Introduction to Computational Finance and Financial Econometrics Probability Review - Part 2

Eric Zivot Spring 2015

Bivariate Probability Distribution - Discrete rv's

- Marginal distributions
- Conditional distributions
- 2 Bivariate Probability Distribution Continuous rv's
- **3** Independence
- 4 Covariance and Correlation

Example: Two discrete rv's X and Y

Bivariate pdf				
		Y		
	%	0	1	$\Pr(X)$
	0	1/8	0	1/8
X	1	2/8	1/8	3/8
	2	1/8	2/8	3/8
	3	0	1/8	1/8
	$\Pr(Y)$	4/8	4/8	1

 $p(x, y) = \Pr(X = x, Y = y) =$ values in table $e.g., \ p(0, 0) = \Pr(X = 0, Y = 0) = \ 1/8$

$$S_{XY} = \{(0,0), (0,1), (1,0), (1,1), (2,0), (2,1), (3,0), (3,1)\}$$
$$p(x,y) \ge 0 \text{ for } x, y \in S_{XY}$$
$$\sum_{x,y \in S_{XY}} p(x,y) = 1$$

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$$p(x) = \Pr(X = x) = \sum_{y \in S_Y} p(x, y)$$

= sum over columns in joint table

$$p(y) = \Pr(Y = y) = \sum_{x \in S_X} p(x, y)$$

= sum over rows in joint table

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Conditional Probability

Suppose we know Y = 0. How does this knowledge affect the probability that X = 0, 1, 2, or 3? The answer involves conditional probability. **Example:**

$$\Pr(X = 0 | Y = 0) = \frac{\Pr(X = 0, Y = 0)}{\Pr(Y = 0)}$$

$$= \frac{\text{joint probability}}{\text{marginal probability}} = \frac{1/8}{4/8} = 1/4$$

Remark:

$$Pr(X = 0 | Y = 0) = 1/4 \neq Pr(X = 0) = 1/8$$
$$\implies X \text{ depends on } Y$$

The marginal probability, Pr(X = 0), ignores information about Y.

• The conditional pdf of X given Y = y is, for all $x \in S_X$,

$$p(x|y) = \Pr(X = x|Y = y) = \frac{\Pr(X = x, Y = y)}{\Pr(Y = y)}$$

• The conditional pdf of Y given X = x is, for all values of $y \in S_Y$

$$p(y|x) = \Pr(Y = y|X = x) = \frac{\Pr(X = x, Y = y)}{\Pr(X = x)}$$

$$\mu_{X|Y=y} = E[X|Y=y] = \sum_{x \in S_X} x \cdot \Pr(X=x|Y=y),$$

$$\mu_{Y|X=x} = E[Y|X=x] = \sum_{y \in S_Y} y \cdot \Pr(Y=y|X=x).$$

$$\sigma_{X|Y=y}^2 = \operatorname{var}(X|Y=y) = \sum_{x \in S_X} (x - \mu_{X|Y=y})^2 \cdot \Pr(X=x|Y=y),$$

$$\sigma_{Y|X=x}^2 = \operatorname{var}(Y|X=x) = \sum_{y \in S_Y} (y - \mu_{Y|X=x})^2 \cdot \Pr(Y=y|X=x).$$

Example

$$E[X] = 0 \cdot \frac{1}{8} + 1 \cdot \frac{3}{8} + 2 \cdot \frac{3}{8} + 3 \cdot \frac{1}{8} = \frac{3}{2}$$

$$E[X|Y = 0] = 0 \cdot \frac{1}{4} + 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} + 3 \cdot 0 = 1,$$

$$E[X|Y = 1] = 0 \cdot 0 + 1 \cdot \frac{1}{4} + 2 \cdot \frac{1}{2} + 3 \cdot \frac{1}{4} = 2,$$

$$var(X) = (0 - \frac{3}{2})^2 \cdot \frac{1}{8} + (1 - \frac{3}{2})^2 \cdot \frac{3}{8} + (2 - \frac{3}{2})^2 \cdot \frac{3}{8} + (3 - \frac{3}{2})^2 \cdot \frac{1}{8} = \frac{3}{4},$$

$$var(X|Y = 0) = (0 - 1)^2 \cdot \frac{1}{4} + (1 - 1)^2 \cdot \frac{1}{2} + (2 - 1)^2 \cdot \frac{1}{2} + (3 - 1)^2 \cdot 0 = \frac{1}{2},$$

$$var(X|Y = 1) = (0 - 2)^2 \cdot 0 + (1 - 2)^2 \cdot \frac{1}{4} + (2 - 2)^2 \cdot \frac{1}{2} + (3 - 2)^2 \cdot \frac{1}{4} = \frac{1}{2}.$$

Independence

Let X and Y be discrete rvs with pdfs p(x), p(y), sample spaces S_X, S_Y and joint pdf p(x, y). Then X and Y are independent rv's if and only if:

$$p(x,y) = p(x) \cdot p(y)$$

for all values of $x \in S_X$ and $y \in S_Y$

Result: If X and Y are independent rv's then,

$$p(x|y) = p(x)$$
 for all $x \in S_X, y \in S_Y$

$$p(y|x) = p(y)$$
 for all $x \in S_X, y \in S_Y$

Intuition:

- Knowledge of X does not influence probabilities associated with Y
- \bullet Knowledge of Y does not influence probablities associated with X

Bivariate Probability Distribution - Discrete rv's

2 Bivariate Probability Distribution - Continuous rv's • Marginal and conditional distributions

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Bivariate Distributions - Continuous rv's

The joint pdf of X and Y is a non-negative function f(x, y) such that:

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

Let $[x_1, x_2]$ and $[y_1, y_2]$ be intervals on the real line. Then,

 $\Pr(x_1 \le X \le x_2, y_1 \le Y \le y_2)$ = $\int_{x_1}^{x_2} \int_{y_1}^{y_2} f(x, y) dx dy$ = volume under probability surface over the intersection of the intervals $[x_1, x_2]$ and $[y_1, y_2]$

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Marginal and conditional distributions

The marginal pdf of X is found by integrating y out of the joint pdf f(x, y) and the marginal pdf of Y is found by integrating x out of the joint pdf:

$$f(x) = \int_{-\infty}^{\infty} f(x, y) dy,$$

$$f(y) = \int_{-\infty}^{\infty} f(x, y) dx.$$

The conditional pdf of X given that Y = y, denoted f(x|y), is computed as,

$$f(x|y) = \frac{f(x,y)}{f(y)},$$

and the conditional pdf of Y given that X = x is computed as,

$$f(y|x) = \frac{f(x,y)}{f(x)}.$$

The conditional means are computed as:

$$\mu_{X|Y=y} = E[X|Y=y] = \int x \cdot p(x|y)dx,$$
$$\mu_{Y|X=x} = E[Y|X=x] = \int y \cdot p(y|x)dy$$

and the conditional variances are computed as,

$$\sigma_{X|Y=y}^2 = \operatorname{var}(X|Y=y) = \int (x - \mu_{X|Y=y})^2 p(x|y) dx,$$

$$\sigma_{Y|X=x}^2 = \operatorname{var}(Y|X=x) = \int (y - \mu_{Y|X=x})^2 p(y|x) dy.$$

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3 Independence

4 Covariance and Correlation

Let X and Y be continuous random variables. X and Y are independent iff:

$$f(x|y) = f(x), \text{ for } -\infty < x, y < \infty,$$

$$f(y|x) = f(y), \text{ for } -\infty < x, y < \infty.$$

Result: Let X and Y be continuous random variables. X and Y are independent iff:

$$f(x,y) = f(x)f(y)$$

The result in the above proposition is extremely useful in practice because it gives us an easy way to compute the joint pdf for two independent random variables: we simple compute the product of the marginal distributions.

Example

Example: Bivariate standard normal distribution Let $X \sim N(0, 1)$, $Y \sim N(0, 1)$ and let X and Y be independent. Then,

$$f(x,y) = f(x)f(y) = \frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}x^2}\frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}y^2}$$
$$= \frac{1}{2\pi}e^{-\frac{1}{2}(x^2+y^2)}.$$

To find Pr(-1 < X < 1, -1 < Y < 1) we must solve,

$$\int_{-1}^{1}\int_{-1}^{1}\frac{1}{2\pi}e^{-\frac{1}{2}(x^{2}+y^{2})}dxdy$$

which, unfortunately, does not have an analytical solution. Numerical approximation methods are required to evaluate the above integral. See R package mytnorm.

Result: If the random variables X and Y (discrete or continuous) are independent then the random variables g(X) and h(Y) are independent for any functions $g(\cdot)$ and $h(\cdot)$.

For example, if X and Y are independent then X^2 and Y^2 are also independent.

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Covariance: Measures direction but not strength of linear relationship between 2 rv's

$$\sigma_{XY} = E[(X - \mu_X)(Y - \mu_Y)]$$

= $\sum_{x.y \in S_{XY}} (x - \mu_X)(y - \mu_Y) \cdot p(x, y) \text{ (discrete)}$
= $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - \mu_X)(y - \mu_Y)f(x, y)dxdy \text{ (cts)}$

Example: For the data in Table 2, we have

$$\sigma_{XY} = \operatorname{Cov}(X, Y) = (0 - 3/2)(0 - 1/2) \cdot 1/8$$
$$+ (0 - 3/2)(1 - 1/2) \cdot 0 + \cdots$$
$$+ (3 - 3/2)(1 - 1/2) \cdot 1/8 = 1/4$$

$$Cov(X, Y) = Cov(Y, X)$$

$$Cov(aX, bY) = a \cdot b \cdot Cov(X, Y) = a \cdot b \cdot \sigma_{XY}$$

$$Cov(X, X) = Var(X)$$

$$X, Y \text{ independent } \implies Cov(X, Y) = 0$$

$$Cov(X, Y) = 0 \Rightarrow X \text{ and } Y \text{ are independent}$$

$$Cov(X, Y) = E[XY] - E[X]E[Y]$$

Correlation: Measures direction and strength of linear relationship between 2 rv's

$$\rho_{XY} = \operatorname{Cor}(X, Y) = \frac{\operatorname{Cov}(X, Y)}{\operatorname{SD}(X) \cdot \operatorname{SD}(Y)}$$
$$= \frac{\sigma_{XY}}{\sigma_X \cdot \sigma_Y} = \text{scaled covariance}$$

Example: For the Data in Table 2

$$\rho_{XY} = \operatorname{Cor}(X, Y) = \frac{1/4}{\sqrt{(3/4) \cdot (1/2)}} = 0.577$$

$$-1 \leq \rho_{XY} \leq 1$$

$$\rho_{XY} = 1 \text{ if } Y = aX + b \text{ and } a > 0$$

$$\rho_{XY} = -1 \text{ if } Y = aX + b \text{ and } a < 0$$

$$\rho_{XY} = 0 \text{ if and only if } \sigma_{XY} = 0$$

$$\rho_{XY} = 0 \Rightarrow X \text{ and } Y \text{ are independent in general}$$

$$\rho_{XY} = 0 \implies \text{ independence if } X \text{ and } Y \text{ are normal}$$

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Let X and Y be distributed bivariate normal. The joint pdf is given by:

$$f(x,y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} \times \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\left(\frac{x-\mu_X}{\sigma_X}\right)^2 + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2 - \frac{2\rho(x-\mu_X)(y-\mu_Y)}{\sigma_X\sigma_Y}\right]\right\}$$

where $E[X] = \mu_X$, $E[Y] = \mu_Y$, $SD(X) = \sigma_X$, $SD(Y) = \sigma_Y$, and $\rho = cor(X, Y)$.

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Let X and Y be rv's. Define a new rv Z that is a linear combination of X and Y:

Z = aX + bY

where a and b are constants. Then,

$$\mu_Z = E[Z] = E[aX + bY]$$
$$= aE[X] + bE[Y]$$
$$= a \cdot \mu_X + b \cdot \mu_Y$$

and,

$$\sigma_Z^2 = \operatorname{Var}(Z) = \operatorname{Var}(a \cdot X + b \cdot Y)$$

= $a^2 \operatorname{Var}(X) + b^2 \operatorname{Var}(Y) + 2a \cdot b \cdot \operatorname{Cov}(X, Y)$
= $a^2 \sigma_X^2 + b^2 \sigma_Y^2 + 2a \cdot b \cdot \sigma_{XY}$

If $X \sim N(\mu_X, \sigma_X^2)$ and $Y \sim N(\mu_Y, \sigma_Y^2)$ then $Z \sim N(\mu_Z, \sigma_Z^2)$.

Example: Portfolio returns

- R_A = return on asset A with $E[R_A] = \mu_A$ and $Var(R_A) = \sigma_A^2$
- R_B = return on asset B with $E[R_B] = \mu_B$ and $Var(R_B) = \sigma_B^2$

•
$$\operatorname{Cov}(R_A, R_B) = \sigma_{AB}$$
 and $\operatorname{Cor}(R_A, R_B) = \rho_{AB} = \frac{\sigma_{AB}}{\sigma_A \cdot \sigma_B}$

Portfolio:

- x_A = share of wealth invested in asset A, x_B = share of wealth invested in asset B
- $x_A + x_B = 1$ (exhaust all wealth in 2 assets)
- $R_P = x_A \cdot R_A + x_B \cdot R_B = \text{portfolio return}$

Portfolio Problem: How much wealth should be invested in assets A and B?

Portfolio expected return (gain from investing):

$$E[R_P] = \mu_P = E[x_A \cdot R_A + x_B \cdot R_B]$$
$$= x_A E[R_A] + x_B E[R_B]$$
$$= x_A \mu_A + x_B \mu_B$$

Portfolio variance (risk from investing):

$$\operatorname{Var}(R_P) = \sigma_P^2 = \operatorname{Var}(x_A R_A + x_B R_B)$$
$$= x_A^2 \operatorname{Var}(R_A) + x_B^2 \operatorname{Var}(R_B) +$$
$$2 \cdot x_A \cdot x_B \cdot \operatorname{Cov}(R_A, R_B)$$
$$= x_A^2 \sigma_A^2 + x_B^2 \sigma_B^2 + 2x_A x_B \sigma_{AB}$$
$$\operatorname{SD}(R_P) = \sqrt{\operatorname{Var}(R_P)} = \sigma_P$$
$$= \left(x_A^2 \sigma_A^2 + x_B^2 \sigma_B^2 + 2x_A x_B \sigma_{AB}\right)^{1/2}$$

Let X_1, X_2, \dots, X_N be rvs and let a_1, a_2, \dots, a_N be constants. Define,

$$Z = a_1 X_1 + a_2 X_2 + \dots + a_N X_N = \sum_{i=1}^N a_i X_i$$

$$\mu_Z = E[Z] = a_1 E[X_1] + a_2 E[X_2] + \dots + a_N E[X_N]$$

$$= \sum_{i=1}^{N} a_i E[X_i] = \sum_{i=1}^{N} a_i \mu_i$$

Linear Combination of N rv's cont.

For the variance,

$$\sigma_Z^2 = \operatorname{Var}(Z) = a_1^2 \operatorname{Var}(X_1) + \dots + a_N^2 \operatorname{Var}(X_N) + 2a_1 a_2 \operatorname{Cov}(X_1, X_2) + 2a_1 a_3 \operatorname{Cov}(X_1, X_3) + \dots + 2a_2 a_3 \operatorname{Cov}(X_2, X_3) + 2a_2 a_4 \operatorname{Cov}(X_2, X_4) + \dots + 2a_{N-1} a_N \operatorname{Cov}(X_{N-1}, X_N)$$

Note: N variance terms and $N(N-1) = N^2 - N$ covariance terms. If N = 100, there are $100 \times 99 = 9900$ covariance terms!

Result: If X_1, X_2, \dots, X_N are each normally distributed random variables then,

$$Z = \sum_{i=1}^{N} a_i X_i \sim N(\mu_Z, \sigma_Z^2)$$

Example

Example: Portfolio variance with three assets

- R_A, R_B, R_C are simple returns on assets A, B and C
- x_A, x_B, x_C are portfolio shares such that $x_A + x_B + x_C = 1$

•
$$R_p = x_A R_A + x_B R_B + x_C R_C$$

Portfolio variance,

$$\sigma_P^2 = x_A^2 \sigma_A^2 + x_B^2 \sigma_B^2 + x_C^2 \sigma_C^2$$

$$+2x_A x_B \sigma_{AB} + 2x_A x_C \sigma_{AC} + 2x_B x_C \sigma_{BC}$$

Note: Portfolio variance calculation may be simplified using matrix layout,

$$egin{array}{ccccc} x_A & x_B & x_C \ x_A & \sigma_A^2 & \sigma_{AB} & \sigma_{AC} \ x_B & \sigma_{AB} & \sigma_B^2 & \sigma_{BC} \ x_C & \sigma_{AC} & \sigma_{BC} & \sigma_C^2 \end{array}$$

Example

Example: Multi-period continuously compounded returns and the square-root-of-time rule

$$r_t = \ln(1 + R_t) = \text{ monthly cc return}$$

 $r_t \sim N(\mu, \sigma^2) \text{ for all } t$
 $\operatorname{Cov}(r_t, r_s) = 0 \text{ for all } t \neq s$

Annual return,

$$r_t(12) = \sum_{j=0}^{11} r_{t-j}$$
$$= r_t + r_{t-1} + \dots + r_{t-11}$$

$$E[r_t(12)] = \sum_{j=0}^{11} E[r_{t-j}]$$
$$= \sum_{j=0}^{11} \mu \ (E[r_t] = \mu \text{ for all } t)$$
$$= 12\mu \ (\mu = \text{mean of monthly return})$$

Example cont.

And,

$$\operatorname{Var}(r_t(12)) = \operatorname{Var}\left(\sum_{j=0}^{11} r_{t-j}\right)$$
$$= \sum_{j=0}^{11} \operatorname{Var}(r_{t-j}) = \sum_{j=0}^{11} \sigma^2$$
$$= 12 \cdot \sigma^2 \ (\sigma^2 = \text{monthly variance})$$
$$\operatorname{SD}(r_t(12)) = \sqrt{12} \cdot \sigma \ (\text{square root of time rule})$$

$$r_t(12) \sim N(12\mu, 12\sigma^2)$$

For example, suppose:

$$r_t \sim N(0.01, (0.10)^2)$$

$$E[r_t(12)] = 12 \times (0.01) = 0.12$$

Var $(r_t(12)) = 12 \times (0.10)^2 = 0.12$
SD $(r_t(12)) = \sqrt{0.12} = 0.346$
 $r_t(12) \sim N(0.12, (0.346)^2)$

And,

$$(q_{\alpha}^{r})^{A} = 12 \times \mu + \sqrt{12} \times \sigma \times z_{\alpha}$$
$$= 0.12 + 0.346 \times z_{\alpha}$$
$$(q_{\alpha}^{R})^{A} = e^{(q_{\alpha}^{r})^{A}} - 1 = e^{0.12 + 0.346 \times z_{\alpha}} - 1$$

Covariance between two linear combinations of random variables

Consider two linear combinations of two random variables:

$$X = X_1 + X_2$$
$$Y = Y_1 + Y_2$$

$$I = I_1 + I_1$$

Then,

$$cov(X, Y) = cov(X_1 + X_2, Y_1 + Y_2)$$

= cov(X_1, Y_1) + cov(X_1, Y_2)
+ cov(X_2, Y_1) + cov(X_2, Y_2)

The result generalizes to linear combinations of ${\cal N}$ random variables in the obvious way.

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