

Statistics 521, Problem Set 5

Wellner; 10/26/2016

Reading: Shorack, PfS, Chapter 3, section 3.5, pages 52-63;
Shorack, PfS, Chapter 4, sections 4.1 - 4.4, pages 65-85.
Durrett, *Probability*, pages 412 - 416.

Due: Wednesday, November 2, 2016.

1. PfS, Exercise 3.4.2, page 48: Show that $\rho = 1$ if and only if $X - \mu_X = a(Y - \mu_Y)$ for some $a > 0$; and $\rho = -1$ if and only if $X - \mu_X = a(Y - \mu_Y)$ for some $a < 0$. Thus ρ measures linear dependence, not dependence.
2. PfS, Exercise 3.4.3, page 48: (Littlewood's inequalities) Let $\mu_r \equiv E|X|^r$. For $r \geq s \geq t \geq 0$ we have $\mu_r^{s-t} \mu_t^{r-s} \geq \mu_s^{r-t}$. In particular, $\mu_2^3 \leq \mu_1^2 \mu_4$.
3. Suppose that $\epsilon_1, \dots, \epsilon_n$ are i.i.d. random variables with $P(\epsilon_i = \pm 1) = 1/2$, and let $a_i \in \mathbb{R}$, $i = 1, \dots, n$. Khintchine's inequalities say that for each $p > 0$

$$A_p \left(\sum_{i=1}^n a_i^2 \right)^{1/2} \leq \left(E \left| \sum_{i=1}^n a_i \epsilon_i \right|^p \right)^{1/p} \leq B_p \left(\sum_{i=1}^n a_i^2 \right)^{1/2}.$$

for some constants A_p and B_p . Prove the above inequalities when $p = 1$.

Hint: The inequality on the right side is easy. Use the previous exercise to prove the inequality on the left side by showing that for $Z \equiv \sum_{i=1}^n a_i \epsilon_i$, we have $E|Z|^4 \leq 3(E(Z^2))^2$.

4. PfS, Exercise 3.5.3, page 55: Consider a probability measure P . (a) Let $Y \geq 0$ have df F . Show that $EY = \int_0^\infty P(Y \geq y)dy = \int_0^\infty [1 - F(y)]dy$. [Hint: prove the claimed formula for simple functions by summing by parts; and then the full claim follows from the MCT. A different proof to come later will use Fubini's theorem.]
(b) use the result of (a) to show that for $Y \geq 0$ and $\lambda \geq 0$ we have

$$\int_{[Y \geq \lambda]} Y dP = \lambda P(Y \geq \lambda) + \int_\lambda^\infty P(Y \geq y)dy.$$

Draw a picture to illustrate this.

(c) Suppose there is a $Y \in \mathcal{L}_1$ such that $P(|X_n| \geq y) \leq P(Y \geq y)$ for all $y > 0$ and all $n \geq 1$. Then use (b) to show that $\{X_n : n \geq 1\}$ is uniformly integrable.

5. (a) Show that if $|X_n| \leq Y$ and Y is integrable, then $\{X_n\}$ is uniformly integrable.

(b) Let $U \sim \text{Uniform}(0, 1)$, and let $X_n \equiv (n/\log n)1_{[0, 1/n]}(U)$ for $n \geq 3$. Show that $\{X_n\}$ is uniformly integrable and $\int X_n dP \rightarrow 0$ even though they are not dominated by any integrable rv Y .

(c) Let $Z_n = n1_{[0, 1/n]}(U) - n1_{[1/n, 2/n]}(U)$. Show that $\{Z_n\}$ is not uniformly integrable, but that $\int Z_n dP \rightarrow 0$.

6. **Optional bonus problem:** PfS, Exercise 3.4.6, page 50 (qualified by “for all $\epsilon \geq 1$ ”): Let $T \sim \text{Binomial}(n, p)$, so $P(T = k) = \binom{n}{k} p^k (1-p)^{n-k}$ for $0 \leq k \leq n$. The measure associated with T has mean np and variance $np(1-p)$. Then use inequality 4.6 with $g(x) = \exp(rx)$ and $r > 0$ to show that

$$P(T/n \geq p\epsilon) \leq \exp(-nph(\epsilon)), \quad \text{where } h(y) \equiv \epsilon(\log(y) - 1) + 1$$

for each $\epsilon > 1$. [Hint: It helps to use $T \stackrel{d}{=} \sum_1^n X_i$ where $X_i \sim \text{Bernoulli}(p)$ are independent, and then apply Theorem 7.1.1 (page 124).]