Looking at the Other Side of Bonferroni

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Multiple Testing: Control the Type I Error Rate

- When analyzing genetic data, one will commonly perform over 1 million (and growing) hypothesis tests.
- In categorical data analysis, one may want to test all pairwise combinations.
- How do we ensure we are properly controlling for the number of false rejections?

2.5 Million Hypothesis Tests



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type I error $\mathbb{P}(\text{reject } H_0 | H_0 \text{ is true}) \leq \alpha$

family-wise error rate $FWER = \mathbb{P}(\# \text{ false pos} \ge 1)$ This is the probability of one or more false positives.

per family error rate $PFER = \mathbb{E}(\# \text{ false pos})$ This is the expected number of false positives.

false discovery rate $FDR = \mathbb{E}(\# \text{ false pos/total } \# \text{ rejected})$ This can be thought of as the average proportion of null hypotheses that are falsely rejected.

How it all fits together

	decide true	decide false	
H ₀ true	U	V	<i>m</i> 0
H_0 false	R	5	$m-m_0$
	m-T	Т	т

- V denotes a type I error.
- The FWER is $\mathbb{P}(V \ge 1)$.
- The PFER is $\mathbb{E}(V)$.
- The FDR is $\mathbb{E}(V/T)$.

- Define A to be the acceptance interval for some summary statistic V_g that gives the desired α risk for one hypothesis test.
- Then, define a rule for all $1 \le g \le m$ such that

 $\begin{array}{ll} \text{if } V_g \in \mathcal{A}, & \text{ can't reject } H_0 \\ \text{if } V_g \notin \mathcal{A}, & \text{ reject } H_0 \end{array}$

► For some test where we know the null hypothesis is true, the probability that our summary measure V_g will fall in the acceptance interval A is

$$\mathbb{P}(V_g \in \mathcal{A}|g \in \mathcal{T}) = \int_{\mathcal{A}} f_0(v) dv$$

= $1 - \alpha$

for all g, where f_0 is the null pdf, and \mathcal{T} is the set of indices of the true null hypotheses.

► We choose A to be the smallest interval that satisfies the above equation.

▶ We can specify the type I error probability for any g to be

$$\alpha = \mathbb{P}(V_g \notin \mathcal{A}|g \in \mathcal{T})$$

Then, we can also see that

- FWER $/m_0 \leq \alpha$.
- The unadjusted significance level can be defined by the desired family wise error rate divided by the total number of tests for which the null hypothesis is true.

- ► If we define A as shown on the previous slide such that FWER $\leq m_0 \alpha$ then
- ► FWER ≤ α, so our definition is consistent with the desired objective.

Finally, we get that

$$PFER = \mathbb{E}(V)$$
$$= \alpha m_0$$
$$\geq FWER$$
$$PFER/m_0 = \alpha$$

where V is the number of false positives, α is the overall significance level and m_0 is the number of true null hypotheses.

The expected number of false positives, the PFER, is equal to the type I error rate divided by the true number of null hypotheses. Bonferroni and Benjamini-Hochberg (BH) procedures

Bonferroni correction calculates

$$\alpha^* = \alpha/m$$

and controls the type I error rate.

BH correction orders the *p*-values in decreasing order, and for each *i* starting at the largest value, finds the point at which

$$p_{(i)} \leq \frac{\alpha i}{m}$$

and controls the FDR.

How does dependence change the FWER?



- Dependence makes the actual type I error less than desired.
- As m and/or ρ increases, this becomes worse.

Plot courtesy of Ken Rice

Simulation Studies

- Simulate 1255 gene expression values, measured for 50 individuals.
- 2 measurements per individual where 125 of the 1255 genes have a different mean.
- Generate a *p*-value for each gene from a standard t-test; 125 of them should be significant.
- Count the number of rejections when using the both the Bonferroni and BH procedures.

Next Steps

With these results, I aim to show when choosing the error rate thresholds appropriately, the Bonferroni and BH procedures are comparable. This will be reproducing the results from the paper [1].

References



Alexander Gordon, Galina Glazko, Xing Qiu, and Andrei Yakovlev.

Control of the mean number of false discoveries, Bonferroni and stability of multiple testing.

The Annals of Applied Statistics, 1(1):179–190, 2007.