# Biostat 551: Homework 5 

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Due December 4, 2014

1. [50 points] In homework 4, the Pedstats software was used to analyze the data files ' $B M I$ BIOST551.ped' and 'BMI BIOST551.dat' for heritability estimation of BMI in European pedigrees for different classes of relative pairs. You will now estimate heritability of BMI considering all pedigree individuals together in a joint analysis using mixed linear models with the QTDT software. The two data files for the European pedigrees previously used for the homework 4 BMI heritability analysis with Pedstats (available on the course website) are also in the required format for the QTDT software.
(a) Provide variance component estimates for BMI using a mixed linear model (MLM) with the following two random effects: additive polygenic and unique (or non-shared) environmental effects. Provide an estimate of the narrow sense heritability of BMI from the variance component estimates from this mixed linear model.

From the 'regress.tbl' file:

```
Family #1 var-covar matrix terms [2]...[[Ve]][[Vg]]
Family #1 regression matrix...
        [linear] =
        [2 x 1] Mu
            2.5 1.000
            2.6 1.000
Some useful information...
                            df : 26412
    log(likelihood) : 76322.95
            variances : 11.895 7.806
                means : 24.844
```

Additive polygenic variance: $=\sigma_{g}^{2}=7.806$
Unique enviromental variance: $=\sigma_{e}^{2}=11.895$
$h^{2}=\frac{\sigma_{g}^{2}}{\sigma_{g}^{2}+\sigma_{e}^{2}}=\frac{7.806}{7.806+11.895}=0.396$.
(b) Now provide variance components estimates for BMI assuming the following three random effects: additive polygenic, shared household, and unique environmental effects. Provide an estimate for narrow sense heritability of BMI with the variance component estimates from this mixed linear model.
From the 'regress.tbl' file:

```
Family #1 var-covar matrix terms [3]...[[Ve]][[Vc]][[Vg]]
Family #1 regression matrix...
        [linear] =
        [2 x 1] Mu
            2.5 1.000
            2.6 1.000
Some useful information...
                            df : 26411
    log(likelihood) : 76240.47
            variances : 14.378 3.304 2.220
                means : 24.841
```

Additive polygenic variance: $=\sigma_{g}^{2}=2.220$
Shared household variance: $=\sigma_{c}^{2}=3.304$
Unique enviromental variance: $=\sigma_{e}^{2}=14.378$
$h^{2}=\frac{\sigma_{g}^{2}}{\sigma_{g}^{2}+\sigma_{e}^{2}}=\frac{2.220}{2.220+3.304+14.378}=0.112$.
(c) Is the shared household effect for BMI significant for the mixed linear model in 1(b) above? Provide evidence to support your answer.
From the output:
he following models will be evaluated...
NULL MODEL
Means $=\mathrm{Mu}$
Variances $=$ Ve + Vg
FULL MODEL
Means $=\mathrm{Mu}$
Variances $=\mathrm{Ve}+\mathrm{Vc}+\mathrm{Vg}$

Testing trait: BMI


Allele $d f(0)-\operatorname{LnLk}(0) \quad d f(V)-\operatorname{LnLk}(V) \quad C h i S q \quad p$ N/A $2641276322.95 \quad 2641176240.47 \quad 164.96 \quad 9 \mathrm{e}-38 \quad$ (26415 probands)

Run completed on Fri Nov 28 10:17:48 2014
1 tests carried out

The most significant result refers to:
Trait: BMI
ChiSq: 164.956
p-value: 9e-38
The likelihood ratio test comparing the model with additive and environmental effects to the model with additive, household, and environmental effects was significant with a p-value of $9 \times 10^{-38}$ and a chisquare statistic of 165 . Therefore, the household effect is significant (we can conclude that there is sufficient evidence to
reject the hypothesis that there is no household effect)
(d) Now compare your BMI heritability estimates from the mixed linear models with variance components in $1(a)$ and $1(b)$ above to the BMI estimates obtained in homework 4 based on different relative pair types. Discuss and provide plausible explanations for any differences and/or similarities that you find in the heritability estimates.
In homework 4 , the heritibility estimates ranged from 0.38 to 0.58 across different relative relationships (with sibling relationships being the highest). The estimate in part (a) (0.396) is in that range which makes sense since both the estimate in part (a) and those from homework 4 assumed no dominance and no household effect. The estimate in part (b) (0.112) accounted for a non-zero household effect, and was found to be lower since $\sigma_{c}^{2}$ was found to be significantly different from zero, as mentioned in part (c).
2. [50 points] A quantitative trait $Y$ is influenced by a single autosomal locus, family effects, and unique environmental effects in a population. The autosomal locus that influences the trait is bi-allelic with alleles $A$ and $a$, and the frequency of allele $A$ is 0.15 . Assume that the population is in Hardy-Weinberg equilibrium (HWE) at the locus. For an individual $j$ from family $i$ in the population, the trait has the following model:

$$
Y_{i j}=1.5+0.4 G_{i j}+c_{i}+\epsilon_{i j}
$$

where $G_{i j}$ is the number of copies of the $A$ allele that individual $j$ from family $i$ has, $c_{i}$ is the family effect on the trait for family $i$, and $\epsilon_{i j}$ is the unique environmental effect for individual $j$ in family $i$. Assume that $G_{i j}, c_{i}$, and $\epsilon_{i j}$ are independent with $c_{i} \sim N\left(0, \sigma_{c}^{2}=0.02\right)$, and $\epsilon \sim N\left(0, \sigma_{e}^{2}=0.09\right)$
(a) Calculate the expected value and variance of the phenotype $Y$ in the population.

$$
\begin{aligned}
E[Y] & =E[1.5+0.4 G+c+\epsilon] \\
& =1.5+0.4 E[G]+E[c]+E[\epsilon] \\
& =1.5+0.4 E[G]+0+0 \\
& =1.5+0.4 \mu_{G} \\
& =1.5+0.4\left(0.15^{2}(2)+2(0.15)(0.85)(1)+0.85^{2}(0)\right) \\
& =1.5+0.4 * 0.3 \\
& =1.62
\end{aligned}
$$

$$
\begin{aligned}
\operatorname{Var}[Y] & =\operatorname{Var}[1.5+0.4 G+c+\epsilon] \\
& =0+0.4^{2} \operatorname{Var}[G]+\operatorname{Var}[c]+\operatorname{Var}[\epsilon] \\
& =0.4^{2} \operatorname{Var}[G]+0.02+0.09 \\
& =0.4^{2}\left(E\left[G^{2}\right]-\mu_{G}^{2}\right)+0.02+0.09 \\
& =0.4^{2}\left(\left(0.15^{2}\left(2^{2}\right)+2(0.15)(0.85)\left(1^{2}\right)+0.85^{2}\left(0^{2}\right)\right)-\mu_{G}^{2}\right)+0.02+0.09 \\
& =0.4^{2}\left(0.345-0.3^{2}\right)+0.02+0.09 \\
& =0.4^{2}(0.255)+0.02+0.09 \\
& =0.151
\end{aligned}
$$

(b) What is the broad sense heritability $\left(H^{2}\right)$ of the phenotype in the population? What is the narrow sense heritability $\left(h^{2}\right)$ of this phenotype in the population?
$\operatorname{Var}[\mathrm{G}]=0.255, \operatorname{Var}[\mathrm{~A}]=0.041, \operatorname{Var}[\mathrm{C}]=0.02, \operatorname{Var}[\epsilon]=0.09, \operatorname{Var}[\mathrm{D}]=0$.

$$
\begin{aligned}
H^{2} & =\frac{\beta_{1}^{2} \sigma_{G}^{2}}{\sigma_{Y}^{2}}=\frac{0.041}{0.151}=0.271 \\
h^{2} & =\frac{\sigma_{A}^{2}}{\sigma_{Y}^{2}}=0.271
\end{aligned}
$$

(c) Consider two individuals $j$ and $k$ from family i. Calculate $\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right)$ (i.e., the covariance of $Y_{i j}$ and $Y_{i k}$ ) for the following relationship types for $j$ and $k$ :

- $j$ and $k$ are mono-zygotic (MZ) twins reared in the same household.
- $j$ and $k$ are full-siblings reared in the same household.
- $j$ and $k$ are half-siblings reared in the same household.
- $j$ and $k$ are half-siblings reared in different households.
- $j$ and $k$ are first-cousins reared in different households.
- $j$ and $k$ are unrelated but reared in the same household (e.g., one of the individuals is adopted).
First,

$$
\begin{aligned}
\sigma_{D}^{2} & =0 \\
\sigma_{e}^{2} & =0.09 \\
\sigma_{A}^{2} & =0.041 \\
\sigma_{C}^{2} & =0.02
\end{aligned}
$$

- j and k are mono-zygotic (MZ) twins reared in the same household.

$$
\begin{aligned}
\operatorname{Cov}(C) & =\sigma_{C}^{2}\left[\begin{array}{ll}
1 & 1 \\
1 & 1
\end{array}\right]=0.02\left[\begin{array}{ll}
1 & 1 \\
1 & 1
\end{array}\right] \\
\operatorname{Cov}(A) & =\sigma_{A}^{2}\left[\begin{array}{cc}
1 & 2 \theta_{j k} \\
2 \theta_{j k} & 1
\end{array}\right]=\sigma_{A}^{2}\left[\begin{array}{cc}
1 & 2(1 / 2) \\
2(1 / 2) & 1
\end{array}\right] \\
\operatorname{Cov}(\epsilon) & =\sigma_{e}^{2}\left[\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right]=0.09\left[\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right] \\
\operatorname{Cov}(D) & =\sigma_{D}^{2}\left[\begin{array}{cc}
1 & \Delta_{7}^{j k} \\
\Delta_{7}^{j k} & 1
\end{array}\right]=\sigma_{D}^{2}\left[\begin{array}{ll}
1 & 1 \\
1 & 1
\end{array}\right] \\
\Omega & =2 \Theta \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}+\sigma_{c}^{2} \Phi_{C}+\sigma_{E}^{2} I \\
& =\sigma_{A}^{2}\left[\begin{array}{cc}
1 & 2(1 / 2) \\
2(1 / 2) & 1
\end{array}\right]+\sigma_{D}^{2}\left[\begin{array}{ll}
1 & 1 \\
1 & 1
\end{array}\right]+0.02\left[\begin{array}{ll}
1 & 1 \\
1 & 1
\end{array}\right]+0.09\left[\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right] \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =\sigma_{A}^{2}+\sigma_{D}^{2}+\sigma_{C}^{2} \\
& =0.0408+0+0.02 \\
& =0.0608
\end{aligned}
$$

- j and k are full-siblings reared in the same household.

$$
\begin{aligned}
\Omega^{j k} & =2 \Theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+\sigma_{E}^{2} I \\
\omega^{j k} & =2 \theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+0 \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =2\left(\frac{1}{4}\right) \sigma_{A}^{2}+\sigma_{D}^{2}\left(\frac{1}{4}\right)+\sigma_{c}^{2} \\
& =\frac{1}{2} \sigma_{A}^{2}+\frac{1}{4} \sigma_{D}^{2}+\sigma_{c}^{2} \\
& =0.0204+0+0.02 \\
& =0.0404
\end{aligned}
$$

- j and k are half-siblings reared in the same household.

$$
\begin{aligned}
\Omega^{j k} & =2 \Theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+\sigma_{E}^{2} I \\
\omega^{j k} & =2 \theta^{j k} \sigma_{A}^{2}+0+\sigma_{c}^{2} \Phi_{C}^{j k}+0 \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =2\left(\frac{1}{8}\right) \sigma_{A}^{2}+\sigma_{c}^{2} \\
& =\frac{1}{4} \sigma_{A}^{2}+\sigma_{c}^{2} \\
& =0.0102+0.02 \\
& =0.0302
\end{aligned}
$$

- j and k are half-siblings reared in different households.

$$
\begin{aligned}
\Omega^{j k} & =2 \Theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+\sigma_{E}^{2} I \\
\omega^{j k} & =2 \theta^{j k} \sigma_{A}^{2}+0+0+0 \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =2\left(\frac{1}{8}\right) \sigma_{A}^{2} \\
& =\frac{1}{4} \sigma_{A}^{2} \\
& =0.0102
\end{aligned}
$$

- j and k are first-cousins reared in different households.

$$
\begin{aligned}
\Omega^{j k} & =2 \Theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+\sigma_{E}^{2} I \\
\omega^{j k} & =2 \theta^{j k} \sigma_{A}^{2}+0+0+0 \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =2\left(\frac{1}{16}\right) \sigma_{A}^{2} \\
& =\frac{1}{8} \sigma_{A}^{2} \\
& =0.0051
\end{aligned}
$$

- j and k are unrelated but reared in the same household (e.g., one of the individuals is adopted).

$$
\begin{aligned}
\Omega^{j k} & =2 \Theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+\sigma_{E}^{2} I \\
\omega^{j k} & =0+0+\sigma_{c}^{2} \Phi_{C}^{j k}+0 \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =\sigma_{c}^{2} \\
& =0.02
\end{aligned}
$$

(d) Now consider two individuals $j$ and $k$ who are from two different (unrelated) families $i$ and $l$, respectively, where individual $j$ is from family $i$ and individual $k$ is from family l. Calculate $\operatorname{Cov}\left(Y_{i j}, Y_{l k}\right)$.

$$
\begin{aligned}
\Omega^{j k} & =2 \Theta^{j k} \sigma_{A}^{2}+\sigma_{D}^{2} \Delta_{7}^{j k}+\sigma_{c}^{2} \Phi_{C}^{j k}+\sigma_{E}^{2} I \\
\omega^{j k} & =0+0+0+0 \\
\operatorname{Cov}\left(Y_{i j}, Y_{i k}\right) & =0
\end{aligned}
$$

## Appendix: Qtdt output and $R$ code

```
### 1 a-c LR Test Output ###
testarossa:qtdt-2.6.1 brendaprice$ /Users/brendaprice/Documents/QTDT/qtdt-2.6.1/qtdt
-d BMI_BIOST551.dat -p BMI_BIOST551.ped -a- -weg -vegc
QTDT - Quantitative TDT 2.6.1
(c) 1998-2007 Goncalo Abecasis (goncalo@umich.edu)
This program implements tests described by
Abecasis et al, Am J Hum Genet 66:279-292 (2000)
Abecasis et al, Eur J Hum Genet 8:545-551 (2000)
and others
The following parameters are in effect:
    QTDT Data File : BMI_BIOST551.dat (-dname)
    QTDT Pedigree File : BMI_BIOST551.ped (-pname)
    QTDT IBD Status File
        Missing Value Code
                        Covariates
        Association Model
    Full Model Variances :
 POLYGENIC 
POLYGENIC
                                    NONE (-o[f|t|m|p|-])
                                    0 (-m9999)
                    123456 (-r9999)
    NELDER AND MEAD (-n[f|n|p])
    FULL PEDIGREE (-t[n|p])
Additional Options
    --dominance, --snp, --multi-allelic, --deviates, --references,
    --exclude-founder-phenotypes, --p-values, --no-regress-tbl
Online documentation http://www.sph.umich.edu/csg/abecasis/QTDT
Comments, bugs: goncalo@umich.edu
The following models will be evaluated..
    NULL MODEL
        Means = Mu
    Variances = Ve + Vg
    FULL MODEL
        Means = Mu
    Variances = Ve + Vc + Vg
Testing trait: BMI
\begin{tabular}{rlllrrr} 
Allele & \(d f(0)\) & \(-\operatorname{LnLk}(0)\) & \(d f(V)\) & \(-\operatorname{LnLk}(V)\) & \(\operatorname{ChiSq}\) & \(p\) \\
N/A & 26412 & 76322.95 & 26411 & 76240.47 & 164.96 & \(9 \mathrm{e}-38\)
\end{tabular}
Run completed on Fri Nov 28 10:17:48 2014
1 \text { tests carried out}
The most significant result refers to:
    Trait: BMI
    ChiSq: 164.956
    p-value: 9e-38
Trait: BMI
==================================================================
Total Probands: 26415
Family #1 Phenotypes - scores : 23.841 24.936
```

```
Family #1 var-covar matrix terms [2]...[[Ve]][[Vg]]
Family #1 regression matrix...
    [linear] =
    [2 x 1] Mu
            2.5 1.000
            2.6 1.000
Some useful information...
                df : 26412
    log(likelihood) : 76322.95
    variances : 11.895 7.806
        means : 24.844
FULL HYPOTHESIS
Family #1 var-covar matrix terms [3]...[[Ve]][[Vc]][[Vg]]
Family #1 regression matrix...
    [linear] =
    [2 x 1] Mu
            2.5 1.000
            2.6 1.000
Some useful information...
                df : 26411
    log(likelihood) : 76240.47
            variances : 14.378 3.304 2.220
            means : 24.841
```

```
####### R code #2
sigc <- 0.02
sige <- 0.09
siga <- 2*(0.15)*(0.85)*0.4^2
muG = (0.15^2*(2) + 2*(0.15)*(0.85)*(1) + 0.85^2*(0))
Ey = 1.5 +0.4* muG
E2G <- (0.15^2*(2^2) + 2*(0.15)*(0.85)*(1^2) + 0.85^2*(0^2) )
VarG <- (E2G-muG^2)
VarY= 0.4^2 *(E2G-muG^2)+ 0.02 + 0.09
#VarY <- siga/(0.8)
H2 <- 0.16*VarG/VarY
sigd <- 0
#sigd <- VarG - siga
h2 <-siga/VarY
```

