

## Statistics 522, Problem Set 8 Solutions

Wellner; 3/5/99

1. Exercise 11.7.2, PFS page 241: If  $F_n \rightarrow_d F$  for a continuous df  $F$ , then  $\|F_n - F\| \rightarrow 0$ .

**Solution:** Let  $M$  be a (large) positive integer, and set  $x_{M,j} \equiv x_j = F^{-1}(j/(M+1))$  for  $j = 1, \dots, M$ , and let  $x_{M,0} \equiv -\infty$ ,  $x_{M,M+1} \equiv \infty$ . Then for  $x \in [x_{j-1}, x_j]$  we have

$$F_n(x) - F(x) \begin{cases} \leq F_n(x_j) - F(x_{j-1}) \leq F_n(x_j) - F(x_j) + 1/M \\ \geq F_n(x_{j-1}) - F(x_j) \geq F_n(x_{j-1}) - F(x_{j-1}) - 1/M \end{cases} ,$$

and hence

$$\begin{aligned} \|F_n - F\| &\leq \max_{1 \leq j \leq M+1} \sup_{x_{j-1} \leq x \leq x_j} |F_n(x) - F(x)| \\ &\leq \max_{1 \leq j \leq M} |F_n(x_j) - F(x_j)| + 1/M \\ &\rightarrow 0 + 1/M \end{aligned}$$

since  $F_n(x) \rightarrow F(x)$  at all  $x$  in view of  $F$  being continuous. But since  $M$  is arbitrary, this can be made arbitrarily small; i.e.  $\|F_n - F\| \rightarrow 0$ .

2. Exercise 11.7.3, PFS page 241: Suppose that  $X_n \sim F_n$ . Show that  $\{F_n : n \geq 1\}$  is tight if either:  
 (a)  $\limsup_n E|X_n|^r < \infty$  for some  $r > 0$ . (b)  $F_n \rightarrow_d F$ .

**Solution:** (a) Let  $\epsilon > 0$ . Now by Markov's inequality

$$\limsup_{n \rightarrow \infty} P(|X_n| > M) \leq \frac{\limsup_{n \rightarrow \infty} E|X_n|^r}{M^r} < \epsilon$$

for  $M > M(r, \epsilon) \equiv (\limsup_{n \rightarrow \infty} E|X_n|^r / \epsilon)^{1/r}$ . Thus there is an  $N \equiv N(\epsilon, r)$  such that  $\sup_{n > N} P(|X_n| > M) \leq 2\epsilon$ . But since  $F_1, \dots, F_N$  are df's, there exists a  $K = K_\epsilon$  so large that  $\max_{1 \leq n \leq N} P(|X_n| > K) < 2\epsilon$ . Taking  $R = R_\epsilon = \max\{M, K\}$  we have

$$\sup_{1 \leq n < \infty} P(|X_n| > R) < 2\epsilon.$$

Thus  $\{F_n\}$ , the family of distributions of  $\{X_n\}$ , is tight.

(b) Let  $\epsilon > 0$ . Let  $r = r(\epsilon) \in C_F$  be so large that  $1 - F(r) < \epsilon/4$ , and let  $l = l(\epsilon) \in C_F$  be so small that  $F(l) < \epsilon/4$ . Now there exists an  $N = N_\epsilon$  so large that

$$|F_n(r) - F(r)| < \epsilon/4 \quad \text{for all } n > N$$

and

$$|F_n(l) - F(l)| < \epsilon/4 \quad \text{for all } n > N.$$

Then we have

$$\begin{aligned} \sup_{n > N} F_n([l, r]^c) &\leq F(l) + (1 - F(r)) \\ &\quad + |F_n(l) - F(l)| + |F_n(r) - F(r)| \\ &\leq \epsilon/4 + \epsilon/4 + \epsilon/4 + \epsilon/4 = \epsilon \end{aligned}$$

But since  $F_1, \dots, F_N$  are distribution functions, we can easily find an interval  $[l', r']$ , such that

$$\max_{1 \leq n \leq N} F_n([l', r']^c) < \epsilon,$$

and hence the interval  $[a, b] \equiv [l \wedge l', r \vee r']$  satisfies

$$\sup_{1 \leq n < \infty} F_n([a, b]^c) < \epsilon;$$

i.e.  $\{F_n\}$  is tight.

3. Exercise 11.8.3, PfS page 241: Suppose that  $\log X \sim N(0, 1)$ ; thus the density of  $X$  with respect to Lebesgue measure is

$$f_X(x) = x^{-1} \exp(-(\log x)^2/2) / \sqrt{2\pi}, \quad x > 0.$$

For  $|a| \leq 1$ , let  $Y$  have the density function

$$f_Y(y) = f_Y(y; a) = f_X(y)[1 + a \sin(2\pi \log y)], \quad y > 0.$$

Show that  $X$  and  $Y$  have exactly the same moments.

**Solution:** It is clear that  $Y$  has the same moments as  $X$  if and only if

$$\int_0^{\infty} y^k f_X(y) \sin(2\pi \log y) dy = 0$$

for  $k = 1, 2, \dots$ . Changing variables, it follows that

$$\begin{aligned} \int_0^{\infty} y^k f_X(y) \sin(2\pi \log y) dy &= \frac{1}{\sqrt{2\pi}} \int_0^{\infty} y^{k-1} \exp(-(\log y)^2/2) \sin(2\pi \log y) dy \\ &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{kx} e^{-x^2/2} \sin(2\pi x) dx. \end{aligned}$$

Now the function in the integrand,  $g(x) = e^{kx} e^{-x^2/2} \sin(2\pi x)$ , is an entire function in the complex plane, and it follows from Cauchy's formula that the integral  $\int_{\Gamma} g(z) dz = 0$  for any closed path  $\Gamma$  in the complex plane. In particular, taking the closed path to consist of the  $x$  axis from  $-R$  to  $R$ , followed by the semicircle of radius  $R$  in the lower half plane connecting  $(+R, 0)$  to  $(-R, 0)$ , then  $\int_{\Gamma} g(z) dz = 0$ . Now for  $z = -Re^{i\theta}$ ,  $\theta \in [0, \pi]$ ,

$$\begin{aligned} g(z) &= \exp(-kRe^{i\theta}) \exp(-R^2/2) \exp(i2\pi Re^{i\theta}) \\ &= \exp(-kR(\cos \theta + i \sin \theta)) \exp(-R^2/2) \exp(i2\pi R(\cos \theta + i \sin \theta)) \\ &= \exp(-kR(\cos \theta + i \sin \theta)) \exp(-R^2/2) \exp(-2\pi R \sin \theta) \exp(i2\pi R \cos \theta) \end{aligned}$$

so that

$$|g(z)| \leq \exp(-R(k \cos \theta + 2\pi \sin \theta)) \exp(-R^2/2) \rightarrow 0$$

as  $R \rightarrow \infty$ . Thus taking the limit on  $R$  across the identity  $\int_{\Gamma(R)} g(z) dz = 0$  yields

$$\int_{-\infty}^{\infty} e^{kx} e^{-x^2/2} \sin(2\pi x) dx = 0.$$

4. Exercise 11.8.5, PFS page 241. Show that the  $N(0, 1)$  distribution is uniquely determined by its moments (by appealing to the previous proposition).

**Solution:** The moments of the  $N(0, 1)$  distribution are  $E(Z^k) = 0$  for  $k$  odd, and  $E(Z^k) = \int z^k \phi(z) dz = (k-1)(k-3)\cdots 3 \cdot 1$ .  $k$  even.

Thus  $\mu_{2k} = (2k-1)(2k-3)\cdots 3\cdot 1$  where there are  $k$  factors. Thus  $\mu_{2k} \leq (2k-1)^k$ , and  $\mu_{2k}^{1/2k} \leq (2k-1)^{1/2}$ , so that

$$\limsup_{k \rightarrow \infty} \frac{1}{k} \mu_{2k}^{1/2k} \leq \limsup_{k \rightarrow \infty} \frac{(2k-1)^{1/2}}{k}.$$

Thus by part (c) of exercise 11.8.4, the  $N(0,1)$  df is the unique d.f. with the moments  $\mu_k$ .

5. Exercise 11.1.1, PfS page 222: Let rv's  $X$  and  $\Delta$  be arbitrary. Let  $\epsilon > 0$ . Suppose that  $G$  has derivative  $g$ . Then

$$\|F_{X+\Delta} - G\| \leq \|F_X - G\| + \epsilon \|g\| + P(|\Delta| > \epsilon).$$

**Solution:** First write

$$\begin{aligned} F_{X+\Delta}(x) - G(x) &= \int_{(-\infty, x]} d(F_{X+\Delta} - G) \\ &= P(X + \Delta \leq x) - G(x) \\ &= P(X + \Delta \leq x, |\Delta| \leq \epsilon) + P(X + \Delta \leq x, |\Delta| > \epsilon) - G(x) \\ &\quad \begin{cases} \leq P(X \leq x + \epsilon) - G(x) + P(|\Delta| > \epsilon) \\ \geq P(X \leq x - \epsilon) - G(x) \end{cases}. \end{aligned}$$

It follows that

$$\begin{aligned} |F_{X+\Delta}(x) - G(x)| &\leq \max\{|F_X(x + \epsilon) - G(x + \epsilon)|, |F_X(x - \epsilon) - G(x - \epsilon)|\} \\ &\quad + \max\{|G(x + \epsilon) - G(x)|, |G(x - \epsilon) - G(x)|\} \\ &\quad + P(|\Delta| > \epsilon). \end{aligned}$$

From this it follows easily that the asserted inequality holds.