

Econ 582

Forecasting

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Forecasting

Let $\{y_t\}$ be a covariance stationary and ergodic process, e.g. an ARMA(p, q) process with Wold representation

$$\begin{aligned}y_t &= \mu + \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j}, \quad \varepsilon_t \sim WN(0, \sigma^2) \\ &= \mu + \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \dots\end{aligned}$$

Let $I_t = \{y_t, y_{t-1}, \dots\}$ denote the information set available at time t . Recall,

$$\begin{aligned}E[y_t] &= \mu \\ \text{var}(y_t) &= \sigma^2 \sum_{j=0}^{\infty} \psi_j^2\end{aligned}$$

Goal: Using I_t produce optimal forecasts of y_{t+h} for $h = 1, 2, \dots, s$

$$y_{t+h} = \mu + \varepsilon_{t+h} + \psi_1 \varepsilon_{t+h-1} + \dots \\ + \psi_{h-1} \varepsilon_{t+1} + \psi_h \varepsilon_t + \psi_{h+1} \varepsilon_{t-1} + \dots$$

Define $y_{t+h|t}$ as the forecast of y_{t+h} based on I_t with known parameters. The forecast error is

$$\varepsilon_{t+h|t} = y_{t+h} - y_{t+h|t}$$

and the mean squared error of the forecast is

$$MSE(\varepsilon_{t+h|t}) = E[\varepsilon_{t+h|t}^2] \\ = E[(y_{t+h} - y_{t+h|t})^2]$$

Theorem: The minimum MSE forecast (best forecast) of y_{t+h} based on I_t is

$$y_{t+h|t} = E[y_{t+h}|I_t]$$

Proof: See Hamilton pages 72-73.

Remarks

1. The computation of $E[y_{t+h}|I_t]$ depends on the distribution of $\{\varepsilon_t\}$ and may be a very complicated nonlinear function of the history of $\{\varepsilon_t\}$. Even if $\{\varepsilon_t\}$ is an uncorrelated process (e.g. white noise) it may be the case that

$$E[\varepsilon_{t+1}|I_t] \neq 0$$

2. If $\{\varepsilon_t\}$ is independent white noise, then $E[\varepsilon_{t+1}|I_t] = 0$ and $E[y_{t+h}|I_t]$ will be a simple linear function of $\{\varepsilon_t\}$

$$y_{t+h|t} = \mu + \psi_h \varepsilon_t + \psi_{h+1} \varepsilon_{t-1} + \dots$$

Linear Predictors

$$y_{t+h} = \mu + \varepsilon_{t+h} + \psi_1 \varepsilon_{t+h-1} + \dots \\ + \psi_{h-1} \varepsilon_{t+1} + \psi_h \varepsilon_t + \psi_{h+1} \varepsilon_{t-1} + \dots$$

A linear predictor of $y_{t+h|t}$ is a linear function of the variables in I_t .

Theorem: The minimum MSE linear forecast (best linear predictor) of y_{t+h} based on I_t is

$$y_{t+h|t} = \mu + \psi_h \varepsilon_t + \psi_{h+1} \varepsilon_{t-1} + \dots$$

Proof. See Hamilton page 74.

The forecast error of the best linear predictor is

$$\begin{aligned}\varepsilon_{t+h|t} &= y_{t+h} - \hat{y}_{t+h|t} \\ &= \mu + \varepsilon_{t+h} + \psi_1 \varepsilon_{t+h-1} + \cdots \\ &\quad + \psi_{h-1} \varepsilon_{t+1} + \psi_h \varepsilon_t + \cdots \\ &\quad - (\mu + \psi_h \varepsilon_t + \psi_{h+1} \varepsilon_{t-1} + \cdots) \\ &= \varepsilon_{t+h} + \psi_1 \varepsilon_{t+h-1} + \cdots + \psi_{h-1} \varepsilon_{t+1}\end{aligned}$$

and the MSE of the forecast error is

$$\text{MSE}(\varepsilon_{t+h|t}) = \sigma^2(1 + \psi_1^2 + \cdots + \psi_{h-1}^2)$$

Remarks

1. $E[\varepsilon_{t+h|t}] = 0$
2. $\varepsilon_{t+h|t}$ is uncorrelated with any element in I_t
3. The form of $y_{t+h|t}$ is closely related to the IRF
4. $MSE(\varepsilon_{t+h|t}) = \text{var}(\varepsilon_{t+h|t}) \leq \text{var}(y_t)$
5. $\lim_{h \rightarrow \infty} y_{t+h|t} = \mu$
6. $\lim_{h \rightarrow \infty} MSE(\varepsilon_{t+h|t}) = \text{var}(y_t)$

Example: BLP for MA(1) process

$$y_t = \mu + \varepsilon_t + \theta\varepsilon_{t-1}, \quad \varepsilon_t \sim \text{WN}(0, \sigma^2)$$

Here

$$\psi_1 = \theta, \quad \psi_h = 0 \text{ for } h > 1$$

Therefore,

$$y_{t+1|t} = \mu + \theta\varepsilon_t$$

$$y_{t+2|t} = \mu$$

$$y_{t+h|t} = \mu \text{ for } h > 1$$

The forecast errors and MSEs are

$$\varepsilon_{t+1|t} = \varepsilon_{t+1}, \quad \text{MSE}(\varepsilon_{t+1|t}) = \sigma^2$$

$$\varepsilon_{t+2|t} = \varepsilon_{t+2} + \theta\varepsilon_{t+1}, \quad \text{MSE}(\varepsilon_{t+2|t}) = \sigma^2(1 + \theta^2)$$

Prediction Confidence Intervals

If $\{\varepsilon_t\}$ is Gaussian then

$$y_{t+h}|I_t \sim N(y_{t+h|t}, \sigma^2(1 + \psi_1^2 + \dots + \psi_{h-1}^2))$$

A 95% confidence interval for the h -step prediction has the form

$$y_{t+h|t} \pm 1.96 \cdot \sqrt{\sigma^2(1 + \psi_1^2 + \dots + \psi_{h-1}^2)}$$

Predictions with Estimated Parameters

Let $\hat{y}_{t+h|t}$ denote the BLP with estimated parameters:

$$\hat{y}_{t+h|t} = \hat{\mu} + \hat{\psi}_h \hat{\varepsilon}_t + \hat{\psi}_{h+1} \hat{\varepsilon}_{t-1} + \dots$$

where $\hat{\varepsilon}_t$ is the estimated residual from the fitted model. The forecast error with estimated parameters is

$$\begin{aligned} \hat{\varepsilon}_{t+h|t} &= y_{t+h} - \hat{y}_{t+h|t} \\ &= (\mu - \hat{\mu}) + \varepsilon_{t+h} + \psi_1 \varepsilon_{t+h-1} + \dots + \psi_{h-1} \varepsilon_{t+1} \\ &\quad + (\psi_h \varepsilon_t - \hat{\psi}_h \hat{\varepsilon}_t) + (\psi_{h+1} \varepsilon_{t-1} - \hat{\psi}_{h+1} \hat{\varepsilon}_{t-1}) \\ &\quad + \dots \end{aligned}$$

Obviously,

$$\text{MSE}(\hat{\varepsilon}_{t+h|t}) \neq \text{MSE}(\varepsilon_{t+h|t}) = \sigma^2(1 + \psi_1^2 + \dots + \psi_{h-1}^2)$$

Note: Most software computes

$$\widehat{\text{MSE}}(\varepsilon_{t+h|t}) = \hat{\sigma}^2(1 + \hat{\psi}_1^2 + \dots + \hat{\psi}_{h-1}^2)$$

Computing the Best Linear Predictor

The BLP $y_{t+h|t}$ may be computed in many different but equivalent ways. The algorithm for computing $y_{t+h|t}$ from an AR(1) model is simple and the methodology allows for the computation of forecasts for general ARMA models as well as multivariate models.

Example: AR(1) Model

$$\begin{aligned}y_t - \mu &= \phi(y_{t-1} - \mu) + \varepsilon_t \\ \varepsilon_t &\sim WN(0, \sigma^2) \\ \mu, \phi, \sigma^2 &\text{ are known}\end{aligned}$$

In the Wold representation $\psi_j = \phi^j$. Starting at t and iterating forward h periods gives

$$\begin{aligned}y_{t+h} &= \mu + \phi^h(y_t - \mu) + \varepsilon_{t+h} + \phi\varepsilon_{t+h-1} + \cdots \\ &\quad + \phi^{h-1}\varepsilon_{t+1} \\ &= \mu + \phi^h(y_t - \mu) + \varepsilon_{t+h} + \psi_1\varepsilon_{t+h-1} + \cdots \\ &\quad + \psi_{h-1}\varepsilon_{t+1}\end{aligned}$$

Based on information at time t , the best forecast for $\varepsilon_{t+1}, \dots, \varepsilon_{t+h}$ is zero because $\varepsilon_t \sim WN(0, \sigma^2)$. Hence,

$$y_{t+h|t} = \mu + \phi^h(y_t - \mu), \quad h = 1, 2, \dots$$

The best linear forecasts of $y_{t+1}, y_{t+2}, \dots, y_{t+h}$ can be recursively computed using the *chain-rule of forecasting* (law of iterated projections)

$$y_{t+1|t} = \mu + \phi(y_t - \mu)$$

$$y_{t+2|t} = \mu + \phi(y_{t+1|t} - \mu) = \mu + \phi(\phi(y_t - \mu))$$

$$= \mu + \phi^2(y_t - \mu)$$

⋮

$$y_{t+h|t} = \mu + \phi(y_{t+h-1|t} - \mu) = \mu + \phi^h(y_t - \mu)$$

The corresponding forecast errors are

$$\varepsilon_{t+1|t} = y_{t+1} - y_{t+1|t} = \varepsilon_{t+1}$$

$$\varepsilon_{t+2|t} = y_{t+2} - y_{t+2|t} = \varepsilon_{t+2} + \phi\varepsilon_{t+1}$$

$$= \varepsilon_{t+2} + \psi_1\varepsilon_{t+1}$$

⋮

$$\begin{aligned}\varepsilon_{t+h|t} &= y_{t+h} - y_{t+h|t} = \varepsilon_{t+h} + \phi\varepsilon_{t+h-1} + \cdots \\ &\quad + \phi^{h-1}\varepsilon_{t+1}\end{aligned}$$

$$= \varepsilon_{t+h} + \psi_1\varepsilon_{t+h-1} + \cdots + \psi_{h-1}\varepsilon_{t+1}$$

The forecast error variances are

$$\begin{aligned} \text{var}(\varepsilon_{t+1|t}) &= \sigma^2 \\ \text{var}(\varepsilon_{t+2|t}) &= \sigma^2(1 + \phi^2) = \sigma^2(1 + \psi_1^2) \\ &\vdots \\ \text{var}(\varepsilon_{t+h|t}) &= \sigma^2(1 + \phi^2 + \dots + \phi^{2(h-1)}) = \sigma^2 \frac{1 - \phi^{2h}}{1 - \phi^2} \\ &= \sigma^2(1 + \psi_1^2 + \dots + \psi_{h-1}^2) \end{aligned}$$

Clearly,

$$\begin{aligned} \lim_{h \rightarrow \infty} y_{t+h|t} &= \mu = E[y_t] \\ \lim_{h \rightarrow \infty} \text{var}(\varepsilon_{t+h|t}) &= \frac{\sigma^2}{1 - \phi^2} \\ &= \sigma^2 \sum_{h=0}^{\infty} \psi_h^2 = \text{var}(y_t) \end{aligned}$$

AR(p) Models

Consider the AR(p) model

$$\begin{aligned}\phi(L)(y_t - \mu) &= \varepsilon_t, \varepsilon_t \sim WN(0, \sigma^2) \\ \phi(L) &= 1 - \phi_1 L - \dots - \phi_p L^p\end{aligned}$$

The forecasting algorithm for the AR(p) models is essentially the same as that for AR(1) models once we put the AR(p) model in state space form. Let $X_t = y_t - \mu$. The AR(p) in state space form is

$$\begin{pmatrix} X_t \\ X_{t-1} \\ \vdots \\ X_{t-p+1} \end{pmatrix} = \begin{pmatrix} \phi_1 & \phi_2 & \cdots & \phi_p \\ \mathbf{1} & \mathbf{0} & \cdots & \mathbf{0} \\ & \cdots & & \vdots \\ \mathbf{0} & & \mathbf{1} & \mathbf{0} \end{pmatrix} \begin{pmatrix} X_{t-1} \\ X_{t-2} \\ \vdots \\ X_{t-p} \end{pmatrix} + \begin{pmatrix} \varepsilon_t \\ \mathbf{0} \\ \vdots \\ \mathbf{0} \end{pmatrix}$$

or

$$\begin{aligned}\xi_t &= \mathbf{F}\xi_{t-1} + \mathbf{w}_t \\ \text{var}(\mathbf{w}_t) &= \Sigma_w\end{aligned}$$

Starting at t and iterating forward h periods gives

$$\boldsymbol{\xi}_{t+h} = \mathbf{F}^h \boldsymbol{\xi}_t + \mathbf{w}_{t+h} + \mathbf{F} \mathbf{w}_{t+h-1} + \cdots + \mathbf{F}^{h-1} \mathbf{w}_{t+1}$$

Then the best linear forecasts of $y_{t+1}, y_{t+2}, \dots, y_{t+h}$ are computed using the *chain-rule of forecasting* are

$$\begin{aligned} \boldsymbol{\xi}_{t+1|t} &= \mathbf{F} \boldsymbol{\xi}_t \\ \boldsymbol{\xi}_{t+2|t} &= \mathbf{F} \boldsymbol{\xi}_{t+1|t} = \mathbf{F}^2 \boldsymbol{\xi}_t \\ &\vdots \\ \boldsymbol{\xi}_{t+h|t} &= \mathbf{F} \boldsymbol{\xi}_{t+h-1|t} = \mathbf{F}^h \boldsymbol{\xi}_t \end{aligned}$$

The forecast for y_{t+h} is given by μ plus the first row of $\boldsymbol{\xi}_{t+h|t} = \mathbf{F}^h \boldsymbol{\xi}_t$:

$$\boldsymbol{\xi}_{t+h|t} = \mathbf{F}^h \boldsymbol{\xi}_t = \begin{pmatrix} \phi_1 & \phi_2 & \cdots & \phi_p \\ 1 & 0 & \cdots & 0 \\ & \ddots & & \vdots \\ 0 & & 1 & 0 \end{pmatrix}^h \begin{pmatrix} y_t - \mu \\ y_{t-1} - \mu \\ \vdots \\ y_{t-p+1} - \mu \end{pmatrix}$$

The forecast errors are given by

$$\begin{aligned}
 \mathbf{w}_{t+1|t} &= \xi_{t+1} - \xi_{t+1|t} = \mathbf{w}_{t+1} \\
 \mathbf{w}_{t+2|t} &= \xi_{t+2} - \xi_{t+2|t} = \mathbf{w}_{t+2} + \mathbf{F}\mathbf{w}_{t+1} \\
 &\vdots \\
 \mathbf{w}_{t+h|t} &= \xi_{t+h} - \xi_{t+h|t} = \mathbf{w}_{t+h} + \mathbf{F}\mathbf{w}_{t+h-1} + \cdots \\
 &\quad + \mathbf{F}^{h-1}\mathbf{w}_{t+1}
 \end{aligned}$$

and the corresponding forecast MSE matrices are

$$\begin{aligned}
 \text{var}(\mathbf{w}_{t+1|t}) &= \text{var}(\mathbf{w}_{t+1}) = \Sigma_w \\
 \text{var}(\mathbf{w}_{t+2|t}) &= \text{var}(\mathbf{w}_{t+2}) + \mathbf{F}\text{var}(\mathbf{w}_{t+1})\mathbf{F}' \\
 &= \Sigma_w + \mathbf{F}\Sigma_w\mathbf{F}' \\
 &\vdots \\
 \text{var}(\mathbf{w}_{t+h|t}) &= \sum_{j=0}^{h-1} \mathbf{F}^j \Sigma_w \mathbf{F}^{j'}
 \end{aligned}$$

Notice that

$$\text{var}(\mathbf{w}_{t+h|t}) = \Sigma_w + \mathbf{F}\text{var}(\mathbf{w}_{t+h-1|t})\mathbf{F}'$$