

Statistics 581, Problem Set 9 Solutions

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1. Lehmann and Casella, TPE, problem 6.3.22, page 503, reworded as follows. (In other words, prove (vi) of theorem 1.2, pages 5-6, chapter 4 notes). Suppose that X_1, \dots, X_n are i.i.d. with density p_θ , $\theta \in \Theta \subset R^k$, satisfying the hypotheses of theorem 4.1, page 463 (the Cramér conditions given in (A) - (D) on pages 462-463). Show that the following Local Asymptotic Normality (LAN) result holds for the (local) log-likelihood ratios: with

$$L_n(\theta) \equiv \log\left(\prod_{i=1}^n p_\theta(X_i)\right) = \sum_{i=1}^n \log p_\theta(X_i),$$

for a fixed $\theta_0 \in \Theta$,

$$\begin{aligned} L_n(\theta_0 + n^{-1/2}\underline{t}) - L_n(\theta_0) &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \underline{t}^T \dot{\mathbf{l}}_\theta(X_i) - \frac{1}{2} \underline{t}^T I(\theta_0) \underline{t} + o_p(1) \\ &\rightarrow_d N(0, \underline{t}^T I(\theta_0) \underline{t}) - \frac{1}{2} \underline{t}^T I(\theta_0) \underline{t} \\ &\stackrel{d}{=} N(-\sigma^2/2, \sigma^2) \end{aligned}$$

under P_{θ_0} where $\sigma^2 \equiv \underline{t}^T I(\theta_0) \underline{t}$. (The convergence in the last display actually holds under the considerably weaker hypothesis of Hellinger differentiability of p_θ at θ_0 , as stated in Corollary 3 of section 3.3, page 28, of the Chapter 3 notes.)

Solution: I will use the notation of the Chapter 4 notes: in particular, $l_n(\theta) = \log(\prod_{i=1}^n p_\theta(X_i)) = \sum_{i=1}^n \log p_\theta(X_i)$. As in the proof of Theorem 1.3 of the Chapter 4 notes, we can write, with $\theta_n = \theta_0 + n^{-1/2}\underline{t}$,

$$\begin{aligned} l_n(\theta_n) - l_n(\theta_0) &= (\theta_n - \theta_0)^T \dot{\mathbf{l}}_n(\theta_0) - \frac{1}{2} (\theta_n - \theta_0)^T \left(-\ddot{\mathbf{l}}_n(\theta_0) \right) (\theta_n - \theta_0) \\ &\quad + \frac{1}{6} \sum_{j=1}^d \sum_{k=1}^d \sum_{l=1}^d (\theta_{nj} - \theta_{0j})(\theta_{nk} - \theta_{0k})(\theta_{nl} - \theta_{0l}) (\ddot{\mathbf{l}}_{n,jkl}(\theta_n^*)) \end{aligned}$$

where $|\theta_n^* - \theta_0| \leq |\theta_n - \theta_0| = n^{-1/2}|\underline{t}|$. Note that

$$\dot{\mathbf{l}}_n(\theta_0) = \sum_{i=1}^n \dot{\mathbf{l}}_\theta(X_i; \theta_0), \quad \text{and} \tag{0.1}$$

$$\ddot{\mathbf{l}}_n(\theta_0) = \sum_{i=1}^n \ddot{\mathbf{l}}_{\theta, \theta}(X_i; \theta_0). \tag{0.2}$$

Furthermore, by the hypothesis A3 (ii),

$$\ddot{\mathbf{l}}_{n,jkl}(\theta_n^*) = \sum_{i=1}^n \ddot{\mathbf{l}}_{j,k,l}(\theta_n^*) = \sum_{i=1}^n \gamma_{j,k,l}(X_i) M_{j,k,l}(X_i)$$

where $0 \leq |\gamma_{j,k,l}(x)| \leq 1$ and

$$\frac{1}{n} \sum_{i=1}^n M_{j,k,l}(X_i) \rightarrow_{p,a.s.} E_{\theta_0} M_{j,k,l}(X_1) \quad (0.3)$$

for all $1 \leq j, k, l \leq d$. Thus we can write

$$\begin{aligned} l_n(\theta_n) - l_n(\theta_0) &= t^T n^{-1/2} \sum_{i=1}^n \dot{\mathbf{i}}_{\theta}(X_i; \theta_0) - \frac{1}{2} t^T I(\theta_0) t \\ &\quad + \frac{1}{2} t^T I(\theta_0) t - \frac{1}{2} t^T \left(-n^{-1} \ddot{\mathbf{i}}_n(\theta_0) \right) t \\ &\quad + \frac{1}{6} \sum_{j=1}^d \sum_{k=1}^d \sum_{l=1}^d t_j t_k t_l n^{-3/2} \sum_{i=1}^n \gamma_{j,k,l}(X_i) M_{j,k,l}(X_i) \\ &= t^T n^{-1/2} \sum_{i=1}^n \dot{\mathbf{i}}_{\theta}(X_i; \theta_0) - \frac{1}{2} t^T I(\theta_0) t \\ &\quad + \frac{1}{2} t^T I(\theta_0) t - \frac{1}{2} t^T \left(-n^{-1} \ddot{\mathbf{i}}_n(\theta_0) \right) t \\ &\quad + \frac{1}{6} \sum_{j=1}^d \sum_{k=1}^d \sum_{l=1}^d t_j t_k t_l n^{-3/2} \sum_{i=1}^n \gamma_{j,k,l}(X_i) M_{j,k,l}(X_i) \\ &\equiv t^T n^{-1/2} \sum_{i=1}^n \dot{\mathbf{i}}_{\theta}(X_i; \theta_0) - \frac{1}{2} t^T I(\theta_0) t \\ &\quad + R_{n,1}(\theta_0, t) + R_{n,2}(\theta_0, t) \end{aligned}$$

where, in view of (0.2) and (0.3)

$$R_{n,1}(\theta_0, t) \rightarrow_p 0 \quad \text{since} \quad -n^{-1} \sum_{i=1}^n \ddot{\mathbf{i}}_{\theta, \theta}(X_i; \theta_0) \rightarrow_p I(\theta_0),$$

and

$$|R_{n,2}(\theta_0, t)| \leq \frac{1}{6} n^{-1/2} \sum_{j=1}^d \sum_{k=1}^d \sum_{l=1}^d |t_j| |t_k| |t_l| n^{-1} \sum_{i=1}^n M_{j,k,l}(X_i) \rightarrow_p 0.$$

2. Lehmann and Casella, problem 6.2.14, page 501. **[Hint:** We did this in class on Friday 11/20 using the first display on page 26 of the Chapter 3 notes. What remains to prove is uniform integrability or some other justification of the interchange of limit and integration.]

First Solution: From the first display on page 26 of Chapter 3, for $\theta_n = cn^{-1/2}$,

$$\begin{aligned} \sqrt{n}(T_n - \theta_n) &\stackrel{d}{=} Z 1_{\|Z+c\| > n^{1/4}} + (aZ + c(a-1)) 1_{\|Z+c\| \leq n^{1/4}} \\ &\equiv Z_n \equiv A_n + B_n \\ &\rightarrow_{a.s.} aZ + c(a-1) \sim N(c(a-1), a^2). \end{aligned}$$

Moreover, by the c_r -inequality (Proposition 2.1.3, page 5, chapter 2 notes) applied twice,

$$\begin{aligned} E|Z_n|^4 &\leq 2^3 (E|A_n|^4 + E|B_n|^4) \\ &\leq 2^3 (E|Z|^4 + E|aZ + c(a-1)|^4) \\ &\leq 2^3 (3 + 2^3(E|aZ|^4 + c^4(1-a)^4)) \\ &= 2^3 (3 + 2^3(3a^4 + c^4(1-a)^4)). \end{aligned}$$

This implies that the sequence $\{|Z_n|^2\}_{n \geq 1}$ is uniformly integrable, and hence by Vitali's theorem (Theorem 2.1.2, page 4, chapter 2 notes),

$$R_n(\theta_n) = E_{\theta_n}\{n[T_n - \theta_n]^2\} = EZ_n^2 \rightarrow E(aZ + c(a-1))^2 = a^2 + c^2(1-a)^2.$$

As we showed in class, this limiting risk exceeds 1, the risk of \bar{X}_n , if $|c| > \sqrt{(1+a)/(1-a)}$.

Second Solution: Now $Z_n \rightarrow_{a.s.} aZ + c(a-1)$ where $Z \sim N(0,1)$ and by the c_r -inequality with $r = 2$

$$|Z_n|^2 \leq 2\{Z^2 + (aZ + c(a-1))^2\} \equiv Y$$

where $EY < \infty$. Thus it follows from the dominated convergence theorem that $E(Z_n^2) \rightarrow E((aZ + c(a-1))^2)$.

3. (a) Exercise 2.1.6, page 10, chapter 2 notes.
 (b) Exercise 2.1.7, page 10, chapter 2 notes.

Solution: (a) From the proof of proposition 1.13, chapter 2 notes, page 9, we see that

$$\begin{aligned} d_{TV}(P, Q) &= \frac{1}{2} \int |p - q| d\mu = \int_{[p \geq q]} (p - q) d\mu = \int_{[p \geq q]} p d\mu - \int_{[p \geq q]} p \wedge q d\mu \\ &= \int_{[p \geq q]} p d\mu + \int_{[p < q]} p d\mu - \int_{[p \geq q]} p \wedge q d\mu - \int_{[p < q]} p d\mu \\ &= \int p d\mu - \int_{[p \geq q]} p \wedge q d\mu - \int_{[p < q]} p \wedge q d\mu \\ &= 1 - \int p \wedge q d\mu \equiv 1 - \eta(P, Q). \end{aligned}$$

Alternatively, use the identity $|a - b| = a + b - 2(a \wedge b)$ for all $a, b \in \mathbb{R}$ to deduce that

$$|p(x) - q(x)| = p(x) + q(x) - 2p(x) \wedge q(x)$$

for each fixed x , and hence

$$\begin{aligned} d_{TV}(P, Q) &= \frac{1}{2} \int |p - q| d\mu = \frac{1}{2} \left(\int p d\mu + \int q d\mu - 2 \int p \wedge q d\mu \right) \\ &= 1 - \int p \wedge q d\mu \equiv 1 - \eta(P, Q). \end{aligned}$$

(b) To see the first inequality, note that $H^2(P, Q) = 1 - \rho(P, Q)$ where

$$\rho(P, Q) = \int \sqrt{pq} \, d\mu \geq \int p \wedge q \, d\mu \equiv \eta(P, Q)$$

since $\sqrt{p(x)q(x)} \geq p(x) \wedge q(x)$ for all x . Thus we have

$$H^2(P, Q) = 1 - \rho(P, Q) \leq 1 - \eta(P, Q) = d_TV(P, Q).$$

For the second inequality, write $|p - q| = |(\sqrt{p} - \sqrt{q})(\sqrt{p} + \sqrt{q})|$ and then apply the Cauchy-Schwarz inequality: thus

$$\begin{aligned} 2d_{TV}(P, Q) &= \int |p - q| \, d\mu = \int |(\sqrt{p} - \sqrt{q})(\sqrt{p} + \sqrt{q})| \, d\mu \\ &\leq \left(\int |\sqrt{p} - \sqrt{q}|^2 \, d\mu \right)^{1/2} \left(\int |\sqrt{p} + \sqrt{q}|^2 \, d\mu \right)^{1/2} \\ &= \sqrt{2}H(P, Q) \left\{ \int (p + 2\sqrt{pq} + q) \, d\mu \right\}^{1/2} \\ &= \sqrt{2}H(P, Q) \{2 + 2\rho(P, Q)\}^{1/2} \\ &= 2H(P, Q) \{1 + \rho(P, Q)\}^{1/2}, \end{aligned}$$

and this yields the claimed inequality. The third inequality is easy since $\rho(P, Q) \leq 1$ by Cauchy-Schwarz again.