## Stat 425 HW2 Solutions

## Fritz Scholz

## Chapter 1, Section 4, Problem 46.:

To test the effectiveness of vitamin  $B_1$  in stimulating growth in mushrooms, vitamin  $B_1$  was applied to 13 mushrooms selected at random from a group of 24, while the remaining 11 did not receive this treatment. The weights of the mushrooms at the end of the period of observation were<sup>1</sup> (in milligrams)

Controls:1814.513.512.52324211718.59.514Treated:273420.529.520282026.52224.53435.519

1. Give the sorted vector of midranks for all observations.

[1] 1.0 2.0 3.0 4.0 5.0 6.0 7.0 8.0 9.0 10.5 10.5 12.0 13.0 14.0 15.0 [16] 16.0 17.0 18.0 19.0 20.0 21.0 22.5 22.5 24.0

- 2. What is the observed value of  $W_s^*$ ? 19+22.5+12+21+10.5+20+10.5+18+14+17+22.5+24+9 = 220
- 3. Use the normal approximation to find the significance probability of these results. Since we expect stimulated growth (higher weight) under the treatment, the *p*-value is  $P_{H_0}(W_s^* \ge 220)$ .

With  $E_{H_0}(W_s^*) = 162.5$  and  $\operatorname{var}_{H_0}(W_s^*) = 297.6576$  the normal approximation gives

$$P_{H_0}(W_s^* \ge 220) \approx 1 - \Phi\left(\frac{220 - 162.5}{\sqrt{297.6576}}\right)$$

 $= 1 - \Phi(3.332801) = 1 - \texttt{pnorm}(3.332801) = 0.0004298825$ 

- 4. For the exact null distribution of  $W_s^*$  how many midrank sums need to be computed? choose (24, 13) = 2496144, almost 2.5 million.
- 5. Compute the exact *p*-value for the observed value of  $W_s^*$  (this may take a few minutes). Proportion of all 2496144 possible sums  $W_s^* \ge 220 = 0.0001842842$
- 6. Compute an estimated *p*-value based on N<sub>sim</sub> = 100,000 simulations. (execute set.seed(35) just prior to running the simulation). This may take a few minutes.
  Proportion of simulated W<sup>\*</sup><sub>s</sub> ≥ 220 =0.00018. Your answer would be .00012 if you set up your midrank vector in your code as Z=rank(c(control,treatment)) rather than Z=rank(c(treatment,control)) as I did. Think about why you would get different results even though you start out with the same seed. So much for having all students getting the same result. It had me puzzled for some time.

<sup>&</sup>lt;sup>1</sup>From Linder, *Statistische Methoden*, 2d ed., Birkhäuser, Basel, 1951, p. 91. Original data from Schopfer and Blumer, "Zur Wirkstoffphysiologie von Trichophyton album Sab.," *Ber. Schweiz. Botan Ges.* **53**:409–456 (1943).

7. Discuss the merits of the three *p*-value calculations, in terms of accuracy, computation time, and general usability.

The normal approximation for the *p*-value is off by a factor of 2.33. However, it would have been judged significant for most practical purposes. The simulation comes as close as it can get for  $N_{\rm sim} = 100,000$ . That may be an accident, but other simulation results would not be far off. Try different set.seed(...) prior to simulation. While the calculation of the exact *p*-value will encounter its limit of computability for slightly larger problems, the same is not true for the simulation approach. As *m* and *n* get larger the normal approximation should get better, but we don't know how much.

The code for the two functions that get the above answers is given below.

```
Problem46=function() {
treatment=c(27, 34, 20.5, 29.5, 20, 28, 20, 26.5, 22, 24.5, 34, 35.5, 19)
control=c(18,14.5, 13.5, 12.5, 23, 24, 21, 17, 18.5, 9.5, 14)
Z=rank(c(treatment,control))
n=length(treatment)
m=length(control)
N=m+n
Ws.star=sum(Z[1:n])
mean.Ws=mean(Z) *n
var.Ws=var(Z) *n*(N-n)/N
combs=choose(m+n,n)
out=combn(Z,n,FUN=sum)
pval=mean(out>=Ws.star)
pval.norm=1-pnorm((Ws.star-mean.Ws)/sqrt(var.Ws))
out=list(m=m,n=n,Z.sort=sort(Z),combs=combs,Ws.star=Ws.star,
mean.Ws=mean.Ws,var.Ws=var.Ws,pval=pval,pval.norm=pval.norm)
out
}
Problem46.sim=function(Nsim=100000) {
treatment=c(27, 34, 20.5, 29.5, 20, 28, 20, 26.5, 22, 24.5, 34, 35.5, 19)
control=c(18,14.5, 13.5, 12.5, 23, 24, 21, 17, 18.5, 9.5, 14)
Z=rank(c(treatment,control))
n=length(treatment)
m=length(control)
Ws.star=sum(Z[1:n])
out=rep(0,Nsim) # This gives a much faster simulation than out=NULL, see HW3.
set.seed(35)
for(i in 1:Nsim) {
out[i]=sum(sample(Z,n,replace=F))
}
pval.sim=mean(out>=Ws.star)
out=list(Nsim,Ws.star,pval.sim)
names(out) = c("Nsim", "Ws.star", "pval.sim")
out}
```

## Chapter 1, Section 4, Problem 49.:

In the context of Prob. 42, suppose that m = n = 10 and the data are given in the following table:

	Very Poor	Poor	Indifferent	Good	Very Good
Control	2	2	5	1	0
Treatment	0	2	4	3	1

- 1. Find the observed value of  $W_s^*$ .  $W_s^* = 2*4.5 + 4*11 + 3*17.5 + 20 = 125.5$
- 2. Give the exact *p*-value for this observed value of  $W_s^*$ . Based on the problem description high values of  $W_s^*$  are significant. Thus the *p*-value is  $P_{H_0}(W_s^* \ge 125.5) = 11922/184756 = 0.06452835$ .
- 3. Find the normal approximation for this *p*-value. We have  $E_{H_0}(W_s^*) = 105$  and  $\operatorname{var}_{H_0}(W_s^*) = 156.4474$  and thus

 $P_{H_0}(W_s^* \ge 125.5) \approx 1 - \text{pnorm}((125.5 - 105)/\text{sqrt}(156.4474)) = 0.05061027.$ 

not very close to 0.06452835.

4. Find the critical value *c* giving the significance level closest to .01, using the exact distribution vector developed in 2. Here it helps to look at the sorted unique values of the exact distribution vector (use unique (...)) and compute some of the corresponding upper tail probabilities.

```
> out49=Problem49() # see below for code of Problem49
> sort(unique(out49$out))
[1] 65.0 71.5 74.0 74.5 78.0 80.5 81.0 83.5 84.0 84.5 87.0 87.5
[13] 90.0 90.5 91.0 93.0 93.5 94.0 96.5 97.0 97.5 99.5 100.0 100.5
[25] 103.0 103.5 104.0 106.0 106.5 107.0 109.5 110.0 110.5 112.5 113.0 113.5
[37] 116.0 116.5 117.0 119.0 119.5 120.0 122.5 123.0 125.5 126.0 126.5 129.0
[49] 129.5 132.0 135.5 136.0 138.5 145.0
> mean(out49$out>=135.5)
[1] 0.007047132
> mean(out49$out>=132)
[1] 0.02185585
```

shows that c = 135.5 gives tail probability closest to .01.

- 5. Compare this  $P(W_s^* \ge c)$  (closest to .01) with its normal approximation. 1-pnorm((135.5-105)/sqrt(156.4474)) = 0.007374992, pretty close to 0.007047132.
- 6. Use the normal approximation of the above *c* corresponding to .01, rounding it to the nearest multiple of .5 (Due to midranks  $W_s^*$  takes only values that are multiples of .5). The normal approximation gives

$$P_{H_0}(W_s^* \ge c) \approx 1 - \Phi\left(\frac{c - 105}{\sqrt{(156.4474)}}\right) = .01 \implies \texttt{qnorm}(.99) = (c - 105)/\texttt{sqrt}(156.4474)$$

c = qnorm(.99) \* sqrt(156.4474) + 105 = 134.0977

which rounds to c = 134, which is a bit off from 135.5. Of course, the latter had tail probability .007. Note that there is no possible value between 132 and 135.5.

The code for Problem49 follows:

```
Problem49=function() {
y=c(0,2,4,3,1); x=c(2,2,5,1,0)
d=x+y; n=sum(y); m=sum(x); ell=length(y); N=m+n
midrank=cumsum(d)-(d-1)/2
midrank.vec=NULL
for(j in 1:ell){
midrank.vec=c(midrank.vec,rep(midrank[j],d[j]))}
out=combn(midrank.vec,n,FUN=sum)
out=sort(out)
Ws.star=sum(y*midrank)
pval=mean(out>=Ws.star)
mean.Ws=n*mean(midrank.vec)
var.Ws=var(midrank.vec)*n*(N-n)/N
pval.norm=1-pnorm((Ws.star-mean.Ws)/sqrt(var.Ws))
list(out=out,Ws.star=Ws.star,mean.Ws=mean.Ws,var.Ws=var.Ws,
        pval=pval,pval.norm=pval.norm)
}
```