

**1. Permutations and combinations**

There are  $n! = \prod_{i=1}^n i = 1.2.3.4. \dots n$  permutations of  $n$  objects.

There are  $\binom{n}{k} = n!/(k!(n-k)!)$  ways of choosing a given  $k$  objects from  $n$ .

**2. Joint and conditional probabilities**

If  $C$  and  $D$  are any events:  $P(C \cup D) = P(C) + P(D) - P(C \cap D)$ .

The conditional probability of  $C$  given  $D$  is  $P(C | D) = P(C \cap D) / P(D)$ .

$C$  and  $D$  are independent if  $P(C \cap D) = P(C).P(D)$ .

**3. Laws and theorems**

Suppose  $E_1, \dots, E_k$  is a partition of  $\Omega$ . That is  $E_i \cap E_j$  is empty, and  $E_1 \cup E_2 \cup \dots \cup E_k = \Omega$ .

The law of total probability states that:  $P(D) = \sum_{j=1}^k P(D \cap E_j) = \sum_{j=1}^k P(D | E_j) P(E_j)$

Bayes' Theorem states that:  $P(E_i | D) = P(D | E_i) P(E_i)/P(D)$

**4. Random variables and distributions**

	discrete (mass)	continuous (density)
Probability mass/density function	pmf: $P(X = x) = p_X(x)$	pdf: $f_X(x)$
Cumulative dist. func. CDF, $P(X \leq x)$	$F_X(x) = \sum_{w \leq x} p_X(w)$	$F_X(x) = \int_{-\infty}^x f_X(w)dw$
Expectation:	$E(X) = \sum_x x P(X = x)$	$\int_{-\infty}^{\infty} x f_X(x)dx$
	$E(g(X)) = \sum_x g(x) P(X = x)$	$\int_{-\infty}^{\infty} g(x) f_X(x)dx$
	<b>provided</b> the sum/integral converges absolutely.	

**For any random variables  $X$  and  $Y$ :**

Variance:  $\text{var}(X) = E((X - E(X))^2) = E(X^2) - (E(X))^2$

Always:  $E(aX + b) = aE(X) + b$ ,  $E(X + Y) = E(X) + E(Y)$ ,  $\text{var}(aX + b) = a^2 \text{var}(X)$ .

If  $X$  and  $Y$  are independent:  $\text{var}(X + Y) = \text{var}(X) + \text{var}(Y)$

If  $X_1, \dots, X_n$  are i.i.d,  $\bar{X}_n \equiv (1/n) \sum_{i=1}^n X_i$  then  $E(\bar{X}_n) = E(X_1)$ ,  $\text{var}(\bar{X}_n) = \text{var}(X_1)/n$ .

**5. Basics of estimation of parameter  $\theta$**

(a) An  $n$ -sample is a set of  $n$  independent data random variables,  $Y_1, \dots, Y_n$  all from the same distribution (i.i.d), indexed by unknown parameter(s)  $\theta$ .

(b) A *statistic*,  $T$ , is any function of  $Y_1, \dots, Y_n$ .

(c) An *estimator* of  $g(\theta)$  is a statistic used to estimate  $g(\theta)$ .

(d) The *estimate* of  $g(\theta)$  is the value taken by the estimator in any particular instance.

(e) The Method of Moments (MoM) estimator of one parameter  $\theta$  is found by solving  $\bar{Y}_n = (1/n) \sum_{i=1}^n Y_i = E_{\theta}(Y)$ , giving say estimator  $W$  for  $\theta$ . Then estimator of  $h(\theta)$  is  $h(W)$ .

**6. Properties of estimators**

(a) *Bias* of estimator  $T$  of  $\theta$  is  $b_T(\theta) = E_{\theta}(T) - \theta$ .

(b) *MSE* of estimator  $T$  of  $\theta$  is  $mse_{\theta}(T) = E((T - \theta)^2) = \text{var}_{\theta}(T) + (b_T(\theta))^2$ .

(c) An estimator  $T_n$  based on an  $n$ -sample is *asymptotically unbiased* if  $b_{T_n}(\theta) \rightarrow 0$  as  $n \rightarrow \infty$ .

(d) An estimator  $T_n$  based on an  $n$ -sample is *consistent* if, for all  $\epsilon > 0$ ,  $P(|T_n - \theta| > \epsilon) \rightarrow 0$  as  $n \rightarrow \infty$ .

(e) Chebychev's inequality: consistency and MSE are related by the fact that, for all  $\epsilon > 0$ ,

$$P(|T_n - \theta| > \epsilon) \leq E((T_n - \theta)^2)/\epsilon^2.$$

## 7. Standard distributions:

	pmf or pdf	mean	variance
(a) Binomial; $B(n, p)$ index $n$ , parameter $p$	$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$ $k = 0, 1, 2, \dots, n$	$np$	$np(1-p)$
(b) Geometric; $Geo(p)$ ; parameter $p$	$P(X = k) = p(1-p)^{k-1}$ $k = 1, 2, 3, 4, \dots$	$1/p$	$(1-p)/p^2$
(c) Neg. Binomial; $NegB(r, p)$ ; index $r$ , parameter $p$	$P(X = k) = \binom{k-1}{r-1} p^r (1-p)^{k-r}$ $k = r, r+1, r+2, \dots$	$r/p$	$r(1-p)/p^2$
(d) Poisson; $Po(\mu)$	$P(X = k) = \exp(-\mu) \mu^k / k!, \quad k = 0, 1, 2, \dots$	$\mu$	$\mu$
(e) Uniform on $(a, b)$ ; $U(a, b)$ ;	$f_X(x) = 1/(b-a), \quad a < x < b$	$(b+a)/2$	$(b-a)^2/12$
(f) Normal, $N(\mu, \sigma^2)$	$f_X(x) = (1/\sqrt{2\pi\sigma^2}) \exp(-(x-\mu)^2/2\sigma^2)$	$\mu$	$\sigma^2$
(g) Exponential, $\mathcal{E}(\lambda)$ rate parameter $\lambda$	$f_X(x) = \lambda \exp(-\lambda x)$ $0 \leq x < \infty$	$1/\lambda$	$1/\lambda^2$