

## Statistics 581, Problem Set 10

Wellner; 12/2/2009

**Reading:** Chapter 4, Sections 1-4;

Ferguson, ACLST, Chapter 20, pages 133-139; Chapter 22, pages 144-150;

Lehmann and Casella, Chapter 6, especially section 6.5, pages 461-468.

**Due:** Wednesday, December 9, 2009.

1. Suppose that (as in Lemma 5.2, page 38, Chapter 3 Notes)  $P$  and  $Q$  are two probability measures on a measurable space  $(\mathcal{X}, \mathcal{A})$  with densities  $p$  and  $q$  with respect to a  $\sigma$ -finite dominating measure  $\mu$ , and  $P^n$  and  $Q^n$  denote the corresponding product measures on  $(\mathcal{X}^n, \mathcal{A}_n)$  (of  $X_1, \dots, X_n$  i.i.d. as  $P$  or  $Q$  respectively).
  - (a) What is the relationship between  $K(P^n, Q^n)$  and  $K(P, Q)$ , if any?
  - (b) If  $P$  is the Normal( $0, \sigma^2$ ) distribution and  $Q$  is the Normal( $\mu, \sigma^2$ ) distribution, compute  $K(P, Q)$ ,  $\rho(P, Q) = \int \sqrt{pq} d\mu$ , and  $H^2(P, Q)$ .
  - (c) Use the results of (a) and (b) together with Lemma 5.2 to calculate  $K(P^n, Q^n)$ ,  $\rho(P^n, Q^n)$ , and  $H^2(P^n, Q^n)$  when  $P$  and  $Q$  are as in (b).
  - (d) Find a sequence  $\mu_n$  so that, with  $Q_n$  being the Normal distribution with mean  $\mu_n$ , the quantities  $K(P^n, Q_n^n)$ ,  $\rho(P^n, Q_n^n)$ , and  $H^2(P^n, Q_n^n)$  converge to finite limits as  $n \rightarrow \infty$ .
2. Consider the Weibull family of example 3.2.5:  $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$  with  $\Theta \subset R^{+2}$  given by the (Lebesgue) densities

$$p_\theta(x) = \frac{\beta}{\alpha} \left(\frac{x}{\alpha}\right)^{\beta-1} \exp\left(-\left(\frac{x}{\alpha}\right)^\beta\right) 1_{[0, \infty)}(x)$$

where  $\theta \equiv (\alpha, \beta) \in (0, \infty) \times (0, \infty) \subset R^2$ . Suppose that  $X, X_1, \dots, X_n$  are i.i.d. with density function  $p_\theta$ .

- (a) If  $X \sim P_\theta \in \mathcal{P}$ , show that the distributions of  $-\log X$  form a location and scale family from a Gumbel (extreme value) density on  $R$ .
- (b) Use the result of (a) to construct method of moments estimators or quantile based estimators  $\bar{\theta}_n$  of  $\theta = (\alpha, \beta)$ .
- (c) Show that the method of moments or quantile estimators  $\bar{\theta}_n$  of  $\theta$  are asymptotically normal, and find the asymptotic distribution; i.e. show that

$$\sqrt{n}(\bar{\theta}_n - \theta) \rightarrow_d N_2(0, \Sigma) \quad \text{for some } \Sigma.$$

[We will use these estimators as “starting points” approximate (or one-step) maximum likelihood estimators in the next problem.]

3. (Problem #2, continued).
  - (a) Does a maximum likelihood estimate of  $\hat{\theta} = (\hat{\alpha}, \hat{\beta})$  exist? Is it unique? (See Lehmann and Casella, Example 6.1, page 468.)
  - (b) Compute an approximate (one - step) maximum likelihood estimate  $\check{\theta}$  of  $\theta$  using

the method of moment (or quantile) estimators  $\bar{\theta}_n$  as the preliminary estimators based on the following data (with  $n = 12$ ):

1, 1.3, 1.7, 3.2, 10.7, 24.3, 51.2, 77.1, 93.7, 105, 111, 305.

[These are failure times in seconds for “breakdown” of an insulating fluid between two electrodes subject to a voltage of 40 kV. – from Nelson, *Applied Life Data Analysis*, page 252, but with some modifications or “recording errors”.]

(c) Compute the maximum likelihood estimator  $\hat{\theta}_n$ , and compare it with the one step estimator computed in E.

4. Suppose that we want to model the survival of twins with a common genetic defect, but with one of the two twins receiving some treatment. Let  $X$  represent the survival time of the untreated twin and let  $Y$  represent the survival time of the treated twin. One (overly simple) preliminary model might be to assume that  $X$  and  $Y$  are independent with  $\text{Exponential}(\eta)$  and  $\text{Exponential}(\nu\eta)$  distributions, respectively:

$$f_{\nu,\eta}(x, y) = \eta e^{-\eta x} \eta \nu e^{-\eta \nu y} 1_{(0,\infty)}(x) 1_{(0,\infty)}(y)$$

A. One crude approach to estimation in this problem is to reduce the data to  $W = X/Y$ , the maximal invariant for the group of scale changes  $g(x, y) = (cx, cy)$  with  $c > 0$ . Find the distribution of  $W$ , and compute the Cramér-Rao lower bound for unbiased estimates of  $\nu$  based on  $W_1, \dots, W_n$  with  $W_i = X_i/Y_i$  and  $(X_i, Y_i)$  i.i.d. as  $(X, Y)$ .

B. Find the information bound for estimation of  $\nu$  based on observation of  $(X, Y)$  pairs when  $\eta$  is known and unknown.

C. Compare the bounds you computed in A and B and discuss the pros and cons of reducing to estimation based on the ratio  $W = X/Y$ .

D. Find the MLE  $\hat{\nu}_n$  of  $\nu$  based on  $(X_i, Y_i)$ ,  $i = 1, \dots, n$ , and the MLE  $\hat{\nu}_n^{(r)}$  of  $\nu$  based on the reduced data  $W_1, \dots, W_n$ . What are the limiting distributions of  $\sqrt{n}(\hat{\nu}_n - \nu)$  and  $\sqrt{n}(\hat{\nu}_n^{(r)} - \nu)$ ?

5. **Optional bonus problem:**

(a) Ferguson, ACLST, page 139, problem 3.

(b) What if Ferguson’s density  $f(x|\theta)$  with  $\theta \in (0, 1)$  is replaced by  $\theta = (\gamma, \eta) \in (0, 1) \times (0, \infty)$  and

$$f(x|\theta) \equiv f(x|\gamma, \eta) = \{(1 - \gamma)e^{-x} + \gamma\eta^2 x \exp(-\eta x)\} 1_{[0,\infty)}(x)?$$

Can you estimate  $\gamma$  and  $\eta$  by the method of moments? Can you improve method of moment estimators via one-step estimators?