

Chapter 8: NPMLE, Censoring, and EM

8.1 Estimating an arbitrary F

(i) X_1, \dots, X_n i.i.d $\sim F$.

(ii) **Problem:** no dominating measure.

(One) **solution:** assume a dominating measure which is counting measure on the discrete values in $\{x_1, \dots, x_n\}$.

Problem: dominating measure changes with $X^{(n)}$.

(iii) **Then**

$$L_{X^{(n)}}(F) = \prod_1^n p_i = \prod_1^k p_j^{n_j}$$

where $p_i = P_F(X = x_i)$, $n_j = \#x_i$ equal to x_j , and (wlog) x_1, \dots, x_k are distinct.

(iv) $\ell(F) = \sum_1^{k-1} n_j \log p_j + n_k \log(1 - \sum_1^{k-1} p_j)$, gives $\widehat{p}_j = n_j/n$.

(v) That is, $\widehat{F} = F_n$, the empirical cdf. Despite (ii), F_n has “good” properties.

(vi) **Glivenko-Cantelli:** $\sup_x |F_n(x) - F(x)| \rightarrow_{a.s.} 0$.

(vii) **Donsker’s Thm:** $U_n \sim U(0, 1)$ with edf G_n . $X_j \equiv F^{-1}(U_j) \sim F$. Then

$\sqrt{n}(G_n(u) - u)$ and $\sqrt{n}(G_n^{-1}(u) - u)$ each converges to Brownian Bridge process, B . And $\sqrt{n}(F_n - F)$ converges to the process $B(F)$.

(i.e. have usual \sqrt{n} convergence, like parametric MLEs)

8.2 Right-censoring and the Kaplan-Meier estimator

(i) (X_i, U_i) i.i.d. $X_i \sim f$, $U_i \sim G$.

We observe only $Y_i = \min(X_i, U_i)$ and $\delta_i = I(X_i \leq U_i)$.

(ii) We want to estimate F : G normally not of interest.

If we could observe all the X_i , F_n would be NPMLE of F .

(iii) For simplicity, assume Y_i distinct, and (notational convenience) $y_1 < y_2 < \dots < y_n < y_+$. We construct an NPMLE putting mass only on $\{y_1, \dots, y_n, y_+\}$.

(iv) $\ell(F) = \sum_1^n (\delta_i \log f(y_i) + (1 - \delta_i) \log(1 - F(y_i)))$

(v) Let $k_1 < \dots < k_{m+1}$ be indices of uncensored ($\delta_i = 1$), obsv., with $y_{k_{m+1}} = y_+$ if $\delta_n = 0$, and $y_{k_{m+1}} = y_n$ if $\delta_n = 1$.

(vi) Let $p_{k_j} = f(y_{k_j})$, and $n_j = \#x_i$ equal to y_{k_j} . Now do EM, with complete-data X_1, \dots, X_n :

$$\ell_c(F) = \sum_1^n \log f(x_i) = \sum_1^{m+1} n_j \log p_{k_j}$$

$$\text{Let } e_j = E(n_j | Y^{(n)}, \delta^{(n)}) = \sum_1^n P(X_i = y_{k_j} | \delta_i) \quad \tilde{p}_{k_j} = e_j/n$$

(vii) Now with $F(t) = \sum_{j: y_{k_j} \leq t} p_{k_j}$, $F(t) = E(F_n(t) | Y^{(n)}, \delta^{(n)})$

(viii) But now we find that a stationary point of EM is

$$\widehat{p}_{k_i} / \sum_{j=i}^{m+1} \widehat{p}_{k_j} = 1/(n - k_i + 1) \quad i = 1, \dots, m.$$

(ix) Then $\widehat{F}(t) = \sum_{i: y_{k_i} \leq t} \widehat{p}_{k_i}$ is NPMLE. This is Kaplan-Meier estimate, although not in usual form.

(x) Consider

$$\prod_{i: y_{k_i} \leq t} \left(1 - \frac{1}{n - k_i + 1}\right) = \dots = \dots = 1 - \widehat{F}(t)$$

$n - k_i + 1$ is “population at risk” just before failure at y_{k_i} .

8.3 Current status data

(i) As above, failure times X_i , but now we observe only times U_i , and $\delta_i = I(X_i \leq U_i)$ (i alive/dead at U_i).

(ii) Again, (X_i, U_i) i.i.d, with X_i indep U_i , $X_i \sim F$, $U_i \sim G$.
 $L(F) = \sum_1^n (\delta_i \log F(u_i) + (1 - \delta_i) \log(1 - F(u_i)))$

(iii) Wlog, $u_1 < u_2 < \dots < u_n < u_+ \equiv u_{n+1}$. We will put probability mass on (a subset of) u_1, \dots, u_n and maybe on u_+ . Then need to find $p_k = P_F(X = u_k)$, $k = 1, \dots, n$.

(iv) Suppose at EM step m we have estimate $p_i^{(m)}$, $i = 1, \dots, n + 1$, giving probs $Q_{ik}^{(m)} = P^{(m)}(X_i = u_k | \delta_i)$.

(v) Now $\ell_c(F) = \sum_1^n \log f(X_i)$ so

$$\begin{aligned} E_m(\ell_c(F) | U^{(n)}, \delta^{(n)}) &= \sum_{i=1}^n E_m(\log f(X_i) | U^{(n)} = u, \delta^{(n)}) \\ &= \sum_{i=1}^n \sum_{k=1}^{n+1} \log p_k P^{(m)}(X_i = u_k | \delta_i) = \sum_{k=1}^{n+1} \left(\log p_k \left(\sum_{i=1}^n Q_{ik}^{(m)} \right) \right) \end{aligned}$$

(vi) Maximizing (M-step): $p_k^{(m+1)} = n^{-1} \sum_{i=1}^n Q_{ik}^{(m)}$;

but $Q_{ik}^{(m)} = \delta_i p_k^{(m)} / F^{(m)}(u_i)$ if $u_i \geq u_k$, and

$(1 - \delta_i) p_k^{(m)} / (1 - F^{(m)}(u_i))$ if $u_i < u_k$.

(viii) Thus $p_k^{(m+1)} = p_k^{(m)} S_{ik}^{(m)}$ where

$$S_{ik}^{(m)} = \frac{1}{n} \sum_{i=1}^n \left(\frac{\delta_i I(u_i \geq u_k)}{F^{(m)}(u_i)} + \frac{(1 - \delta_i) I(u_i < u_k)}{1 - F^{(m)}(u_i)} \right)$$

(ix) Either $p_k^{(m)} \rightarrow \widehat{p}_k > 0$; then $\widehat{S}_{ik} = 1$,

or $p_k^{(m)} \rightarrow 0$, and then $\widehat{S}_{ik} \leq 1$.

8.4 The Cusum Diagram

(i) Define points in \mathfrak{R}^2 , $P_0 = (0, 0)$, $P_k = (k, \sum_1^k \delta_i)$.

(ii) $F(u_i) = P(\text{failed by } u_i) \approx (1/k) \sum_1^k \delta_i = \text{slope of } (P_0, P_k)$.
BUT F must be non-decreasing. So we take largest convex fn $\leq \{P_k\}$ and let $F(\widehat{u}_i)$ be slope of this function at i .

(iii) **E.g.** $\delta = (1, 0, 0, 1, 0, 1)$, then $F(\widehat{u}_1) = F(\widehat{u}_2) = F(\widehat{u}_3) = 1/3$,
 $F(\widehat{u}_4) = F(\widehat{u}_5) = 1/2$, $F(\widehat{u}_6) = 1$.

(iv) Note, change in slope at $k \Rightarrow \delta_{k+1} = 1$, so we have prob mass at failure observation times.

(v) Suppose slope changes are at $k_1 - 1, k_2 - 1, \dots$, with $1 \leq k_1 < k_2 < \dots < k_+ \leq n$, so $p_{k_i} > 0$.

(vi) If $k_j \leq l \leq k_{j+1} - 1$, $F(\widehat{u}_l) = \Delta_j / (k_{j+1} - k_j)$,
where $\Delta_j = \sum_{i=k_j}^{k_{j+1}-1} \delta_i$.

(vii) If $k_j > m$, $\sum_{i=k_j}^{k_{j+1}-1} \delta_i / F(\widehat{u}_i) = (k_{j+1} - k_j)$.

If $k_{j+1} \leq m$, $\sum_{i=k_j}^{k_{j+1}-1} (1 - \delta_i) / (1 - F(\widehat{u}_i)) = (k_{j+1} - k_j)$.

(viii) If $k_{j+1} = m$, $S_m^* = n^{-1}(k_1 + (k_2 - k_1) + \dots + (n - k_+)) = 1$.

If $k_j < m < k_{j+1}$,

$$\begin{aligned} S_m^* &= n^{-1}(k_1 + \dots + (k_j - k_{j-1}) + G_{j,m} + (k_{j+2} - k_{j+1}) \\ &\quad + \dots + (n - k_+)) \\ &= n^{-1}(n + G_{j,m} - (k_{j+1} - k_j) \quad \text{where} \\ G_{j,m} &= \frac{k_{j+1} - k_j}{k_{j+1} - k_j - \Delta_j} (m - k_j + 1 - D_m) \\ &\quad + \frac{k_{j+1} - k_j}{\Delta_j} (\Delta_j - D_m) \leq k_{j+1} - k_j \end{aligned}$$

where $D_m = \sum_{i=k_j}^m \delta_i$. So $S_m^* \leq 1$.

QED!!!