Sparse Permutation Invariant Covariance Estimation: Motivation, Background and Key Results

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Sparse permutation invariant covariance estimation

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- Reductionism vs. Dynamism
- Gene-gene interaction networks
- High-dimensional setting, i.e. n << p
- Two key questions:
 - Which gene products are directly dependent (yes/no for each pair of genes)?
 - What is the strength and direction of this dependence (numeric for each pair of genes)?
- Roadmap for lab research

The Gaussian graphical models approach

• Multivariate normality assumption (with standardization)

 $ec{X} \sim N_{
m p}(0,\Sigma)$

• For the ith and jth elements of \vec{X} , we know:

$$\Sigma_{i,j} = cov(X_i, X_j)$$

By normality, a zero here implies indepedence, but what does a non-zero imply?

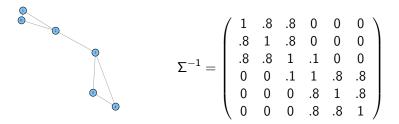
• Consider then $\Omega = \Sigma^{-1}$. A result from Lauritzen (1996):

 $\Omega_{i,j} = 0 \Leftrightarrow X_i$ and X_j are conditionally independent

A non-zero here implies dependence conditional on all other variables

The Gaussian graphical models approach

- The inverse covariance matrix (Question 2) can be thought of as implying a graphical model (Question 1)
- Hence the graph reflects conditional dependence



Answering question 1:

- Multiple testing procedures (Drton and Perlman, 2008)
- Multiple regressions (Meinshausen and Buehlmann, 2006)
- Graphical Lasso aka GLASSO (Friedman et al., 2007)
- Thresholding the sample covariance matrix

Answering question 2 (and therefore also 1):

- Sample covariance, inconsistent in high-dimensional setting (Johnstone, 2001)
- Shrink eigenvalues of covariance matrix (Ledoit and Wolf, 2003)
- Add structure to the covariance matrix, e.g. banding think AR-1. (Bickel and Levina, 2008).
- ℓ_q penalized likelihood (d'Aspremont et al. and Yuan and Lin)

The sparse permutation invariant covariance estimator aka SPICE

$$\hat{\Omega}_{\lambda} = \arg\min_{\Omega \succ 0} \{ tr(\Omega \hat{\Sigma}) - \log det(\Omega) + \lambda |\Omega^{-}|_{1} \}$$

where :

• $\Omega = \Sigma^{-1}$

•
$$\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^{n} (X_i - \bar{X}) (X_i - \bar{X})^T$$

- $\Omega^- = \Omega diagonal(\Omega)$
- λ is the tuning parameter

• Assumptions:

- the maximum and minimum of the eigenvalues of the true covariance matrix are bounded
- the number of non-zero inverse covariance elements are bounded (sparsity)

$$\frac{\lambda}{\sqrt{\left(\frac{\log p}{n}\right)}} \to c, \text{ where c is a constant}$$

• Theorem 1:

$$\|\hat{\Omega}_{\lambda} - \Omega_0\|_F = O_P(\sqrt{rac{(p+s)\log p}{n}})$$

• Theorem 2:

$$\|\Omega_{\lambda} - \Omega_0\| = O_P(\sqrt{rac{(s+1)\log p}{n}})$$

• The algorithm uses the Cholesky decomposition to solve the minimization problem, i.e. that positive definite matrices can be written as:

$$\hat{\Omega}_{\lambda} = T^T T$$

where T is a lower triangular matrix

- The three terms from the SPICE estimator can then be rewritten as f(T)
- A quadratic approximation of the penalty term is used and each element of T is minimized one at a time (cyclical coordinate descent)
- Convergence because of cyclical coordinate descent and smooth functions (Bazaraa, 2006)

Simulation study

- Uses four different inverse covariance matrix generating mechanisms (all sparse)
- Compares sample inverse covariance matrix and Ledoit-Wolf estimator using Kullback-Leibler loss and true positive and true negative rates

Colon tumor classification example

- Compares classification error rates
- Uses Naive Bayes, Ledoit-Wolf and SPICE with three ways of deriving the tuning parameter

This quarter

- Prove key results 1, 2, 3 (potentially starting with 3)
- Linear algebra review (and new learning)
- Impact of standardization of covariance or inverse covariance matrix

In the future

- Nearly all approaches make the MVN assumption, if this does not hold what do the results mean if we use this or similar methods?
- Contrast the algorithm in this paper with the GLASSO algorithm
- Applied papers review
- Penalization requires training set to determine tuning parameter value