Analysis of Longitudinal Data

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LDA Progress!

- During the last couple of decades statistical methods have been developed (ie. LMM, GEE) that can analyze longitudinal data with:
  - Unequal number of observations per person \( (n_i) \)
  - Unequally spaced observations \( (t_{ij}) \)
  - Time-varying covariates \( (x_{ij}) \)

- Regression questions:
  \[
  \mu_i(t) = E[Y_i(t) \mid X_i(t)]
  \]

- **Q:** When should we directly apply these now standard longitudinal methods to data with the features listed above?
Session Eight Outline

- **Examples**
  - Cystic Fibrosis Foundation (CFF)
  - Maternal Stress and Child Morbidity (MSCM)
  - United States Renal Data System (USRDS)

- **Time-varying Covariate Processes**
  - Exogenous
    - Lagged covariates
  - Endogenous
    - Fixed vs Dynamic exposure

- **Analysis with Death**
  - Specification of model
  - Inference
Repeated Measures Data

Cystic Fibrosis Data

- $N = 23,530$ subjects, $4,772$ deaths, 1986-2000
- $n = 160,005$ longitudinal observations
- Longitudinal measurements: FEV1, weight, height
- Goal: identify factors associated with decline in pulmonary function.
- (Another Goal: predict mortality; transplantation selection)
CFF Survival

Time

Survival

0 10 20 30 40 50
0.00.20.40.60.81.0

CFF Survival

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Example: Scientific Goals & CF

- Parad RB, Gerard CJ, Zurakowski D, Nichols DP, Pier GB
  “Pulmonary outcome in cystic fibrosis is influenced primarily by mucoid Pseudomonas aeruginosa infection and immune status and only modestly by genotype.”

- Variables:
  - Measurement time: $t_{ij}$
  - Pulmonary function: $Y_i(t_{ij})$
  - Time-dependent covariate: $X_i(t_{ij})$ – infection status
  - Death: $D_i(t)$ counting process for $T_i$
CFF Data and Visit Times

Age−Age0

Subject
0 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15

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CFF Data and Visit Times -- CASES
Example 2: Time-dependent covariates

- daily indicators of stress (maternal), and illness (child)
- primary outcome: illness, utilization
- covariates: employment, stress
- \textbf{Q}: association between employment, stress and morbidity?
- \textbf{Q}: Does stress cause morbidity?
Fig. 1. Determinants of episodic illness care utilization.
USRDS Data: Safety of ESAs?

- End Stage Renal Disease (ESRD)
  - Poor kidney function
  - Dialysis
  - Fail to stimulate formation of red blood cells
- Epoetin
  - Anemia treatment
  - $3 billion Medicare / year
- Studies show an association between high dose and risk of death
  - Adverse outcomes?
  - Confounding by indication?
USRDS Dialysis Data

ID = 69366

Event dose

Hematocrit

ID = 69366

Event dose

Hematocrit
USRDS Dialysis Data

ID = 71650

[Graphs showing EpoDose and Hematocrit over time]

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The Processes

Primary Process

\[ Y_i(t) \] The response process

Secondary Processes

\[ X_i(t) \] The covariate process
\[ S_i(t) \] The scheduling process (not today)
\[ R_i(t) \] The recording process
\[ D_i(t) \] The death process
\[ B_i(t) \] The birth process (not today)
LDA and Regression

- Most statistical representations focus on discussion of

\[ \mu_i(t) = E[Y_i(t) | X_i(t)] \]

- But what about the other processes? Do we mean:

CFF : \( E[Y_i(t) | X_i(t), X_i(s), S_i(t) = 1, R_i(t) = 1, D_i(t) = 0] \)

USRDS : \( E[Y_i(t) | X_i(t), X_i(s), R_i(t) = 1, D_i(t) = 0] \)

MSCM : \( E[Y_i(t) | X_i(t), X_i(s), R_i(t) = 1] \)
Motivation: Hospitalization and EPO Dose?

- **Background:**
  - **NEJM – November 2006**
    * RCTs target high versus low hemoglobin
    * Higher target $\rightarrow$ higher Epo dose
    * Higher target associated with AEs
  - **FDA – March 2007**
    Issued a “black box warning” which indicated that aggressive use of erythropoiesis-stimulating agents to raise hemoglobin to a target of 12 g/dL or higher was associated with “serious and life-threatening side-effects and/or death.”

- **General Question:**
  - **Q**: Are higher doses of EPO associated with greater rates of adverse events such as hospitalization?
Motivation: Full Data History

• **Regression:**

\[ E[\text{Hosp}(t) \mid \text{Dose}(t - 1), \text{Dose}(t - 2), \ldots, X] \]

• **Statistical Issues:**

▷ What aspects of exposure history are associated with current hosp?

▷ What is the role of the outcome history \( \text{Hosp}(t - 1), \text{Hosp}(t - 2), \ldots \)?

▷ What is the role of intermediate history \( \text{Hem}(t - 1), \text{Hem}(t - 2), \ldots \)?
Time-dependent Covariates: Lagged Covariates

- **Exogenous** – future covariates are not influenced by current / past outcomes.

\[ [X(t + 1) \mid Y(t), X(t)] \sim [X(t + 1) \mid X(t)] \]

- **Analysis Issues:**
  - Include single lagged covariates (current, cumulative)
    - **MSCM:** \( E[\text{Sick}(t) \mid \text{Stress}(t - k)] \)
    - **USRDS:** \( E[\text{Hosp}(t) \mid \text{Dose}(t - k)] \)
  - Include multiple lagged covariates
    - **MSCM:** \( E[\text{Sick}(t) \mid \text{Stress}(t - 1), \text{Stress}(t - 2)] \)
    - **USRDS:** \( E[\text{Hosp}(t) \mid \text{Dose}(t - 1), \text{Dose}(t - 2)] \)
Multivariate models with different lags

Time Lag (days) coefficient (log odds ratio)
1 2 3 4 5 6 7
-0.5 0.0 0.5 1.0

lag=1:7
lag=1:6
lag=1:5
lag=1:4
lag=1:3
lag=1:2

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**OZONE AND MORTALITY IN US URBAN COMMUNITIES**

*Figure 1. Percentage Change in Daily Mortality for a 10-ppb Increase in Ozone for Total and Cardiovascular Mortality, for Single-Lag and Distributed-Lag Models*

The single-lag model reflects the percentage increase in mortality for a 10-ppb increase in ozone on a single day. The distributed-lag model reflects the percentage change in mortality for a 10-ppb increase in ozone during the previous week. Error bars indicate 95% posterior intervals.
Endogenous: Analysis

- **Definition**: The covariate is influenced by past outcomes (or intermediate variables)

\[ Y(t) \to X(t + 1) \]

- **Implication**: 

\[ E[Y_i(t) \mid X_i(1), \ldots, X_i(n)] \]

depends on \( X_i(s) \) for \( s > t \) (future values of covariate).

- Role for causal inference concepts.
Causal Targets of Inference

- Longitudinal Treatment

\[ \text{vec}(X_0) \equiv [X(1) = 0, X(2) = 0, \ldots, X(n) = 0] \]
\[ \text{vec}(X_1) \equiv [X(1) = 1, X(2) = 1, \ldots, X(n) = 1] \]

- Population Means

  ▶ Mean of population if all subjects had \( X = 1 \) at all times, and similar population mean if \( X = 0 \) at all times.

\[ \mu_0(n) \equiv E[Y(n) \mid \text{vec}(X_0)] \]
\[ \mu_1(n) \equiv E[Y(n) \mid \text{vec}(X_1)] \]
Endogenous Covariates

treatment / exposure  response
Model / Estimation

- G-computation

▷ **Model:** model **outcome** given past outcomes / exposure.

\[
P[Y(t) \mid X(t), \{Y(s), X(s)\} \ s < t] : \text{outcome}
\]

\[
P[X(t) \mid \{Y(s), X(s)\} \ s < t] : \text{exposure}
\]

▷ **Compute:** compute means of interest by allowing intermediate effects, \(Y(s)\), to occur naturally, but controlling exposure.

\[
\mu_1(t) = E_t \big\{ E_s[Y(t) \mid X(t)=1, \{Y(s), X(s)=1\} \ s < t] \big\}
\]
Model / Estimation

- **Marginal Structural Models**
  - **Model:** model exposure given past outcomes / exposure.
    \[ X(t) \mid \{Y(s), X(s)\} \ s < t \]
  - **Compute:** compute a regression of the outcome using inverse probability weights (IPW) to control for exposure selection bias.
Table 1: Regression of stress, $S_{it}$, on illness, $I_{it-k}$ $k = 0, 1$, and previous stress, $S_{it-k}$ $k = 1, 2, 3, 4+$ using GEE with working independence.

<table>
<thead>
<tr>
<th></th>
<th>est.</th>
<th>s.e.</th>
<th>Z</th>
</tr>
</thead>
<tbody>
<tr>
<td>(Intercept)</td>
<td>-1.88</td>
<td>(0.36)</td>
<td>-5.28</td>
</tr>
<tr>
<td>$I_{it}$</td>
<td>0.50</td>
<td>(0.17)</td>
<td>2.96</td>
</tr>
<tr>
<td>$I_{it-1}$</td>
<td>0.08</td>
<td>(0.17)</td>
<td>0.46</td>
</tr>
<tr>
<td>$S_{it-1}$</td>
<td>0.92</td>
<td>(0.15)</td>
<td>6.26</td>
</tr>
<tr>
<td>$S_{it-2}$</td>
<td>0.31</td>
<td>(0.14)</td>
<td>2.15</td>
</tr>
<tr>
<td>$S_{it-3}$</td>
<td>0.34</td>
<td>(0.14)</td>
<td>2.42</td>
</tr>
<tr>
<td>mean($S_{it-k}$, $k \geq 4$)</td>
<td>1.74</td>
<td>(0.24)</td>
<td>7.27</td>
</tr>
<tr>
<td>employed</td>
<td>-0.26</td>
<td>(0.13)</td>
<td>-2.01</td>
</tr>
<tr>
<td>married</td>
<td>0.16</td>
<td>(0.12)</td>
<td>1.34</td>
</tr>
<tr>
<td>maternal health</td>
<td>-0.19</td>
<td>(0.07)</td>
<td>-2.83</td>
</tr>
<tr>
<td>child health</td>
<td>-0.09</td>
<td>(0.07)</td>
<td>-1.24</td>
</tr>
<tr>
<td>race</td>
<td>0.03</td>
<td>(0.12)</td>
<td>0.21</td>
</tr>
<tr>
<td>education</td>
<td>0.42</td>
<td>(0.13)</td>
<td>3.21</td>
</tr>
<tr>
<td>house size</td>
<td>-0.16</td>
<td>(0.12)</td>
<td>-1.28</td>
</tr>
</tbody>
</table>
Table 2: MSM estimation of the effect of stress, $S_{it-k} \geq 1$, on illness, $I_{it}$.

<table>
<thead>
<tr>
<th></th>
<th>est.</th>
<th>s.e.</th>
<th>Z</th>
</tr>
</thead>
<tbody>
<tr>
<td>(Intercept)</td>
<td>-0.71</td>
<td>(0.40)</td>
<td>-1.77</td>
</tr>
<tr>
<td>$S_{it-1}$</td>
<td>0.15</td>
<td>(0.14)</td>
<td>1.03</td>
</tr>
<tr>
<td>$S_{it-2}$</td>
<td>-0.19</td>
<td>(0.18)</td>
<td>-1.05</td>
</tr>
<tr>
<td>$S_{it-3}$</td>
<td>0.18</td>
<td>(0.15)</td>
<td>1.23</td>
</tr>
<tr>
<td>mean($S_{it-k}, k \geq 4$)</td>
<td>0.71</td>
<td>(0.43)</td>
<td>1.65</td>
</tr>
<tr>
<td>employed</td>
<td>-0.11</td>
<td>(0.21)</td>
<td>-0.54</td>
</tr>
<tr>
<td>married</td>
<td>0.55</td>
<td>(0.17)</td>
<td>3.16</td>
</tr>
<tr>
<td>maternal health</td>
<td>-0.13</td>
<td>(0.10)</td>
<td>-1.27</td>
</tr>
<tr>
<td>child health</td>
<td>-0.34</td>
<td>(0.09)</td>
<td>-3.80</td>
</tr>
<tr>
<td>race</td>
<td>0.72</td>
<td>(0.21)</td>
<td>3.46</td>
</tr>
<tr>
<td>education</td>
<td>0.34</td>
<td>(0.22)</td>
<td>1.57</td>
</tr>
<tr>
<td>house size</td>
<td>-0.80</td>
<td>(0.18)</td>
<td>-4.51</td>
</tr>
<tr>
<td>method</td>
<td>logOR</td>
<td></td>
<td></td>
</tr>
<tr>
<td>-------------------------------------------------------</td>
<td>-------</td>
<td></td>
<td></td>
</tr>
<tr>
<td>GEE cross-sectional association</td>
<td>0.66</td>
<td></td>
<td></td>
</tr>
<tr>
<td>GEE with seven days lagged</td>
<td>1.38</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Transition model (direct effect)</td>
<td>0.50</td>
<td></td>
<td></td>
</tr>
<tr>
<td>G-computation</td>
<td>0.80</td>
<td></td>
<td></td>
</tr>
<tr>
<td>MSM</td>
<td>0.85</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
Summary of Endogenous

- Interest in exposure over time – more than simply the acute (most recent) exposure.

- A variable (perhaps outcome) is both a consequence of exposure at early times, and a cause of exposure at later times.

- Intermediate and confounder.

- G-computation

- MSM

- Interest in outcomes under a controlled and static treatment plan.
EPO: November 2006 NEJM

- **Drüeke** CREATE
  - Control **Hemoglobin** rather than fix the dose.
    - Low group (11.0-12.5)
    - Normal group (13.0-15.0)

- **Singh** CHOIR
  - Control **Hemoglobin** rather than fix the dose.
    - Low group (11.3)
    - Normal group (13.5)

- **Research Question(s)**
  - **Q**: What target hemoglobin should be used? How to use observational data to compare different targets and/or compare mortality experience to RCT data?
Analysis of Dynamic Treatment

• **Note:** The guidelines for Epo do not suggest a static dose be administered. Rather, dose is driven by the state of the intermediate (Hb):

\[
X(t + 1) = \begin{cases} 
1.25 \times X(t) & \text{if } Z(t) \leq 11 \\
X(t) & \text{if } 11 < Z(t) \leq 13 \\
0.75 \times X(t) & \text{if } Z(t) > 13
\end{cases}
\]

• This corresponds to a **dynamic treatment guideline**, \( G_1 \).

• **Q:** How to formulate DOSE questions in this setting?
  
  ▶ \( G_1 \) corresponds to correction of ±25% at Hb=(11,13).
  
  ▶ Compare to a \( G_2 \) which uses alternative target Hb threshold(s).
USRDS Data (2003 sample)

25% or more Change -- LOW Epo

25% or more Change -- HIGH Epo
LDA with Death

- Different than drop-out

- **With Drop-out:**

  \[ E[Y_i(t) | X_i] = E[Y_i(t) | X_i, R_i(t) = 1] \times P[R_i(t) = 1 | X_i] + E[Y_i(t) | X_i, R_i(t) = 0] \times P[R_i(t) = 0 | X_i] \]

- Linear Mixed Models (LMM) applied to the observed data where \( R_i(t) = 1 \) can validly estimate parameters in the mean \( E[Y_i(t) | X_i] \) when data are MAR.

- **With Death:**

  \[ E[Y_i(t) | X_i] = E[Y_i(t) | X_i, D_i(t) = 0] \times P[D_i(t) = 0 | X_i] + E[Y_i(t) | X_i, D_i(t) = 1] \times P[D_i(t) = 1 | X_i] \]
LDA with Death: Analysis

- Analysis conditional on death information:
  - Full (future) stratification:
    \[ E[Y_i(t) \mid X_i(t), T_i = s] \quad s > t \]
    * See: Pauler, McCoy & Moinpour (2003)
  - Partial (current status) conditioning:
    \[ E[Y_i(t) \mid X_i(t), T_i > t] \]
    * See: Kurland and Heagerty (2004)
  - Conditional on principal strata (potential status):
    \[ E[Y_i(t \mid 1) - Y_i(t \mid 0) \mid \{T_i(0) > t, T_i(1) > t}\} \]
    * See Frangakis and Rubin (2002), Rubin (2007)
LDA with Death: Comments on Analysis

- Full stratification using \([T_i = s] \ s > t\)
  - Compares groups defined by \(X_i\) comparable in terms of death.
  - Conditions on future (not yet observed) information.

- Partial (current status) conditioning: \([T_i > t]\)
  - Conditions on observed vital status.
  - Compares groups defined by \(X_i\) after selection by death.

- Principal stratification: \([\{T_i(0) > t, T_i(1) > t\}]\)
  - Compares subgroups defined by \(X_i\) comparable in terms of death.
  - Conditions on unobservable potential status.
Some recommendations

- In applications we should identify factors that influence the secondary stochastic processes and choose appropriate statistical techniques in order to validly answer the scientific question.

- In statistical research reports we should be explicit about the assumptions we are making regarding the secondary stochastic processes.

- For time-dependent covariates ask about associations with both past and future covariate values – consider the factors that drive the covariate.
Thanks!