Initial Review of *Maximum Likelihood Estimation of Misspecified Models* by Halbert White

Jim Harmon

University of Washington, Department of Statistics

April 17, 2012

Jim Harmon (University of Washington, Depa

MLE of Misspecified Models

April 17, 2012 1 / 16

Maximum Likelihood Estimation of Misspecified Models

by Halbert White

Econometrica, Vol. 50, No. 1 (Jan., 1982), pp. 1-25

Jim Harmon (University of Washington, Depa

MLE of Misspecified Models

April 17, 2012 2 / 16

Obligatory Demotivational Quote

"This is a very frustrating paper." - V. Minin, 2012



• • • • • • • • • • • •

Correct model specification = efficient inference

Especially helpful for small sample size

But how do you **know**?

Answers questions about the following in a unified framework:

- does MLE converge? (interpretation?)
- if yes, is MLE asymptotically normal?
- can properties of MLE determine model truth?

That Which Hath Come Before

Consistency of MLE estimator

Interpretation of such

Asymptotic normality

Inference

White's real contributions:

- A unified approach
- More restrictive, yet more intuitive and easily checked conditions
- Some new statistics and tests

Theorem

Given assumptions, $\sqrt{n}(\hat{\theta}_{MLE} - \theta_{KLIC}) \rightarrow_d N(0, \mathbf{C}(\theta_{KLIC}))$

- $\hat{\theta}_{MLE}$ converges to the Kullback-Leibler Information Criterion-minimizing parameter
- Form of the asymptotic covariance matrix?

A Sandwich Estimator!



Jim Harmon (University of Washington, Depa

MLE of Misspecified Models

April 17, 2012 9 / 16

Covariance Matrix Formula

 $U_t = \text{data vectors}$ $\theta = \text{parameter vector of model family}$ $f(U_t, \theta) = \text{model density function}$ $\mathbf{A}(\theta) = E\left[\frac{\partial^2 \log(f(U_t, \theta))}{\partial \theta_i \partial \theta_j}\right]$ $\mathbf{B}(\theta) = E\left[\frac{\partial \log(f(U_t, \theta))}{\partial \theta_i}\frac{\partial \log(f(U_t, \theta))}{\partial \theta_j}\right]$ $\mathbf{C}(\theta) = \mathbf{A}(\theta)^{-1}\mathbf{B}(\theta)\mathbf{A}(\theta)^{-1}$

Jim Harmon (University of Washington, Depa

April 17, 2012 10 / 16

▲□▶ ▲□▶ ▲□▶ ▲□▶ □ のの⊙

Theorem

Given some assumptions, the elements of ${\bf A}$ and ${\bf B}$ can be used to derive a

statistic that indicates model misspecification

Idea: In a correct model $\mathbf{A} + \mathbf{B} = \mathbf{0}$.

The further you are from this, the more likely your model is misspecified.

Sensitive to misspecifications which invalidate usual inference.

Some Mathy Specifics

 θ is a *p*-dimensional vector.

$$\begin{split} d_l(U_t,\theta) &= \partial \log(f(U_t,\theta))/\partial \theta_i \cdot \partial \log(f(U_t,\theta))/\partial \theta_j \\ &+ \partial^2 \log(f(U_t,\theta))/\partial \theta_i \partial \theta_j \\ \dim(d) &= q \times 1 \text{ with } q \leq p(p+1)/2 \\ \nabla D(\theta) &= E[\partial d(U_t,\theta)/\partial \theta_k] \\ W &= d(U_t,\theta) - \nabla D(\theta) \mathbf{A}(\theta)^{-1} \nabla \log(f(U_t,\theta)) \\ \mathbf{V}(\theta) &= E[W \cdot W^T] \end{split}$$

Theorem

Given some assumptions, $\sqrt{n}D_n(\hat{\theta}_n) \rightarrow_d N(0, \mathbf{V}(\theta_0)); \mathbf{V}_n(\hat{\theta}_n) \rightarrow_{a.s.} \mathbf{V}(\theta_0);$ $\Im_n = nD_n(\hat{\theta}_n)'\mathbf{V}_n(\hat{\theta}_n)^{-1}D_n(\hat{\theta}_n) \rightarrow_d \chi_q^2$ Two more tests for model misspecification.

Sensitive to inconsistent estimator misspecifications.

I'm still working these out.

3

What Do We Do With All This???

THE METHOD

- Step 1: Perform first test.
- Step 2a: If you "do not reject", MLE away!
- Step 2b: If you "reject", perform the other two tests
- Step 3a: If you "do not reject", use sandwich inference
- Step 3b: If you "reject", reconsider your model choice.

Review Proofs

Simulation Study

Study with Actual Data??

3

The End

(What the title says)

Jim Harmon (University of Washington, Depa

MLE of Misspecified Models

April 17, 2012 16 / 16

3

<ロ> (日) (日) (日) (日) (日)