## Homework 1: Outline solutions

1. (a) Draw a picture: easiest to separate into x < y, v < 0 and  $x \ge y$ ,  $v \ge 0$ . The the range of values (v, w) of (V, W) is

$$\{(v,w): \ -1 \leq v < 0, \ -v \leq w \leq 1\} \ \cup \ \{(v,w); \ 0 \leq v \leq 1, \ v \leq w \leq 1\}$$

(b) If transforming the density, again separate into x > y and x < y. The e.g. on x > y inverse transformation is X = W, Y = W - V, note this mapping is continuously differentiable, and don't forget the Jacobian (even though it turns out to be 1).

Easier, is direct use of the distribution function

$$\begin{split} P(V \le v, \ W \le w) &= \Pr(X - Y \le v, \ \max(X, Y) \le w) \\ &= \Pr(Y \ge X - v, X \le w, Y \le W) \\ &= \begin{cases} w^2 - \frac{1}{2}(w - v)^2, & \text{if } 0 \le v \le w \le 1 \\ \frac{1}{2}(v + w)^2, & \text{if } -1 \le v < 0, \ 0 < -v \le w \le 1. \end{cases} \end{split}$$

Use your picture!

Differentiating w.r.t v and w gives the uniform density 1 on the range of (V, W).

(c) No, they cannot be independent since the ranges are not separable: we know  $|v| \leq w$ .

2. (a) The sequence of random variables  $V_n$  converges in distribution to U(0,1) since

 $\Pr(V_n \leq x) = \lfloor nx \rfloor / n \rightarrow x \text{ as } n \rightarrow \infty$ 

for 0 < x < 1. (Draw a picture!)

For convergence in probability there is insufficient information to say. We do not know the joint dsn of the  $V_n$  nor even that they are defined on the same probability space.

(b) There are various ways to do this. This one comes from Ferguson.

Let  $\epsilon > 0$  and  $k > 2/\epsilon$  (an integer). Since F is continuous, can choose  $x_j$  s.t.  $F(x_j) = j/k, j = 1, ..., k - 1$ . Since  $F_n(x_j) \to F(x_j), \exists N_j$  s.t.  $|F_n(x_j) - F(x_j)| < 1/k$  for all  $n > N_j$ . Now consider  $n > N = \max(N_1, ..., N_{k-1})$ , and  $x_j \le x < x_{j+1}$ , with  $x_k = \infty$ :

$$F_n(x) \leq F_n(x_{j+1}) \leq F(x_{j+1}) + 1/k \leq F(x) + 2/k$$
  

$$F_n(x) \geq F_n(x_j) \geq F(x_j) - 1/k \geq F(x) - 2/k$$

Hence  $|F_n(x) - F(x)| \leq 2/k < \epsilon$  for all n > N and for all x.

3. (a) Checking the integral

$$\int_0^1 \int_0^1 f_{\theta}(v, w) dv dw = 1 \quad \text{for all real } \theta$$

we see that all we need is  $f_{\theta}(v, w) \ge 0$  for all (v, w) in unit square, which by monotonicity of (1 - 2v) will hold iff it holds at the corners (0,0), (0,1), (1,0) and (1,1) of the square, or for  $|\theta| \le 1$ .

(b) The corresponding distribution function  $F_{\theta}$  is given by

$$F_{\theta}(v,w) = \int_0^v \int_0^w f_{\theta}(x,y) dx dy$$
  
=  $vw(1 + \theta(1-v)(1-w))$ 

Note that

$$F_{\theta}(v,1) = v$$
 and  $F_{\theta}(1,w) = w$ ,

so each of V and W has a uniform U(0,1) marginal distribution.

(c) Now E(V) = E(W) = 1/2 and var(V) = var(W) = 1/12, Then direct integration gives

$$Cov(V,W) = \int_0^1 \int_0^1 vw f_{\theta}(v,w) dv dw - (1/2)^2 = \theta/36$$

and hence

$$\rho(V, W) = \frac{\operatorname{Cov}(V, W)}{\sqrt{\operatorname{var}(V)\operatorname{var}(W)}} = \theta/3$$

Thus  $|\rho(V, W)| \leq 1/3$ , which limits the range of correlation between two U(0, 1) random variables that can be provided by this simple family.

4. (a)  $E(Y_n) = \theta/2$  and  $var(Y_n) = \theta^2/12$ , hence immendiately by the CLT,  $V_n$  does converge in distribution, and it converges to  $N(0, \theta^2/12)$ .

(b) 
$$\Pr(Y_n) > y) = 1 - y/\theta$$
, so  $\Pr(\min(Y_n) > y) = (1 - y/\theta)^n$ . Then  

$$P(W_n > w) = \Pr(\min(Y_n) > y/n) = (1 - y/n\theta)^n$$

$$\to \exp(-y/\theta) \text{ as } n \to \infty$$

Thus  $W_n$  does converge in distribution. It converges to an exponential random variable with mean  $\theta$ .