#### CSSS 569 · Visualizing Data

# EXPLORATORY DATA ANALYSIS: BETWEEN DATA & MODEL

#### Christopher Adolph

Department of Political Science and Center for Statistics and the Social Sciences

University of Washington, Seattle





## **Exploratory Data Analysis**

Exploratory Data Analysis refers to a approach pioneered by statistician John Tukey Emphasis on

- letting the data speak for itself
- non-parametric models
- tools for exploring high-dimensional datasets

Relatively little used in social science, where we prefer parametric models

Dangers: sometimes parametric models are fragile, and EDA could help show this

Without preliminary EDA, finding the right parametric specification may be harder

### **EDA:** Some example techniques

The lattice package implements a set of EDA techniques pioneered by Bell Labs/Bill Cleveland.

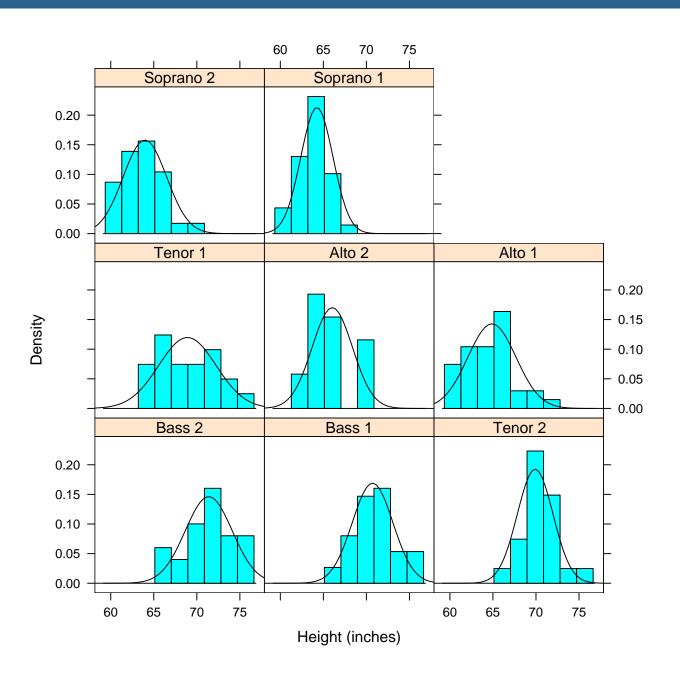
Basic idea: small multiples that show relations between x and y conditioning on z, and perhaps w, etc.

Lattice plots consist of multiple panels of plotted data

The panels are linked to strips which identify a conditioning variable

Let's see how this works with *histograms* and *scatterplots* 

# **Lattice in action**



### **Key lattice options**

dev.off()

Notice two trademark elements of lattice:

- the use of a formula to input the data
- the presence of a customizable panel function

#### Lattice

Key parameters for lattice plots often hide in panel.XXX() where XXX() is the function of interest

Example: the key parameter for 3D plots (how to spin them) is screen, which is documented in panel.cloud() only

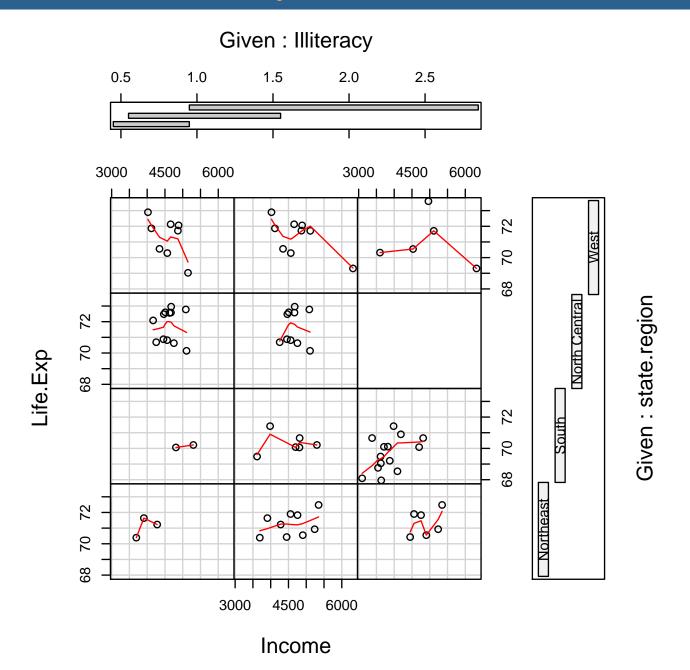
par() doesn't work for lattice.

Use trellis.par.get() and trellis.par.set() to modify lattice parameters

What are the lattice parameters? Mostly undocumented!

print(trellis.par.get()) gives a list of them, for what it's worth

## Another example, this time from base



### Lattice-like graphics in base

Notice the use of two conditioning variables

Notice the smoother added by panel

## **Basic Exploratory Data Analysis**

We've now covered some of the basic tools of EDA:

- scatterplot matrices
- conditional plots (lattice/trellis)
- histogram-like plots

Let's take a close look at two EDA topics:

the border between EDA & modeling

EDA for high dimensional data

## Why is this a topic for visualization?

Simple answer: only way to understand some fits is visually

Deeper answer: visual EDA complements and supports better statistical modeling

Henceforth, our goal will be to use VDQIs to improve our statistical modeling and inference

Complement to your other coursework

### The border between EDA and modeling

Models make simplifying assumptions

The precision of model estimates comes from these assumptions

Wanted: Assumptions "pretty close" to the behavior of the data

How do we check? non-parametric & semi-parametric EDA

Approach: partially relax modeling assumptions, and see if data support simplification

E.g., let the line wiggle if it wants; then check for approximate linearity

Introducing graphical techniques for a wide variety of statistical methods

 $\rightarrow$  We need a language to refer to diverse probability models

Most models have a stochatic component:

$$\mathbf{y} \sim f_{\mathcal{D}}(\boldsymbol{\mu}, \boldsymbol{\alpha})$$

and a systematic component

$$\mu = g(\mathbf{X}, \boldsymbol{\beta})$$

y is the data vector of interest

X is a matrix of covariates

 $f_{\mathcal{D}}$  is a probability density function for distribution  $\mathcal{D}$ 

 $\mu$  is (usually) the expected value

lpha is a "nuisance" parameter vector

 $oldsymbol{eta}$  is a parameter vector associated with the covariates

$$\mathbf{y} \sim f_{\mathcal{D}}(\mu, \alpha)$$
  
 $\mu = g(\mathbf{X}, \beta)$ 

nests most (if not all) models you know.

Linear regression: 
$${f y} \sim f_{
m Normal}({m \mu},\sigma^2)$$
 (continuous data)  ${m \mu} = {f X}{m eta}$ 

Logit: 
$$\mathbf{y} \sim f_{\mathrm{Bernoulli}}(\boldsymbol{\mu})$$
 (binary data)  $\boldsymbol{\mu} = \mathrm{logit}^{-1}(\mathbf{X}\boldsymbol{\beta})$ 

Poisson: 
$$\mathbf{y} \sim f_{\mathrm{Poisson}}(\boldsymbol{\lambda})$$
 (count data)  $\boldsymbol{\lambda} = \exp(\mathbf{X}\boldsymbol{\beta})$ 

and so on

$$\mathbf{y} \sim f_{\mathcal{D}}(\mu, \alpha)$$
  
 $\mu = g(\mathbf{X}, \beta)$ 

nests most (if not all) models you know.

Linear regression: 
$$\mathbf{y} \sim f_{\mathrm{Normal}}(\boldsymbol{\mu}, \sigma^2)$$
 (continuous data)  $\boldsymbol{\mu} = \mathbf{X}\boldsymbol{\beta}$ 

Logit: 
$$\mathbf{y} \sim f_{\mathrm{Bernoulli}}(\boldsymbol{\mu})$$
 (binary data)  $\boldsymbol{\mu} = [1 - \exp(-\mathbf{X}\boldsymbol{\beta})]^{-1}$ 

Poisson: 
$$\mathbf{y} \sim f_{\mathrm{Poisson}}(\boldsymbol{\lambda})$$
 (count data)  $\boldsymbol{\lambda} = \exp(\mathbf{X}\boldsymbol{\beta})$ 

and so on

Note the parallel to the notation of Generalized Linear Models (GLMs)

For example, we can write logit equivalently

$$\mathbf{y} \sim f_{\mathrm{Bernoulli}}(\boldsymbol{\mu})$$
  $\mathbf{y} \sim f_{\mathrm{Bernoulli}}(\boldsymbol{\mu})$   $egin{aligned} \boldsymbol{\mu} &= g(\mathbf{X}\boldsymbol{\beta}) \end{aligned}$   $g^{-1}(\boldsymbol{\mu}) = \mathbf{X}\boldsymbol{\beta}$   $egin{aligned} \boldsymbol{\mu} &= [1 - \exp(-\mathbf{X}\boldsymbol{\beta})]^{-1} \end{aligned}$   $\log[\boldsymbol{\mu}/(1 - \boldsymbol{\mu})] = \mathbf{X}\boldsymbol{\beta}$ 

The framework on the left applies to just about any distribution you will encounter It is equivalent to the form on the right, which is customary for GLMs

 $g(\cdot)$  is called the link function in the GLM context

GLMs are a class of models for which  $f_{\mathcal{D}}$  is a member of the exponential family, which includes the Normal, Binomial, Gamma, and Poisson; in R, see ?family()

Nice aspects of the framework:

- General: works for most models & most estimation methods (MLE, Bayes, etc.)
- Focuses on the data of interest, y
- Reduces attention to  $\beta$ , which is just a cog in the machine that turns X into y (For different models,  $\beta$  has different, usually non-obvious interpretations)
- For any given counterfactual set of covariate values  $\mathbf{x}_c$ , the conditional expecation  $\mathrm{E}(y_c c | \mathbf{x}_c)$  and the expected first difference  $\mathrm{E}(\mathbf{y}_c \mathbf{y}_d | \mathbf{x}_c, \mathbf{x}_d)$  have simple, substantively interesting interpretations

Under this framework, our problem is simply to explain how y (or an interesting function of y,) shifts as we vary  $\mathbf{x}_c$ 

That's usually what our model is for, so that's what we want to visualize

## Unpacking the framework

Let's focus on

$$\mu = g(\mathbf{X}, \boldsymbol{\beta})$$

Here are some typical specifications of the RHS of this equation in LS models

$$\boldsymbol{\mu} = \beta_0 + \beta_1 \mathbf{x}_1 + \beta_2 \mathbf{x}_2 + \dots$$

$$\boldsymbol{\mu} = \beta_0 + \beta_1 \mathbf{x}_1^2 + \beta_2 \mathbf{x}_2 + \dots$$

$$\boldsymbol{\mu} = \beta_0 + \beta_1 \mathbf{x}_1 + \beta_2 \log(\mathbf{x}_2) + \dots$$

A general form for these transformations:

$$\boldsymbol{\mu} = \boldsymbol{\eta} + h_1(\beta_1, \mathbf{x}_1) + h_2(\beta_2, \mathbf{x}_2) \dots$$

## Unpacking the framework

A general form for these transformations:

$$\boldsymbol{\mu} = \boldsymbol{\eta} + h_1(\beta_1, \mathbf{x}_1) + h_2(\beta_2, \mathbf{x}_2) \dots$$

which we can write compactly as:

$$oldsymbol{\mu} = oldsymbol{\eta} + \sum_{k=1}^p h_k(eta_k, \mathbf{x}_k)$$

 $h_k(\cdot)$  could be any arbitrary function.

We could multiply  $\beta$  by the square of  $\mathbf{x}_k$ 

We could multiply  $\beta$  by the log of  $\mathbf{x}_k$ 

We could multiply  $\beta$  by the average of the nearest t neighboring  $x_{ki}$ 's

## Unpacking the framework

$$\mu = \eta + \sum_{k=1}^{p} h_k(\beta_k, \mathbf{x}_k)$$

What if we don't even need a  $\beta_k$  for some of our  $h_k(\cdot)$ 's?

E.g., for time series y, we could use a running average of (say) the last 3 values y itself as the sole predictors:

$$E(y_i) = \sum_{t=1}^{3} y_{i-t}/3$$

Then we get a nonparametric specification – there are no  $\beta$ 's or other unknown parameters!

Note that nonparametric does not mean choiceless—we still had to choose 3 rather than 5 values to average

But there is no population parameter to *estimate*, just a modeling choice to *calibrate* 

### Non-parametric models: smooths

Non-parametric "smooths", like the moving average, fall on the border between plots of the data and traditional regression models

Scatterplots leave model "fitting" up to the viewer's eyes

Regression models, such as this linear regression on polynomials:

$$y_i = \beta_0 + \beta x_{1i} + \beta_1 x_{1i}^2 + \varepsilon_i$$

or this logit regression on linear terms

$$\Pr(y_i|x_i) = \text{logit}^{-1}(\beta_0 + \beta_1 x_{1i})$$

make assumptions the viewer must take on faith, esp. the specification

Choices about whether terms should enter as linear, cubed, logged, etc are often arbitrary

Smooths reduce the influence of the model on functional form, in favor of the data

### Non-parametric models: smooths

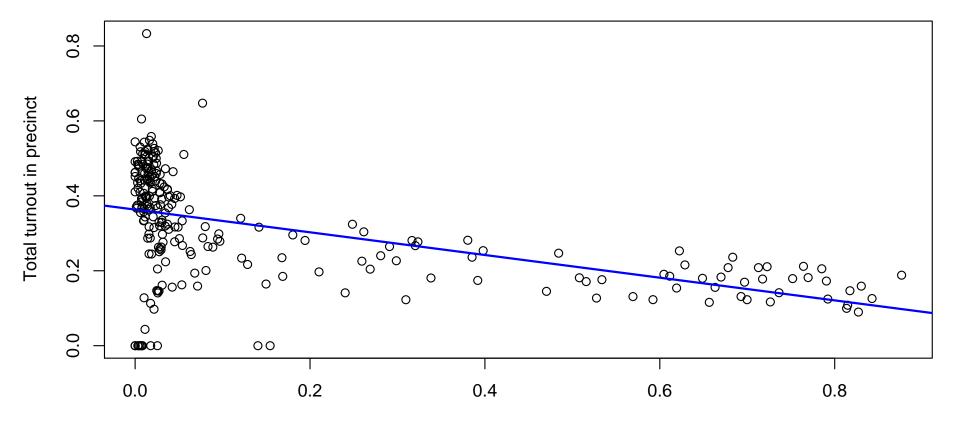
Some commonly used smoothers:

- running averages
- running medians
- loess
- splines
- kernel density (mainly useful for smoothing histograms)

Let's look at how a variety of smoothers deal with a simple dataset:

Turnout (turnout) as a function of %Hispanic voters (hisp) in a Pennsylvania State Senate election

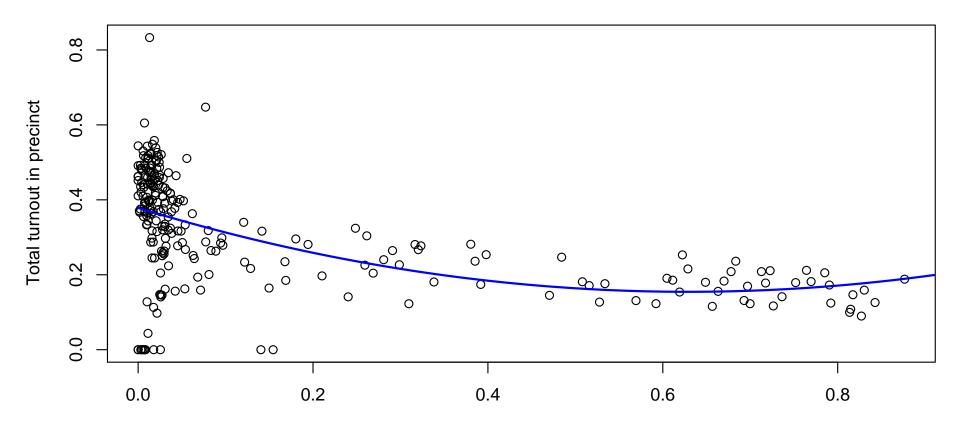
#### **Least squares fit**



Percent Hispanics among Voting Age Population, by precinct

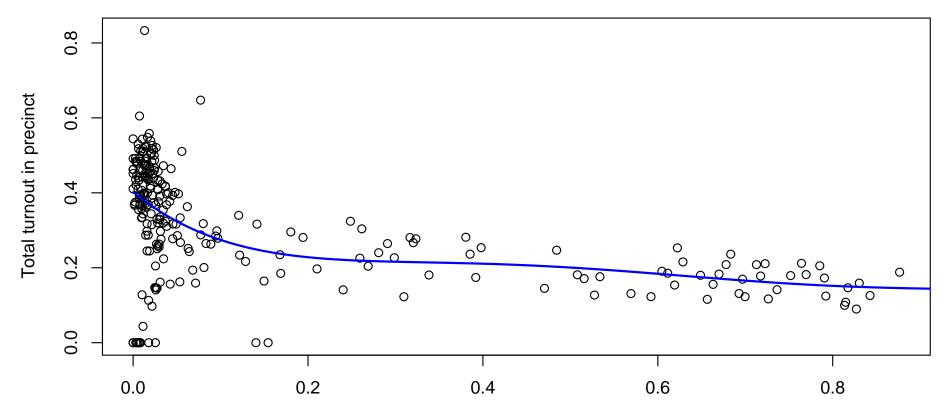
```
res.lm <- lm(turnout~hisp)
abline(coef(res.lm)[1], coef(res.lm)[2], col="blue", lwd=2)</pre>
```

#### Quadratic (LS) fit



Percent Hispanics among Voting Age Population, by precinct

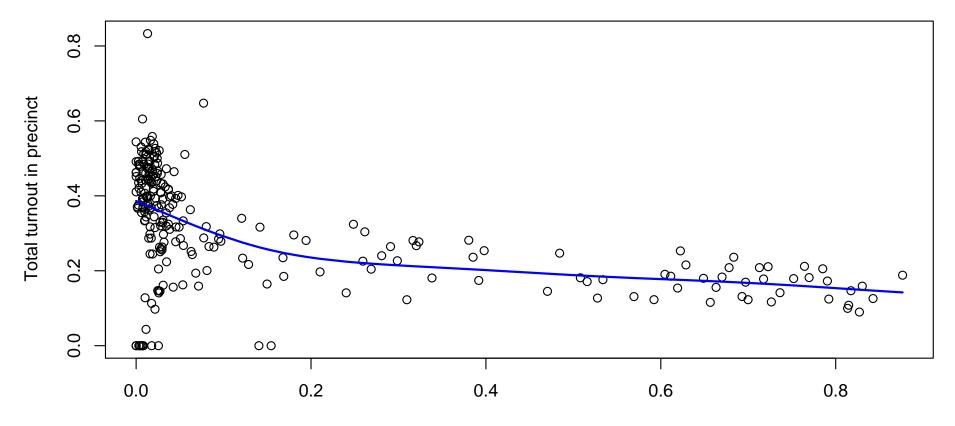
#### 5th order polynomial (LS) fit



Percent Hispanics among Voting Age Population, by precinct

res.q <- lm(t~hisp + I(hisp^2) + I(hisp^3) + I(hisp^4) + I(hisp^5))
etc. Danger of overfitting with any polynomial fit.</pre>

#### Smoothing splines (Smoothing determined by cross-validation)



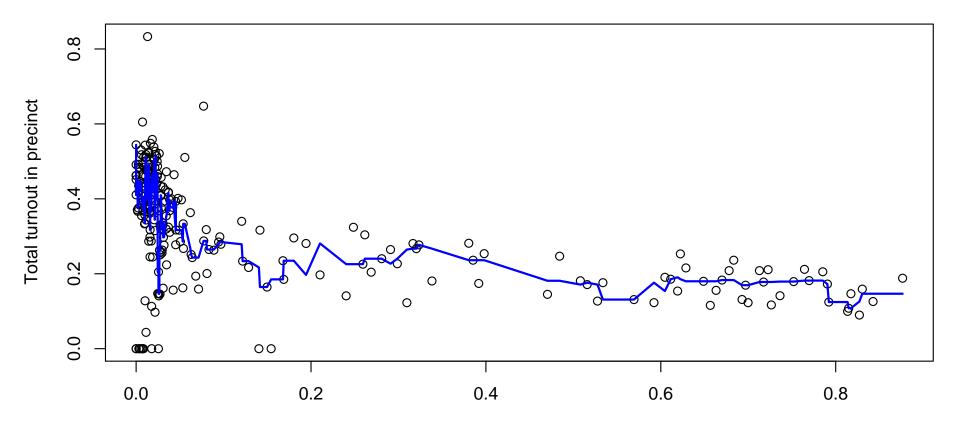
Percent Hispanics among Voting Age Population, by precinct

lines(smooth.spline(y=turnout, x=hisp), col="blue", lwd=2)

We can approximate high order polynomial fits with *smoothing splines* 

This function tries to avoid overfitting by selecting the "wiggliness" based on cross-validation

#### Running median (window of 5 observations)



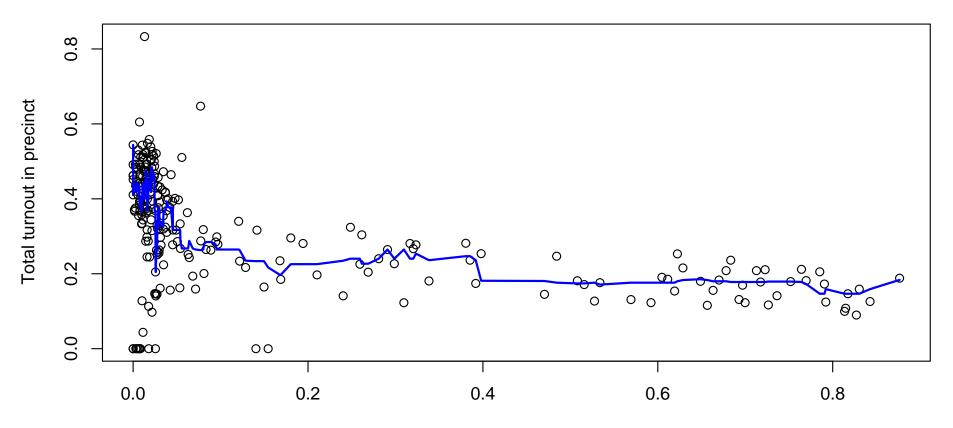
Percent Hispanics among Voting Age Population, by precinct

This works only if turnout and hisp both sorted by hisp first:

lines(x=hisp, y=runmed(turnout, k=5), lwd=2, col="blue")

Maximum robustness: not very smooth though

#### **Running median (window of 11 observations)**



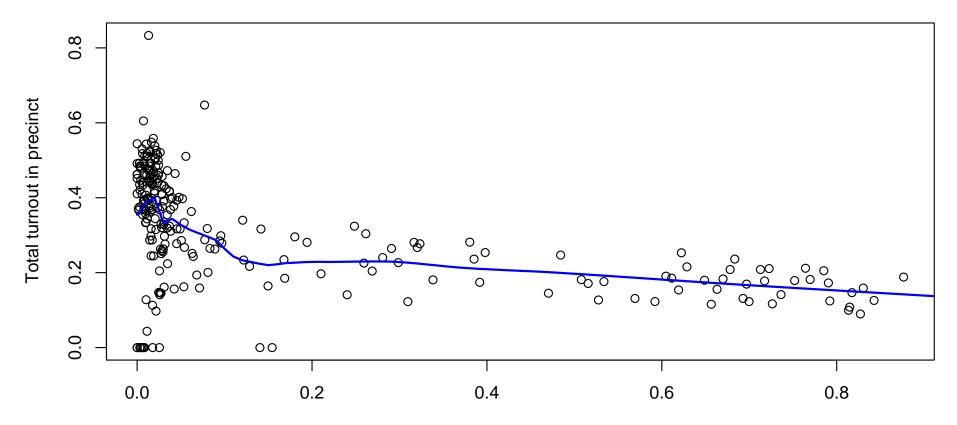
Percent Hispanics among Voting Age Population, by precinct

This works only if t and x both sorted by x *first*:

lines(x=hisp, y=runmed(turnout, k=11), lwd=2, col="blue")

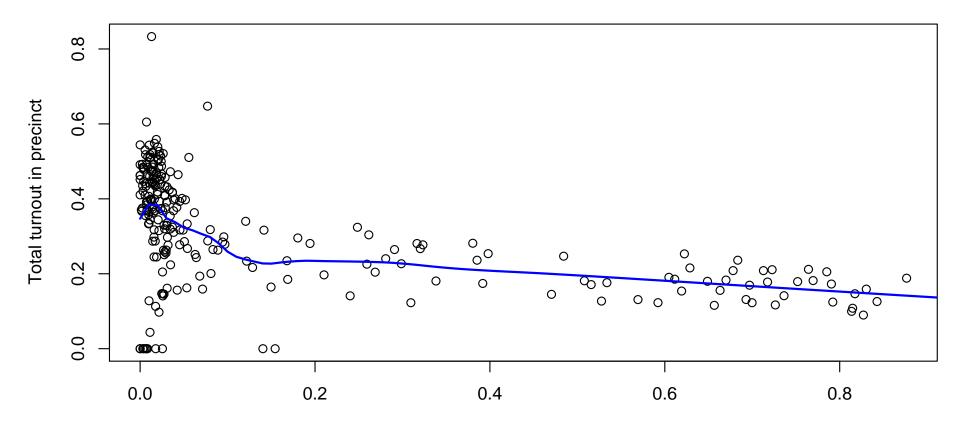
Maximum robustness: A bit smoother, but we'd like another alternative

#### Loess fit (bandwidth=0.25, order=1)



Percent Hispanics among Voting Age Population, by precinct

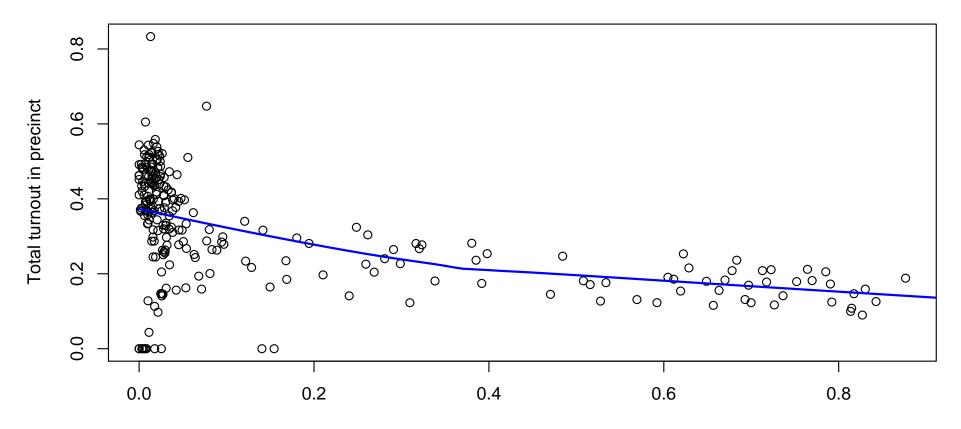
#### Loess fit (bandwidth=0.5, order=1)



Percent Hispanics among Voting Age Population, by precinct

 $Higher\ bandwidth\ \rightarrow \ a\ smoother\ plot$ 

#### Loess fit (bandwidth=0.95, order=1)



Percent Hispanics among Voting Age Population, by precinct

This is about as high as we go. Note that we find a kink.

## Non-parametric models: smooths

Lots of mathematical details here.

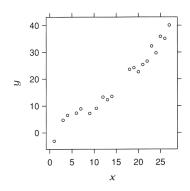
But key issue is how much flexibility to allow

In each model, there is a choice of how flexible the fit should be

- for loess and kernel, the bandwidth;
- for splines, the degree of smoothing and number of knots
- for running averages, the number and period of the averages

How do we perform this fit? And why is there no  $\beta$ ?

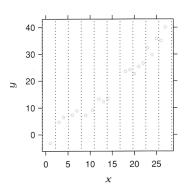
$$\hat{\mathbf{y}} = f_{\text{loess}}(\mathbf{x})$$



Adapted from: Cleveland, Visualizing Data

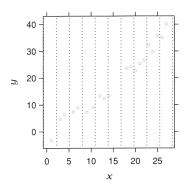
Let's fit a loess curve to the data at right

 Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*



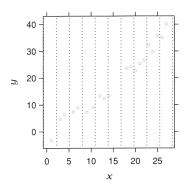
Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\alpha$



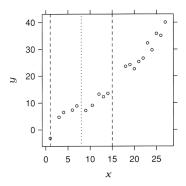
Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\boldsymbol{\alpha}$
- 3. Choose a polynomial order  $\lambda$ ; usually 1 or 2



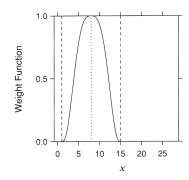
Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\alpha$
- 3. Choose a polynomial order  $\lambda$ ; usually 1 or 2
- 4. For each element in  $\mathbf{x}^*$ ,  $x_k^*$ ,
  - i. Select the an/2 nearest neighbors to  $x_k^*$ ,



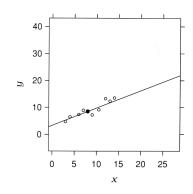
Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\alpha$
- 3. Choose a polynomial order  $\lambda$ ; usually 1 or 2
- 4. For each element in  $\mathbf{x}^*$ ,  $\mathbf{x}_k^*$ ,
  - i. Select the  $\alpha n/2$  nearest neighbors to  $x_k^*$ ,
  - ii. Apply weights descending in distance from  $x_k^*$



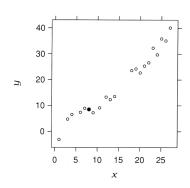
Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\alpha$
- 3. Choose a polynomial order  $\lambda$ ; usually 1 or 2
- 4. For each element in  $\mathbf{x}^*$ ,  $x_k^*$ ,
  - i. Select the  $\alpha n/2$  nearest neighbors to  $x_k^*$ ,
  - ii. Apply weights descending in distance from  $x_k^*$
  - iii. Fit y<sub>selected</sub> on the λth-order polynomial of x<sub>selected</sub> by WLS



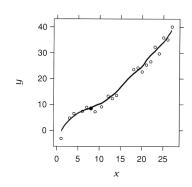
Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\alpha$
- 3. Choose a polynomial order  $\lambda$ ; usually 1 or 2
- 4. For each element in  $\mathbf{x}^*$ ,  $x_k^*$ ,
  - i. Select the  $\alpha n/2$  nearest neighbors to  $x_k^*$ ,
  - ii. Apply weights descending in distance from  $x_k^*$
  - iii. Fit  $\mathbf{y}_{selected}$  on the  $\lambda$ th-order polynomial of  $\mathbf{x}_{selected}$  by WLS
  - iv. Record only the WLS prediction of  $\hat{y}_k^*$  (discard the rest of the line)



Adapted from: Cleveland, Visualizing Data

- Choose a sequence x\* of k equally spaced points at which to calculate loess fits y\*
- 2. Choose a smoothing parameter  $\alpha$
- 3. Choose a polynomial order  $\lambda$ ; usually 1 or 2
- 4. For each element in  $\mathbf{x}^*$ ,  $\mathbf{x}_k^*$ ,
  - i. Select the  $\alpha n/2$  nearest neighbors to  $x_k^*$ ,
  - ii. Apply weights descending in distance from  $x_k^*$
  - iii. Fit  $\mathbf{y}_{selected}$  on the  $\lambda$ th-order polynomial of  $\mathbf{x}_{selected}$  by WLS
  - iv. Record only the WLS prediction of  $\hat{y}_k^*$  (discard the rest of the line)
- 5. Connect the fitted  $\hat{y}_k^*$ 's to make a curve.



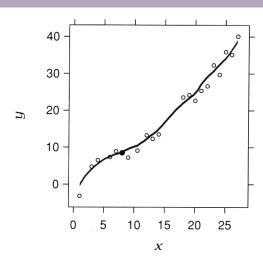
Adapted from: Cleveland, Visualizing Data

Key issue here is the bandwidth parameter,  $\alpha \in [0, 1]$ 

Lower values make a jerkier line; higher values a smoother one

If your data seem to curve, use a quadratic fit within loess ( $\lambda=2$ ) to get a better fit

Note: Our loess example is an artist's impression; let's look at some real ones at demonstrations.wolfram.com/



Adapted from: Cleveland, Visualizing Data

### What the heck is a "spline"?

Splines are a concept from carpentry, of all places

Splines let us summarize very complex curves with a few numbers

Basic idea:

- Imagine a flexible piece of wood.
- We pick it up and bend in many places;
- then tack it down (at "knots") to a board.

# The idea behind splines

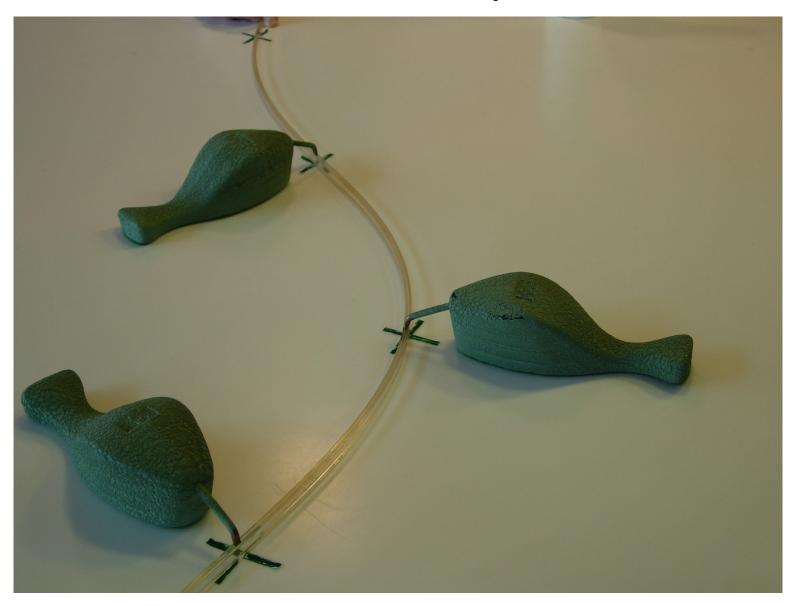


Photo of draftman's spline from Carl de Boor www.cs.wisc.edu/deboor/draftspline.html

## **Splines** in statistics

Many similar shapes can be approximated by local cubic polynomials, which we will call cubic splines

(Note: there are *many* kinds of splines.)

Even more than loess, Cubic splines rely on simple local structure to create global flexibility

- 1. Start with a few data points.
- 2. Connect every point to its nearest neighbor with a (twice differentiable) cubic polynomial
- 3. To make a "smoothing" or approximate spline, take the weighted average of the spline and the least squares fit
- 4. Choose knots & amt of smoothing subject to a penalized likelihood criterion, e.g.,

 $\max(2 \times \text{log-likelihood} - \text{trade-off} \times \text{smoothness penalty})$ 

# **Spline examples**

Using demonstration software

http://www.particleincell.com/blog/2012/bezier-splines/

## More complex example: Democracy & Development

The relationship between democracy, dictatorship, and economic development is well-explored in political science

Key recent work: Przeworksi, Alvarez, Cheibub & Limongi. Democracy and Development: Political Insitutions and Well-being in the World, 1950–1990

We'll borrow their data, but run much cruder models

In particular, PACL investigate selection bias and the difference between creation and sustenance of democracies.

We'll ignore these issues, and consider covariates of

the degree of civil liberties (CIVLIB: 1-7 scale)

the presence of democracy (REG: 0-1 binary)

### More complex example: Democracy & Development

### Our candidate covariates

GDPW GDP per capita in real international prices

EDT average years of education

ELF60 ethnolinguistic fractionalization

MOSLEM percentage of Muslims in country

OIL whether oil accounts for 50+% of exports

STRA count of recent regime transitions

NEWC whether county was created after 1945

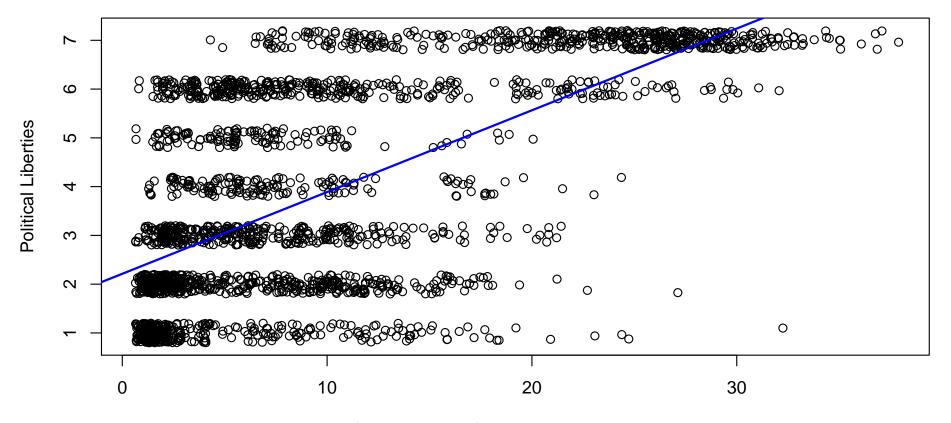
BRITCOL whether country was a British colony

Let's start with the simplest model we can run.

Is there a relationship between civil liberties and economic development?

Let's pretend linear models are appropriate for this categorical variable

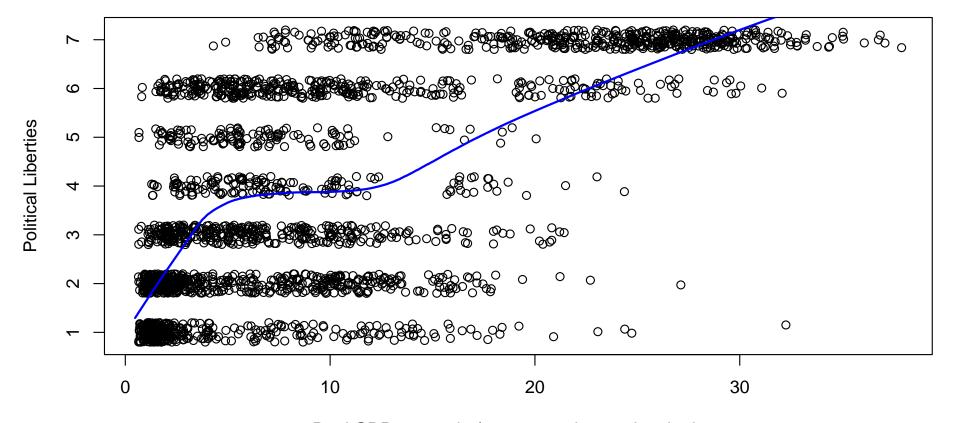
### Linear fit



Real GDP per capita/1000, 1985 international prices

Leaving aside the categorical nature of the data, does this fit look "right"?

### Loess fit (bandwidth=0.5, order=1)



Real GDP per capita/1000, 1985 international prices

Loess reveals "thresholds".

A problem: We haven't controlled for anything

Smoothers won't be much use if we're restricted to bivariate models

Fortunately, there is a generalization to the multivariate case. . .

## Generalized Additive Models (GAMs)

Generalized Additive Models are a generalization of GLMs

incorporate smooths into multiple regression, and into logit, probit, etc.

GAMs take the following form:

$$g^{-1}(\mu) = \alpha + \sum_{j=1}^{p} \beta_j X_j + \sum_{k=1}^{q} f_k(Z_k)$$

where y is a response,  $X_j$  and  $Z_k$  are covariates, and  $f_k$  is a smoothing function

Note we can estimate parametric relations for some covariates, and non-parametric for others

 $f_k$  are often splines or loess fits

### **GAM** for Democracy

Let's control for the rest of our variables while letting the degree of political liberties remain a smooth function of the level of development

### Sample code:

### Notes:

- This is the gam function in mgcv.
   There is another gam in library gam; slightly different
- Without specifying a distribution family, gam defaults to a linear Normal model
- We could specify more details of the smooth;
   this code just let's R find the best smoothing spline by cross-validation

### Output of summary(res.gam1)

Family: gaussian

Link function: identity

### Formula:

POLLIB ~ s(GDPW) + EDT + NEWC + BRITCOL + STRA + ELF60 + OIL + MOSLEM

### Parametric coefficients:

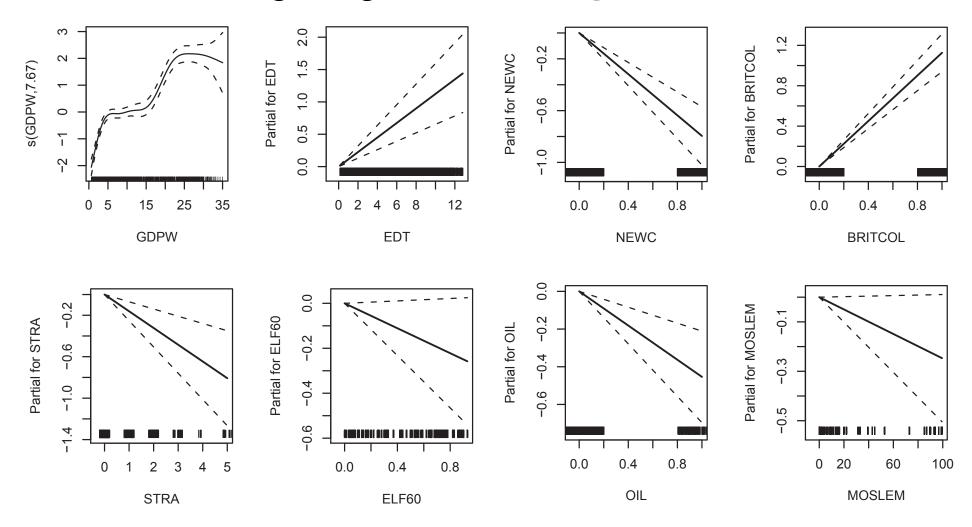
	Estimate	std. err.	t ratio	Pr(> t )
(Intercept)	3.8475	0.1552	24.8	< 2.22e-16
EDT	0.11236	0.02354	4.773	1.9840e-06
NEWC	-0.79545	0.1125	-7.07	2.3322e-12
BRITCOL	1.1263	0.09354	12.04	< 2.22e-16
STRA	-0.16153	0.04582	-3.525	0.00043583
ELF60	-0.2774	0.1521	-1.824	0.068387
OIL	-0.45368	0.1214	-3.737	0.00019278
MOSLEM	-0.0024823	0.001295	-1.917	0.055408

### Approximate significance of smooth terms:

	$\operatorname{edf}$	chi.sq	p-value
s(GDPW)	7.666	328.44	< 2.22e-16

R-sq.(adj) = 0.638 Deviance explained = 64.1% GCV score = 1.8114 Scale est. = 1.7934 n = 1580

# mgcv's gam has preset graphics



Above made with: plot(res.gam1,pages=1,all.terms=T) plus editing in Illustrator

### We can smooth over two dimensions

Suppose we wanted to smooth over both development & education:

One way is to let each smooth be additive to the response:

which lets more of the above plots be smooths

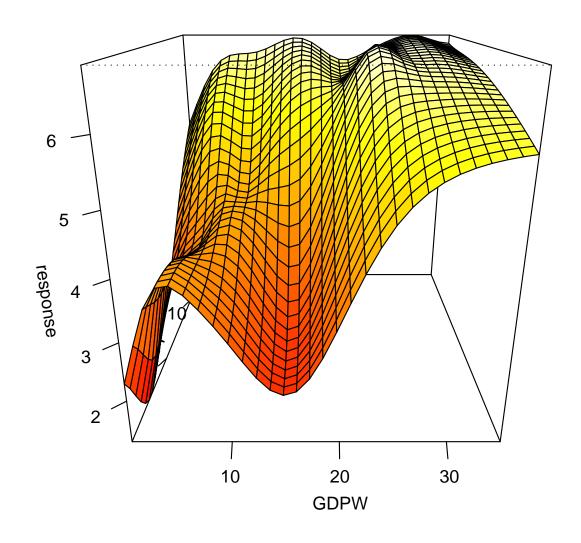
Another way is to let the smooths "interact." Sample code:

```
library(mgcv)
res.gam2 <- gam(POLLIB~s(GDPW,EDT)+NEWC+BRITCOL+STRA+ELF60+OIL+MOSLEM)
summary(res.gam2)</pre>
```

Now there is a smooth *surface* relating the joint levels of GDPW & EDT to POLLIB

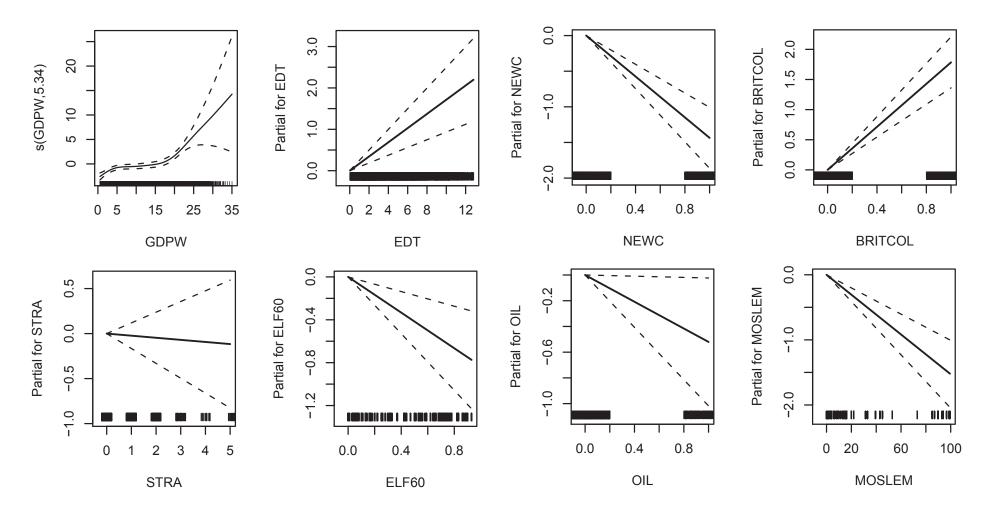
We can visualize this with

# A smooth over two covariates together



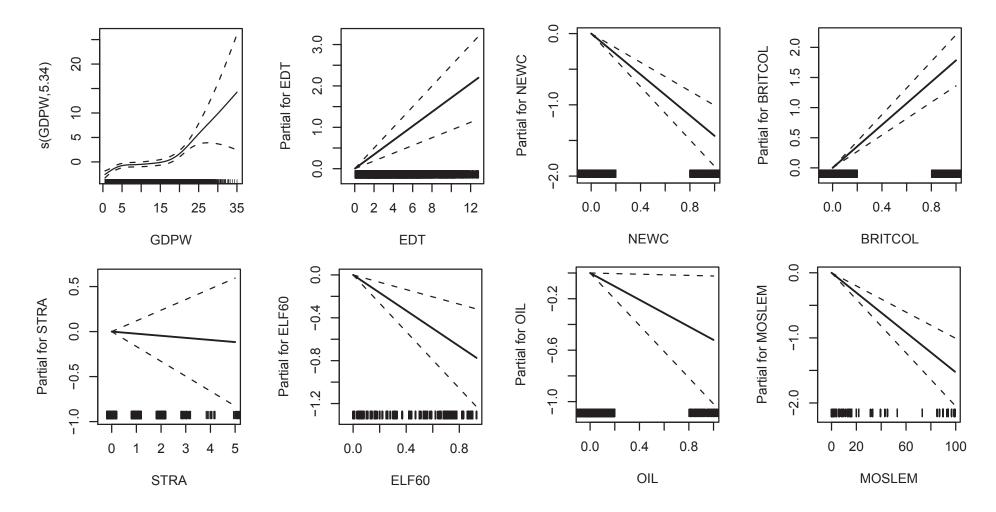
Interpretation is challenging, but this may be occassionally enlightening

## A logit example using GAM

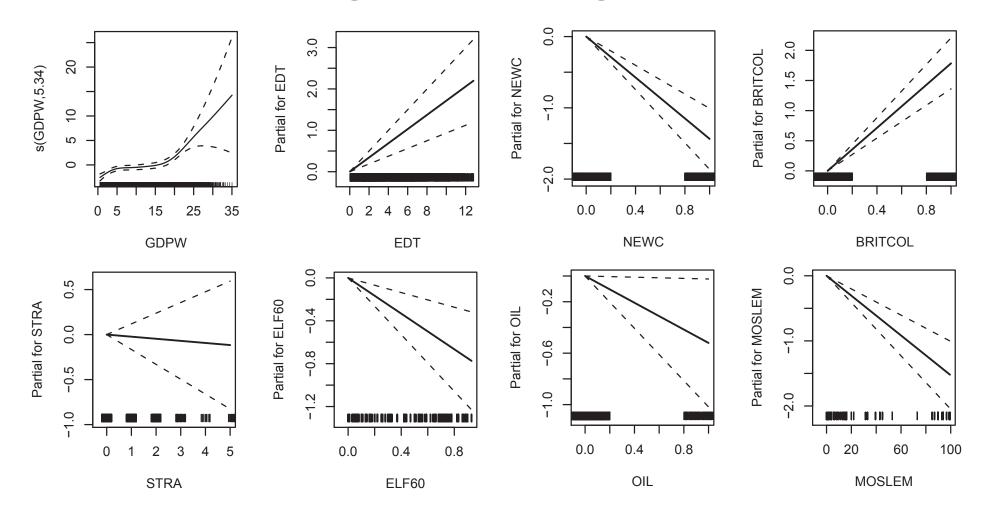


We can use GAM to fit GLM type models. What are the correlates of Democracy?

### A logit example using GAM

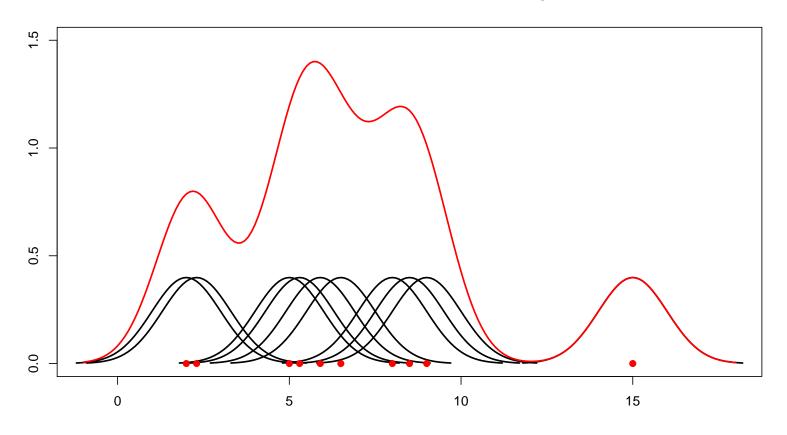


# A logit example using GAM



When the GAM is a nonlinear model (e.g., logit), the scale is difficult to interpret It would be better if the y-axes were in the units of Pr(REG)

### One last smooth: kernel density estimation



(Source for KDE diagrams: Stefanie Scheid - Introduction to Kernel Smoothing)

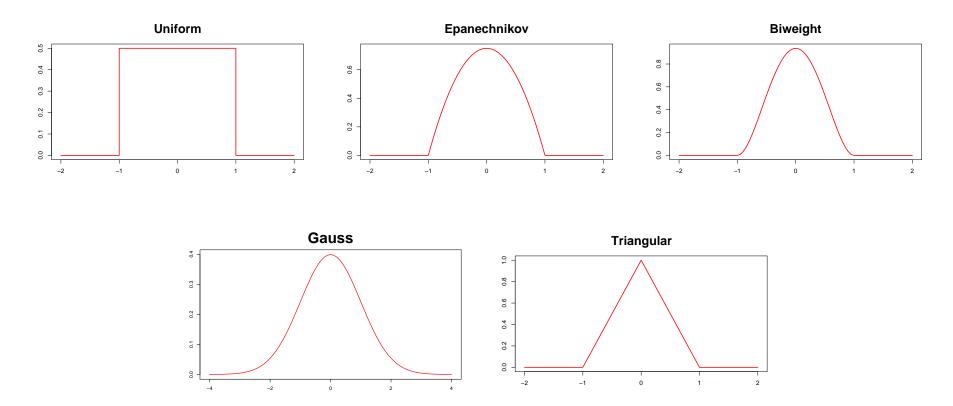
Another popular smoother is kernel density estimation (KDE)

KDE treats each data point as the center of a "kernel", and adds those kernels up

This let's the effect of each datapoint "smooth out".

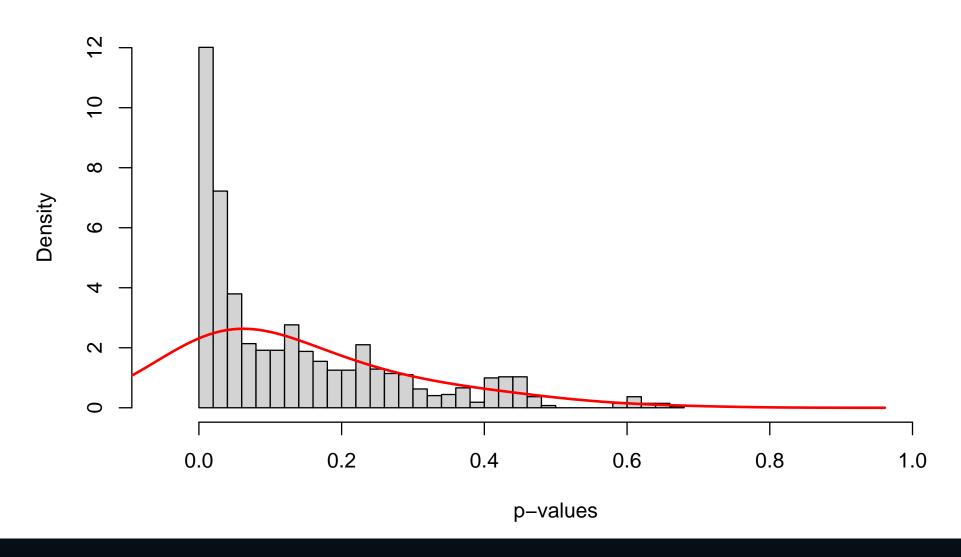
Shape of smoothing out is the kernel; degree of smoothing is the bandwidth

# Choices of kernels to place around datapoints

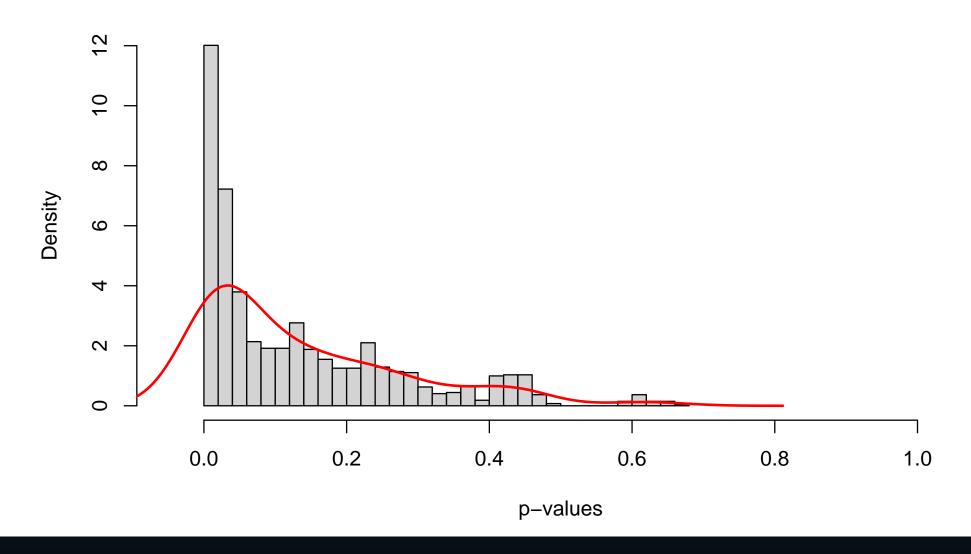


Epanechnikov minimizes error. But not usually important which you choose.

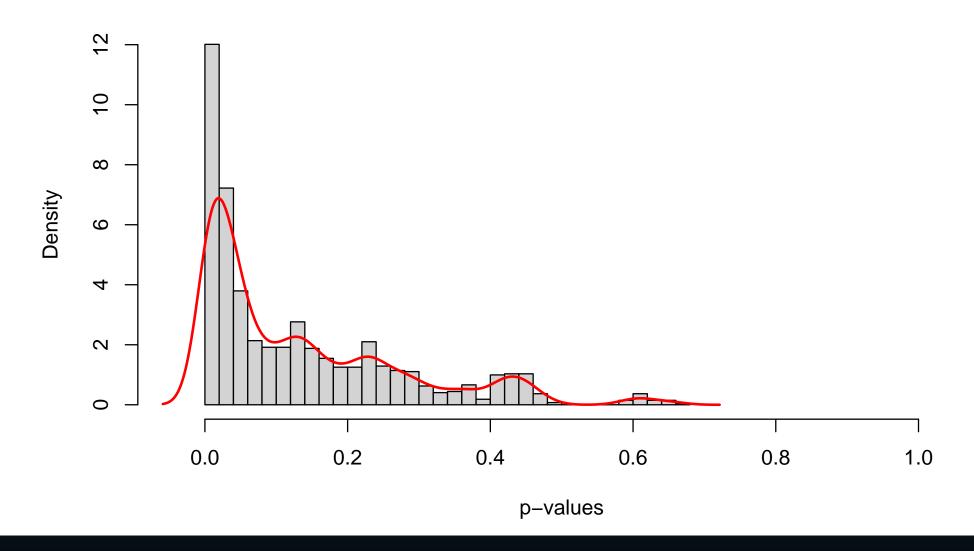




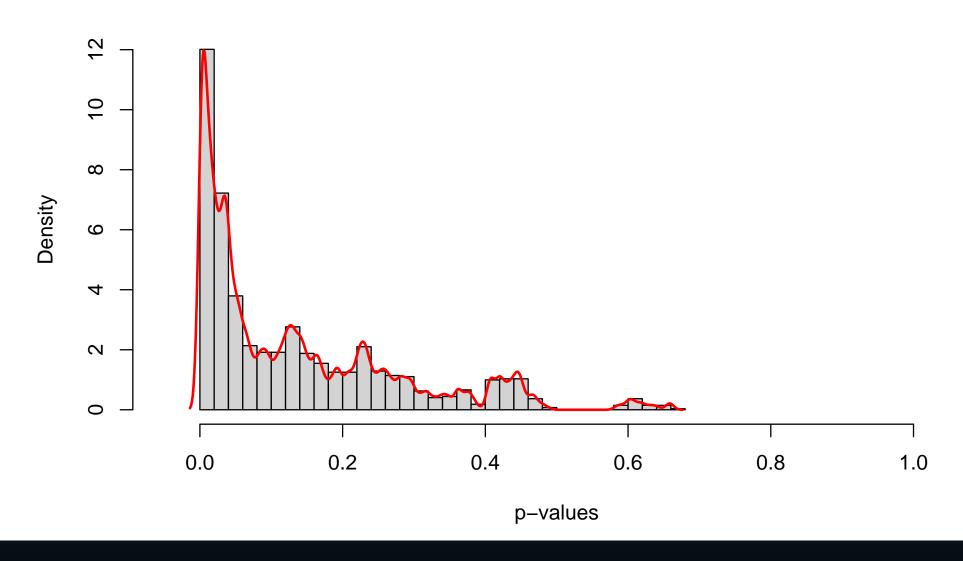
### KDE with b=0.05



### KDE with b=0.02



### **KDE** with **b=0.005**



### Uses of KDE

KDE is very useful for smoothing histograms of all kinds

If you draw from the predictive or posterior distribution of a model, you can smooth with KDE

This works even in two dimensions (e.g., you want the joint confidence region of two parameters)

Generally, if you have a 2D histogram, and you want contours, consider KDE

In R, see the density() command for one dimensional KDE

See kde2d in the MASS library for 2D version

### The curse of dimensionality

The biggest problem in visual display is that humans can only perceive 3 dimensions Most data, esp in social sciences, have many variables; potentially many dimensions Paper & computer screens are even more limiting: 2D

Three approachs to fighting the curse of dimensionality:

- Cleverly display all the data.
- Use models to reduce the number of dimensions
- Use time & motion

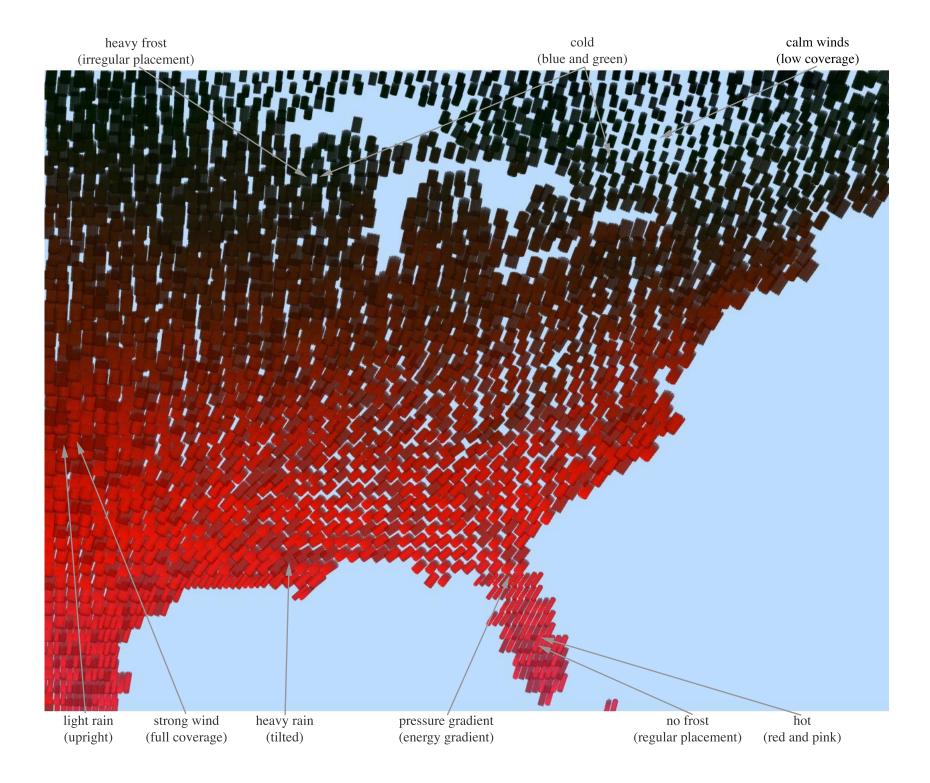
## Clever display of the whole dataset

This is what we've done all quarter:

- 1. Small multiples
- 2. Glyphs
- 3. Layer pre-attentive elements

We've seen lots of examples.

Here is a final example that pushes the envelope





Source: Christopher G. Healey, "Combining Perception and Impressionist Techniques for Nonphotorealistic Visualization of Multidimensional Data", SIGGRAPH 2001 Course 32: Nonphotorealistic Rendering in Scientific Visualization 2001, http://www.csc.ncsu.edu/faculty/healey/download/sig-course.01.pdf

Does the 5-dimensional version work "better" than the small multiples?

For look-up, no. But for gestalt impressions, perhaps worth looking at both.

I've experimented with the same 5 dimensions independently. Probably near the limit for a single diagram. Challenging to process.

So let's try to reduce dimensions while keeping most of the data

#### Principal Components Analysis (PCA)

PCA is related to factor analysis and multidimensional scaling

Details involve lots of linear algebra; instead, a conceptual summary

Imagine a dataset with k continuous variables plotted in k-space

#### Principal Components Analysis (PCA)

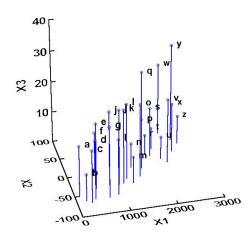
PCA is related to factor analysis and multidimensional scaling

Details involve lots of linear algebra; instead, a conceptual summary

Imagine a dataset with *k* continuous variables plotted in *k*-space

For example, k = 3

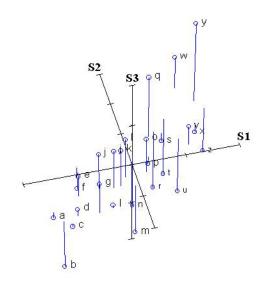
Note the different scales for each variable



Source: Michael Palmer, ordination.okstate.edu/PCA.htm

It will help to normalize these variables to have mean zero and unit variance

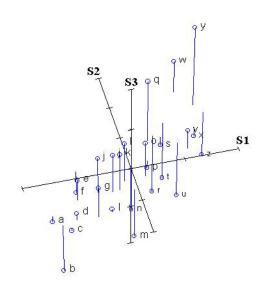
PCA finds a new set of dimensions,  $\leq k$ , that best explain and separate the variance in these data



It will help to normalize these variables to have mean zero and unit variance

PCA finds a new set of dimensions,  $\leq k$ , that best explain and separate the variance in these data

We search for new orthogonal axes (components), which we label:

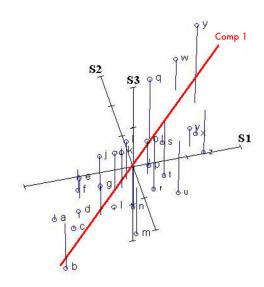


It will help to normalize these variables to have mean zero and unit variance

PCA finds a new set of dimensions,  $\leq k$ , that best explain and separate the variance in these data

We search for new orthogonal axes (components), which we label:

Component 1



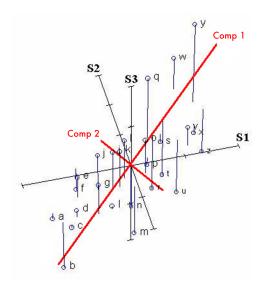
It will help to normalize these variables to have mean zero and unit variance

PCA finds a new set of dimensions,  $\leq k$ , that best explain and separate the variance in these data

We search for new orthogonal axes (components), which we label:

Component 1
Component 2

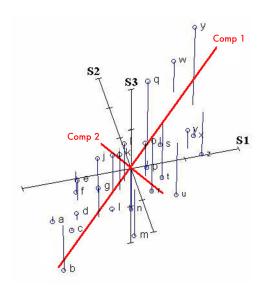
etc.



The first principal component is the axis that explains the most variation in the data

The second (third, etc.) principal component is the line orthogonal to the prior components that explains the greatest part of the remaining variation

Two principal components are often (not always) a nearly complete summary of variation on the k original dimensions

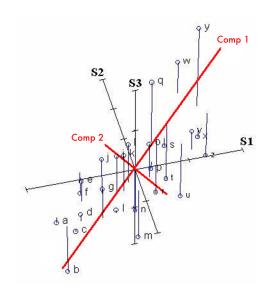


Each principal component can be seen as a linear combination of the original *k* axes

Very useful lower dimensional replacements for multidimensional data

Widely used to create "index" variables

Less arbitrary – and more informative – than averaging the underlying variables

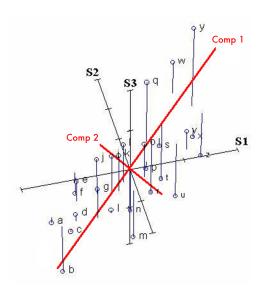


Note the graph at right has two sets of coordinate systems

The original variables are currently the plot coordinates, while the PCA coordinates are "plotted" as if they were data

What if we switched these coordinate systems, so the principal components are the plot coordinates?

And the original data and original axes are "plotted" in PCA space?



Example: FY 1992 state budget priorities

E.g., suppose a state divided its budget:

8% highways

30% education

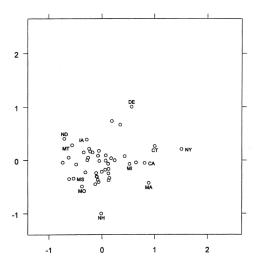
35% health

12% corrections

3% law enforcement

12% welfare

We wish to plot one point for each of the 50 states, but that takes 6 dimensions!

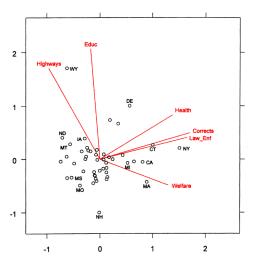


Solution: Use PCA to reduce the 6 budget dimensions to 2 principal components

Plot the 50 states locations in 2d PCA space

Source: William G. Jacoby, Statistical Graphics for Visualizing

Multivariate Data, Sage Paper 07-120



Source: William G. Jacoby, Statistical Graphics for Visualizing

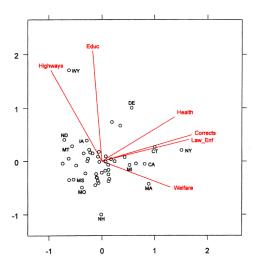
Multivariate Data, Sage Paper 07-120

Solution: Use PCA to reduce the 6 budget dimensions to 2 principal components

Plot the 50 states locations in 2d PCA space

Then add to the plot the projection into 2d PCA space of the original 6 budget dimensions

This simultaneous plot of dimensions and data is called the **biplot** 

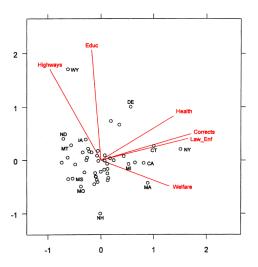


Tips for interpreting biplots:

Overlapping dimensions are perfectly correlated

Source: William G. Jacoby, Statistical Graphics for Visualizing

Multivariate Data, Sage Paper 07-120

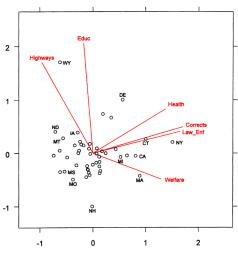


Tips for interpreting biplots:

Overlapping dimensions are perfectly correlated

Dimensions form an acute angle  $\Rightarrow$  somewhat positive correlation

Source: William G. Jacoby, Statistical Graphics for Visualizing
Multivariate Data, Sage Paper 07-120



Source: William G. Jacoby, Statistical Graphics for Visualizing

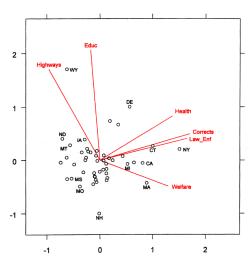
Multivariate Data, Sage Paper 07-120

Tips for interpreting biplots:

Overlapping dimensions are perfectly correlated

Dimensions form an acute angle ⇒ somewhat positive correlation

Dimensions form a 90° angle ⇒ orthogonal



Tips for interpreting biplots:

Overlapping dimensions are perfectly correlated

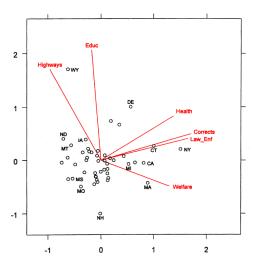
Dimensions form an acute angle  $\Rightarrow$  somewhat positive correlation

Dimensions form a 90 $^{\circ}$  angle  $\Rightarrow$  orthogonal

Dimensions form an obtuse angle ⇒ somewhat negative correlation

Source: William G. Jacoby, Statistical Graphics for Visualizing

Multivariate Data, Sage Paper 07-120



Source: William G. Jacoby, Statistical Graphics for Visualizing

Multivariate Data, Sage Paper 07-120

Tips for interpreting biplots:

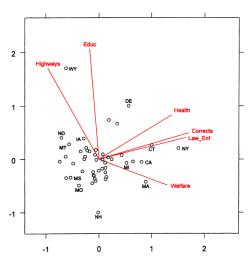
Overlapping dimensions are perfectly correlated

Dimensions form an acute angle  $\Rightarrow$  somewhat positive correlation

Dimensions form a 90 $^{\circ}$  angle  $\Rightarrow$  orthogonal

Dimensions form an obtuse angle  $\Rightarrow$  somewhat negative correlation

Dimensions form a 180 $^{\circ}$  line  $\Rightarrow$  perfect inverse correlation



Source: William G. Jacoby, Statistical Graphics for Visualizing

Multivariate Data, Sage Paper 07-120

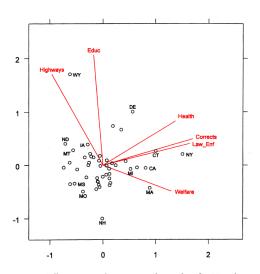
More tips for interpreting biplots:

Dimensions meet at the origin (0,0). (Why?)

"Short" dimensions load poorly on components. (Why?)

Data far from the origin are outliers

Distances between points are Mahalanobis distances (variables standardized to have unit variances)



How to do this in R:

res <- princomp(~ health

+ correct

+ lawenf + welfare

+ highways

+ educ)

biplot(res)

biplot() has several options, but is surprisingly inflexible

Source: William G. Jacoby, Statistical Graphics for Visualizing

Multivariate Data, Sage Paper 07-120

Emily Kalah Gade
(UW-Political Science)
provides data on the number
of documents published on
federal websites mentioning
topics related to climate
change

We have counts of the number of documents mentioning each topic by each agency by year

How could we visualize these

#### Topics (Dimensions)

freshwater pollution **IPCC** global warming food security climate change natural disaster greenhouse gas anthropogenic desertification forest conservation security of food climate research unit ocean acidification anthropocene climategate

#### Agencies (Cases)

usda.gov house.gov ed.gov hhs.gov doi.gov senate.gov dot.gov whitehouse.gov usdoj.gov va.gov dol.gov state.gov treasury.gov energy.gov dhs.gov commerce.gov defense.gov dod.gov other

#### Topics (Dimensions) Agencies (Cases) freshwater usda.gov pollution house.gov **IPCC** ed.gov global warming hhs.gov food security doi.gov climate change senate.gov natural disaster dot.gov greenhouse gas whitehouse.gov anthropogenic usdoj.gov desertification va.gov forest conservation dol.gov security of food state.gov climate research unit treasury.gov ocean acidification energy.gov anthropocene dhs.gov climategate commerce.gov defense.gov dod.gov other

1. Could make an agency  $\times$  topic contingency table.

Hard to read

#### Topics (Dimensions)

freshwater pollution IPCC

global warming

food security climate change natural disaster greenhouse gas

anthropogenic desertification

forest conservation security of food climate research unit ocean acidification

anthropocene

climategate

#### Agencies (Cases)

usda.gov house.gov

ed.gov

hhs.gov doi.gov

senate.gov
dot.gov

whitehouse.gov

usdoj.gov va.gov

dol.gov state.gov

treasury.gov

energy.gov dhs.gov

commerce.gov
defense.gov

dod.gov

1. Could make an agency imes topic contingency table.

Hard to read

2. Could make a heatmap, as we did with trade flows.

#### Topics (Dimensions)

freshwater pollution **IPCC** global warming food security climate change natural disaster greenhouse gas anthropogenic desertification forest conservation security of food climate research unit ocean acidification anthropocene climategate

#### Agencies (Cases)

usda.gov house.gov ed.gov hhs.gov doi.gov senate.gov dot.gov whitehouse.gov usdoj.gov va.gov dol.gov state.gov treasury.gov energy.gov dhs.gov commerce.gov defense.gov dod.gov other

1. Could make an agency imes topic contingency table.

Hard to read

2. Could make a heatmap, as we did with trade flows.

Can we instead use a biplot to reduce the number of plotted dimensions to 2?

YES, using correspondence analysis, an analog of PCA for contingency tables

#### Topics (Dimensions)

freshwater pollution **IPCC** global warming food security climate change natural disaster greenhouse gas anthropogenic desertification forest conservation security of food climate research unit ocean acidification anthropocene climategate

#### Agencies (Cases)

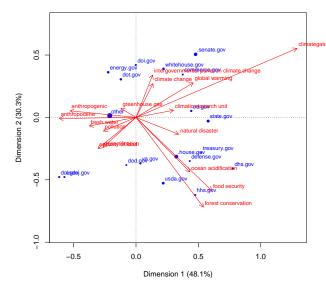
usda.gov house.gov ed.gov hhs.gov doi.gov senate.gov dot.gov whitehouse.gov usdoj.gov va.gov dol.gov state.gov treasury.gov energy.gov dhs.gov commerce.gov defense.gov dod.gov other

Just as with PCA, a correspondence analysis biplot shows a 2D view of a high dimensional space Here we see a cloud of points relative to 14(!) axes

Which topics (dimensions) are similar? Which agencies (points) are similar?

Which agencies load strongly on which dimensions? How sure are you of this?

#### Climate topics in official reports dated 2013



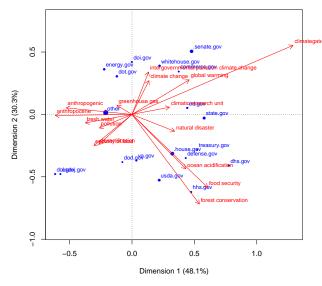
Data: Emily Kalah Gade (UW)

Figure: Will Lowe (Univ. of Mannheim)

#### Climate topics in official reports dated 2013

```
# Roughly, code is:
res <- ca(data)
plot(res,
     mass=c(TRUE, TRUE)
 but see plot.ca()
  for more options
  -- better than
     biplot()
```

How to do it in R:



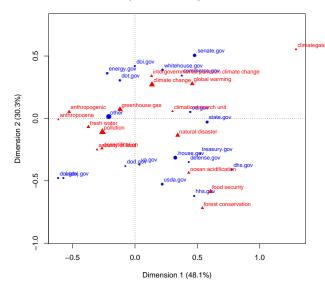
Data: Emily Kalah Gade (UW)
Figure: Will Lowe (Univ. of Mannheim)

#### Climate topics in official reports dated 2013

With large numbers of dimensions, the arrows can get distracting

Once you understand the biplot, you don't need them

Not an easy plot to explain, but a powerful way to explore many dimensions of data at once



Data: Emily Kalah Gade (UW)
Figure: Will Lowe (Univ. of Mannheim)