

## Statistics 523, Problem Set 4 Solutions

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1. PfS Course Notes, Exercise 12.3.1, page 309:  
 Let  $Z \sim N(0, 1)$  and the Brownian bridges  $\mathbb{V}$ ,  $\mathbb{U}^{(1)}$ , and  $\mathbb{U}^{(2)}$  be independent. Fix  $a > 0$ . Show that:
  - (11)  $\mathbb{S}(t) = \mathbb{V}(t) + tZ$ ,  $0 \leq t \leq 1$  is a Brownian motion.
  - (12)  $\mathbb{S}(at)/\sqrt{a}$ ,  $0 \leq t < \infty$  is a Brownian motion.
  - (13)  $\mathbb{S}(t+a) - \mathbb{S}(a)$ ,  $t \geq 0$ , is a Brownian motion.
  - (14)  $\sqrt{1-a}\mathbb{U}^{(1)} \pm \sqrt{a}\mathbb{U}^{(2)}$  is a Brownian bridge if  $0 \leq a \leq 1$ .
  - (15)  $\mathbb{Z}(t) \equiv \{\mathbb{U}^{(1)}(t) + \mathbb{U}^{(2)}(1-t)\}/\sqrt{2}$ ,  $0 \leq t \leq 1/2$  is a Brownian bridge.
  - (16)  $\mathbb{U}(t) = (1-t)\mathbb{S}(t/(1-t))$ ,  $0 \leq t \leq 1$ , is a Brownian bridge; use the LIL at infinity to show that this  $\mathbb{U}$  converges to 0 at  $t = 1$ .
  - (17)  $t\mathbb{S}(1/t)$ ,  $0 \leq t < \infty$  is a Brownian motion; apply the LIL of (10) to verify that these sample paths converge to 0 at  $t = 0$ .

**Solution:** (11) Now  $\mathbb{S}$  is clearly Gaussian with  $E\mathbb{S}(t) = E\mathbb{V}(t) + tE(Z) = 0 + t \cdot 0 = 0$  and, by independence of  $\mathbb{V}$  and  $Z$ ,

$$E\mathbb{S}(s)\mathbb{S}(t) = E\mathbb{V}(s)\mathbb{V}(t) + stE(Z^2) = s \wedge t - st + st = s \wedge t$$

for  $0 \leq s, t \leq 1$ , so  $\mathbb{S}$  is a Brownian motion process on  $[0, 1]$ .

(12) Again  $\tilde{\mathbb{S}} \equiv \mathbb{S}(at)/\sqrt{a}$  is a Gaussian process with  $E\tilde{\mathbb{S}}(t) = 0$  and  $E\tilde{\mathbb{S}}(s)\tilde{\mathbb{S}}(t) = \{(as) \wedge (at)\}/a = s \wedge t$ , so again  $\tilde{\mathbb{S}}$  is a Brownian motion process on  $[0, \infty)$ .

(13)  $\tilde{\mathbb{S}}(t) \equiv \mathbb{S}(t+a) - \mathbb{S}(a)$  is clear Gaussian with mean 0 and

$$\begin{aligned} E\tilde{\mathbb{S}}(s)\tilde{\mathbb{S}}(t) &= E\{(\mathbb{S}(s+a) - \mathbb{S}(a))(\mathbb{S}(t+a) - \mathbb{S}(a))\} \\ &= (s+a) \wedge (t+a) - (s+a) \wedge a - (t+a) \wedge a + a \\ &= (s+a) - a - a + a = s \quad \text{if } s \leq t, \\ &= s \wedge t \quad \text{for arbitrary } s, t \geq 0. \end{aligned}$$

Hence  $\tilde{\mathbb{S}}$  is again Brownian motion on  $[0, \infty)$ .

(14) Now  $\mathbb{V}(t) \equiv \sqrt{1-a}\mathbb{U}^{(1)} \pm \sqrt{a}\mathbb{U}^{(2)}$  with  $\mathbb{U}^{(1)}$  and  $\mathbb{U}^{(2)}$  independent

and  $a \in [0, 1]$  is clearly Gaussian with mean 0 since both  $\mathbb{U}^{(1)}$  and  $\mathbb{U}^{(2)}$  are Gaussian with 0 mean. Furthermore

$$\begin{aligned} E\mathbb{V}(s)\mathbb{V}(t) &= (1-a)E\mathbb{U}^{(1)}(s)\mathbb{U}^{(1)}(t) + aE\mathbb{U}^{(2)}(s)\mathbb{U}^{(2)}(t) \\ &= (1-a)(s \wedge t - st) + a(s \wedge t - st) = s \wedge t - st. \end{aligned}$$

Thus  $\mathbb{V}$  is a standard Brownian bridge process on  $[0, 1]$ .

(15) Again,  $\mathbb{Z}$  is a mean 0 Gaussian process on  $[0, 1]$  and

$$\begin{aligned} E\{\mathbb{Z}(s)\mathbb{Z}(t)\} &= 2^{-1}\{E\mathbb{U}^{(1)}(s)\mathbb{U}^{(1)}(t) + E\mathbb{U}^{(2)}(1-s)\mathbb{U}^{(2)}(1-t)\} \\ &= 2^{-1}\{s \wedge t - st + (1-s) \wedge (1-t) - (1-s)(1-t)\} \\ &= 2^{-1}\{s - st + (1-t) - (1-s)(1-t)\} \text{ if } s \leq t \\ &= 2^{-1}\{s - st + (1-t)(1 - (1-s))\} = s(1-t) \text{ if } s \leq t \\ &= s \wedge t - st \text{ for } 0 \leq s, t \leq 1. \end{aligned}$$

Thus  $\mathbb{Z}$  is a standard Brownian bridge process.

(16) Now  $\mathbb{U}$  is a mean zero Gaussian process since  $\mathbb{S}$ , and we compute

$$\begin{aligned} E\mathbb{U}(s)\mathbb{U}(t) &= E\{(1-s)\mathbb{S}(s/(1-s))(1-t)\mathbb{S}(t/(1-t))\} \\ &= (1-s)(1-t)\frac{s}{1-s} \wedge \frac{t}{1-t} \\ &= (1-s)(1-t)\frac{s}{1-s} \text{ if } s \leq t \\ &= s(1-t) \text{ if } s \leq t \\ &= s \wedge t - st \text{ for general } 0 \leq s, t \leq 1. \end{aligned}$$

Moreover, note that

$$\begin{aligned} &\lim_{t \nearrow 1} (1-t)|\mathbb{S}(t/(1-t))| \\ &\leq \lim_{t \nearrow 1} (1-t)\sqrt{2(t/(1-t) \log \log(t/(1-t)))} \limsup_{t \rightarrow 1} \frac{|\mathbb{S}(t/(1-t))|}{\sqrt{2(t/(1-t) \log \log(t/(1-t)))}} \\ &= \lim_{t \nearrow 1} \sqrt{2(t(1-t) \log \log(t/(1-t)))} \limsup_{s \rightarrow \infty} \frac{|\mathbb{S}(s)|}{\sqrt{2s \log \log s}} \\ &= 0 \cdot 1 = 0 \text{ almost surely} \end{aligned}$$

by the LIL for BM at  $\infty$ . Hence  $\mathbb{U}$  is a standard Brownian bridge process on  $[0, 1]$ .

(17) Now  $\tilde{\mathbb{S}}(t) \equiv t\mathbb{S}(1/t)$  is a 0–mean Gaussian process with

$$\begin{aligned} E\{\tilde{\mathbb{S}}(s)\tilde{\mathbb{S}}(t)\} &= stE\{\mathbb{S}(1/s)\mathbb{S}(1/t)\} \\ &= st((1/s) \wedge (1/t)) = st \cdot (1/t) \quad \text{if } s \leq t \\ &= s \quad \text{if } s \leq t \\ &= s \wedge t \quad \text{for any } 0 \leq s, t < \infty. \end{aligned}$$

Furthermore,

$$\begin{aligned} \lim_{t \searrow 0} |\tilde{\mathbb{S}}(t)| &\leq \lim_{t \searrow 0} t \sqrt{2t \log \log(1/t)} \cdot \limsup_{t \searrow 0} \frac{|\mathbb{S}(1/t)|}{\sqrt{2t \log \log(1/t)}} \\ &= \lim_{t \searrow 0} \sqrt{2t \log \log(1/t)} \cdot \limsup_{t \searrow 0} \frac{|\mathbb{S}(1/t)|}{\sqrt{2(1/t) \log \log(1/t)}} \\ &= \lim_{t \searrow 0} \sqrt{2t \log \log(1/t)} \cdot \limsup_{r \nearrow \infty} \frac{|\mathbb{S}(r)|}{\sqrt{2(r) \log \log(r)}} \\ &= 0 \cdot 1 = 0 \quad \text{a.s.} \end{aligned}$$

by the LIL for  $\mathbb{S}$  at  $+\infty$ . Thus  $\tilde{\mathbb{S}}$  is standard Brownian motion on  $[0, \infty)$ .

2. Pfs Course Notes, Exercise 12.3.6, page 310:

- (a) Suppose that  $h$  and  $\tilde{h}$  on  $[0, 1]$  are in  $L_2[0, 1]$ . View *white noise* as an operator  $d\mathbb{S}$  that takes the function  $h$  into a random variable  $\int_{[0,1]} h(t)d\mathbb{S}(t)$  in the sense of  $\rightarrow_{L_2}$ . Define this integral first for step functions, and then use Exercise 4.4.6 to define it in general. Then show that  $\int_{[0,1]} h(t)d\mathbb{S}(t)$  exists as such an  $L_2$ –limit for all  $h \in L_2[0, 1]$ .
- (b) In case  $h$  has a bounded derivative  $h'$  on  $[0, 1]$ , show that

$$\int_{[0,1]} h(t)d\mathbb{S}(t) = h\mathbb{S}|_{0+}^{1-} - \int_{[0,1]} \mathbb{S}(t)h'(t)dt.$$

- (c) Determine the joint distribution of  $\int_{[0,1]} h(t)d\mathbb{S}(t)$  and  $\int_{[0,1]} \tilde{h}(t)d\mathbb{S}(t)$ .

**Solution:** (a) First consider  $h(t) = 1_{(a,b]}(t)$  for  $0 \leq a < b \leq 1$ . Then we define  $\mathbb{S}(h) \equiv \int_{[0,1]} h d\mathbb{S} = \int_{(a,b]} d\mathbb{S}(t) = \mathbb{S}(b) - \mathbb{S}(a)$ . Now for any function  $h$  of the form  $h_m(t) = \sum_{j=1}^m c_j 1_{(t_{j-1}, t_j]}(t)$  with  $c_j$  real and

$\sum_{j=1}^m (t_{j-1}, t_j] = (0, 1]$ , define

$$\mathbb{S}(h_m) \equiv \int_{[0,1]} h_m d\mathbb{S} = \sum_{j=1}^m c_j (\mathbb{S}(t_j) - \mathbb{S}(t_{j-1})).$$

Then  $\mathbb{S}(h_m)$  is clearly Gaussian and we have

$$E(\mathbb{S}(h_m)^2) = \sum_{j=1}^m c_j^2 (t_j - t_{j-1}) = \int_0^1 h_m^2(t) dt = \|h_m\|^2$$

where  $\|h_m\|_2$  denotes the norm in  $L_2([0, 1], \lambda)$  for Lebesgue measure  $\lambda$ . Thus the map  $h \rightarrow \mathbb{S}(h) = \int_{[0,1]} h d\mathbb{S}$  defines a linear isometry from the collection of step functions in  $L_2([0, 1], \lambda)$  into  $L_2(\Omega, \mathcal{A}, P)$ . Since the set of step functions is dense in  $L_2([0, 1], \lambda)$  by Exercise 4.4.6, we may extend this isometry by continuity to all  $h \in L_2([0, 1], \lambda)$ , and by linearity the whole process  $\{\mathbb{S}(h) : h \in L_2([0, 1])\}$  is Gaussian. Note that for  $g, h \in L_2[0, 1]$  we have

$$4\langle g, h \rangle = \|g + h\|^2 - \|g - h\|^2$$

(this is known as a *polarization identity*), and hence inner products are also preserved: for all  $g, h \in L_2[0, 1]$

$$E(\mathbb{S}(g)\mathbb{S}(h)) = \langle g, h \rangle = \int_0^1 gh d\lambda.$$

(b) When  $h$  has a bounded derivative  $h'$ , we instead define

$$\begin{aligned} \mathbb{S}(h) &\equiv h(t)\mathbb{S}(t)|_{0+}^{1-} - \int_0^1 h'(u)\mathbb{S}(u)du \\ &= h(1)\mathbb{S}(1) - \int_0^1 h'(u)\mathbb{S}(u)du \end{aligned}$$

since  $|h(t)| = |\int_0^t h'(s)ds| \leq \|h'\|_\infty t$  for  $0 \leq t \leq 1$ . Our goal will be to show for this alternative definition that we also have  $E(\mathbb{S}(h)^2) = \int_0^1 h^2(t)dt$  and hence the same isometry holds as in (a) for the functions in  $C^1[0, 1]$  as a subset of  $L_2[0, 1]$ . Since  $C^1[0, 1]$  is also dense in  $L_2[0, 1]$ , the definition agrees with the definition given in (a). We calculate as

follows. First note that it suffices by linearity to consider  $h$  monotone non-decreasing and bounded by  $\|h'\|_\infty$ . For monotone  $h$  we compute (using Fubini's theorem frequently)

$$\begin{aligned}
ES^2(h) &= E\left(h(1)\mathbb{S}(1) - \int_0^1 \mathbb{S}(s)h'(s)dx\right)^2 \\
&= h^2(1) - 2h(1)E\{\mathbb{S}(1) \int_0^1 h'(s)\mathbb{S}(s)ds\} + E\left(\int_0^1 h'\mathbb{S}d\lambda\right)^2 \\
&= h^2(1) - 2h(1) \int_0^1 sh'(s)ds + \int_0^1 \int_0^1 (s \wedge t)h'(s)h'(t)dsdt \\
&= h^2(1) - 2h(1) \int_0^1 sh'(s)ds + \left(\int_0^1 sh'(s)ds\right)^2 \\
&\quad + \int_0^1 \int_0^1 (s \wedge t - st)h'(s)h'(t)dsdt \\
&= \left(h(1) - \int_0^1 sh'(s)ds\right)^2 + \int_0^1 \int_0^1 (s \wedge t - st)h'(s)h'(t)dsdt \\
&= \left(\int_0^1 h(s)ds\right)^2 + \int_0^1 \int_0^1 (s \wedge t - st)dh(s)dh(t) \\
&= \left(\int_0^1 h(s)ds\right)^2 + \text{Var}(h(\xi)) \quad \text{where } \xi \sim U(0,1) \\
&\quad \text{by using (6.4.13) on page 113 of PfS, course notes} \\
&= \left(\int_0^1 h(s)ds\right)^2 + \int_0^1 h^2(t)dt - \left(\int_0^1 h(t)dt\right)^2 \\
&= \int_0^1 h^2(t)dt.
\end{aligned}$$

(c) Now for any step functions  $g_m, h_m$  in  $L_2[0, 1]$  with  $g_m \rightarrow_2 g$  and  $h_m \rightarrow_2 h$  we see that

$$\int_{[0,1]} g_m d\mathbb{S} = \sum_{j=1}^m c_j (\mathbb{S}(t_j) - \mathbb{S}(t_{j-1})) \sim N(0, \|g_m\|^2)$$

by the independent and stationary increments property of  $\mathbb{S}$ , and sim-

ilarly

$$\begin{aligned} E \left( \int_{[0,1]} g_m d\mathbb{S} \right) \left( \int_{[0,1]} h_m d\mathbb{S} \right) &= \sum_{j=1}^m c_j d_j (t_j - t_{j-1}) = \int_0^1 g_m(t) h_m(t) dt \\ &= \langle g_m, h_m \rangle \rightarrow \langle g, h \rangle \end{aligned}$$

and we conclude (again) that

$$E(\mathbb{S}(g)\mathbb{S}(h)) = \langle g, h \rangle = \int_0^1 gh d\lambda.$$

Thus  $(\mathbb{S}(g), \mathbb{S}(h)) \sim N_2(0, \Sigma)$  where

$$\Sigma = \begin{pmatrix} \|g\|^2 & \langle g, h \rangle \\ \langle g, h \rangle & \|h\|^2 \end{pmatrix}.$$

3. PfS Course Notes, Exercise 12.9.2, page 332:

Let  $V_n(r) \equiv \sum_{k=1}^n |\mathbb{S}(t_{nk}) - \mathbb{S}(t_{n,k-1})|^r$  and  $\|\mathcal{P}_n\| \equiv \sup_{1 \leq k \leq n} |t_{nk} - t_{n,k-1}|$ .

( $\alpha$ ) Let  $Z \sim N(0, 1)$ . Let  $r > 0$ . Show that:

(a)  $C_r \equiv E|Z|^r = 2^{r/2} \Gamma((r+1)/2) / \sqrt{\pi}$ .

(b)  $E|\mathbb{S}(t_{n,k-1}, t_{n,k})|^r = C_r |t_{n,k} - t_{n,k-1}|^{r/2}$ ,

$$\text{Var}(|\mathbb{S}(t_{n,k-1}, t_{n,k})|^r) = (C_{2r} - C_r^2) |t_{n,k} - t_{n,k-1}|^r.$$

( $\beta$ ) Now show that  $EV_n(2) = 1$  and  $\text{Var}(V_n(2)) \leq (C_{2r} - C_r^2) \|\mathcal{P}_n\|$  and hence

$$\sum_{n=1}^{\infty} P(|V_n(2) - 1| \geq \epsilon) \leq \epsilon^{-2} (C_{2r} - C_r^2) \sum_{n=1}^{\infty} \|\mathcal{P}_n\|.$$

( $\gamma$ ) Finally prove that  $V_n(2) \rightarrow_{a.s.} 1$  if  $\sum_{n=1}^{\infty} \|\mathcal{P}_n\| < \infty$ .

**Solution:** ( $\alpha$ ) First, if  $Z \sim N(0, 1)$  and  $r > 0$

$$\begin{aligned} E|Z|^r &= 2 \int_0^{\infty} z^r (2\pi)^{-1/2} \exp(-z^2/2) dz \\ &= \sqrt{\frac{2}{\pi}} \int_0^{\infty} (2v)^{(r-1)/2} e^{-v} dv \\ &= \sqrt{\frac{2}{\pi}} 2^{(r-1)/2} \Gamma((r+1)/2) = 2^{r/2} \Gamma((r+1)/2) / \sqrt{\pi} \equiv C_r. \end{aligned}$$

Thus  $C_{2k} = 1 \cdot 3 \cdots (2k - 1)$ . Then note that  $\mathbb{S}(t_{n,k}) - \mathbb{S}(t_{n,k-1}) \sim N(0, t_{n,k} - t_{n,k-1})$  and hence

$$E|\mathbb{S}(t_{n,k-1}, t_{n,k})|^r = |t_{n,k} - t_{n,k-1}|^{r/2} E|Z|^r = C_r |t_{n,k} - t_{n,k-1}|^{r/2}.$$

and

$$\begin{aligned} \text{Var}(\mathbb{S}(t_{n,k-1}, t_{n,k})^2) &= E|\mathbb{S}(t_{n,k-1}, t_{n,k})|^4 - \{E|\mathbb{S}(t_{n,k-1}, t_{n,k})|^2\}^2 \\ &= (C_4 - C_2^2) |t_{n,k} - t_{n,k-1}|^2 \end{aligned}$$

as claimed (and with  $r = 2$ ).

( $\beta$ ) Now it follows from ( $\alpha$ ) that

$$EV_n(2) = C_2 \sum_{k=1}^n (t_{n,k} - t_{n,k-1}) = 1 \cdot \sum_{k=1}^n (t_{n,k} - t_{n,k-1}) = 1$$

since  $C_2 = 1$ , and using the independence of the increments of Brownian motion,

$$\begin{aligned} \text{Var}(V_n(2)) &= (C_4 - C_2^2) \sum_{k=1}^n (t_{n,k} - t_{n,k-1})^2 \\ &\leq (3 - 1) \|\mathcal{P}_n\| \sum_{k=1}^n (t_{n,k} - t_{n,k-1}) = 2 \|\mathcal{P}_n\| \end{aligned}$$

as claimed (but with  $r = 2$ ). Thus by Markov's inequality,

$$\sum_{n=1}^{\infty} P(|V_n(2) - 1| \geq \epsilon) \leq 2\epsilon^{-2} \sum_{n=1}^{\infty} \|\mathcal{P}_n\|.$$

( $\gamma$ ) This last display together with the Borel-Cantelli lemma shows that  $P(|V_n(2) - 1| \geq \epsilon \text{ i.o.}) = 0$  for every  $\epsilon > 0$ , and hence  $V_n(2) \rightarrow_{a.s.} 1$  if  $\sum_{n=1}^{\infty} \|\mathcal{P}_n\| < \infty$ .

4. PfS Course Notes, Exercise 12.9.3, page 332: Prove theorem 9.1(b) when all  $t_{nk} = k/2^n$ . That is, with  $V_n(r) \equiv \sum_{k=1}^{2^n} |\mathbb{S}(t_{nk}) - \mathbb{S}(t_{n,k-1})|^r$ , show that  $V_n(1) \rightarrow_{a.s.} \infty$  if  $\|\mathcal{P}_n\| \equiv \sup_{1 \leq k \leq 2^n} |t_{nk} - t_{n,k-1}| \rightarrow 0$ .

**Solution:** By the Paley-Zygmund inequality, for any  $\lambda \in (0, 1)$  we have

$$P(V_n(1) \geq \lambda EV_n(1)) \geq (1 - \lambda)^2 \frac{\{EV_n(1)\}^2}{EV_n^2(1)}$$

where, by  $(\alpha)$  of the previous problem,

$$EV_n(1) = C_1 \sum_{k=1}^{2^n} \left| \frac{k}{2^n} - \frac{k-1}{2^n} \right|^{1/2} = C_1 2^n \cdot 2^{-n/2} = C_1 2^{n/2}.$$

and

$$\text{Var}(V_n(1)) = (C_2 - C_1^2) \sum_{k=1}^{2^n} \left| \frac{k}{2^n} - \frac{k-1}{2^n} \right| = C_2 2^n \cdot 2^{-n} = (1 - C_1^2).$$

It follows that

$$\begin{aligned} EV_n^2(1) &= \text{Var}(V_n(1)) + \{EV_n(1)\}^2 \\ &= (1 - C_1^2) + C_1^2 2^n \end{aligned}$$

where  $EV_n(1) = C_1 2^{n/2} \rightarrow \infty$  as  $n \rightarrow \infty$ . Thus

$$\liminf_{n \rightarrow \infty} P(V_n(1) \geq \lambda EV_n(1)) \geq (1 - \lambda)^2$$

for every  $\lambda \in (0, 1)$ . This implies that  $V_n(1) \rightarrow_{a.s.} 1$ .

5. Let  $\mathbb{S}$  denote a standard Brownian motion process on  $[0, \infty)$ , and define

$$\mathbb{X}(t) = e^{-t} \mathbb{S}(e^{2t}) \quad \text{for } t \in (-\infty, \infty).$$

- (a) Compute  $E\mathbb{X}(t)\mathbb{X}(s)$  and  $E(\mathbb{X}(t)\mathbb{X}(t+h))$ .
- (b) Is  $\mathbb{X}$  a Gaussian process?
- (c) Is  $\mathbb{X}$  a stationary process?

**Solution:** (a) Now

$$\begin{aligned} E\mathbb{X}(t)\mathbb{X}(s) &= e^{-s} e^{-t} E\mathbb{S}(e^{2s})\mathbb{S}(e^{2t}) \\ &= e^{-s} e^{-t} \{e^{2s} \wedge e^{2t}\} = e^{-s} e^{-t} e^{2s} \quad \text{if } s \leq t \\ &= e^{s-t}, \quad \text{if } s \leq t, \\ &= e^{-|t-s|}, \quad \text{for general } s, t \in -\infty, \infty. \end{aligned}$$

- (b)  $\mathbb{X}$  is clearly a mean 0 Gaussian process since  $\mathbb{S}$  is mean-zero and Gaussian. (c) Note that for any  $t_1 < t_2 < \dots < t_k$  with  $t_j \in \mathbb{R}$  we have

$$(\mathbb{X}(t_1), \dots, \mathbb{X}(t_k)) \sim N_k(0, \Sigma)$$

where  $\Sigma = (\exp(-|t_k - t_j|))_{j,k=1}^m$ . But if  $h > 0$  we see that

$$(\mathbb{X}(t_1 + h), \dots, \mathbb{X}(t_k + h)) \sim N_k(0, \Sigma)$$

as well since the new covariance matrix has entries  $\exp(-|(t_j + h) - (t_k + h)|) = \exp(-|t_j - t_k|)$ . Thus  $\mathbb{X}$  is stationary. The process  $\mathbb{X}$  is known as an *Ornstein-Uhlenbeck process*; see Breiman (1968), pages 347-350. It turns out that  $\mathbb{X}$  is also a Markov process, and that if a process is Gaussian, stationary, and Markov with continuous sample paths, then it is an Ornstein-Uhlenbeck process up to its mean.