

RANDOM VARIABLES

RVS ARE NUMERICAL OUTCOMES OF EXPERIMENTS

$(S, \mathcal{F}, P)$

S = REAL LINE, COMPLEX PLANE, N -SPACE

$\mathcal{F}$ = SUBSETS $\{ \underline{x} \leq x \}$

P = probability measure, parameterized by x ,
cdf

cdf $P(x)$

- $0 \leq P(x) \leq 1$
- $P(-\infty) = 0, P(+\infty) = 1$
- P MONOTONIC
NONDECREASING
- $P(x) = \int_{-\infty}^x p(x) dx$
- often useful to work
with $\bar{P}(x) = 1 - P(x)$

$$P(x) = \Pr[\underline{x} \leq x]$$

$$\bar{P}(x) = \Pr[\underline{x} > x]$$

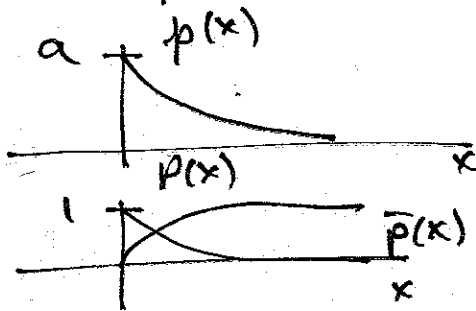
pdf $p(x)$

- $p(x) \geq 0$
- $\int_{-\infty}^{+\infty} p(x) dx = 1$
- $p(x) = \frac{d}{dx} P(x)$

$$p(x) dx = \Pr[x < \underline{x} \leq x + dx]$$

EX: exponential distribution $p(x) = a e^{-ax} u(x)$

$$a > 0$$



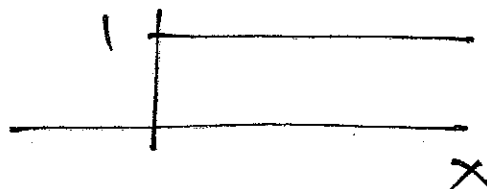
check NORMALIZATION

$$\begin{aligned} \int_{-\infty}^{+\infty} p(x) dx &= 1 \\ &= \int_0^{\infty} a e^{-ax} dx = -e^{-ax} \Big|_0^{\infty} = 1 \end{aligned}$$

SOME USEFUL FUNCTIONS

• UNIT STEP $u(x)$

$$u(x) = \begin{cases} 1 & x \geq 0 \\ 0 & x < 0 \end{cases}$$



• DELTA FUNCTION (CT) $\delta(x)$

FOR A CONTINUOUS $f(x)$

$$\int_{-\infty}^{\infty} \delta(x-x_0) f(x) dx = f(x_0)$$

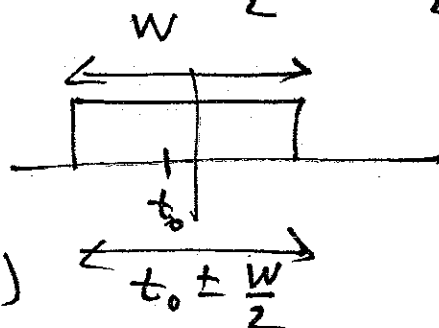
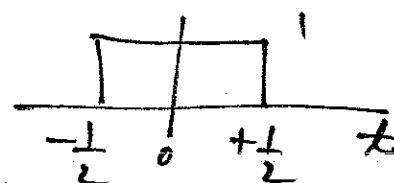


• RECTANGLE $\text{RECT}(x)$

$$\text{RECT}\left(\frac{x-x_0}{W}\right)$$

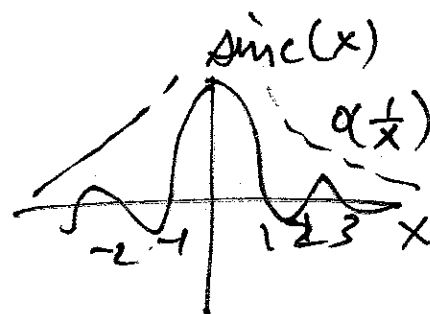
x_0 = center of mass

W = TOTAL (2-sided) WIDTH



• SINC I DEFINE

$$\text{sinc}(x) = \frac{\sin(\pi x)}{\pi x}$$



• zero crossings at $x = \pm 1, \pm 2, \dots$ (integers)

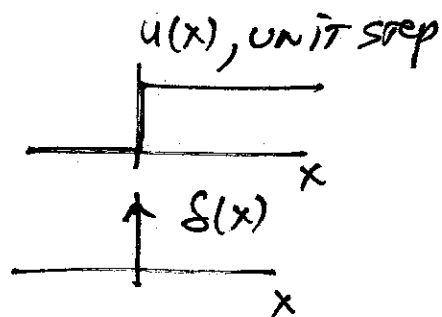
• $\text{sinc}(0) = 1$ (use L'Hôpital's rule)

PDFS

DISCRETE RVS

$\approx$

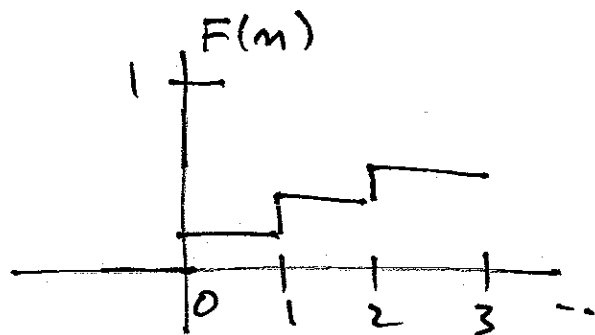
Recall $\frac{d}{dx} u(x) = \delta(x)$



Since $\int_{-\infty}^x \delta(x) dx = \begin{cases} 1 & x > 0 \\ 0 & x < 0 \end{cases}$

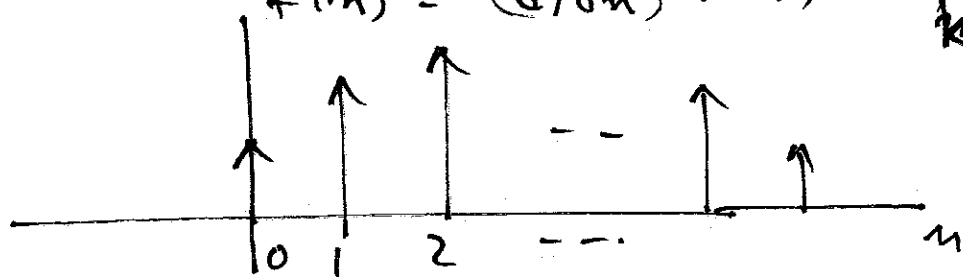
Consider the RV $\approx$, $\sum_{n=0}^{\infty} p_n = 1$
 $p_n = \text{PR}[\approx = n] = \frac{\lambda^n e^{-\lambda}}{n!}$ for some $\lambda > 0$

POISSON RV, $n! = n(n-1) \dots 3 \cdot 2 \cdot 1$
 $= n \text{ FACTORIAL}$



STAIRCASE

$$f(n) = (d/dn) F(n) = \sum_{k=0}^{\infty} p_k \delta(k-n)$$



We use the Dirac Delta to mark the location of $\approx$.

CHECK $f(n) \geq 0$ ✓

$$\int_{-\infty}^{\infty} f(n) dn = 1$$

A RV whose pdf contains only δ 's is called discrete.
Poisson is very useful in telecommunications.

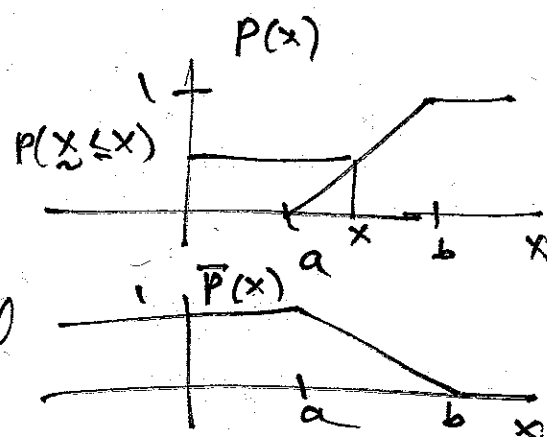
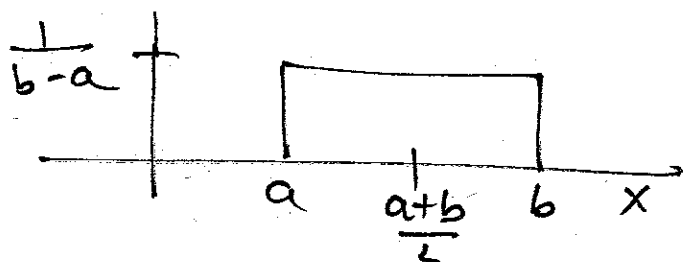
EX

UNIFORM DISTRIBUTION

$$X \sim \text{unif}(a, b)$$

$$b > a$$

$$p(x) = \frac{1}{b-a} \text{rect} \left(\frac{x - \frac{a+b}{2}}{b-a} \right)$$

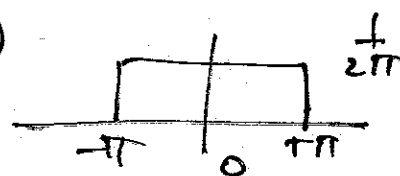


We say X is uniformly distributed on $(a, b]$.

Typical example

$$\Theta \sim \text{unif}(-\pi, +\pi)$$

is a statistical model for a random phase angle.

EX: GAUSSIAN OR NORMAL

$$X \sim N(\mu, \sigma^2)$$

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$P(x) = \int_{-\infty}^x p(x) dx = \int_{-\infty}^x \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(u-\mu)^2}{2\sigma^2}} du$$

$$\text{let } t = \frac{u-\mu}{\sigma} \quad dt = \frac{du}{\sigma}$$

$$P(x) = \Pr[X \leq x] = \int_{-\infty}^{\frac{x-\mu}{\sigma}} \underbrace{\frac{1}{\sqrt{2\pi}} e^{-t^2/2}}_{p(t)} dt = \text{ERF}\left(\frac{x-\mu}{\sigma}\right)$$

here $t \sim N(0, 1)$ since

$$p(t) = \frac{1}{\sqrt{2\pi}} e^{-t^2/2}$$

So if $t \sim N(0, 1)$ $X = \sigma t + \mu \sim N(\mu, \sigma^2)$

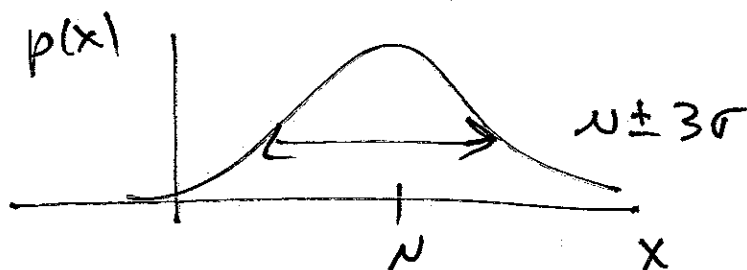
GAUSSIAN RV

A continuous RV X with pdf

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[-\frac{(x-\mu)^2}{2\sigma^2} \right]$$

is said to have $X \sim N[\mu, \sigma^2]$

A NORMAL or gaussian distribution with mean μ and variance σ^2 .



$$\int_{-\infty}^{\infty} p(x) dx = 1$$

$P(x) = \int_{-\infty}^x p(x) dx$ is a tabulated integral or can be done numerically.
is the cdf. (ERROR FUNCTION)

THIS IS THE MOST IMPORTANT pdf IN signal processing. DUE TO THE CENTRAL LIMIT THEOREM, THE BELL SHAPED CURVE IS MOST IMPORTANT IN STATISTICS.

MATLAB: `RANDN (NR, NC)` GENERATES A SAMPLE OF $NR \cdot NC = N$ $N(0,1)$ RVS ORGANIZED IN A MATRIX OF NR ROWS, NC COLUMNS.

`RAND (MR, MC)` IS FOR $UNIF(0,1)$ RVS.

- EXAMINE some iid time series AND histograms using MATLAB.
- Develop a plot of pdf, cdf and tabulated values using `ERFC.M`

MATHEMATICAL ASIDE ON THE GAUSSIAN

Prove THAT
$$I = \int_{-\infty}^{\infty} \frac{e^{-x^2/2}}{\sqrt{2\pi}} dx = 1$$

P.F. Let
$$I^2 = I \times I$$

$$= \int_{-\infty}^{\infty} e^{-x^2/2} \frac{dx}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-y^2/2} \frac{dy}{\sqrt{2\pi}}$$

$$= \frac{1}{2\pi} \iint_{-\infty}^{\infty} e^{-\frac{(x^2+y^2)}{2}} dx dy$$

Let $r^2 = x^2 + y^2$ $\tan(\theta) = y/x$
 $dx dy = r dr d\theta$ in polar coordinates

$$I^2 = \frac{1}{2\pi} \int_0^{\infty} \int_0^{2\pi} e^{-\frac{r^2}{2}} r dr d\theta$$

$$= \frac{1}{2\pi} \times 2\pi \times \left[-e^{-r^2/2} \right]_0^{\infty} = 1$$

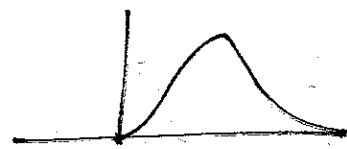
$\therefore I^2 = 1 \rightarrow I = 1$ ✓

evaluate ① $\int_0^{\infty} e^{-t^2/2} dt = \frac{\sqrt{2\pi}}{2}$; ② $\int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \frac{dx}{\sqrt{2\pi\sigma^2}} = 1$

③ $\int_{-\infty}^{\infty} e^{-at^2} dt = \int_{-\infty}^{\infty} e^{-\frac{t^2}{2\sigma^2}} \frac{dt}{\sqrt{2\pi\sigma^2}} \Big|_{a = \frac{1}{2\sigma^2}}$
 ④ $\int_0^{\infty} e^{-at^2} dt = \frac{1}{2} \sqrt{\pi/a} = \sqrt{\frac{2\pi}{a}} = \sqrt{\frac{\pi}{a}}$

RAYLEIGH

$$p(r) = \frac{r}{\sigma^2} e^{-r^2/2\sigma^2} u(r)$$



$$P(r) = 1 - e^{-r^2/2\sigma^2}$$

$$\bar{P}(r) = e^{-r^2/2\sigma^2}$$

$$P(r) = \int_0^r \frac{r}{\sigma^2} e^{-r^2/2\sigma^2} dr$$

let $w = r^2$
 $dw = 2r dr$

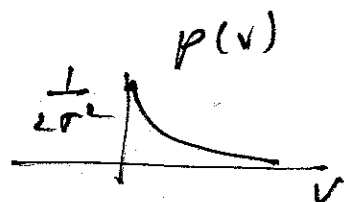
$$P(r) = \frac{1}{2\sigma^2} \int_0^{r^2} e^{-w/2\sigma^2} dw = \int_0^{r^2} a e^{-aw} dw$$

[Recall exponential $a e^{-ax} u(x)$.]

$$a = \frac{1}{2\sigma^2}$$

$$\therefore P(R < r) = P(X < r^2)$$

where $V \sim \text{exponential}$



THIS SUGGESTS

$$V \stackrel{\sim}{=} R^2$$

$$R \stackrel{\sim}{=} \sqrt{V}$$

WHICH IS TRUE.

AND WE NEED TO CONSIDER TRANSFORMATIONS OF IRV.

How does the pdf change?

If $X \sim (p, P, \bar{P}) \leftarrow \text{NOTATION}$

and $g(\cdot)$ is given, what is $(f, F, \bar{F})$

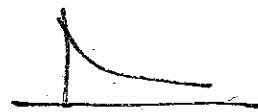
for $Y = g(X)$

g is the TRANSFORMATION.

EXAMPLES OF NORMALIZATION

EXPONENTIAL

$$(1) \quad f(t) = c e^{-at} \quad t > 0$$



$$\text{Then } \int_0^{\infty} f(t) dt = 1$$

$$c \int_0^{\infty} e^{-at} dt = \frac{c}{a} = 1 \quad \underline{c=a}$$

and $f(t) = a e^{-at}$ is a p.d.f

$$E(t) = \frac{1}{a} \quad (\text{CHECK THIS})$$

(2) MIXTURES OF EXPONENTIALS

$$\text{Let } a_1, a_2, \dots, a_k, \dots, a_n > 0$$

$$\text{and } f(t) = \sum_{k=1}^n w_k a_k e^{-a_k t}$$

UNDER WHAT CONDITIONS OF $\{w_k\}$ IS $f(t)$ a p.d.f.

$$(1) \quad \int_0^{\infty} f(t) dt = 1$$

$$(2) \quad f(t) \geq 0 \text{ everywhere}$$

$$\sum_{k=1}^n w_k \int_0^{\infty} a_k e^{-a_k t} dt = \sum_{k=1}^n w_k = 1$$

$$\therefore (i) \quad w_k \geq 0 \quad k=1, 2, \dots, n$$

$$\text{and (ii) } \sum_{k=1}^n w_k = 1$$

THESE WEIGHTS
MUST SUM TO 1

THE GAMMA FUNCTION

$$\Gamma(x) := \int_0^{\infty} t^{x-1} e^{-t} dt \quad ; \quad 0 < x$$

- GENERALIZES THE FACTORIAL TO NON INTEGER n

$$n! = \Gamma(n+1) = \int_0^{\infty} t^n e^{-t} dt$$

- VERY USEFUL IN PROBABILITY FOR NORMALIZATION INTEGRALS.

- M-FILE IN MATLAB GAMMA.M

- SIMILAR TO A LAPLACE TRANSFORM $\{ \mathcal{L}[t^{x-1}] \}$

- EVALUATION FOR INTEGER $x-1 = n$

$$\int_0^{\infty} t^n e^{-t} dt = \int_0^{\infty} t^n e^{-st} dt \Big|_{s=+1}$$

(CHECK THIS YOURSELF)

$$= (-1)^n \left(\frac{d}{ds} \right)^n \int_0^{\infty} e^{-st} dt \Big|_{s=+1}$$

$$= (-1)^n \left(\frac{d}{ds} \right)^n \frac{(-1)^n n!}{s^{n+1}} \Big|_{s=+1} = n!$$

- EXAMPLE IN NORMALIZATION

$$f(t) = C t^a e^{-t^b} \quad t > 0$$

DETERMINE C

NORMALIZATION

$$C \Rightarrow \int_0^{\infty} f(t) dt = 1$$

and change variable

NORMALIZATION EXAMPLE

$$f(t) = C t^a e^{-\frac{t}{b}} \quad ; \quad t > 0, f(t) \geq 0$$

$$\int_0^{\infty} f(t) dt = 1 \quad \text{condition for a pdf}$$

$$= C \int_0^{\infty} t^a e^{-\frac{t}{b}} dt$$

Let $u = \frac{t}{b}$
 $du = \frac{1}{b} dt$
 $t = u b$
 $dt = b du$

$$= C \int_0^{\infty} (u b)^a e^{-u} \frac{1}{b} du$$

$$= \frac{C}{b} \int_0^{\infty} u^{\frac{a+1}{b}-1} e^{-u} du$$

$$1 = \frac{C}{b} \Gamma\left(\frac{a+1}{b}\right)$$

$$\left[\text{since } \Gamma(x) = \int_0^{\infty} u^{x-1} e^{-u} du \right]$$

SOLVE

$$C = \frac{b}{\Gamma\left(\frac{a+1}{b}\right)}$$

and

$$f(t) = \frac{b}{\Gamma\left(\frac{a+1}{b}\right)} t^a e^{-\frac{t}{b}} \quad t > 0$$

is a pdf for $b > 0$

$$\frac{a+1}{b} \neq 0, -1, -2, \dots$$

THIS IS CALLED THE GENERALIZED GAMMA PDF
 AND USUALLY WRITTEN

$$f_{\text{GR}}(*; a, b, k) = \frac{C(\frac{t}{a})}{b^{k-1} - (t/a)^b} e^{-\frac{t}{a}} \quad t > 0$$