

## Statistics 394, Problem Set 1 Solutions

Wellner; 1/12/2000

1. K, 1.1, # 5, but assume that the dice are not fair; instead, if  $Y$  denotes the outcome of tossing a single die, assume that  $P(Y = i)$  is proportional to  $i$  for  $i = 1, \dots, 6$ . First, find these six probabilities. Then solve questions a(ii), a(iii), a(v), and b.

**Solution:** Since  $P(Y = i) = Ci$  for  $i = 1, \dots, 6$ , we must have

$$1 = \sum_{i=1}^6 P(Y = i) = C(1 + 2 + 3 + 4 + 5 + 6) = C6 \cdot 7/2 = 21 \cdot C,$$

and hence  $C = 1/21$ .

a(ii):

$$\begin{aligned} P(T = 6) &= P(Y_1 = 1)P(Y_2 = 5) + P(Y_1 = 2)P(Y_2 = 4) + P(Y_1 = 3)P(Y_2 = 3) \\ &\quad + P(Y_1 = 4)P(Y_2 = 2) + P(Y_1 = 5)P(Y_2 = 1) \\ &= \{1 \cdot 5 + 2 \cdot 4 + 3 \cdot 3 + 4 \cdot 2 + 5 \cdot 1\} / (21)^2 \\ &= 35 / (21)^2 \approx .0794. \end{aligned}$$

a(iii):

$$\begin{aligned} P(\text{double}) &= P(Y_1 = 1, Y_2 = 1) + P(Y_1 = 2, Y_2 = 2) + P(Y_1 = 3, Y_2 = 3) + P(Y_1 = 4, Y_2 = 4) \\ &= P(Y_1 = 1)P(Y_2 = 1) + P(Y_1 = 2)P(Y_2 = 2) + P(Y_1 = 3)P(Y_2 = 3) \\ &\quad + P(Y_1 = 4)P(Y_2 = 4) + P(Y_1 = 5)P(Y_2 = 5) + P(Y_1 = 6)P(Y_2 = 6) \\ &= \{1^2 + 2^2 + 3^2 + 4^2 + 5^2 + 6^2\} / (21)^2 \\ &= 91 / (21)^2 \approx .2063. \end{aligned}$$

a(v):

$$\begin{aligned} P(\text{at least one of } Y_1, Y_2 > 4) &= 1 - P(Y_1 \leq 4, Y_2 \leq 4) \\ &= 1 - P(Y_1 \leq 4)P(Y_2 \leq 4) = 1 - (10/21)^2 \approx .7732. \end{aligned}$$

b: With  $X = Y_1 + Y_2$ , the events  $[X = x]$  decompose as follows:

$$\begin{aligned} [X = 2] &= [Y_1 = 1, Y_2 = 1] \\ [X = 3] &= [Y_1 = 1, Y_2 = 2] \cup [Y_1 = 2, Y_2 = 1] \\ [X = 4] &= [Y_1 = 1, Y_2 = 3] \cup [Y_1 = 2, Y_2 = 2] \cup [Y_1 = 3, Y_2 = 1] \\ [X = 5] &= [Y_1 = 1, Y_2 = 4] \cup [Y_1 = 2, Y_2 = 3] \cup [Y_1 = 3, Y_2 = 2] \cup [Y_1 = 4, Y_2 = 1] \\ [X = 6] &= [Y_1 = 1, Y_2 = 5] \cup [Y_1 = 2, Y_2 = 4] \cup [Y_1 = 3, Y_2 = 3] \\ &\quad \cup [Y_1 = 4, Y_2 = 2] \cup [Y_1 = 5, Y_2 = 1] \\ [X = 7] &= [Y_1 = 1, Y_2 = 6] \cup [Y_1 = 2, Y_2 = 5] \cup [Y_1 = 3, Y_2 = 4] \\ &\quad \cup [Y_1 = 4, Y_2 = 3] \cup [Y_1 = 5, Y_2 = 2] \cup [Y_1 = 6, Y_2 = 1] \\ [X = 8] &= [Y_1 = 2, Y_2 = 6] \cup [Y_1 = 3, Y_2 = 5] \cup [Y_1 = 4, Y_2 = 4] \\ &\quad \cup [Y_1 = 5, Y_2 = 3] \cup [Y_1 = 6, Y_2 = 2] \\ [X = 9] &= [Y_1 = 3, Y_2 = 6] \cup [Y_1 = 4, Y_2 = 5] \cup [Y_1 = 5, Y_2 = 4] \cup [Y_1 = 6, Y_2 = 3] \\ [X = 10] &= [Y_1 = 4, Y_2 = 6] \cup [Y_1 = 5, Y_2 = 5] \cup [Y_1 = 6, Y_2 = 4] \\ [X = 11] &= [Y_1 = 5, Y_2 = 6] \cup [Y_1 = 6, Y_2 = 5] \\ [X = 12] &= [Y_1 = 6, Y_2 = 6]. \end{aligned}$$

Straightforward computation as in a(ii) above yields the following table:

$k$	2	3	4	5	6	7	8	9	10	11	12
$(21)^2 \cdot p_X(x)$	1	4	10	20	35	56	70	76	73	60	36
$p_X(x) \approx$	.002	.009	.023	.045	.079	.127	.159	.172	.166	.136	.082

2. K, 1.1, # 8, but do with *five* slips of paper numbered 1, 2, 3, 4, 5.

**Solution:** Now there are two draws, without replacement, from a hat containing slips numbered 1, 2, 3, 4, 5. Let  $X_1 \equiv$  the number drawn on the first draw, and  $X_2 \equiv$  the number drawn on the second draw. Then we have:

$$\text{a(i): } P(\text{slip 2 drawn}) = P([Y_1 = 2] \cup [Y_2 = 2]) = P(Y_1 = 2) + P(Y_2 = 2) = \frac{1}{5} + \frac{1}{5} = \frac{2}{5}.$$

a(ii):

$$\begin{aligned} P(\text{slip 3 drawn, slip 4 not drawn}) &= P((Y_1 = 3) \cap [Y_2 \neq 4]) \cup ([Y_1 \neq 4] \cap [Y_2 = 3]) \\ &= P([Y_1 = 3] \cap [Y_2 \neq 4]) + P([Y_1 \neq 3 \text{ or } 4] \cap [Y_2 = 3]) \\ &= \frac{1}{5} \cdot \frac{3}{4} + \frac{3}{5} \cdot \frac{1}{4} \\ &= \frac{6}{20}. \end{aligned}$$

a(iii):

$$\begin{aligned} &P(\text{max of two numbers is not more than 3}) \\ &= P(\max\{Y_1, Y_2\} \leq 3) = P(Y_1 \leq 3, Y_2 \leq 3) \\ &= \frac{3}{5} \cdot \frac{2}{4} = \frac{3}{10} = \frac{\binom{3}{2} \binom{2}{0}}{\binom{5}{2}}. \end{aligned}$$

b: If  $X = Y_1 + Y_2$ , then  $X$  has possible values 3, ..., 9, and we calculate, for example,

$$\begin{aligned} P(X = 6) &= P(Y_1 = 1, Y_2 = 5) + P(Y_1 = 2, Y_2 = 4) + P(Y_1 = 4, Y_2 = 2) + P(Y_1 = 5, Y_2 = 1) \\ &= \frac{1}{5} \cdot \frac{1}{4} + \frac{1}{5} \cdot \frac{1}{4} + \frac{1}{5} \cdot \frac{1}{4} + \frac{1}{5} \cdot \frac{1}{4} = 4/20 = 1/5. \end{aligned}$$

Thus we find:

$k$	3	4	5	6	7	8	9
$10 \cdot p_X(x)$	1	1	2	2	2	1	1

c(i):  $P(Y_1 = 3) = 1/5$ ,  $P(Y_2 = 3) = 1/5$ . To see the second probability, note that we can write it as

$$P(Y_2 = 3) = \sum_{j=1, j \neq 3}^5 P(Y_2 = 3 | Y_1 = j) P(Y_1 = j) = \frac{1}{4} \cdot \frac{1}{5} + \frac{1}{4} \cdot \frac{1}{5} + \frac{1}{4} \cdot \frac{1}{5} + \frac{1}{4} \cdot \frac{1}{5} = \frac{1}{5}.$$

3. K, 1.2, # 5, but assume that  $P(HEAD) = 2/3$ , not  $1/2$ .

**Solution:** (a). The possible outcomes are  $H, TH, TTH, TTTH, \dots$ ; thus a general outcome for the experiment consists of  $i - 1$  tails ( $T$ ), followed by a head ( $H$ ). Assuming that the probability of a head on each toss of the coin is  $2/3$ , these various outcomes have corresponding probabilities  $2/3, (1/3)(2/3) = 2/9, (1/3)^2(2/3), (1/3)^3(2/3), \dots$ . Thus the general outcome of  $i - 1$  tails followed by a head has probability  $(1/3)^{i-1}(2/3) = 2/3^i, i = 1, 2, 3, \dots$ . Note that these do sum to 1:

$$\sum_{i=1}^{\infty} (2/3)(1/3)^{i-1} = (2/3) \sum_{j=0}^{\infty} \frac{1}{3^j} = \frac{2/3}{1 - 1/3} = 1.$$

(b). Let  $Y \equiv$  the number of tosses needed to get a head ( $H$ ). Then by (a),  $P(Y = i) = (1/3)^{i-1}(2/3)$  for  $i = 1, 2, 3, \dots$ . Thus

$$\begin{aligned} P(\text{more than } k \text{ tosses needed to get a head}) \\ &= P(Y > k) = \sum_{i=k+1}^{\infty} (1/3)^{i-1}(2/3) \\ &= (1/3)^k(2/3) + (1/3)^{k+1}(2/3) + \dots \\ &= (1/3)^k(2/3) \{1 + (1/3) + (1/3)^2 + \dots\} \\ &= (1/3)^k(2/3) \frac{1}{1 - 1/3} = (1/3)^k. \end{aligned}$$

Thus  $P(Y > 2) = 1/9, P(Y > 3) = 1/27, P(Y > 4) = 1/81, \dots$ .

(c).

$$\begin{aligned} P(Y \text{ is odd}) &= P([Y = 1] \text{ or } [Y = 3] \text{ or } [Y = 5] \text{ or } \dots) \\ &= \sum_{j=1}^{\infty} (1/3)^{2j-2}(2/3) \\ &= (2/3) \sum_{m=0}^{\infty} (1/9)^m \\ &= (2/3) \frac{1}{1 - 1/9} = \frac{2}{3} \cdot \frac{9}{8} = \frac{3}{4}. \end{aligned}$$

4. K, 1.2, # 19, but assume that only 10 people favor A while the other 1235 favor B.

**Solution:** (a) The probability of getting at most one supporter of A in a sample of size 7 without replacement, if there are only 10 people favoring A in the population, is:

$$\begin{aligned} &\frac{\binom{10}{0} \binom{1235}{7}}{\binom{1245}{7}} + \frac{\binom{10}{1} \binom{1235}{6}}{\binom{1245}{7}} \\ &= \frac{1235}{1245} \cdot \frac{1234}{1244} \cdot \frac{1233}{1243} \cdot \frac{1232}{1242} \cdot \frac{1231}{1241} \cdot \frac{1230}{1240} \cdot \frac{1229}{1239} \\ &\quad + \frac{70}{1239} \cdot \frac{1235}{1245} \cdot \frac{1234}{1244} \cdot \frac{1233}{1243} \cdot \frac{1232}{1242} \cdot \frac{1231}{1241} \cdot \frac{1230}{1240} \\ &\approx .94498 + .05382 = .9988. \end{aligned}$$

(b) If the sampling is carried out with replacement, then the probability of at most one supporter of A in the sample is

$$\begin{aligned} & \binom{7}{0} \left(\frac{10}{1245}\right)^0 \left(\frac{1235}{1245}\right)^7 + \binom{7}{1} \left(\frac{10}{1245}\right)^1 \left(\frac{1235}{1245}\right)^6 \\ &= \left(\frac{1235}{1245}\right)^7 + 7 \left(\frac{10}{1245}\right) \left(\frac{1235}{1245}\right)^6 \\ &\approx .94511 + .05357 = .9987. \end{aligned}$$

Thus the probabilities under sampling with and without replacement differ only by about .0001.

5. K, 1.3, # 7: justify your answers.

**Solution:** Yes: c and g. Not possible: e and f. Possible to find a constant to make them densities: a,b,d. Here's why:

(a)  $x^4 \geq 0$  for  $0 \leq x \leq 1$  and  $\int_0^1 x^4 dx = (1/5)x^5|_0^1 = (1/5)$ , so  $f(x) = 5x^4 1_{[0,1]}(x)$  is a density function.

(b)  $x^4 \geq 0$  for  $0 \leq x \leq 2$ , and  $\int_0^2 x^4 dx = (1/5)x^5|_0^2 = 32/5$ , so  $f(x) = (5/32)x^4 1_{[0,2]}(x)$  is a density.

(c)  $x^{-2} \geq 0$  for  $x \geq 1$ , and  $\int_1^\infty x^{-2} dx = (-1)^{-1}x^{-1}|_1^\infty = 1$ . Thus  $f(x) = x^{-2} 1_{[1,\infty)}(x)$  is a density function.

(d)  $\sin(x) \geq 0$  for  $0 \leq x \leq \pi$  and

$$\int_0^\pi \sin(x) dx = -\cos(x)|_0^\pi = -(-1 - 1) = 2.$$

Thus  $f(x) = (1/2) \sin(x) 1_{[0,\pi]}(x)$  is a density function.

(e)  $\cos(x) < 0$  for  $\pi/2 < x \leq \pi$ , so it is not possible to normalize this function to be a density function.

(f)  $1/x \geq 0$  for  $x > 0$ , but

$$\int_0^\infty \frac{1}{x} dx = \log(x)|_1^\infty = \infty - (-\infty) = \infty.$$

Thus this  $f$  cannot be renormalized to be a density function.

(g)  $x \exp(-x^2/2) \geq 0$  for  $x \geq 0$ , and, by changing to the new variable  $y = x^2/2$ ,

$$\int_0^\infty x \exp(-x^2/2) dx = \int_0^\infty e^{-y} dy = -e^{-y}|_0^\infty = -(0 - 1) = 1.$$

Thus this function  $f$  is a density function on  $[0, \infty)$ .

6. Let  $\Omega = [0, 1)$  with the uniform distribution as in example 1.3.1. Define the random variable  $Z$  by  $Z(\omega) = -\log(\omega)$ . In part e, name the distribution.

**Solution:** (a)  $P(Z \leq 2) = P(-\log(\omega) \leq 2) = P(\log(\omega) \geq -2) = P(\omega \geq e^{-2}) = 1 - e^{-2}$ .

(b)  $P(Z > 3) = P(-\log(\omega) > 3) = P(\log(\omega) < -3) = P(\omega < e^{-3}) = e^{-3}$ .

(c)

$$\begin{aligned} P(3 \leq Z \leq 4) &= P(3 \leq -\log(\omega) \leq 4) = P(-3 \geq \log(\omega) \geq -4) \\ &= P(e^{-3} \geq \omega \geq e^{-4}) = e^{-3} - e^{-4}. \end{aligned}$$

(d)  $P(a \leq Z \leq b) = e^{-a} - e^{-b} = 1 - e^{-b} - (1 - e^{-a})$ .

(e) Thus  $Z$  has an exponential distribution with parameter 1.

7. K, 1.3, # 14: omit parts a,b,c,d. Instead find the probability of the event that the random point lies within  $1/2$  unit of the center *and* lies in the first quadrant of the disk.

**Solution:** Since the circle of radius 1 has area  $\pi(1)^2 = \pi$ , and the set of points in the first quadrant which are within  $1/2$  of the center of the disk has area  $(1/4)\pi(1/2)^2 = \pi/16$ , the probability of getting a point in the first quadrant within  $1/2$  of the center of the disk is  $(\pi/16)/\pi = 1/16$ .