

Homework 1; Due 8.30 a.m. Wed Jan 14: Stat394 Final Exam

1. (a) Let T be the event of testing positive.

$$P(T) = P(T | G)P(G) + P(T | G^c)P(G^c) = 0.95 \times .01 + 0.05 \times 0.99 = 0.059$$

(b) $P(G | T) = P(T | G)P(G)/P(T) = 0.95 \times .01/0.059 = 0.161,$

$$P(G^c | T) = 1 - P(G | T) = 1 - 0.161 = 0.839. \text{ (Or, } .99 \times 0.05/0.059 = 0.839)$$

(c) $P(G | T^c) = P(T^c | G)P(G)/P(T^c) = .01 \times .05/(1 - 0.059) = .00053.$

2 (a) Note having H and testing positive for H are same thing, since there is no error in testing for H .

$$P(H | T) = P(H | G)P(G | T) + P(H | G^c)P(G^c | T) = .161 \times (3/4) + .839 \times (1/4) = .3305$$

(b) $P(H \cap G^c \cap T) = P(H | (G^c \cap T))P(G^c \cap T) = P(H | G^c)P(G^c | T)P(T) = .25 \times .839 \times .059 = 0.0124$

Or, $P(H \cap G^c \cap T) = P(H | (G^c \cap T))P(G^c \cap T) = P(H | G^c)P(T | G^c)P(G^c) = .25 \times 0.05 \times .99 = 0.0124.$

(c) $P(H \cap G \cap T^c) = P(H | G)P(G | T^c)P(T^c) = 0.75 \times .00053 \times (1 - 0.059) = 0.000375.$

Or $P(H \cap G \cap T^c) = P(H | G)P(T^c | G)P(G) = 0.75 \times 0.05 \times .01 = 0.000375.$

3. (a) $P(\text{bad}) = 1 - P(\text{good}) = 1 - P(0 \text{ faulty}) - P(1 \text{ faulty}) = 1 - (0.9)^{10} - 10 \times 0.1 \times 0.9^9 = 0.264$

(b) $P(2^{\text{nd}} \text{ bad is on trial 9}) = \binom{8}{1} \times .1^2 \times .9^7 = 0.0383. \quad P(\text{first 9 all good}) = 0.9^9 = 0.387.$

$$P(\text{sold} | \text{exactly 9 tested}) = 0.387/(0.387 + 0.0383) = 0.91.$$

(c) $P(\text{gadget sold} | \text{bad}) = P(\text{sold})P(\text{bad} | \text{sold})/P(\text{bad}) =$

$$P(\text{first 5 good}) \times P(2 \text{ or more bad in 2nd 5})/P(\text{bad}) = .59 \times .08/0.264 = 0.1788$$

$$P(\text{gadget discarded} | \text{good}) = P(1 \text{ bad in first 5}) \times P(0 \text{ in untested 5})/P(\text{good})$$

$$= .33 \times .59/(1 - 0.264) = 0.2645. \text{ (Note if more than 1 bad in first 5, gadget cannot be good)}$$

4. (a) Let T be number of winners in any week. T is approx Poisson with mean $E(T) = 40000/20000 = 2.$

$$P(T > 2) = 1 - P(T = 0) - P(T = 1) - P(T = 2) =$$

$$1 - \exp(-2) - 2\exp(-2) - 2^2 \exp(-2)/2 = 1 - 5\exp(-2) = .323.$$

(b) Let X be the number of weeks (out of 208) with more than 2 winners. (Note $X \sim \text{Bin}(208, 0.323).$)

$$E(X) = 208 \times 0.323 = 67.25, \text{ var}(X) = 208 \times 0.323 \times (1 - 0.323) = 45.5.$$

(c) Approximately a standard Normal is $Z = (X - 67.25)/\sqrt{45.5}$ and we want $P(X > 59.5) = P(Z > -1.15).$

From Table $P(Z < 1.15) = 0.875$ and, by symmetry, this is the required probability.

5. (a) $P(T > t) = \int_t^\infty f_T(s)ds = \int_t^\infty \lambda \exp(-\lambda s)ds = [-\exp(-\lambda s)]_t^\infty = \exp(-\lambda t)$

(b) Let the failure times be T_1 and T_2 : $P(T_1 > t) = P(T_2 > t) = \exp(-t)$ (from (a)).

$$\text{By independence: } P(T_1 > t \cap T_2 > t) = \exp(-t) \times \exp(-t) = \exp(-2t).$$

$$\text{Hence from (a), with } \lambda = 2, \text{ (or by differentiating) we have the required pdf } 2\exp(-2t).$$

(c) Suppose first component fails at s :

$$\text{we want } P(T_2 > t + s | T_2 > s) = P(T_2 > (t + s))/P(T_2 > s) = \exp(-(t + s))/\exp(-s) = \exp(-t).$$

(or, one can cite forgetting property of exponential here.)

Hence from (a), or by differentiating we get the required pdf $\exp(-t)$ on $0 < t < \infty.$

(d) After each repair, it runs in full mode for, on average, 0.5 weeks, and then as a copier for, on average, 1 week. So between each repair, it runs on average 1.5 weeks, and for 0.5 weeks in full mode: $0.5/1.5 = 1/3.$

OR: rate of failing when in full mode is twice the rate when running only as a copier,

(Note this is NOT a proof, just a reasonable explanation.)