

Statistics 581, Midterm Exam

Wellner; 11/6/2000

1. (24 points) **Define** any three of the following terms. In each case, provide an appropriate context for your definition.
 - (a) A probability measure P on a measurable space (Ω, \mathcal{A}) .
 - (b) The total variation distance between two probability measures P and Q .
 - (c) Absolute continuity of a measure ν with respect to a measure μ on a measure space (Ω, \mathcal{A}) .
 - (d) Almost sure convergence (of a sequence of random variables).
 - (e) Convergence in distribution (of a sequence of random variables).
 - (f) The definition of a non-central chi-square random variable with n degrees of freedom and noncentrality parameter δ (in terms of central chi-square random variables and a random variable with a Poisson distribution).

2. (24 points) **State** any three of the following results, providing the appropriate context for your statement:
 - (a) The fundamental event identity for the inverse transformation.
 - (b) The dominated convergence theorem.
 - (c) A (joint) central limit theorem (asymptotic normality result) for the distribution of k different sample quantile $(\mathbb{F}_n^{-1}(t_1), \dots, \mathbb{F}_n^{-1}(t_k))$ where $0 < t_1 < \dots < t_k < 1$.
 - (d) The Cramér - Wold device.
 - (e) The Lindeberg-Feller central limit theorem.
 - (f) The Mann-Wald or continuous mapping theorem.

Do **either** problems 3 **or** problem 4.

3. (30 points) (a) **State** the Glivenko-Cantelli theorem. Then **prove** that it holds if it holds for the case of i.i.d. Uniform(0, 1) random variables.
(b) **Prove** the Glivenko-Cantelli theorem for i.i.d. Uniform(0, 1) random variables: if $\xi_1, \dots, \xi_n, \dots$ are i.i.d. Uniform(0, 1) with empirical distribution function

$$\mathbb{G}_n(t) = \frac{1}{n} \sum_{i=1}^n 1_{[0,t]}(\xi_i), \quad \text{then} \quad \sup_{0 \leq t \leq 1} |\mathbb{G}_n(t) - t| \rightarrow_{a.s.} 0.$$

4. (30 points) Suppose that X, X_1, \dots, X_n are i.i.d. with distribution function F given by $P(X > x) = 1 - F(x) = 1/x^6$, $x \geq 1$, $F(x) = 0$, $x \leq 1$.
 - (a) For what values of $r > 0$ is $E|X|^r < \infty$? If they are finite, compute $\mu = E(X)$ and $\sigma^2 = Var(X)$.
 - (b) Which of the following statements are true? (Briefly indicate why or why not.)
 - (i) $\sum_{i=1}^n X_i = O_p(n^{1/2})$.
 - (ii) $n^{1/3}(\bar{X}_n - \mu) = o_p(1)$.
 - (iii) $n^{2/3}(\bar{X}_n - \mu) = O_p(1)$.

- (iv) $\tan(\pi\sqrt{n}(\overline{X}_n - \mu)) = O_p(1)$.
 (v) $g(n^{1/3}(\overline{X}_n - \mu)) = O_p(1)$ with $g(x) = 1/(1 + e^{-x})$.

Do **any two** of problems 5, 6, and 7.

5. (30 points) Let (Ω, \mathcal{A}, P) be a probability space. Let $\{A_n\}$ be a sequence of events, $A_n \subset \Omega$, $A_n \in \mathcal{A}$ for $n = 1, 2, \dots$, and let $X_n = 1_{A_n}$ be the indicator functions of the events A_n .
- (a) Define the sets $\limsup A_n = [A_n \text{ i.o.}]$ and $\liminf A_n = [A_n \text{ a.a.}]$ in terms of the collection $\{A_n\}$.
- (b) What does Fatou's lemma say about $E(\liminf X_n)$? Translate this into a statement relating $P(\liminf A_n)$ to $P(A_n)$.
- (c) Based on your answer to (b), is the inequality $P(\limsup A_n^c) \leq \limsup P(A_n^c)$ true or false? If it is true, explain why. If false, give a counterexample.
6. (30 points) Suppose that $\underline{N} = (N_1, \dots, N_k) \sim \text{Mult}_k(n, \underline{p})$ where $\underline{p} = (p_1, \dots, p_k)$. In class and homework problems we have discussed the chi-square statistic Q_n and the Hellinger distance statistic $4nH_n^2$ as test statistics for testing $H : \underline{p} = \underline{p}_0$ versus $K : \underline{p} \neq \underline{p}_0$. An alternative statistic for testing H versus K is the likelihood ratio statistic $2 \log \lambda_n$ where

$$\lambda_n \equiv \frac{\sup_{\underline{p}} L_n(\underline{p})}{L_n(\underline{p}_0)} = \frac{\prod_{j=1}^k \hat{p}_j^{N_j}}{\prod_{j=1}^k p_{0j}^{N_j}} = \prod_{j=1}^k \left\{ \frac{\hat{p}_j}{p_{0j}} \right\}^{N_j}.$$

- (a) Show that

$$2 \log \lambda_n = 2n \sum_{j=1}^k \hat{p}_j \log \left(\frac{\hat{p}_j}{p_{0j}} \right).$$

- (b) If the alternative hypothesis K is true, so $\underline{p} \neq \underline{p}_0$, show that

$$n^{-1} 2 \log \lambda_n = g(\hat{\underline{p}}) \rightarrow_p g(\underline{p}),$$

and identify $g(\underline{p})$ as a function of \underline{p} and \underline{p}_0 .

- (c) If the alternative hypothesis K is true, so $\underline{p} \neq \underline{p}_0$, show that

$$\sqrt{n}(2n^{-1} \log \lambda_n - g(\underline{p})) = \sqrt{n}(g(\hat{\underline{p}}) - g(\underline{p})) \rightarrow_d N(0, V^2(\underline{p})),$$

and compute $V^2(\underline{p})$. Could you use this to approximate the power of the likelihood-ratio test? How?

7. (30 points) Suppose that X_1, \dots, X_m are i.i.d. F and Y_1, \dots, Y_n are i.i.d. G and independent of the X 's. Consider testing $H : F = G$ (with $F = G$ continuous) versus $K : F \neq G$. Let \mathbb{F}_m denote the empirical distribution function of the X_i 's, and let \mathbb{G}_n denote the empirical distribution function of the Y_j 's, and set $N \equiv m + n$. The *two-sample Kolmogorov statistic* for testing H versus K is

$$D_{m,n} \equiv \sqrt{\frac{mn}{N}} \|\mathbb{F}_m - \mathbb{G}_n\|_\infty = \sqrt{\frac{mn}{N}} \sup_x |\mathbb{F}_m(x) - \mathbb{G}_n(x)|.$$

(a) By introducing two independent samples ξ_1, \dots, ξ_m and ζ_1, \dots, ζ_n of $\text{Uniform}(0, 1)$ random variables, with corresponding empirical distribution functions

$$\mathbb{K}_m(u) \equiv m^{-1} \sum_{i=1}^m 1_{[0,u]}(\xi_i), \quad \text{and} \quad \mathbb{L}_n(u) \equiv n^{-1} \sum_{i=1}^n 1_{[0,u]}(\zeta_i),$$

and uniform empirical processes \mathbb{U}_m and \mathbb{V}_n defined by

$$\mathbb{U}_m(t) = \sqrt{m}(\mathbb{K}_m(t) - t), \quad \text{and} \quad \mathbb{V}_n(t) = \sqrt{n}(\mathbb{L}_n(t) - t),$$

show that under the null hypothesis H

$$D_{m,n} =_d \|\sqrt{1 - \lambda_N} \mathbb{U}_m - \sqrt{\lambda_N} \mathbb{V}_n\|_\infty = \sup_{0 \leq t \leq 1} |\sqrt{1 - \lambda_N} \mathbb{U}_m(t) - \sqrt{\lambda_N} \mathbb{V}_n(t)|.$$

where $\lambda_N \equiv m/N$, $\bar{\lambda}_N = 1 - \lambda_N = n/N$.

(b) Use the result of (a) to show that when $m, n \rightarrow \infty$ with $\lambda_N \rightarrow \lambda \in [0, 1]$,

$$D_{m,n} \rightarrow_d \|\sqrt{1 - \lambda} \mathbb{U} - \sqrt{\lambda} \mathbb{V}\|_\infty =_d \sup_{0 \leq t \leq 1} |\mathbb{U}(t)|$$

where \mathbb{U} and \mathbb{V} denote two independent Brownian bridge processes.

(c) When the alternative hypothesis holds, so that $F \neq G$, show that $D_{m,n}/\sqrt{mn/N} \rightarrow_{a.s.}$ “something”, and find “something”.

(d) Based on our discussions in class and homework, can you suggest any other test statistic for testing H versus K ?