

Statistics 583, Problem Set 9

Wellner; 5/25/2016

Reading: Chapter 8, sections 8.1- 8.4;

van der Vaart, Asymptotic Statistics, chapter 23, pages 326 - 340;

Wasserman, Chapters 2-3, pages 13-41.

Due: Wednesday, June 1, 2016

1. Suppose that $T(F) = Var_F(X)$ so that $T_n \equiv T(\mathbb{F}_n) = n^{-1} \sum_{i=1}^n (X_i - \bar{X})^2$. Show that the jackknife estimate of the variance $\sigma_n^2(F) \equiv Var_F(T_n)$ is

$$\widehat{Var} = \frac{n^2}{(n-1)^3} (\widehat{\mu}_4 - \widehat{\mu}_2^2)$$

where $\widehat{\mu}_k \equiv n^{-1} \sum_{i=1}^n (X_i - \bar{X})^k$ for $k = 1, 2, \dots$. Hence, assuming that $EX^4 < \infty$, the jackknife estimate of variance is consistent for this T :

$$n\widehat{Var} \rightarrow_p \mu_4 - \mu_2^2 = \mu_2^2 \left\{ 2 + \frac{\mu_4}{\mu_2^2} - 3 \right\} = T_2(F)(2 + \gamma_2).$$

2. (a) Wasserman, problem 3.8.9, page 40: Let X_1, \dots, X_n be n distinct observations (no ties). Let X_1^*, \dots, X_n^* denote a bootstrap sample (from the empirical d.f. \mathbb{F}_n of the X_i 's), and let $\bar{X}_n^* = n^{-1} \sum_{i=1}^n X_i^*$. Find: $E\{\bar{X}_n^* | X_1, \dots, X_n\}$, $Var(\bar{X}_n^* | X_1, \dots, X_n)$, and $Var(\bar{X}_n^*)$.
 (b) Wasserman, problem 3.8.13, page 41: Let X_1, \dots, X_n be n distinct observations (no ties). Let X_1^*, \dots, X_n^* denote a bootstrap sample (from the empirical d.f. \mathbb{F}_n of the X_i 's). Let G denote the marginal distribution of X_i^* . Note that $G(x) = P(X_i^* \leq x) = E\{P(X_i^* \leq x | X_1, \dots, X_n)\} = E\{\mathbb{F}_n(x)\} = F(x)$. So it appears that X_i^* and X_i have the same distribution. But in (a) we showed that $Var(\bar{X}_n) \neq Var(\bar{X}_n^*)$. Explain.
3. Consider a V -functional of order $r = 2$ given by $T(P) = \int \int h(x, y) dP(x) dP(y)$ where h is permutation symmetric. Find the jackknife estimate of bias for the (V -statistic) estimator $T(\mathbb{P}_n)$ of $T(P)$. Also find the jackknife estimator of $T(P)$.
4. Wasserman, problem 3.8.11, page 41: Let X_1, \dots, X_n be i.i.d. $Uniform(0, \theta)$. The MLE of θ is $\widehat{\theta}_n \equiv X_{(n)} = \max\{X_1, \dots, X_n\}$.
 (a) Find the distribution of $\widehat{\theta}_n$ and the exact and limiting distribution of $n(\theta - \widehat{\theta}_n)$.
 (b) Compare the true and limiting distribution of $n(\theta - \widehat{\theta}_n)$ with the parametric and nonparametric bootstrap distributions when $\theta = 1$.
 (c) Show that for the parametric bootstrap $P(\widehat{\theta}_n^* = \widehat{\theta}_n) = 0$ but for the nonparametric bootstrap $P(\widehat{\theta}_n^* = \widehat{\theta}_n) = 1 - (1 - 1/n)^n \rightarrow 1 - e^{-1} \approx .632 \dots$

5. **Optional bonus problem 1:** (Bootstrapping a linear regression model a simple way.) Consider bootstrapping a linear regression model

$$Y_i = \mathbf{x}_i^T \beta + \epsilon_i, \quad i = 1, \dots, n$$

where the ϵ_i are i.i.d. mean 0, finite variance, and the \mathbf{x}_i are given p -dimensional vectors, such that there is no constant term in the regression.

(a) Show that the estimated residuals $\hat{\epsilon}^T = (\hat{\epsilon}_1, \dots, \hat{\epsilon}_n)$ satisfy $\hat{\epsilon} - \epsilon = -H\epsilon$ where $H = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ is the “hat matrix” (i.e. the projection matrix onto the column space of \mathbf{X}).

(b) Suppose that $\hat{\epsilon}_1^*, \dots, \hat{\epsilon}_n^*$ is a bootstrap sample (with replacement) from $\{\hat{\epsilon}_1, \dots, \hat{\epsilon}_n\}$. Show that

$$E_*(n^{1/2}(\hat{\beta}^* - \hat{\beta})) = \left(\frac{1}{n} \mathbf{X}^T \mathbf{X}\right)^{-1} \left(\frac{1}{n} \mathbf{X}^T \mathbf{1}\right) Z_n$$

where $Z_n = n^{-1/2} \sum_{i=1}^n \hat{\epsilon}_i$.

(c) Show that if $\max_{1 \leq i \leq n} h_{ii} \rightarrow 0$, and $n^{-1} \mathbf{X}^T \mathbf{X} \rightarrow V$, a positive definite matrix, then

$$\sqrt{n}(\hat{\beta} - \beta) \rightarrow_d N_p(0, \sigma^2 V^{-1})$$

[This is a variant of the result we established in 581 via the Lindeberg - Feller CLT.]

(d) Find the mean and variance of Z_n .

(e) Suppose that:

(i) $n^{-1} \mathbf{X}^T \mathbf{X} \rightarrow V$, a positive definite matrix;

(ii) $\mathbf{X}^T \mathbf{1}/n \rightarrow \mathbf{h}$ with $\mathbf{h}^T V^{-1} \mathbf{h} < 1$;

(iii) $\max_{1 \leq i \leq n} h_{ii} \rightarrow 0$ where h_{ii} are the diagonal elements of the hat matrix H .

Show that if (i) - (iii) hold, then the bootstrap fails in the sense that the random variable Z_n in (b) converges in distribution to a proper random variable rather than to zero.

Hint: show that (iii) implies that $\max_{1 \leq i \leq n} |c_{ni}| \rightarrow 0$ where $\mathbf{c} = n^{-1/2}(I - H)\mathbf{1}$.

6. **Optional bonus problem 2:** Suppose now that the bootstrap residuals are drawn from the collection of *centered* residuals $\hat{\epsilon} - \mathbf{1}(\mathbf{1}^T \hat{\epsilon}/n)$. Compute $E_*(\sqrt{n}(\hat{\beta}^* - \hat{\beta}))$ and $E_*(\sqrt{n}(\hat{\beta}^* - \hat{\beta}))^{\otimes 2}$ for this bootstrap resampling scheme.