Introduction to Super Learning

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Learning Goals

• Conceptual understanding of Super Learning (SL)

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- Conceptual understanding of Super Learning (SL)
- Comfort with the SuperLearner R package

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- Conceptual understanding of Super Learning (SL)
- Comfort with the SuperLearner R package
- Awareness of the mathematical backbone of SL

Outline

- I. Motivation and description of SL (30 minutes)
- II. Lab 1: Vanilla SL for a continuous outcome (30 minutes)
- III. Mathematical presentation of SL (20 minutes)
- IV. Lab 2: Vanilla SL for a binary outcome (30 minutes)

15 minute break

Outline

15 minute break

- V. Bells and whistles: Screens, weights, and CV-SL (30 minutes)
- VI. Lab 3: Binary outcome redux (40 minutes)
- VII. Lab 4: Case-control analysis of Fluzone vaccine (30 minutes)

I. Motivation and description of Super Learning

Notation

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- X is a *p*-variate set of predictors
- We observe *n* independent copies

 $(Y_1, \mathbf{X}_1), \ldots, (Y_n, \mathbf{X}_n)$

from the joint distribution of (Y, \mathbf{X}) .

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- Super Learning can be applied in all of the above settings
- We will focus on estimating the regression function

 $\mu(\mathbf{x}) := E[Y \mid \mathbf{X} = \mathbf{x}].$

1. Exploratory analysis

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- 6. Confirmatory analysis/hypothesis testing (not our goal here)











How do we choose which algorithm to use?

Super Learning is:

An **ensemble method** for combining predictions from many candidate machine learning algorithms

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If we knew MSE(µ̂k), we could choose the µ̂k with the smallest MSE(µ̂k).

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- This estimator will favor
 *µ*_k which are overfit, because
 *µ*_k
 are trained on the same data used to evaluate the MSE.
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- Instead, we estimate MSE using cross-validation.

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1	1	1	1	1	1	1	1	1	1
2	2	2	2	2	2	2	2	2	2
3	3	3	3	3	3	3	3	3	3
4	4	4	4	4	4	4	4	4	4
5	5	5	5	5	5	5	5	5	5
6	6	6	6	6	6	6	6	6	6
7	7	7	7	7	7	7	7	7	7
8	8	8	8	8	8	8	8	8	8
9	9	9	9	9	9	9	9	9	9
10	10	10	10	10	10	10	10	10	10

Fold 1 Fold 2 Fold 3 Fold 4 Fold 5 Fold 6 Fold 7 Fold 8

Schematic of 10-fold cross-validation. Gray: training sets. Yellow: validation sets.

Fold 9 Fold 10

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- 3. Our cross-validated MSE is

$$\widehat{MSE}_{CV}(\hat{\mu}_k) = \frac{1}{V} \sum_{\nu=1}^{V} \frac{1}{|\mathcal{V}_{\nu}|} \sum_{i \in \mathcal{V}_{\nu}} [Y_i - \hat{\mu}_{k,\nu}(\mathbf{X}_i)]^2.$$

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We average the MSEs of the *V* validation sets.

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2	2	2	2	2	2	2	2	2	2	2
3	3	3	3	3	3	3	3	3	3	3
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CV preds.

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(People typically use V = 5 or V = 10.)

"Discrete" Super Learner

· At this point, we have cross-validated MSE estimates

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- We call this the "discrete Super Learner".

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- Super Learner considers as its set of candidate algorithms all convex combinations μ̂_λ := Σ^K_{k=1} λ_kμ̂_k.
- The Super Learner is $\hat{\mu}_{\widehat{\lambda}}$, where

$$\widehat{\boldsymbol{\lambda}} := \argmin_{\boldsymbol{\lambda} \in \mathcal{S}_{\mathcal{K}}} \widehat{MSE}_{CV} \left(\sum_{k=1}^{\mathcal{K}} \lambda_k \widehat{\mu}_k \right)$$

(We use constrained optimization to compute the argmin.)

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Super Learner: steps

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4. Take $\hat{\mu}_{SL} = \sum_{k=1}^{K} \hat{\lambda}_k \hat{\mu}_k$.

II. Lab 1: Vanilla SL for a continuous outcome

III. Into the weeds: a mathematical presentation of SL

Review

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In this section, we generalize this procedure to estimation of any summary of the observed data distribution given an appropriate loss for the summary of interest.

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- Suppose we want to estimate a parameter $\theta : \mathcal{M} \to \Theta$.
- Denote $\theta_0 := \theta(P_0)$ the true parameter value.

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- $R_0(\theta) = E_{P_0}[L(\mathbf{O}, \theta)]$ is called the **oracle risk**.
- These definitions of loss and risk come from the statistical learning literature (see, e.g. Vapnik, 1992, 1999, 2013) and are not to be confused with loss and risk from the decision theory literature (e.g. Ferguson, 2014).

MSE is the oracle risk corresponding to a squared-error loss function

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- $L(\mathbf{O}, \mu) = [Y \mu(\mathbf{X})]^2$ is the squared-error loss.
- $R_0(\mu) = MSE(\mu) = E_{P_0}[Y \mu(\mathbf{X})]^2$.

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- The naive estimator is $\widehat{R}(\widehat{\theta}_k) = \frac{1}{n} \sum_{i=1}^{n} L(\mathbf{O}_i, \widehat{\theta}_k).$
- We instead estimate $R_0(\theta)$ using the cross-validated risk

$$\widehat{R}_{CV}(\widehat{\theta}_k) = \frac{1}{V} \sum_{\nu=1}^{V} \frac{1}{|\mathcal{V}_{\nu}|} \sum_{i \in \mathcal{V}_{\nu}} L(\mathbf{O}_i, \widehat{\theta}_{k,\nu}).$$

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4. Take $\hat{\theta}_{SL} = \sum_{k=1}^{K} \hat{\lambda}_k \hat{\theta}_k$.

Theoretical guarantees

van der Vaart et al. (2006) showed that, under some conditions, the **oracle risk of the SL estimator** is **as good** as the **oracle risk of the oracle minimizer** up to a multiple of $\frac{\log n}{n}$ as long as the number of candidate algorithms is **polynomial in** *n*.

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- For **binary** *Y*, squared-error loss is still valid.

Loss functions for a binary outcome

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- AUC loss.

IV. Lab 2: Vanilla SL for a binary outcome

15 minute break

V. Bells and whistles: Screens, weights, and CV-SL



In this section, we will introduce three of the add-ons to SL that are frequently useful in practice: variable screens, observation weights, and cross-validated SL.

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Screening algorithms allow us to guide the SL using our domain knowledge.

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- We can force certain variables to always be used.

Observation weights

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- In some applications, we need to include observation weights in the procedure – e.g. case-control sampling, or as a simple way to account for loss-to-followup.
- Observation weights can be included directly in a call to SuperLearner, but method.AUC does not make correct use of weights!!!!
- Note that some SuperLearner wrappers might not make use of observation weights.

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- Suppose specimens from all n_{case} cases (Y_i = 1) are assayed.
- A random subset of $N_{control}$ controls ($Y_i = 0$) (out of $n_{control}$ total controls) are assayed.
- We will use this case-control cohort to predict disease status using the results of the assay and other covariates.

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- **Cases** have weight $w_i = 1$.
- **Controls** have weight $w_i = n_{control}/N_{control}$.
- Control weights could also be estimated using a logistic regression of the indicator of inclusion in the control cohort on baseline covariates.

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- *T* is subject to right-censoring by *C*: we observe
 Y = min{*T*, *C*} and Δ = *I*(*T* ≤ *C*).
- We want to estimate

$$\mu(\mathbf{x}) = P(T \leq t_0 \mid \mathbf{X} = \mathbf{x}) = E[Y \mid \mathbf{X} = \mathbf{x}].$$

$$\mu_{0} = \arg\min_{\mu} E_{P_{0}} \left\{ \frac{\Delta}{G_{0}(Y \mid \mathbf{X})} L((Y, \mathbf{X}), \mu) \right\}$$

- Here, $G_0(t \mid \mathbf{x}) = P_0(C > t \mid \mathbf{X} = \mathbf{x}).$
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- Instead, we estimate G₀ and plug in this estimator to obtain an estimated weight.
- If C ⊥ T, we can use a Kaplan-Meier estimator for G₀;
 otherwise we might use a Cox model.

CV-Super Learner

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 - a. Run regular SL on the training set for fold v using V_2 -fold CV.
 - b. Obtain discrete SL and SL predictions for the validation set for fold *v*.
- Combine the validation sets to obtain CV-risks for the discrete SL and SL.

VI. Lab 3: Binary outcome redux

VII. Lab 4: Case-control analysis of Fluzone vaccine

• Health adults aged 18–49 years, Michigan, 2007–2008.

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- Randomly assigned to:
 - Fluzone inactivated influenza vaccine (IIV)
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 - placebo.
- We are only interested in Fluzone vs placebo.
- Followed for one flu season.
- Endpoint = laboratory-confirmed influenza.

Treatment	Group	No.
Placebo	Total	325
	Cases	30
	Controls	295
LAIV	Total	814
	Cases	54
	Controls	760
IIV	Total	813
	Cases	22
	Controls	791

 All 52 cases and 52 random controls were assayed for a variety of markers (HAI, NAI, MN, AM titers, proteins/virus/peptide magnitude/breadth).

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 - Day 30 markers

- All 52 cases and 52 random controls were assayed for a variety of markers (HAI, NAI, MN, AM titers, proteins/virus/peptide magnitude/breadth).
- Measured variables:
 - Demographics: age, vaccinated in last year (EVERVAX)
 - Day 0 markers
 - Day 30 markers
 - Difference markers = Day 30 markers Day 0 markers

1. Demo.

- 1. Demo.
- 2. Demo. + Day 0 markers

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- 2. Demo. + Day 0 markers
- 3. Demo. + Day 30 markers

- 1. Demo.
- 2. Demo. + Day 0 markers
- 3. Demo. + Day 30 markers
- 4. Demo. + Difference markers

- 1. Demo.
- 2. Demo. + Day 0 markers
- 3. Demo. + Day 30 markers
- 4. Demo. + Difference markers
- 5. Demo. + Day 0 markers + EVERVAX \times Day 0 markers

- 1. Demo.
- 2. Demo. + Day 0 markers
- 3. Demo. + Day 30 markers
- 4. Demo. + Difference markers
- 5. Demo. + Day 0 markers + EVERVAX \times Day 0 markers
- 6. Demo. + Day 30 markers + EVERVAX \times Day 30 markers

- 1. Demo.
- 2. Demo. + Day 0 markers
- 3. Demo. + Day 30 markers
- 4. Demo. + Difference markers
- 5. Demo. + Day 0 markers + EVERVAX \times Day 0 markers
- 6. Demo. + Day 30 markers + EVERVAX \times Day 30 markers
- 7. Demo. + Diff. markers + EVERVAX \times Diff. markers

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- 2. Demo. + Day 0 markers
- 3. Demo. + Day 30 markers
- 4. Demo. + Difference markers
- 5. Demo. + Day 0 markers + EVERVAX \times Day 0 markers
- 6. Demo. + Day 30 markers + EVERVAX \times Day 30 markers
- 7. Demo. + Diff. markers + EVERVAX \times Diff. markers
- 8. Demo. + Day 0 + Day 30 + EVERVAX × (Day 0 + Day 30)

- 1. Demo.
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- 3. Demo. + Day 30 markers
- 4. Demo. + Difference markers
- 5. Demo. + Day 0 markers + EVERVAX \times Day 0 markers
- 6. Demo. + Day 30 markers + EVERVAX \times Day 30 markers
- 7. Demo. + Diff. markers + EVERVAX \times Diff. markers
- 8. Demo. + Day 0 + Day 30 + EVERVAX \times (Day 0 + Day 30)
- 9. Demo. + Day 0 + Diff. + EVERVAX \times (Day 0 + Diff.)

Analysis goals

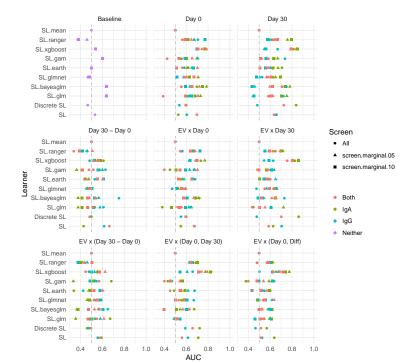
 We want to compare the quality of these nine sets of variables for predicting flu status in the placebo and Fluzone arms separately.

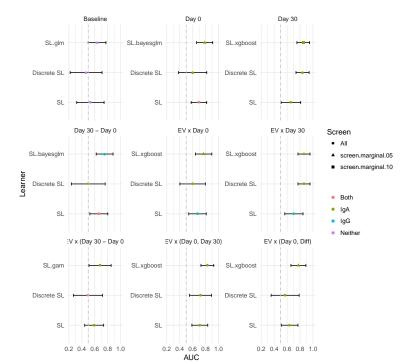
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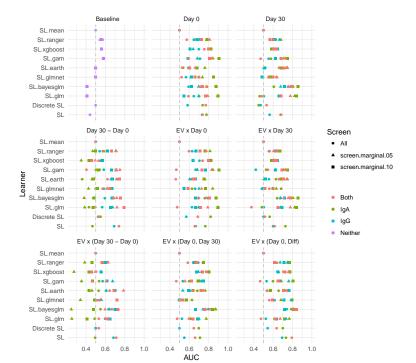
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- We also want to compare the predictive quality of IgA, IgG, and both IgA + IgG measurements.

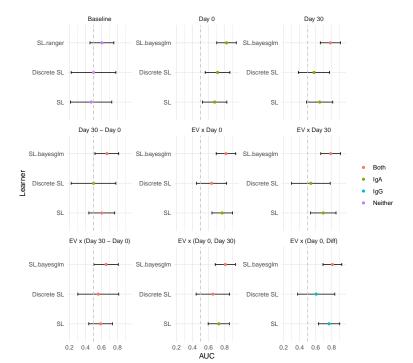
Analysis goals

- We want to compare the quality of these nine sets of variables for predicting flu status in the placebo and Fluzone arms separately.
- We also want to compare the predictive quality of IgA, IgG, and both IgA + IgG measurements.
- We will use cross-validated Super Learning to do this.









Ferguson, T. S. (2014). *Mathematical statistics: A decision theoretic approach*. Academic Press.

van der Vaart, A. W., Dudoit, S., and van der Laan, M. J. (2006). Oracle inequalities for multi-fold cross validation. *Statistics & Decisions*, 24(3):351–371.

 Vapnik, V. (1992). Principles of risk minimization for learning theory. In Advances in Neural Information Processing Systems, pages 831–838.
 Vapnik, V. (2013). The nature of statistical learning theory. Springer Science

& Business Media.

Vapnik, V. N. (1999). An overview of statistical learning theory. *IEEE Transactions on Neural Networks*, 10(5):988–999.