

Statistics 583: Final Exam

Wellner; 6/5/2000

Instructions: This exam is to be done with closed notes and closed books.

1. (30 points). **Define** the following terms, providing an appropriate context for your definition:
 - A. A continuous functional $T(F)$ with respect to a metric d on distributions.
 - B. A Fréchet differentiable functional $T(F)$ with respect to a metric d .
 - C. The cumulative hazard function corresponding to a distribution function F .
 - D. A maximal invariant with respect to a group G .
 - E. An invariant test with respect to a group G .
 - F. A metric d between distribution functions which is “compatible” with respect to the empirical measure.

2. (27 points). **State** any *three* of the following four theorems, providing an appropriate context in each case:
 - A. The Neyman-Pearson lemma.
 - B. A Theorem about UMP G -invariant tests in the presence of monotone likelihood ratio of the G -maximal invariant.
 - C. A theorem about asymptotic normality of the natural empirical estimator $T_n = T(\mathbb{F}_n)$ of a Fréchet - differentiable functional $T(F)$.
 - D. Some version of *Hoeffding's formula* for the distribution of the ranks under alternatives.

3. (40 points). A simple test for asymmetry of a distribution function F is based on the difference of the mean and median: $T(F) = \int x dF(x) - F^{-1}(1/2)$. Note that $T(F) = 0$ if F is symmetric about some point, while $T(F)$ is positive for F skewed to the right, and negative for F skewed to the left. Suppose that X_1, \dots, X_n are i.i.d. with distribution function F , and let \mathbb{F}_n be the empirical distribution of the sample, $\mathbb{F}_n(x) = n^{-1} \sum_{i=1}^n 1_{(-\infty, x]}(X_i)$.
 - A. Show that T is invariant under location shifts: If F_θ is defined by $F_\theta(x) = F(x - \theta)$, then $T(F_\theta) = T(F)$.
 - B. Under appropriate assumptions on F (make these explicit), find the influence function of $T(F)$. [Hint: we already essentially did this in class, since we calculated the influence functions of $F^{-1}(1/2)$ and $\int x dF(x)$.]
 - C. Under appropriate assumptions on F (make these explicit), state a limit theorem for $\sqrt{n}(T(\mathbb{F}_n) - T(F))$ and compute the asymptotic variance $V^2(F)$.
 - D. Consider the bootstrap and jackknife estimators of $V_n(F) \equiv Var_F(\sqrt{n}(T(\mathbb{F}_n) - T(F)))$. Would either or both of these “work” for estimation of $V^2(F)$? Explain why or why not.

4. (40 points). Consider testing $H : F = G$ versus $K : (1 - G) = (1 - F)^\theta$ where $\theta > 1$ based on X_1, \dots, X_m i.i.d. F and Y_1, \dots, Y_n i.i.d. G . We want a test which is invariant with respect to monotone transformations of the data.
- A. Draw a picture to show the relationship of F and G under the alternative hypothesis K . What is the relationship of the corresponding cumulative hazard functions under K ? Draw a picture showing this.
- B. Give an appropriate version of Hoeffding's formula for the distribution of the order Y -ranks $Q_1 < \dots < Q_n$ under the alternative K .
- C. Use the formula in B (or any other way) to compute the complete distribution of \underline{Q} when $m = n = 2$ (so $N = m + n = 4$) and $\theta = 2$. For testing H versus $\theta = 2$, what observed value of \underline{Q} would lead you to reject H at size $\alpha = 1/6$?
- D. What statistic would you use for testing H versus K ? In the case $m = n = 2$ in part C, what are its possible values? What observed value of \underline{Q} leads you to reject H at size $\alpha = 1/6$?
5. (30 points). Suppose that X_{ij} for $j = 1, \dots, n_i$, $i = 1, \dots, I$ are independent, normally distributed random variables with common variance σ^2 , and suppose that $EX_{ij} = \theta_i$. Thus X_{i1}, \dots, X_{i,n_i} is a sample of size n_i from the $N(\theta_i, \sigma^2)$ distribution. Consider testing $H : \theta_1 = \dots = \theta_I$ versus $K : \theta_i \neq \theta_j$ for some $i \neq j$.
- A. Explain briefly the canonical form of this testing problem, and the groups which leave the testing problem invariant.
- B. Find the UMP-invariant test of H versus K , and specify its distribution under the null hypothesis H .
- C. What is the distribution of the test statistic you found in B under the general hypothesis K ?
6. (40 points). Suppose that X_1, \dots, X_m are i.i.d. exponential(λ), and that Y_1, \dots, Y_n are i.i.d. exponential(μ); thus the density of X_1 is $\lambda \exp(-\lambda x) 1_{[0, \infty)}(x)$. Consider testing $H : \lambda \leq \mu$ versus $K : \lambda > \mu$.
- A. Show that this testing problem is invariant with respect to the group of scale changes, G given by $g_c(\underline{x}, \underline{y}) = (c\underline{x}, c\underline{y})$ where $c > 0$.
- B. Find the UMP G -invariant test of H versus K . [Hint: you may use the fact that the family of distributions $\{\delta^{-1}F_{r,s} : \delta > 0\}$ has monotone decreasing likelihood ratio.]
- C. Specify exactly how you would carry out the test derived in B.
- D. Find a UMP - unbiased test of H versus K . Is this the same test as in part B or a different test?