

QUESTIONS

5.6.1. Let X_1, X_2, \dots, X_n be a random sample of size n from the geometric distribution,

$$p_X(k; p) = (1 - p)^{k-1} p, k = 1, 2, \dots. \text{ Show that } \hat{p} = \sum_{i=1}^n X_i \text{ is sufficient for } p.$$

5.6.2. Suppose a random sample of size n is drawn from the pdf,

$$f_Y(y; \theta) = e^{-(y-\theta)}, \quad \theta \leq y$$

(a) Show that $\hat{\theta} = Y_{\min}$ is sufficient for the threshold parameter θ .

(b) Show that Y_{\max} is not sufficient for θ .

5.6.5. Show that $\hat{\sigma}^2 = \sum_{i=1}^n Y_i^2$ is sufficient for σ^2 if Y_1, Y_2, \dots, Y_n is a random sample from a normal pdf with $\mu = 0$.

5.5.2. Let Y_1, Y_2, \dots, Y_n be a random sample from $f_Y(y; \theta) = \frac{1}{\theta} e^{-y/\theta}, y > 0$. Compare the Cramér-Rao lower bound for $f_Y(y; \theta)$ to the variance of the maximum likelihood estimator for $\theta, \hat{\theta} = \frac{1}{n} \sum_{i=1}^n Y_i$. Is \bar{Y} a best estimator for θ ?

5.5.3. Let X_1, X_2, \dots, X_n be a random sample of size n from the Poisson distribution, $p_X(k; \lambda) = \frac{e^{-\lambda} \lambda^k}{k!}, k = 0, 1, \dots$. Show that $\hat{\lambda} = \frac{1}{n} \sum_{i=1}^n X_i$ is an efficient estimator for λ .

5.5.4. Suppose a random sample of size n is taken from a normal distribution with mean μ and variance σ^2 , where σ^2 is known. Compare the Cramér-Rao lower bound for $f_Y(y; \mu)$ with the variance of $\hat{\mu} = \bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$. Is \bar{Y} an efficient estimator for μ ?