Special Exercise

For SNPs that have no true association with the outcome, the p-values for the association test should have a uniform distribution between 0 and 1. In a genome-wide association study, while we hope for some true associations, nearly all the SNPs will not have any association with the outcome, so nearly all the p-values should come from a uniform distribution. When a large fraction of p-values depart from the expected uniform distribution it usually indicates a problem with the data, either poor data quality or confounding by population substructure.

A standard quality-control measure for GWAS analyses is to compare the p-values to a uniform distribution with a quantile-quantile plot. We plot \(-\log_{10}p\), sorted from smallest to largest, against the expected distribution of \(-\log_{10}p\) if the p-values had a uniform distribution. In R

```R
qqplot(-log10(ppoints(length(pvalues))), -log10(pvalues))
```

should lie, for most of the SNPs, along the diagonal line given by

```R
abline(0,1)
```

A numerical summary of the departure from the uniform distribution is the so-called ‘genomic control coefficient’ \(\lambda\). If \(\beta\) and \(se\) are the coefficient estimates and standard errors, respectively, then

```R
lambda <- median((beta/se)^2) / 0.4549
```

measures the departure of the median p-value from its expected position. We expect to see \(\lambda\) close to 1, and this is typically the case in moderate-sized GWAS analyses.

However, \(\lambda\) behaves differently in SNP x environment interaction analyses, and spuriously high values of \(\lambda\) can appear as the usual standard error calculations fail to take account of the randomly-changing observed distribution of the environmental variable. [Voorman et al, PLoS One, May 2011]

You will demonstrate this phenomenon.
To generate a single data set of size 1000, simulate

```r
environ <- rpois(1000, 10)
outcome <- environ*(environ+rnorm(1000))
```

and a SNP variable with minor allele frequency 0.3

```r
gene <- rbinom(1000, 2, 0.3)
```

The linear model for the main effect of SNP is

```r
maineffect <- lm(outcome~gene)
```

and to extract the main effect's coefficient estimate (\texttt{beta}), its estimated standard errors (\texttt{se}) and corresponding p-value, look at the final row of

```r
coef(summary(maineffect))
```

For the interaction between SNP and environment, we similarly fit

```r
interact <- lm(outcome~gene*environ)
```

and extract what we need from the last row of `coef(summary(interact))`

1. In a GWAS study there would be a single environmental variable, a single outcome phenotype variable, and many SNPs. For this exercise we will assume 5000 independent SNPs, and
(a) estimate the main effect of each SNP with the outcome, draw the quantile-quantile plot, and compute lambda.
(b) estimate the interaction between each SNP and environment, draw the quantile-quantile plot, and compute lambda

2. Simulate a reasonable number of repetitions of this study. You will not be able to draw hundreds of quantile-quantile plots, but you will be able to compute \texttt{lambda} for the main effects and interaction analyses. Summarize the results.

3. For enthusiastic people who know more about regression modeling, you might try one of the following extra exercises:
   (a) show that if the relationship between outcome and environmental variable is simulated to be exactly linear the spurious data-quality problem goes away
   (b) use the model-robust standard errors provided by the \texttt{vcovHC()} function in the \texttt{sandwich} package instead of the model-based standard errors provided by \texttt{lm()}, and show that the problem goes away.