

**STATISTICS 581:**  
**Day 1 Quiz Solutions, Fall, 2010**

1. **State** the (classical) Central Limit Theorem.

**Solution:** If  $X_1, \dots, X_n$  are i.i.d. with  $E(X_1) = \mu$  and  $Var(X_1) = \sigma^2 < \infty$ , then

$$\sqrt{n}(\bar{X}_n - \mu) \rightarrow_d Y \sim N(0, \sigma^2).$$

That is,

$$P(\sqrt{n}(\bar{X}_n - \mu) \leq x) \rightarrow P(Y \leq x) = P(\sigma Z \leq x) = \Phi(x/\sigma)$$

where  $Z \sim N(0, 1)$  and  $P(Z \leq z) \equiv \Phi(z) = \int_{-\infty}^z (2\pi)^{-1/2} \exp(-t^2/2) dt$

2. **State** the weak law of large numbers.

**Solution:** If  $X_1, \dots, X_n$  are i.i.d. with  $E|X_1| < \infty$ , then

$$\bar{X}_n = n^{-1} \sum_{i=1}^n X_i \rightarrow_p \mu \equiv E(X_1).$$

3. Suppose that  $X$  has density  $f(x) = 2x^{-3}1_{[1, \infty)}(x)$ .

(a) For what values of  $r \in \mathbb{R}$  is it true that  $E(X^r) < \infty$ ?

(b) For the values of  $r$  for which the expectation is finite, compute it explicitly.

(c) If  $X_1, X_2, \dots$  are i.i.d. with the same density as  $X$ , does the law of large numbers hold?

(d) If  $X_1, X_2, \dots$  are i.i.d. with the same density as  $X$ , does the central limit theorem hold?

**Solution:** (a) We compute

$$\begin{aligned} E(X^r) &= \int_1^\infty x^r f(x) dx = \int_1^\infty 2x^{r-3} dx = \frac{2}{r-2} x^{r-2} \Big|_1^\infty \\ &= \frac{2}{2-r} \text{ if } r < 2. \end{aligned}$$

Thus  $E(X^r) < \infty$  if  $r < 2$ .

(b) The computation in (a) gave  $E(X^r) = 2/(2-r)$ .

(c) The WLLN holds since  $E|X| = E(X) = 2 < \infty$ .

(d) The central limit theorem fails since  $E(X^2) = \infty$ .

4. **Define** what is meant by:

- (a)  $X_n$  converges in distribution to  $X$  for random variables  $X$  and  $X_n, n \geq 1$ .
- (b)  $X_n$  converges in probability to  $X$ .

**Solution:** (a)  $X_n \rightarrow_d X$  if

$$F_{X_n}(x) \equiv P(X_n \leq x) \rightarrow P(X \leq x) \equiv F_X(x)$$

for all  $x \in C_{F_X}$ , the set of continuity points of  $F_X$ .

(b)  $X_n \rightarrow_p X$  if

$$P(|X_n - X| > \epsilon) \rightarrow 0 \text{ as } n \rightarrow \infty$$

for every  $\epsilon > 0$ .

5. Suppose that  $U$  is a random variable with a Uniform(0,1) distribution. For each integer  $n \geq 2$  define  $X_n = n1_{[0,1/n]}(U)$ .

(a) Does  $X_n \rightarrow_d X$ ? (If so, identify the distribution of the limiting variable  $X$ .)

(b) Does  $X_n \rightarrow_p X$ ? (If so, identify the limit variable  $X$ .)

(c) Compute  $E(X_n)$ . Does it converge to  $E(X)$  for some  $X$ ?

**Solution:** (a) & (b)  $X_n \rightarrow_p 0 \equiv X$ . To see this, let  $\epsilon > 0$ . Then

$$P(|X_n| > \epsilon) \leq P(U \leq 1/n) = 1/n \rightarrow 0.$$

Since  $X_n \rightarrow_p 0$ , we also have  $X_n \rightarrow_d 0$ .

(c)

$$\begin{aligned} E(X_n) &= E\{n1_{[0,1/n]}(U)\} = nE1_{[0,1/n]}(U) \\ &= nP(U \leq 1/n) = n(1/n) = 1. \end{aligned}$$

Thus  $E(X_n) \rightarrow 1 \neq 0 = E(X)$  where  $P(X = 0) = 1$ .

6. Suppose that  $X_1, \dots, X_n, \dots$  are independent and identically distributed Exponential( $\theta$ ) random variables (i.e.  $P(X > x) = \exp(-\theta x)$  for  $x \geq 0$ ). Let  $T_n = X_1 + \dots + X_n$ .

(a) What is the distribution of  $T_n$ ?

(b) Does  $\bar{X}_n = n^{-1}T_n \rightarrow_p$  something? If so, what is "something"?

(c) Does  $\sqrt{n}(\bar{X}_n - p) \rightarrow_d$  something? If so, what is "something"?

(d) What is the Cramér-Rao bound for unbiased estimators of  $p$ ?

**Solution:** (a)  $T_n \sim \text{Gamma}(n, \theta)$ .

(b) First, for any  $r > 0$

$$E(X^r) = \int_0^\infty x^r \theta \exp(-\theta x) dx = \theta^{-r} \int_0^\infty t^r e^{-t} dt = \theta^{-r} \Gamma(r+1).$$

Thus for  $r = 1$  we have  $E(X) = \theta^{-1} \Gamma(2) = \theta^{-1}$ , and for  $r = 2$  we have  $E(X^2) = \theta^{-2} \Gamma(3) = 2\theta^{-2}$ . Hence it follows that

$$\text{Var}(X) = E(X^2) - E(X)^2 = 2\theta^{-2} - (\theta^{-1})^2 = \theta^{-2}.$$

(c) From the WLLN  $\bar{X}_n \rightarrow_p E(X_1) = \theta^{-1}$ .

(d) From the CLT,

$$\sqrt{n}(\bar{X}_n - \theta^{-1}) \rightarrow_d N(0, \theta^{-2}).$$

(e)  $p(x; \theta) = \theta \exp(-\theta x) 1_{(0, \infty)}(x)$ , so

$$\log p(x; \theta) = \log \theta - \theta x,$$

and

$$\dot{\mathbf{l}}_{\theta}(x) = \theta^{-1} - x, \quad \ddot{\mathbf{l}}_{\theta, \theta}(x) = -\theta^{-2}.$$

Thus the information for  $\theta$  (in a sample of size = 1) is  $I(\theta) = \theta^{-2}$ , and the information lower bound for unbiased estimators of  $\theta$  is  $1/(nI(\theta)) = \theta^2/n$ .