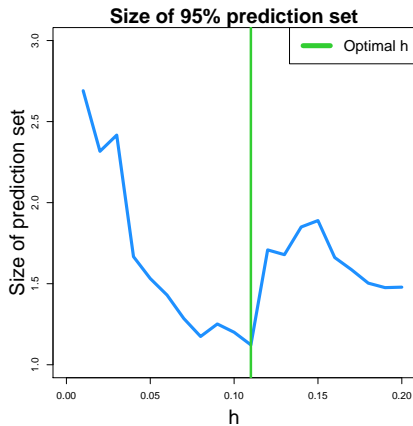
We can choose the smoothing parameter h via minimizing the size of the prediction set.

Namely, we choose

$$h^* = \operatorname{argmin}_{h>0} \operatorname{Vol} \left(\widehat{\mathcal{P}}_{1-\alpha} \right),$$

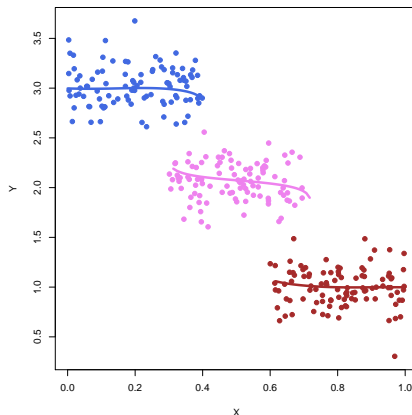
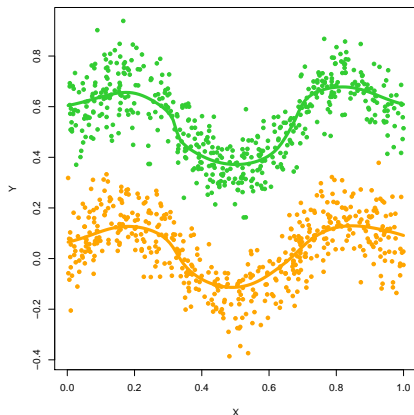
where $\widehat{\mathcal{P}}_{1-\alpha}$ is the prediction set.

Example: Bandwidth Selection



Regression Clustering

- Clustering based on the response Y .
- Clusters as functions of covariates X .

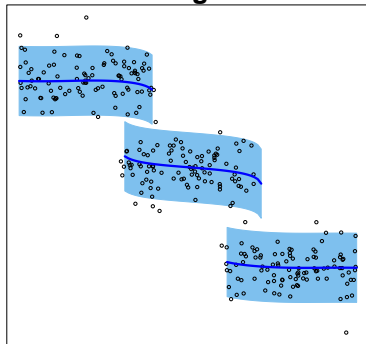


Modal Regression VS Mixture Regression

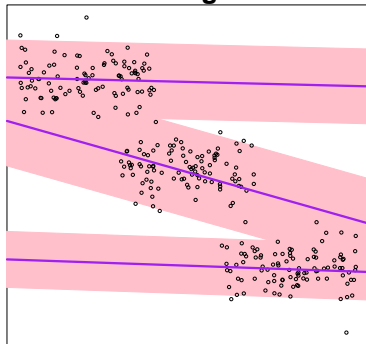
Modal regression and mixture regression are solving different problems.

Here is a case where modal regression gives a better result.

Modal Regression



Mixture Regression



CONCLUDING REMARKS

List of Papers

- Level Sets:

1. **Chen**, Genovese, and Wasserman. "Density Level Sets: Asymptotics, Inference, and Visualization." (2015).

- Ridges:

1. **Chen, Genovese, and Wasserman. "Asymptotic theory for density ridges." The Annals of Statistics (2015).**
2. **Chen**, Genovese, and Wasserman. "Generalized Mode and Ridge Estimation." (2014).
3. **Chen** et al. "Optimal Ridge Detection using Coverage Risk." NIPS (2015).

- Clustering:

1. **Chen**, Genovese, and Wasserman. "A Comprehensive Approach to Mode Clustering." The Electronic Journal of Statistics (2016+).
2. Azizyan and **Chen** et al. "Risk Bounds for Mode Clustering." (2015).

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- Morse-Smale Complex:

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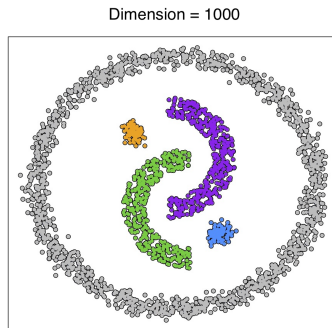
- Astronomy:

1. **Chen** et al. "Detecting Effects of Filaments on Galaxy Properties in Sloan Digital Sky Survey III." (2015).
2. **Chen** et al. "Cosmic Web Reconstruction through Density Ridges: Catalogue." (2015).
3. **Chen** et al. "Investigating Galaxy-Filament Alignment in Hydrodynamic Simulations using Density Ridges." Mon. Not. Roy. Astro. Soc. (2015).
4. **Chen** et al. "Cosmic Web Reconstruction through Density Ridges: Method and Algorithm." Mon. Not. Roy. Astro. Soc. (2015).

Future Work

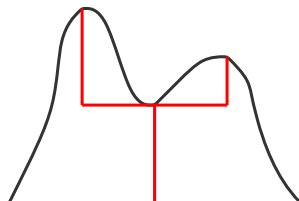
Some future directions:

- More to do in geometric features.
- High-dimensional clustering.
- Influence for visualization tools.
- Interdisciplinary collaborations.



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More details can be found in: <http://www.stat.cmu.edu/~yenchic/>

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Thank you!

1. Chen, Yen-Chi, Christopher R. Genovese, and Larry Wasserman. "Density Level Sets: Asymptotics, Inference, and Visualization." Under review of the Journal of American Statistical Association. arXiv preprint arXiv:1504.05438 (2015).
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- 4. Backups for Density Ridges
 - 4.1 Regularity Conditions
 - 4.2 Bandwidth Selection
 - 4.3 Local Uncertainty
 - 4.4 Why Smoothed Structures?
 - 4.5 General Ridges
 - 4.6 Illustration for Asymptotics
- 5. Backups for Modal Regression
 - 5.1 Regularity Conditions
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 - 5.3 Bifurcation and Merge
 - 5.4 Comparisons
 - 5.5 Theory for Prediction Sets
 - 5.6 More about Confidence Sets

BACKUPS FOR DENSITY RIDGES

Regularity Conditions

- (K1) The kernel function K is \mathbf{BC}^4 and integrable.
- (K2) K satisfies the VC-type class condition.
- (P1) The density p is in \mathbf{BC}^4 .
- (P2) The eigengap $\lambda_1(x) - \lambda_2(x) \geq \beta_0 > 0$ for points around ridges.
- (P3) The orientation of each ridge point is close to the gradient.

Regularity Conditions on Kernel Functions

(K1) The kernel K is in \mathbf{BC}^4 and $\|K\|_{\infty,4}^* < \infty$.

(K2) Let

$$\mathcal{K}_r = \left\{ y \mapsto K^{(\alpha)} \left(\frac{x - y}{h} \right) : x \in \mathbb{R}^d, |\alpha| = r \right\},$$

where $K^{(\alpha)}$ is the α -th derivative and let $\mathcal{K}_l^* = \bigcup_{r=0}^l \mathcal{K}_r$. We assume that \mathcal{K}_4^* is a VC-type class. i.e. there exists constants A, v and a constant envelope b_0 such that

$$\sup_Q N(\mathcal{K}_4^*, \mathcal{L}^2(Q), b_0 \epsilon) \leq \left(\frac{A}{\epsilon} \right)^v, \quad (1)$$

where $N(T, d_T, \epsilon)$ is the ϵ -covering number for an semi-metric set T with metric d_T and $\mathcal{L}^2(Q)$ is the L_2 norm with respect to the probability measure Q .

Regularity Conditions on Distributions

(P1) The density p_h is in \mathbf{BC}^4 .

(P2) There exists constants $\beta_0, \beta_1, \beta_2, \delta_0 > 0$ such that

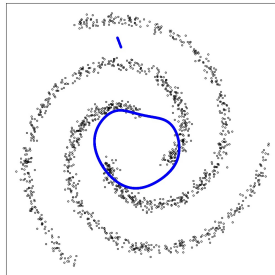
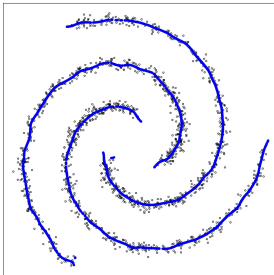
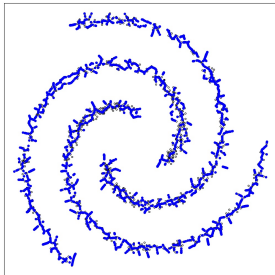
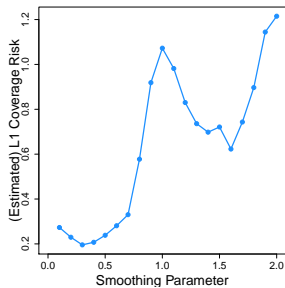
$$\begin{aligned}\lambda_2(x) &\leq -\beta_1 \\ \lambda_1(x) &\geq \beta_0 - \beta_1 \\ \|g_h(x)\| \max_{|\alpha|=3} |p_h^{(\alpha)}(x)| &\leq \beta_0(\beta_1 - \beta_2)\end{aligned}\tag{2}$$

for all $x \in R_h \oplus \delta_0$.

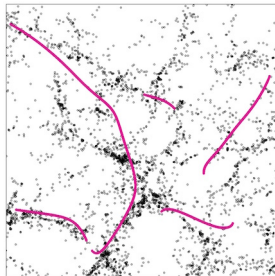
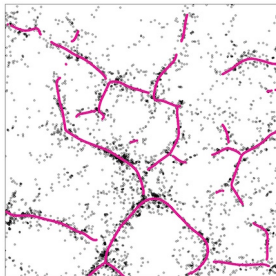
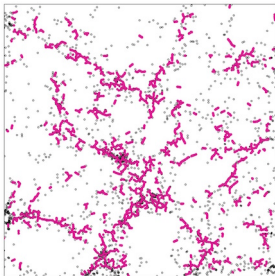
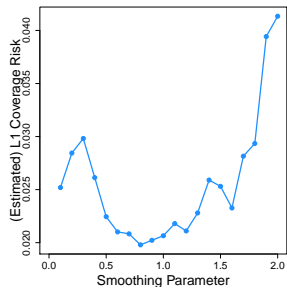
(P3) For each $x \in R_h$, $|e(x)^T g_h(x)|^2 \geq \frac{\lambda_1(x)}{\lambda_1(x) - \lambda_2(x)}$ where $e(x)$ is the direction of R_h at point $x \in R_h$.

(P4) The above assumptions hold for all sufficiently small h .

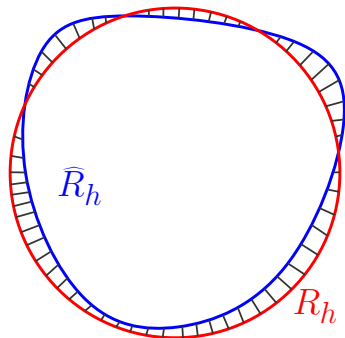
Bandwidth Selection



Bandwidth Selection



Bandwidth Selection

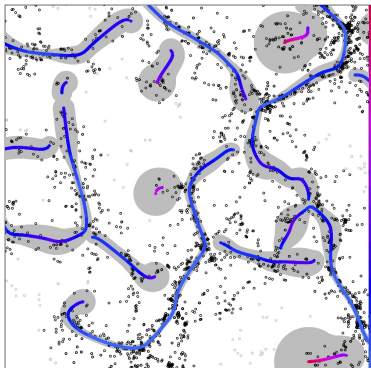


L_1 distance are like the area of the shady regions.

We estimate this distance by data splitting or the bootstrap.

Reference: **Chen** et al. 'Optimal Ridge Detection using Coverage Risk' (NIPS 2015).

Local Uncertainty and Pointwise Confidence Sets

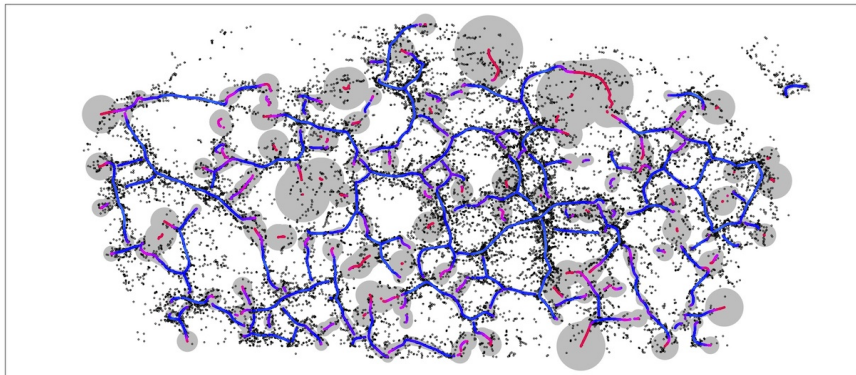


Color denotes the amount of uncertainty.

Red: unstable filaments.

Blue: stable filaments.

Local Uncertainty and Pointwise Confidence Sets



Color denotes the amount of uncertainty.

Red: unstable filaments.

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Why Smoothed Density? - Bias Consideration

We have the following decomposition:

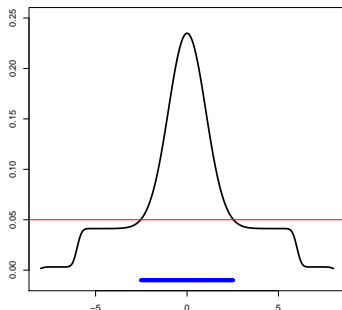
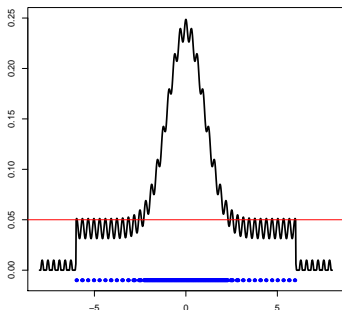
$$\begin{aligned}\text{Haus}(\widehat{R}_h, R) &\leq \text{Haus}(R_h, R) + \text{Haus}(\widehat{R}_h, R) \\ &= O(h^2) + O_P\left(\sqrt{\frac{\log n}{nh^{d+2}}}\right).\end{aligned}$$

Bias + $\sqrt{\text{Variance}}$.

Work on smoothed ridges R_h allows us to avoid the problem of bias.

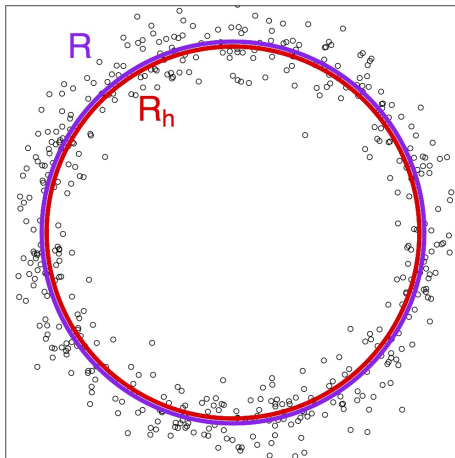
Optimal rate: $O_P\left(\left(\frac{\log n}{n}\right)^{\frac{2}{d+6}}\right)$ when we choose $h = O\left(\left(\frac{\log n}{n}\right)^{\frac{1}{d+6}}\right)$.

Why Smoothed Density? - A Level Set Example



Ridges VS Smoothed Ridges

Radius of ring: $r = 1$. Smoothing bandwidth: $h = 0.25$. Gaussian noise level: $\sigma = 0.1$



General Ridges

We can generalize ridges to higher dimensions. Pick

$$V_r(x) = [v_{r+1}(x), \dots, v_d(x)].$$

We define

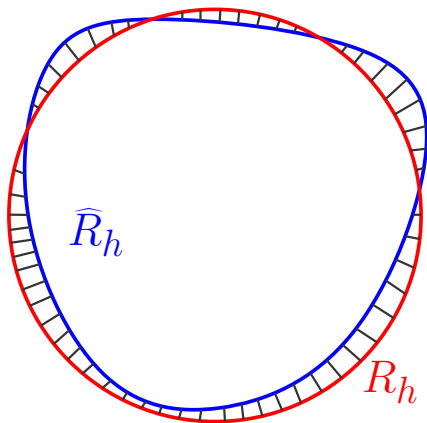
$$r\text{-Ridge}(p) = \{x : V_r(x)V_r(x)^T \nabla p(x) = 0, \lambda_{r+1}(x) < 0\}.$$

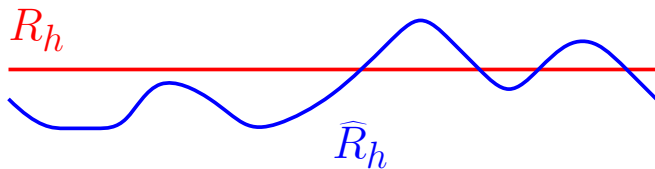
$V_r(x)$ is a $d \times (d - r)$ matrix. There are $d - r$ constraints.

By Implicit Function Theorem, r -ridges are r -manifolds.

In Astronomy, $r = 2$ can be used to model 'Cosmic Sheets (Walls)'.

$r = 0$ coincides with the definition of local modes.





BACKUPS FOR MODAL REGRESSION

Regularity Conditions

- (K1) The kernel function K is \mathbf{BC}^4 and integrable.
- (K2) K satisfies the VC-type class condition.
- (P1) The density p is in \mathbf{BC}^4 .
- (P2) The second derivative along y axis is bounded away from 0 for points on M .
- (P3) M contains L well-separated, connected components.

Regularity Conditions on Kernel Functions

(K1) The kernel K is in \mathbf{BC}^4 and $\|K\|_{\infty,4}^* < \infty$.

(K2) Let

$$\mathcal{K}_r = \left\{ y \mapsto K^{(\alpha)} \left(\frac{x - y}{h} \right) : x \in \mathbb{R}^d, |\alpha| = r \right\},$$

where $K^{(\alpha)}$ is the α -th derivative and let $\mathcal{K}_l^* = \bigcup_{r=0}^l \mathcal{K}_r$. We assume that \mathcal{K}_2^* is a VC-type class. i.e. there exists constants A, v and a constant envelope b_0 such that

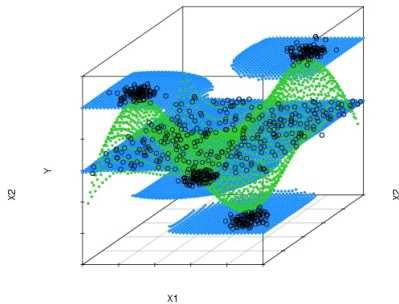
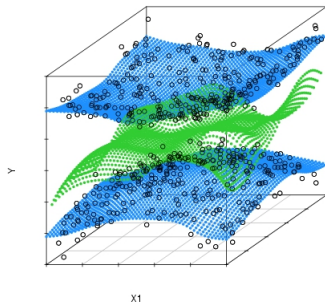
$$\sup_Q N(\mathcal{K}_2^*, \mathcal{L}^2(Q), b_0 \epsilon) \leq \left(\frac{A}{\epsilon} \right)^v, \quad (3)$$

where $N(T, d_T, \epsilon)$ is the ϵ -covering number for an semi-metric set T with metric d_T and $\mathcal{L}^2(Q)$ is the L_2 norm with respect to the probability measure Q .

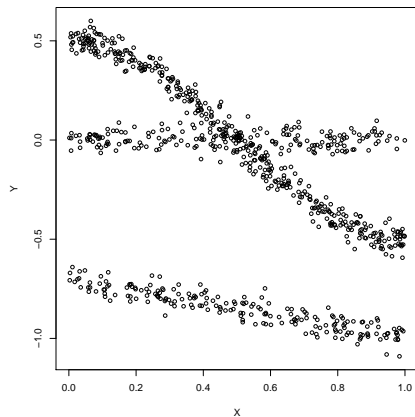
Regularity Conditions on Distributions

- (P₁) The density p is in \mathbf{BC}^4 .
- (P₂) There exists constants $\lambda_0 > 0$ such that for any $(x, y) \in \mathbb{K} \times \mathbb{R}$ with $\frac{\partial}{\partial y} p(x, y) > 0$, the second derivative satisfies $\frac{\partial^2}{\partial^2 y} p(x, y) \leq -\lambda_0 < 0$.
- (P₃) Modal function $M = \cup_{j=1}^L M_j$, where each M_j is a connected component with $M_j \cap M_i = \emptyset$ for $i \neq j$.

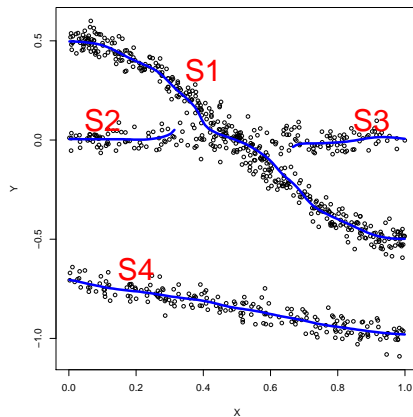
3D Modal Regression



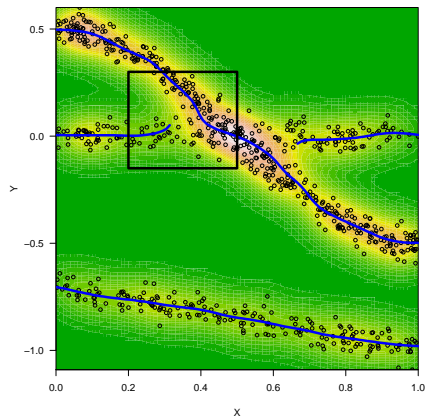
Bifurcation and Merge



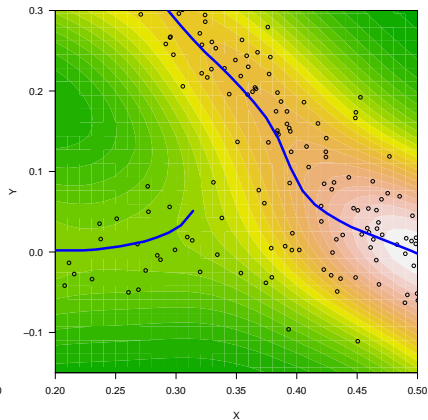
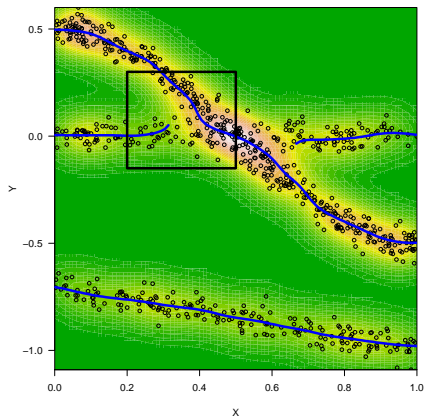
Bifurcation and Merge



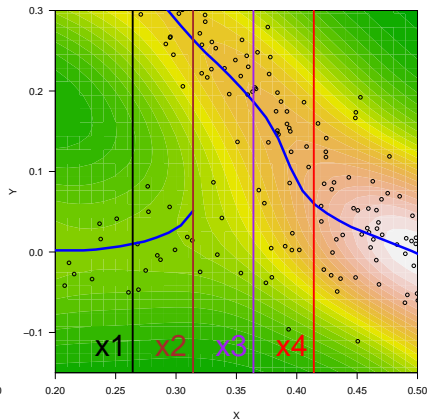
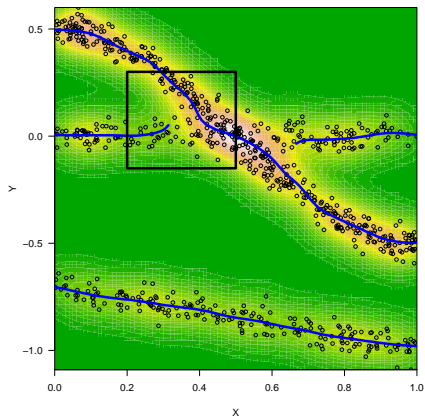
Bifurcation and Merge



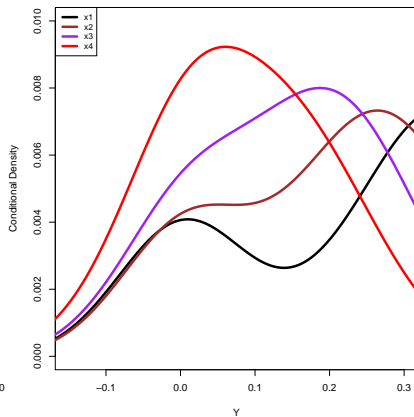
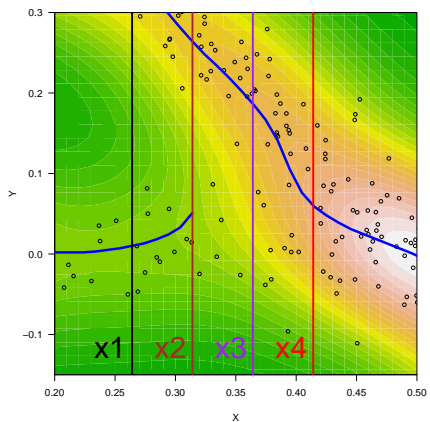
Bifurcation and Merge



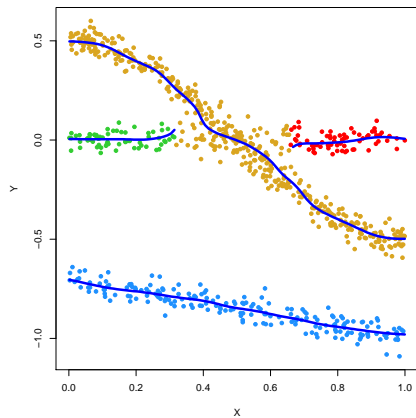
Bifurcation and Merge



Bifurcation and Merge



Bifurcation and Merge



Comments on Mixture Regression

A general model for mixture regression:

$$p(y|x) = \sum_{j=1}^K \pi_j(x) \phi_j(y; \mu_j(x), \sigma_j^2(x)),$$

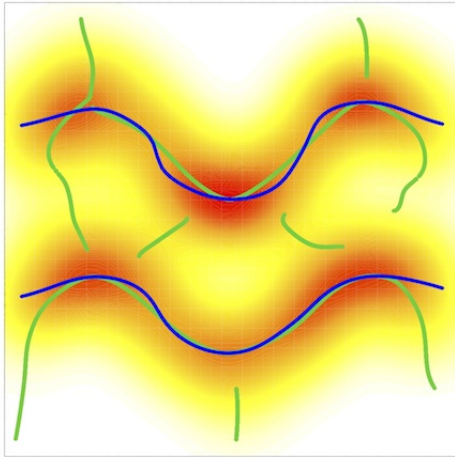
where each $\phi_j(y; \mu, \sigma^2)$ is a density function with mean μ and variance σ^2 .

Common assumptions:

1. $\pi_j(x) = \pi_j$.
2. $\mu_j(x) = \beta_j^T x$.
3. $\sigma_j^2(x) = \sigma_j^2$.
4. ϕ_j is a Gaussian.

Generally, we need to use EM algorithm to estimate the parameters.

Modal Regression VS Density Ridges



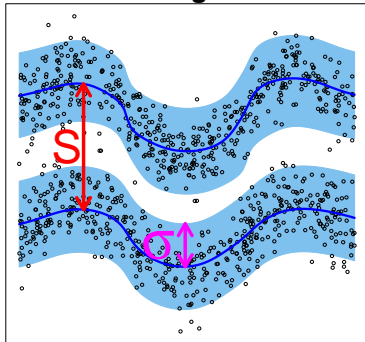
Mixture Inference versus Modal Inference

	Mixture-based	Mode-based
Density estimation	Gaussian mixture	Kernel density estimate
Clustering	K-means	Mean-shift clustering
Regression	Mixture regression	Modal regression
Algorithm	EM	Mean-shift
Complexity parameter	K (number of components)	h (smoothing bandwidth)
Type	Parametric model	Nonparametric model

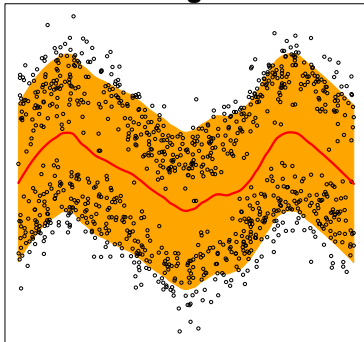
Table: Comparison for methods based on mixtures versus modes.

Theory for Prediction Sets

Modal Regression



Local Regression



Theorem (Chen, Genovese, and Wasserman (2015))

When the signal-to-noise ratio S/σ is sufficiently large, the modal regression has a smaller prediction set than the nonparametric regression.

Confidence Sets

We can construct confidence sets using the uniform loss.

Reason: the uniform loss Δ_n is like an L_∞ metric for modal regression.

Let $t_{1-\alpha}$ be the $1 - \alpha$ quantile of F_n , the CDF of Δ_n .

$\hat{M}_n(x) \pm t_{1-\alpha}$ is a confidence set for $M(x)$ uniformly for all x .

Problem: $t_{1-\alpha}$ cannot be computed.

Solution: the bootstrap.

The Bootstrap

- Bootstrap sample \Rightarrow bootstrap modal function \widehat{M}_n^* .
- Repeat B times, we obtain B bootstrap modal functions $\widehat{M}_n^{*(1)}, \dots, \widehat{M}_n^{*(B)}$.
- Compute $\widehat{\Delta}_n^{*(1)}, \dots, \widehat{\Delta}_n^{*(B)}$ by $\widehat{\Delta}_n^{*(\ell)} = \sup_x \text{Haus}(\widehat{M}_n^{*(\ell)}(x), \widehat{M}_n(x))$.
- Compute the CDF estimator \widehat{F}_n by

$$\widehat{F}_n(t) = \frac{1}{B} \sum_{\ell=1}^B I(\widehat{\Delta}_n^{*(\ell)} < t).$$

- Choose $\widehat{t}_{1-\alpha}$ be the $1 - \alpha$ quantile for \widehat{F}_n .
- $\widehat{M}_n(x) \pm \widehat{t}_{1-\alpha}$ is an asymptotic confidence set uniformly for all x .

Bootstrap consistency follows in the similar way as ridges.

Pointwise Confidence Sets

