

# Notes on Convergence in Law of Maxima

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June 1, 2015

## 1. Introduction

For the basics of convergence in distribution of maxima of independent random variables, see Van der Vaart (1998), section 21.4, pages 312-314, and Ferguson (1996), chapter 14, pages 94-100. For further more recent results in this vein, see Engelke et al. (2015) and Kabluchko (2011a, 2011b, 2014).

## 2. Example 1: maxima of i.i.d. standard normal

Suppose that  $X_1, \dots, X_n$  are i.i.d. with d.f.  $F = \Phi$ , the standard normal distribution. Then as discussed by van der Vaart (1998), section 21.4, pages 312-314,  $X_{(n)} \equiv M_n \equiv \max_{1 \leq i \leq n} X_i$  satisfies

$$G_n \equiv b_n(M_n - a_n) \rightarrow_d G \sim Ev \tag{2.1}$$

where  $Ev(x) = \exp(-\exp(-x))$ ,

$$b_n \equiv \sqrt{2 \log n},$$
$$a_n \equiv \sqrt{2 \log n} - \frac{1}{2} \frac{\log \log n + \log(4\pi)}{\sqrt{2 \log n}}$$

Moreover, the density  $f_{G_n}(x) \rightarrow Ev'(x)$  and  $d_{TV}(P_{G_n}, P_G) \rightarrow 0$ . Figure 1 illustrates the convergence in (2.1) and the claimed convergence of densities is illustrated in Figure 2. The rate of convergence in (2.1) is  $O(1/\log n)$ ; see Hall (1979) and Resnick (1987), page 121.

## 3. Example 2: maxima of i.i.d. standard exponentials

Now suppose that  $X_1, \dots, X_n$  are i.i.d. with distribution function  $F$  given by  $1 - F(x) = \exp(-x)$ . Again let  $M_n \equiv X_{(n)}$ . Then it is easily seen that

$$G_n \equiv M_n - \log n \rightarrow_d G \sim Ev \tag{3.1}$$

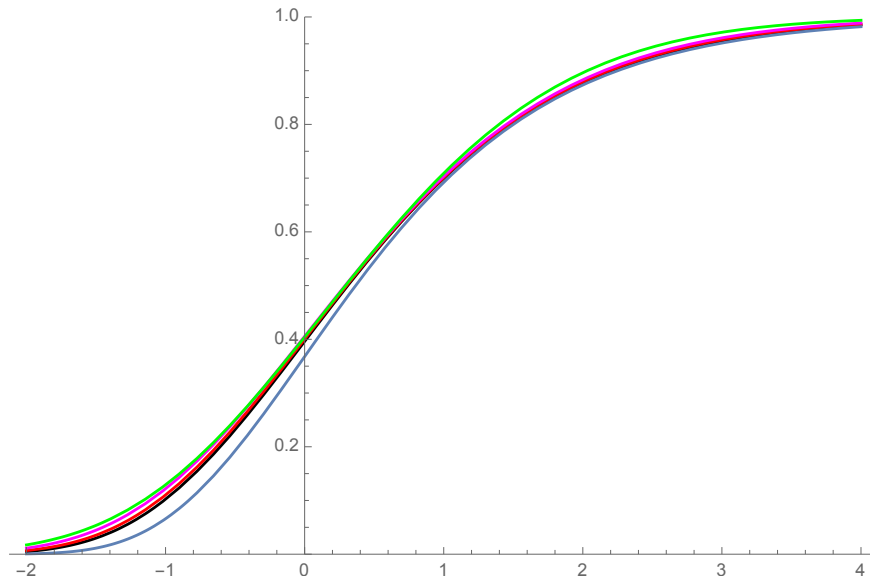


Figure 1: Blue, limit distribution  $Ev$ ; Green,  $n = 10$ ; Magenta,  $n = 100$ ; Red,  $n = 1000$ ; Black,  $n = 10^4$

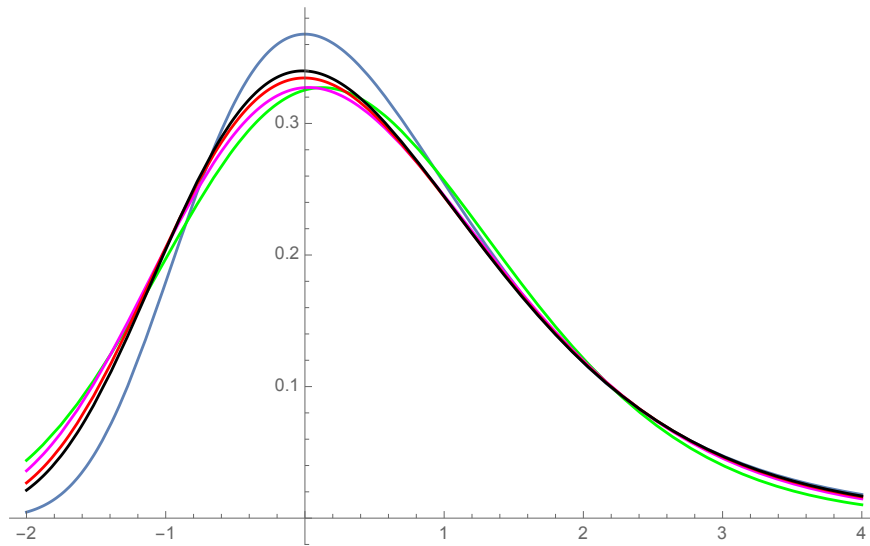


Figure 2: Blue, limit distribution  $Ev$ ; Green,  $n = 10$ ; Magenta,  $n = 100$ ; Red,  $n = 1000$ ; Black,  $n = 10^4$

This follows since

$$P(M_n - \log n \leq x) = P(X_{(n)} \leq x + \log n) = (1 - \exp(-(x + \log n)))^n$$

$$= \left(1 - \frac{e^{-x}}{n}\right)^n \rightarrow \exp(-\exp(-x)) = Ev(x)$$

for every  $x$ . Furthermore,

$$\begin{aligned} f_{G_n}(x) &= n(1 - \exp(-(x + \log n)))^{n-1} \exp(-(x + \log n)) \\ &\rightarrow \exp(-\exp(-x)) \exp(-x) = Ev'(x) \end{aligned}$$

for every  $x$ , and hence by Scheffé's theorem,  $d_{TV}(P_{G_n}, P_G) \rightarrow 0$ . Figure 3 illustrates the convergence in (3.1) and the claimed convergence of densities is illustrated in Figure 4. The rate of convergence in both cases is  $n^{-1}$ ; see Hall & W (1979).

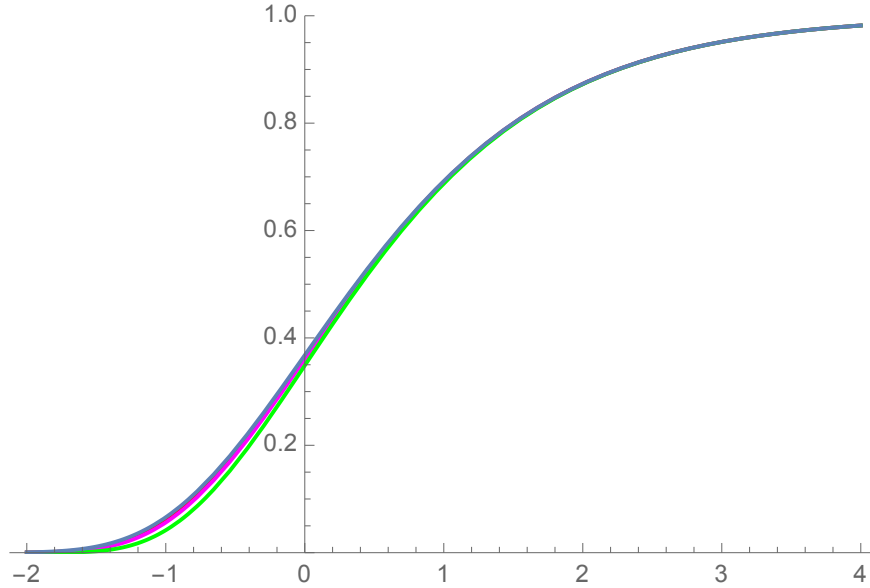


Figure 3: Blue, limit distribution  $Ev$ ; Green,  $n = 10$ ; Magenta,  $n = 25$ ; Red,  $n = 50$

#### 4. Example 3: supremum of a standard kernel estimator

Let  $\hat{f}_n$  be the kernel estimator of a density  $f$  on  $[0, 1]$  based on a kernel  $w$  and the bandwidth  $h_n = n^{-\delta}$  with  $1/5 < \delta < 1/2$ . Bickel and Rosenblatt (1973) show (under hypotheses specified in their paper) that

$$\tilde{M}_n \equiv \sup_{0 < t < 1} \frac{\sqrt{nh_n} |\hat{f}_n(t) - f(t)|}{\sqrt{f(t)}}$$

satisfies the following extreme value convergence:

$$\sqrt{2\delta \log n} \left( \frac{\tilde{M}_n}{\lambda(w)} - d_n \right) \rightarrow_d Ev^2$$

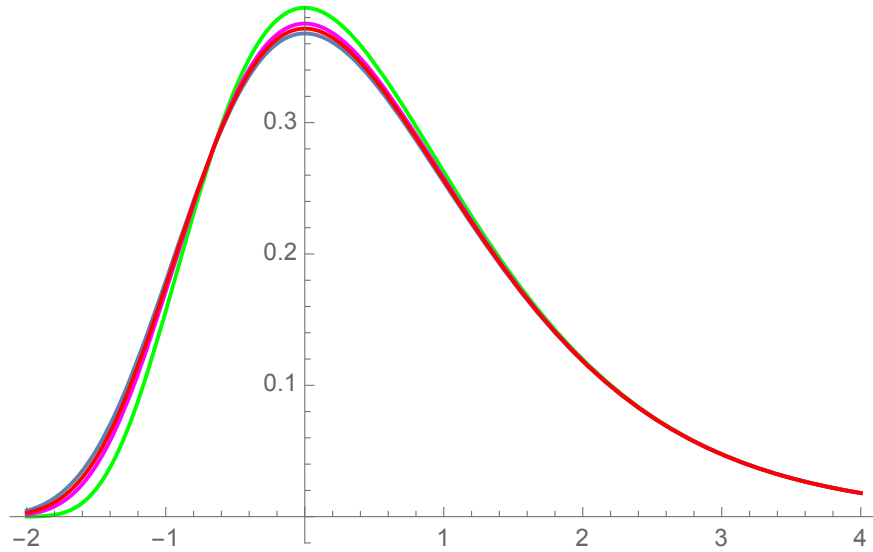


Figure 4: Blue, limit distribution  $Ev$ ; Green,  $n = 10$ ; Magenta,  $n = 25$ ; Red,  $n = 50$

where  $Ev^2(x) = \exp(-2 \exp(-x))$  and where

$$\lambda(w) \equiv \int w^2(t) dt,$$

$$K_1(w) \equiv \frac{w^2(A) + w^2(-A)}{2\lambda(w)},$$

$$K_2(w) \equiv \frac{1}{2\lambda(w)} \int \{w'(t)\}^2 dt,$$

$$d_n = \begin{cases} (2\delta \log n)^{1/2} + \frac{1}{(2\delta \log n)^{1/2}} \left\{ \frac{K_1(w)}{\sqrt{\pi}} - \frac{1}{2}(\log(\delta \log n)) \right\}, & \text{if } K_1(w) > 0, \\ \sqrt{2\delta \log n} + \frac{\log[K_2(w)/(2\pi)]}{\sqrt{2\delta \log n}}, & \text{otherwise.} \end{cases}$$

## 5. Example 4.

Let  $\mathbb{G}_n$  be the empirical distribution function of i.i.d.  $\text{uniform}(0,1)$  random variables and let  $Z_n(t) \equiv \sqrt{n}(\mathbb{G}_n(t) - t)/\sqrt{t(1-t)}$  for  $0 < t < 1$ . Then Jaeschke and Eicker showed that  $\|Z_n\|_\infty \equiv \sup_{0 < t < 1} |Z_n(t)|$  satisfies

$$b_n (\|Z_n\|_\infty - c_n/b_n) \rightarrow_d Ev^4$$

where

$$b_n \equiv \sqrt{2 \log \log n}, \quad c_n \equiv 2 \log \log n + 2^{-1} (\log \log \log n - \log(4\pi)).$$

Here  $Ev^4(x) = \exp(-4 \exp(-x))$ . See Shorack & W (1986), page 600.

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