

Math/Stat 523, Spring 2020



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Lecture 4

Wednesday April 8?)

Outline

- 1: Variation and (non-)differentiability; the Hölder condition
- 2: Structure of the zero set.
- 3: Strassen's functional LIL
- 4: Boundary crossing probabilities

1. Variation and (non-)differentiability

Definition 12.9.1: (Variation) For a sequence of partitions

$$\mathcal{P}_n \equiv \{(t_{n,k-1}, t_{n,k}] : k = 1, \dots, n\} \text{ of } [0, 1]$$

with $0 \leq t_{n,0} < \dots < t_{n,n} \equiv 1$, define the r th *variation* of \mathbb{S} corresponding to \mathcal{P}_n by

$$V_n(r) \equiv \sum_{k=1}^n |\mathbb{S}(t_{n,k}) - \mathbb{S}(t_{n,k-1})|^r.$$

We call these partitions *nested* if $\mathcal{P}_n \subset \mathcal{P}_{n+1}$ for all $n \geq 1$. We also define the *mesh* of the partition to be $\|\mathcal{P}_n\| \equiv \sup_{1 \leq k \leq n} |t_{n,k} - t_{n,k-1}|$.

2. Non-differentiability, variation, and Hölder condition

Theorem 12.9.1: (Non-differentiability)

- (a) Almost every Brownian path is nowhere differentiable.
- (b) In fact, $V_n(1) \rightarrow_{a.s.} \infty$ a.s. if $\|\mathcal{P}_n\| \rightarrow 0$.
- (c) (Finite squared variation) $V_n(2) \rightarrow_2 1$ if $\|\mathcal{P}_n\| \rightarrow 0$.
- (d) (Finite squared variation) $V_n(2) \rightarrow_{a.s.} 1$ if either
 - (i) $\sum_n \|\mathcal{P}_n\| < \infty$ or
 - (ii) the \mathcal{P}_n are nested with $\|\mathcal{P}_n\| \rightarrow 0$.
- (e) (Dudley) $V_n(2) \rightarrow_{a.s.} 1$ if and only if $(\log n)\|\mathcal{P}_n\| \rightarrow 0$.

Partial proof of Theorem 12.9.1:

(a) If $t \mapsto \mathbb{S}(t, \omega)$ is differentiable at some point $s \in [0, 1]$, then

$$A_n \equiv \{|\mathbb{S}(t) - \mathbb{S}(s)| < C(t - s), \quad s < t < s + 5/n\}, \quad \text{for } n \geq m$$

for some $1 \leq C < \infty$ and some $m \geq 1$. But this is a subset of the event

$$B_n \equiv \bigcup_{0 < i \leq n+2} \bigcap_{i < k \leq i+3} \left\{ \left| \mathbb{S}\left(\frac{k}{n}\right) - \mathbb{S}\left(\frac{k-1}{n}\right) \right| < \frac{7C}{n} \right\}.$$

Then, since

$$\mathbb{S}\left(\frac{k}{n}\right) - \mathbb{S}\left(\frac{k-1}{n}\right) \stackrel{d}{=} n^{-1/2}Z \sim N(0, 1/n),$$

it follows that

$$\begin{aligned} P(A_n) \leq P(B_n) &\leq (n+2)P\left(\left|\mathbb{S}\left(\frac{1}{n}\right)\right| < \frac{7C}{n}\right)^3 \\ &= (n+2)P\left(|Z| < 7C/\sqrt{n}\right)^3 \\ &\leq Mn^{-1/2} \rightarrow 0 \end{aligned}$$

for some constant M (and $M = 14C/\sqrt{2\pi}$ works).

(c) For $r = 2$, note that

$$\begin{aligned} EV_n(2) &= E\left\{\sum_{k=1}^n |\mathbb{S}(t_{n,k}) - \mathbb{S}(t_{n,k-1})|^2\right\} \\ &= \sum_{k=1}^n E|\mathbb{S}(t_{n,k}) - \mathbb{S}(t_{n,k-1})|^2 \\ &= \sum_{k=1}^n (t_{n,k} - t_{n,k-1}) = 1. \end{aligned}$$

Thus $E|V_n(2) - 1|^2 = \text{Var}(V_n(2)) \leq (C_4 - C_2^2)\|\mathcal{P}_n\| \rightarrow 0$ if $\|\mathcal{P}_n\| \rightarrow 0$. Here the variance calculation follows easily using $E|Z|^r \equiv C_r = 2^{r/2}\Gamma((r+1)/2)/\sqrt{\pi}$ for $Z \sim N(0, 1)$ and $r > 0$. Hence

$$\begin{aligned} E|\mathbb{S}(t_{n,k-1}, t_{n,k})|^r &= C_r(t_{n,k} - t_{n,k-1})^{r/2}, \\ \text{Var}[|\mathbb{S}(t_{n,k-1}, t_{n,k})|^r] &= (C_{2r} - C_r^2)(t_{n,k} - t_{n,k-1})^r. \end{aligned}$$

Theorem 12.9.2: (Lévy) The following Hölder type condition holds:

$$\limsup_{\substack{0 \leq s < t \leq 1: \\ t-s=a \searrow 0}} \frac{|\mathbb{S}(t) - \mathbb{S}(s)|}{\sqrt{2a \log(1/a)}} = 1 \quad \text{a.s.}$$

2. Structure of the zero set.

Theorem 12.9.3: (The zeros of \mathbb{S} in $[0, 1]$) Define

$$Z(\omega) = \{t \in [0, 1] : \mathbb{S}(t, \omega) = 0\}.$$

For almost all ω the set $Z(\omega)$ is a closed and perfect set of Lebesgue measure zero. [A set A is called *dense in itself* if every point x of A is such that every neighborhood of x contains another point of A ; i.e. no point of A is isolated. A closed set dense in itself is called *perfect*.]

Partial proof: Note that $\lambda(Z(\omega)) = \int_0^1 1_0(\mathbb{S}(t, \omega)) dt$, so

$$E\lambda(Z) = E \int_0^1 1_0(\mathbb{S}(t)) dt = \int_0^1 P(\mathbb{S}(t) = 0) dt = \int_0^1 0 \cdot dt = 0$$

by the Tonelli part of Fubini's theorem. Thus $\lambda(Z(\omega)) = 0$ a.s.

3. Strassen's functional LIL .

Strassen's functional LIL is to the ordinary LIL as Donsker's (functional central limit) theorem is to the ordinary CLT.

Notation:

- Let \mathbb{S} be standard Brownian motion.
- Let $\mathbb{Z}_n(t) \equiv \mathbb{S}(nt)/\sqrt{2n\log\log n}$ for $0 \leq t \leq 1$.
- Let \mathcal{K} denote the set of absolutely continuous functions f on $[0, 1]$ with $f(0) = 0$ and $\int_0^1 [f'(t)]^2 dt \leq 1$; thus

$$\mathcal{K} \equiv \left\{ f \in C[0, 1] : f(0) = 0, f(t) = \int_0^t f'(s) ds, \int_0^1 [f'(s)]^2 ds \leq 1 \right\}.$$

Theorem 12.9.4: (Strassen) For almost all ω the sequence of functions $\{t \mapsto \mathbb{Z}_n(t, \omega) : n = 3, 4, \dots\}$ is relatively compact with limit set \mathcal{K} . That is

$$P \left(\overline{\lim}_n \|\mathbb{Z}_n - \mathcal{K}\| = 0 \right) = 1 \quad \text{and}$$

$$P \left(\bigcap_{f \in \mathcal{K}} [\underline{\lim}_n \|\mathbb{Z}_n - f\| = 0] \right) = 1.$$

We will write this conclusion symbolically as $\mathbb{Z}_n \rightsquigarrow \mathcal{K}$. By an embedding in Brownian motion we can also conclude that $\tilde{\mathbb{Z}}_n \equiv S_n(t)/\sqrt{2\log\log n}$ satisfies $\tilde{\mathbb{Z}}_n \rightsquigarrow \mathcal{K}$.

Corollary: Let $g : C[0, 1] \rightarrow \mathbb{R}$ be continuous with respect to the uniform metric $\|\cdot\|$. Then:

$$\limsup_n g(\mathbb{Z}_n) = \sup_{f \in \mathcal{K}} g(f) \quad a.s. \quad \text{and}$$
$$\liminf_n g(\mathbb{Z}_n) = \inf_{f \in \mathcal{K}} g(f) \quad a.s.$$

Example 1. Fix $t_0 \in (0, 1)$, and let $g(f) = f(t_0)$. Then

$$\limsup_n g(\mathbb{Z}_n) = \limsup_n \mathbb{Z}_n(t_0) = +\sqrt{t_0} \quad a.s. \quad \text{and}$$
$$\liminf_n g(\mathbb{Z}_n) = \liminf_n \mathbb{Z}_n(t_0) = -\sqrt{t_0} \quad a.s..$$

Proof: Now

$$g(f) = f(t_0) = \int_0^{t_0} f'(s) ds = \int_0^1 1_{[0,t_0]}(s) f'(s) ds,$$

so by the Cauchy - Schwarz inequality

$$|g(f)| \leq \left(\int_0^1 1_{[0,t_0]}(s) ds \right)^{1/2} \left(\int_0^1 [f'(s)]^2 ds \right)^{1/2} \leq t_0^{1/2} \cdot 1$$

so

$$\sup_{f \in \mathcal{K}} g(f) = \sqrt{t_0}$$

with equality when $f(t) = (t \wedge t_0) / \sqrt{t_0}$, $0 \leq t \leq 1$. Note that this f is linear on $[0, t_0]$ with slope $t_0^{-1/2}$, is constant on $[t_0, 1]$ at the height $\sqrt{t_0}$, and it satisfies

$$\int_0^1 [f'(s)]^2 ds = 1.$$

4. Boundary Crossing probabilities

Notation: Let \mathbb{S} denote standard Brownian motion starting from 0, and let \mathbb{U} denote a (standard) Brownian bridge process on $[0, 1]$, and let $Z \sim N(0, 1)$.

We will use the strong Markov property of \mathbb{S} repeatedly.

Application 1. Let $a, b > 0$, and let $\tau_b \equiv \inf\{t > 0 : \mathbb{S}(t) = b\}$. Then

$$\begin{aligned} P\left(\sup_{0 \leq t \leq a} \mathbb{S}(t) \geq b\right) &= P(\tau_b \leq a) = 2P(\mathbb{S}(a) \geq b) = P(Z \geq b/\sqrt{a}) \\ &= 2(1 - \Phi(b/\sqrt{a})). \end{aligned}$$

Proof: The first equality is obvious. The second inequality follows from a reflection argument:

$$\begin{aligned} P(\tau_b \leq a) &= P(\tau_b \leq a, \mathbb{S}(a) \geq b) + P(\tau_b \leq a, \mathbb{S}(a) < b) \\ &= P(\tau_b \leq a, \mathbb{S}(a) - \mathbb{S}(\tau_b) \geq 0) + P(\tau_b \leq a, \mathbb{S}(a) - \mathbb{S}(\tau_b) < 0) \\ &= P(\mathbb{S}(a) \geq b) + P(\tau_b \leq a) \cdot P(\mathbb{S}(\tau_b + (a - \tau_b)) - \mathbb{S}(\tau_b) < 0) \\ &= P(\mathbb{S}(a) \geq b) + P(\tau_b \leq a) \cdot (1/2). \end{aligned}$$

Rearranging this last equality yields the second equality of the claim. \square

Application 2. Fix $b \in \mathbb{R}$. Then $\tilde{\mathbb{S}}$ defined by

$$\tilde{\mathbb{S}}(t) = \begin{cases} \mathbb{S}(t), & t < \tau_b, \\ 2b - \mathbb{S}(t), & t \geq \tau_b \end{cases}$$

is a standard Brownian motion on $[0, \infty)$.

Proof: Consider the two processes $Y(t) \equiv \mathbb{S}(t)$, $0 \leq t \leq \tau_b$, and $Z(t) \equiv \mathbb{S}(\tau_b + t) - \mathbb{S}(\tau_b) = \mathbb{S}(\tau_b + t) - b$. By the strong Markov property of BM, Z is a Brownian motion independent of Y ; $-Z$ is also a Brownian motion independent of Y . Thus $(Y, Z) \stackrel{d}{=} (Y, -Z)$. Now $\varphi(Y, Z) = Y(t)1_{[t \leq \tau_b]} + (b + Z(t - \tau_b))1_{[t > \tau_b]}$ yields a continuous process with the same law as $\varphi(Y, -Z)$. But $\text{var}\phi(Y, Z) = \mathbb{S}$ and $\varphi(Y, -Z) = \tilde{\mathbb{S}}$. \square

Application 3. Let $M(t) \equiv \sup_{s \leq t} S(s)$. Then, for $b, y, t \geq 0$,

$$P(M(t) \geq b, S(t) \leq b - y) = P(S(t) \geq b + y).$$

Proof: With \tilde{S} as in Application 2 above,

$$\begin{aligned} P(M(t) \geq b, S(t) \leq b - y) &= P(M(t) \geq b, S(t) \leq b - y) \\ &= P(M(t) \geq b, \tilde{S}(t) \leq b - y) = P(M(t) \geq b, 2b - S(t) \leq b - y) \\ &= P(S(t) \geq b + y). \end{aligned}$$

□

Application 4. Let $\tau_{a,b} \equiv \inf\{t > 0 : S(t) \in (-a, b)^c\}$ for $a, b > 0$. Then for $b > 0$,

$$\begin{aligned} P(\|S\|_0^1 > b) &= P(\tau_{b,b} < 1) \\ &= 4 \sum_{k=1}^{\infty} P((4k-3)b < Z < (4k-1)b) \\ &= 1 - \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{(-1)^k}{2k+1} \exp\left(-\frac{(2k+1)^2\pi^2}{8b^2}\right). \end{aligned}$$

Proof: Let

$$\begin{aligned} A_+ &\equiv [\|S^+\| > b] = [S(t) > b \text{ for some } 0 \leq t \leq 1], \text{ and} \\ A_- &\equiv [\|S^-\| > b] = [S(t) < -b \text{ for some } 0 \leq t \leq 1]. \end{aligned}$$

Note that

$$[\|S\|_0^1 < b] = [A_+ \cup A_-]$$

where A_+ and A_- are not disjoint, so $P(\|S\| > b) \leq P(A_+) + P(A_-)$. Thus we need to consider

$$A_{+,-} = [S(t) < b \text{ for some } 0 \leq t \leq 1, \text{ then } S(t) < -b \text{ for some } t > \tau_b],$$

$$A_{-,+} = [S(t) < -b \text{ for some } 0 \leq t \leq 1, \text{ then } S(t) > b \text{ for some } t > \tau_b].$$

Note that

$$P(A_{+,-}) = P(\|S^+\| > 3b),$$

$$P(A_{-,+}) = P(\|S^-\| > 3b).$$

Thus ...

$$\begin{aligned}
P(\|S\|_0^1 > b) &= P(A_+) - P(A_{+,-}) + P(A_{+,-,+}) \cdots - \\
&\quad + P(A_-) - P(A_{-,+}) + P(A_{-,+,-}) \cdots - \\
&= 2 \left\{ P(A_+) - P(A_{+,-}) + P(A_{+,-,+}) - \cdots \right\} \\
&\quad \text{by symmetry} \\
&= 2 \sum_{k=1}^{\infty} (-1)^k \cdot 2P(S(1) > (2k+1)b) \\
&= 4 \sum_{k=1}^{\infty} P((4k-3)b < Z < (4k-1)b).
\end{aligned}$$

The last equality follows from the Poisson summation formula; cf. Feller, Volume II, pages 342-343, (X.5), and 629 - 630, IXI.5. Note that

$$\begin{aligned}
P(\|S\|_0^a \geq b) &\leq P(\sup_{0 \leq t \leq a} S(t) \geq b) + P(\sup_{0 \leq t \leq a} (-S(t)) \geq b) \\
&\leq 2P(S(a) \geq b) + 2P(S(a) \geq b) = 4P(S(a) \geq b) \\
&= 4P(Z \geq b/\sqrt{a}).
\end{aligned}$$

More boundary crossing probabilities:

Theorem 12.7.2: Let $c \geq 0$, $d > 0$. Then

$$P(S(t) \geq ct + d \text{ for some } t > 0) = e^{-2cd},$$

$$P(|S(t)| \geq ct + d \text{ for some } t > 0) = 2 \sum_{k=1}^{\infty} (-1)^{k+1} e^{-2k^2 cd}.$$

Proof: For any $\theta > 0$ the process

$$\begin{aligned} V(t) &\equiv \exp(\theta(S(t) - \theta^2 t/2)), \quad t \geq 0, \\ &= \exp(\theta(S(t) - \theta t/2)) \equiv \exp(\theta X(t)) \end{aligned}$$

is a martingale. Now define

$$\tau_{a,b} \equiv \inf\{t > 0 : X(t) \in (-a, b)^c\}.$$

Optional sampling gives

$$1 = EV(\tau_{a,b}) = P(X(\tau_{a,b}) = -a)e^{-\theta a} + P(X(\tau_{a,b}) = b)e^{\theta b},$$

so that

Thus

$$\begin{aligned}P(X(\tau_{a,b}) = b) &= \frac{1 - e^{-\theta a}}{e^{\theta b} - e^{-\theta a}} \\&\rightarrow e^{-\theta b} \quad \text{if } \theta > 0 \text{ and } a \rightarrow \infty \\&= e^{-2cd} \quad \text{if } \theta = 2c \text{ and } b = d.\end{aligned}$$

But we also have

$$\begin{aligned}P(X(\tau_{a,b}) = b) &\rightarrow P(X(t) \geq b \text{ for some } t \geq 0) \quad \text{as } a \rightarrow \infty \\&= P(S(t) - \theta t/2 \geq b \text{ for some } t \geq 0) \\&= P(S(t) \geq \theta t/2 + b \text{ for some } t \geq 0) \\&= P(S(t) \geq ct + d \text{ for some } t \geq 0) .\end{aligned}$$

if $\theta = 2c$ and $b = d$.

□

Theorem 12.7.3: Suppose that $b > 0$. Then

$$P(\|\mathbb{U}^\pm\| > b) = e^{-2b^2} \quad \text{for all } b > 0,$$

$$P(\|\mathbb{U}\| > b) = 2 \sum_{k=1}^{\infty} (-1)^{k+1} e^{-2k^2b^2}, \quad b > 0.$$

Proof: First, $\|\mathbb{U}^+\| \stackrel{d}{=} \|\mathbb{U}^-\|$. Furthermore

$$\begin{aligned} P(\|\mathbb{U}^+\| > b) &= P(\mathbb{U}(t) > b \text{ for some } 0 \leq t \leq 1) \\ &= P\left((1-t)\mathbb{S}\left(\frac{t}{1-t}\right) > b \text{ for some } 0 \leq t \leq 1\right) \\ &= P(\mathbb{S}(r) > b(1+r) \text{ for some } r \geq 0) \\ &= e^{-2b^2} \quad \text{by Theorem 12.7.2.} \end{aligned}$$

Similarly,

$$\begin{aligned} P(\|\mathbb{U}\| > b) &= P(|\mathbb{S}(r)| > b(1+r) \text{ for some } r > 0) \\ &= 2 \sum_{k=1}^{\infty} (-1)^{k+1} \exp(-2k^2b^2). \end{aligned}$$